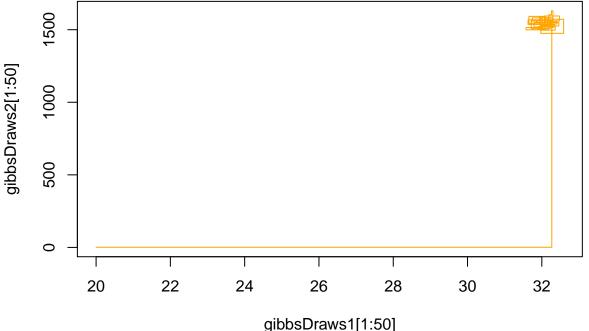
Lab3

Elon Brange, Ludwig Thaung 5/19/2019

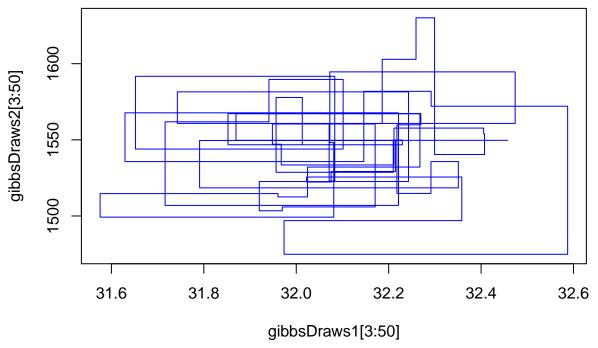
Task 1

(a)

```
rainfallData<-read.table("rainfall.dat",header=TRUE)</pre>
plotColors = list('red', 'green', 'blue', 'yellow', 'black', 'orange')
# Setup
mu1 <- 1
mu2 <- -1
rho <- 0.9
mu <- c(mu1,mu2)
Sigma = matrix(c(1, rho, rho, 1), 2, 2)
nDraws <- 600 # Number of draws
v0 = 2
mu0 = 20
sigma0 = 1
tau0 = 20
currentMu = mu0
currentSigma = sigma0
currentTau = tau0
currentV = v0
n = nrow(rainfallData)
averageRain = sum(rainfallData)/nrow(rainfallData)
sigma = var(rainfallData-averageRain)[1]
gibbsDraws <- matrix(0,nDraws,2)</pre>
for (i in 1:nDraws) {
  currentTau = 1/((n/currentSigma) + 1/tau0)
  w = (n/currentSigma)/(n/currentSigma + 1/tau0)
  currentMu = w*averageRain + (1 - w)*mu0
  currentMu <- rnorm(1, currentMu, currentTau)</pre>
  gibbsDraws[i,1] <- currentMu</pre>
  currentV = n + v0
  currentSigma = (v0*sigma0 + sum((rainfallData - currentMu)^2)/(n + v0))
  currentSigma <- rinvchisq(n = 1, df = currentV, scale = currentSigma)</pre>
  gibbsDraws[i, 2] <- currentSigma</pre>
averages <- matrix(0, nDraws, 2)
for (i in (1:((nDraws-500)/10))) {
  averages[i, 1] = sum(gibbsDraws[(i*10):(i*10+500), 1])/500
  averages[i, 2] = sum(gibbsDraws[(i*10):(i*10+500), 2])/500
}
#plot(averages[1:(nDraws-500)/10, 1], type='l')
```

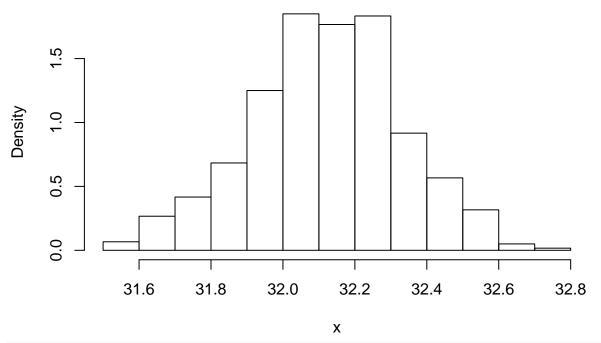


plot(gibbsDraws1[3:50],gibbsDraws2[3:50], type ='s', col ='blue')



```
hist(gibbsDraws[,1], freq = FALSE, main='Gibbs draws', xlab='x')
#hist(gibbsDraws[,1], freq = FALSE, main='Gibbs draws', ylim = c(0,0.5), xlab='x')
lines(seq(-2,4,by=0.01),dnorm(seq(-2,4,by=0.01), mean = 1), col = 'orange',
    lwd = 1)
```

