Stat 102C HW1: Answer Key

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Problem 1.1

$$P(Y \le x) = P(F^{-1}(U) \le x) \tag{1}$$

$$= P(U \le F(x)) \tag{2}$$

$$= \int_0^{F(x)} du \tag{3}$$

$$= F(x) \tag{4}$$

Problem 1.2

Textbook, p.44 $P(F(X) \leq u) = P[X \leq F^{-1}(u)] = F(F^{-1}(u)) = u$

Problem 2.1

The pdf is: $f(x) = \lambda exp(-\lambda x)$ So, the CDF is: $F(x) = \int_0^x f(x) dx = 1 - \exp(-\lambda x)$

Problem 2.2

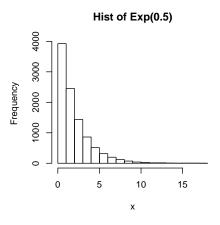
 $u = 1 - \exp(-\lambda x)$ $1 - u = \exp(-\lambda x)$ $\log(1 - u) = -\lambda x$

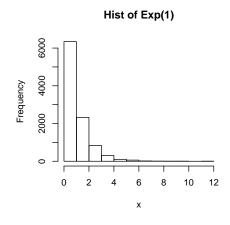
 $x = -\log(1 - u)/\lambda$

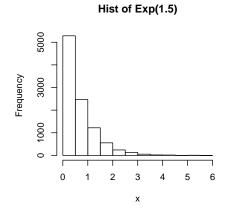
Since $u \in [0,1]$, this is equivalent to $x = -\log(u)/\lambda$. Isn't that prettier?

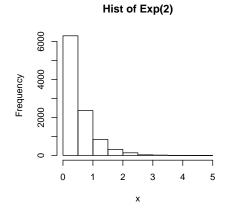
Problem 2.3 and 2.4

```
exp_invcdf = function(n, lambda) {
    u = runif(n)
    return(-log(u)/lambda)
}
par(mfrow = c(2, 2))
lambda_nums = seq(0.5, 2, 0.5)
n = 10000
for (lambda in lambda_nums) {
    hist(exp_invcdf(n, lambda), xlab = "x", main = paste0("Hist of Exp(", lambda, ")"))
}
```









Problem 3.1

This is identical to exercise 2.5. Because X is generated from g, so we integral over g.

$$P(U < f(X)/Mg(X)) = \int_{-\infty}^{\infty} \frac{f(x)}{Mg(x)} g(x) dx$$
 (5)

$$= \int_{-\infty}^{\infty} \frac{f(x)}{M} \mathrm{d}x \tag{6}$$

$$=\frac{1}{M}\tag{7}$$

Problem 3.2

See p.52 in your textbook.

Problem 4.1

I'll tell you the real way to prove this, and the intuitive explanation. Proof:

$$P(X < x | X > c) = \frac{P(X < x, X > c)}{P(X > C)}$$
(8)

$$F(x) = \frac{\int_C^x \phi(x) \, \mathrm{d}x}{1 - \Phi(C)} \tag{9}$$

$$f(x) = \frac{d}{dx} \left[\frac{\int_C^x \phi(x) \, \mathrm{d}x}{1 - \Phi(C)} \right] \tag{10}$$

$$=\frac{\phi(x)}{1-\Phi(C)}\tag{11}$$

Intuitive: The shape of PDF over the interval $[C, \infty]$ remains the same as it was with N(0,1), but now the probability outside of that range is 0. So the pdf over that range must be shifted upward by some constant.

What is that constant? Remember that the area under the new pdf needs to sum to one. $\Phi(x)$ denotes the area under N(0,1) from 0 to x, so $1-\Phi(x)$ denotes the area under N(0,1) from x to ∞ . Since $\phi(x)$ is the pdf of N(0,1), the new pdf is: $f(x) = \phi(x)/(1-\Phi(C))$ for $x \ge C > 0$

Problem 4.2

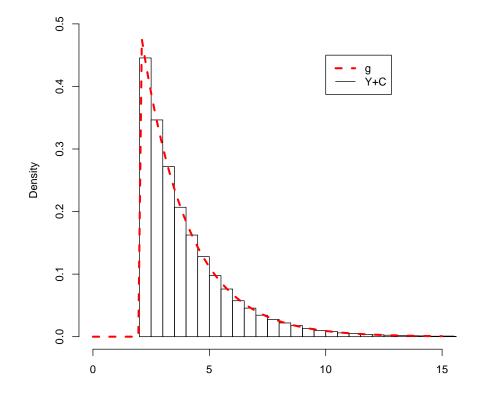
$$G_X(x) = \int_C^x g(x) dx = 1 - \exp[-\lambda(x - C)]$$
 (12)

$$G_Y(y) = P(Y < y) = P(X < y + C) = G_X(y + C) = 1 - \exp(-\lambda y)$$
 (13)

Hence, $Y \sim \mathcal{E}xp(\lambda)$.

Just to show this graphically...

Comparison of Random Variables



Problem 4.3

Let $\frac{\partial f(x)/g(x)}{\partial x} = 0$, we will get $x = \lambda$. Thus,

$$M(\lambda) = \frac{1}{\lambda\sqrt{2\pi}[1 - \Phi(C)]} \exp\{\frac{\lambda^2}{2} - \lambda C\}$$

Let $\frac{\partial M(\lambda)}{\partial \lambda}=0$, we will get $\lambda=\frac{C+\sqrt{C^2+4}}{2}$. As the code below shows, if C=1 the optimal lambda $\lambda=1.618$, and the M at that point is M = 1.141.