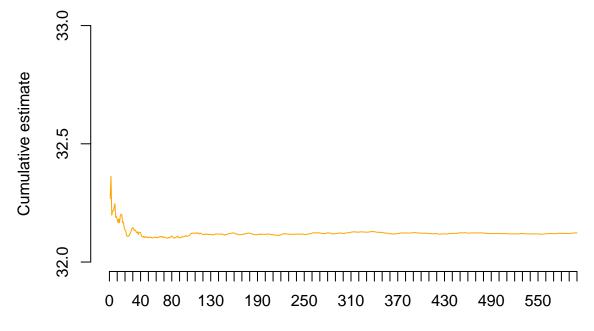
Gibbs draws



cusumData = cumsum(gibbsDraws[,1])/seq(1,nDraws)
minY = floor(min(cusumData))

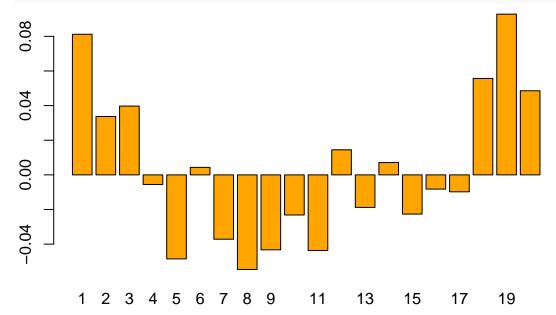
```
maxY = ceiling(max(cusumData))
plot(1:nDraws, cusumData, type = "l", col = 'orange', ylab='Cumulative estimate',
    lwd = 1, axes=FALSE, xlab = 'MCMC iteration', xlim = c(0,nDraws),
    ylim = c(minY,maxY), main = 'Cusum - Gibbs')
lines(seq(1,nDraws),1*matrix(1,1,nDraws),col = 'orange', lwd=1)
axis(side = 1, at = seq(0, nDraws, by = 10))
axis(side = 2, at = seq(minY, maxY, by = 0.5))
```

Cusum - Gibbs



MCMC iteration

```
a = acf(gibbsDraws[,1], main='Gibbs draws', lag.max = 20, plot = F)
barplot(height = a$acf[-1], names.arg=seq(1,20), col = 'orange')
```



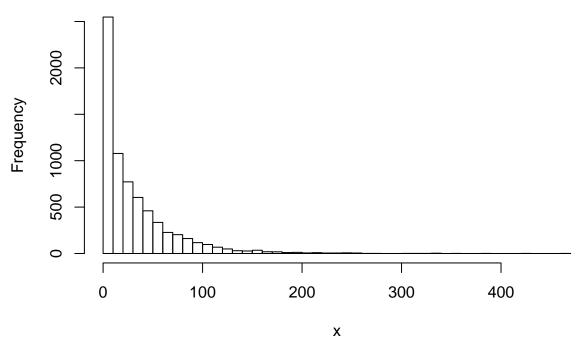
(b)

```
# Estimating a simple mixture of normals
# Author: Mattias Villani, IDA, Linkoping University. http://mattiasvillani.com
#########
              BEGIN USER INPUT ###############
# Data options
rawData <- read.table("rainfall.dat",header=TRUE)</pre>
x <- as.matrix(rawData)</pre>
# Model options
nComp <- 2
              # Number of mixture components
# Prior options
alpha <- 10*rep(1,nComp) # Dirichlet(alpha)</pre>
muPrior <- rep(30,nComp) # Prior mean of mu
tau2Prior <- rep(10,nComp) # Prior std of mu
sigma2_0 <- rep(var(x),nComp) # s20 (best guess of sigma2)</pre>
nu0 <- rep(4,nComp) # degrees of freedom for prior on sigma2
# MCMC options
nIter <- 10 # Number of Gibbs sampling draws
# Plotting options
plotFit <- TRUE
lineColors <- c("blue", "green", "magenta", 'yellow')</pre>
sleepTime <- 0 # Adding sleep time between iterations for plotting
###### Defining a function that simulates from the
rScaledInvChi2 <- function(n, df, scale){
  return((df*scale)/rchisq(n,df=df))
}
###### Defining a function that simulates from a Dirichlet distribution
rDirichlet <- function(param){</pre>
  nCat <- length(param)</pre>
  piDraws <- matrix(NA,nCat,1)</pre>
  for (j in 1:nCat){
    piDraws[j] <- rgamma(1,param[j],1)</pre>
  piDraws = piDraws/sum(piDraws) # Diving every column of piDraws by the sum of the elements in that co
  return(piDraws)
# Simple function that converts between two different representations of the mixture allocation
S2alloc <- function(S){
  n \leftarrow dim(S)[1]
  alloc \leftarrow rep(0,n)
  for (i in 1:n){
    alloc[i] <- which(S[i,] == 1)</pre>
  return(alloc)
```

```
# Initial value for the MCMC
nObs <- length(x)
S <- t(rmultinom(nObs, size = 1 , prob = rep(1/nComp,nComp))) # nObs-by-nComp matrix with
mu <- quantile(x, probs = seq(0.2,0.8,length = nComp))
sigma2 <- rep(var(x),nComp)
probObsInComp <- rep(NA, nComp)

# Setting up the plot
xGrid <- seq(min(x)-1*apply(x,2,sd),max(x)+1*apply(x,2,sd),length = 100)
xGridMin <- min(xGrid)
xGridMax <- max(xGrid)
mixDensMean <- rep(0,length(xGrid))
effIterCount <- 0
ylim <- c(0,2*max(hist(x, n = 50)$density))</pre>
```

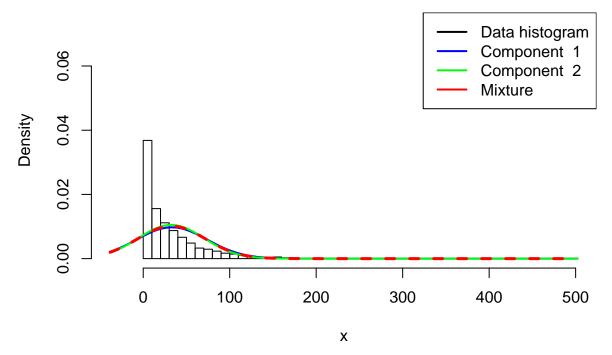
Histogram of x



```
for (k in 1:nIter){
    message(paste('Iteration number:',k))
    alloc <- S2alloc(S) # Just a function that converts between different representations of the group al
    nAlloc <- colSums(S)
    print(nAlloc)
    # Update components probabilities
    pi <- rDirichlet(alpha + nAlloc)

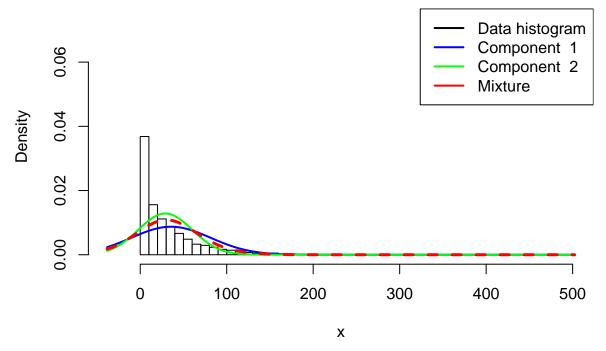
# Update mu's
    for (j in 1:nComp){
        precPrior <- 1/tau2Prior[j]
        precData <- nAlloc[j]/sigma2[j]</pre>
```

```
precPost <- precPrior + precData</pre>
    wPrior <- precPrior/precPost</pre>
    muPost <- wPrior*muPrior + (1-wPrior)*mean(x[alloc == j])</pre>
    tau2Post <- 1/precPost
    mu[j] <- rnorm(1, mean = muPost, sd = sqrt(tau2Post))</pre>
  # Update sigma2's
  for (j in 1:nComp){
    sigma2[j] \leftarrow rScaledInvChi2(1, df = nu0[j] + nAlloc[j], scale = (nu0[j]*sigma2_0[j] + sum((x[alloc = nu0[j] + nAlloc[j], scale = (nu0[j] + sigma2_0[j] + sigma2_0[j]))
  # Update allocation
  for (i in 1:n0bs){
    for (j in 1:nComp){
      probObsInComp[j] <- pi[j]*dnorm(x[i], mean = mu[j], sd = sqrt(sigma2[j]))</pre>
    S[i,] <- t(rmultinom(1, size = 1 , prob = probObsInComp/sum(probObsInComp)))
  # Printing the fitted density against data histogram
  if (plotFit && (k\\1 ==0)){
    effIterCount <- effIterCount + 1</pre>
    hist(x, breaks = 50, freq = FALSE, xlim = c(xGridMin,xGridMax), main = paste("Iteration number",k),
    mixDens <- rep(0,length(xGrid))</pre>
    components <- c()
    for (j in 1:nComp){
      compDens <- dnorm(xGrid,mu[j],sd = sqrt(sigma2[j]))</pre>
      mixDens <- mixDens + pi[j]*compDens</pre>
      lines(xGrid, compDens, type = "1", lwd = 2, col = lineColors[j])
      components[j] <- paste("Component ",j)</pre>
    }
    mixDensMean <- ((effIterCount-1)*mixDensMean + mixDens)/effIterCount
    lines(xGrid, mixDens, type = "1", lty = 2, lwd = 3, col = 'red')
    legend("topright", box.lty = 1, legend = c("Data histogram",components, 'Mixture'),
            col = c("black",lineColors[1:nComp], 'red'), lwd = 2)
    Sys.sleep(sleepTime)
  }
}
## Iteration number: 1
## [1] 3427 3492
```