Problem 4.4

```
f = function(x, C) dnorm(x)/(1 - pnorm(C))
g = function(x, lambda, C) lambda * exp(-lambda * (x -
    C))
M = function(lambda, C) {
    f(lambda, C)/g(lambda, lambda, C)
```

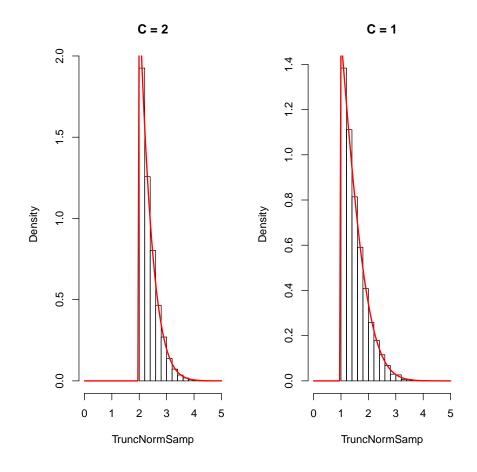
```
samplefrom = function(C) {
   lambda = (C + sqrt(C^2 + 4))/2
   m = M(lambda, C)
   y = rexp(1, lambda)
   x = y + C
   u = runif(1)
   i = 1
   # u*M*g(x) > f(x)
   while (u * m * g(x, lambda, C) > f(x, C)) {
        y = rexp(1, lambda)
        x = y + C
       u = runif(1)
        i = i + 1
   res = list(counter = i, sample = x)
   return(res)
```

Problem 4.5

```
Nsim = 10000
library(truncnorm)
par(mfrow = c(1, 2))

TruncNormSamp = replicate(Nsim, samplefrom(2)$sample)
hist(TruncNormSamp, freq = FALSE, main = "C = 2", xlim = c(0, 5))
curve(dtruncnorm(x, a = 2), col = "red", lwd = 2, add = T)

TruncNormSamp = replicate(Nsim, samplefrom(1)$sample)
hist(TruncNormSamp, freq = FALSE, main = "C = 1", xlim = c(0, 5))
curve(dtruncnorm(x, a = 1), col = "red", lwd = 2, add = T)
```

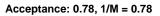


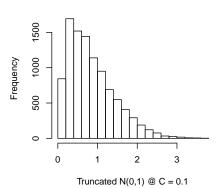
```
rate = function(Nsim, C) {
   j = 0
   freq = 0
   for (j in 1:Nsim) freq = freq + samplefrom(C)$counter
   lambda = (C + sqrt(C^2 + 4))/2
   m = M(lambda, C)
   res = list(rate = Nsim/freq, prob = 1/m)
   return(res)
rate(Nsim, 2) # see acceptance rate when C = 2
## $rate
## [1] 0.9341
##
## $prob
## [1] 0.9336
rate(Nsim, 1) # see acceptance rate when C = 2
## $rate
## [1] 0.8754
##
## $prob
## [1] 0.8765
```

An alternative way for Problem 4.4 and 4.5

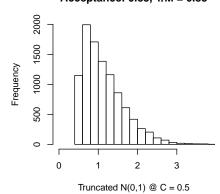
```
f = function(x, C) {
    ifelse(x >= C, dnorm(x)/(1 - pnorm(C)), 0)
}
g = function(x, lambda, C) {
    # ifelse(x >= C, lambda * exp(-lambda * (x-C)), 0)
    ifelse(x >= C, dexp((x - C), lambda), 0)
}
M_funct = function(lambda, C) {
    optimize(f = function(x) {
        f(x, C)/g(x, lambda, C)
    }, interval = c(C, C + (20 * lambda)), maximum = TRUE)$objective
}
gen_truncNorm = function(nsamples, Mfunct, C = 1) {
    # finds lambda and M
    optimum = optimize(f = function(L) M_funct(L, C),
        interval = c(0.1, 7), maximum = FALSE)
    lambda = optimum$minimum
    M_min = optimum$objective
```

```
nsim = nsamples
    # start with num_samples, add 1 for each additional
    # simulation
   target = numeric(nsamples)
   for (i in 1:nsamples) {
        u = runif(1) * M_min
       y = rexp(1, lambda) + C
        while (u > f(y, C)/g(y, lambda, C)) {
            u = runif(1) * M_min
            y = rexp(1, lambda) + C
           nsim = nsim + 1
        target[i] = y
   return(list(x = target, accept = nsamples/nsim,
       M = M_{\min})
N = 10000
Cvals = c(0.1, 0.5, 1, 2)
par(mfrow = c(2, 2))
for (Cstar in Cvals) {
   truncNorm = gen_truncNorm(N, M_funct, C = Cstar)
   hist(truncNorm$x, xlim = c(0, max(truncNorm$x)),
        xlab = paste0("Truncated N(0,1) @ C = ", Cstar),
        main = pasteO("Acceptance: ", round(truncNorm$accept,
            2), ", 1/M = ", round(1/truncNorm$M, 2)))
```

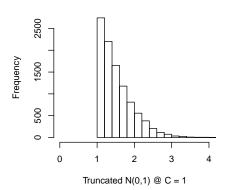




Acceptance: 0.83, 1/M = 0.83



Acceptance: 0.87, 1/M = 0.88



Acceptance: 0.93, 1/M = 0.93