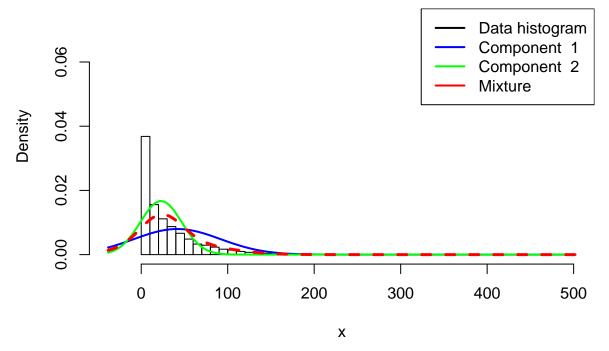


- ## [1] 3363 3556
- ## Iteration number: 3

Iteration number 2



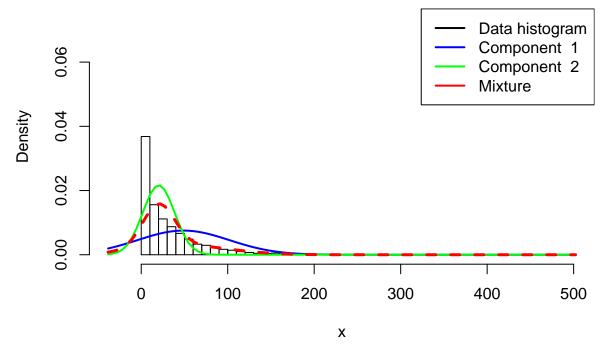
- ## [1] 3108 3811
- ## Iteration number: 4



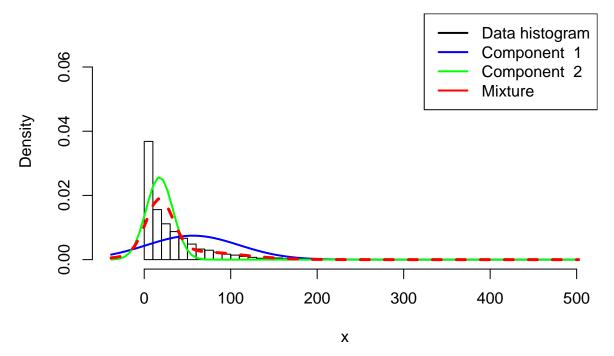
[1] 2761 4158

Iteration number: 5

Iteration number 4



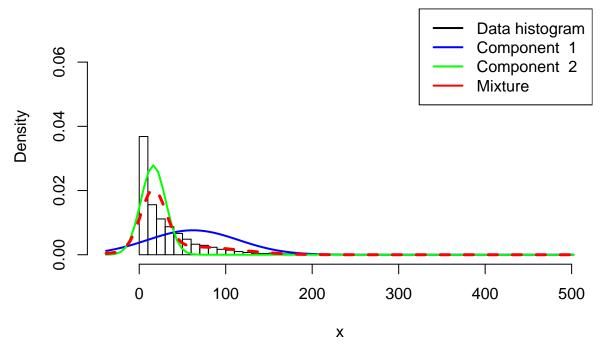
[1] 2360 4559



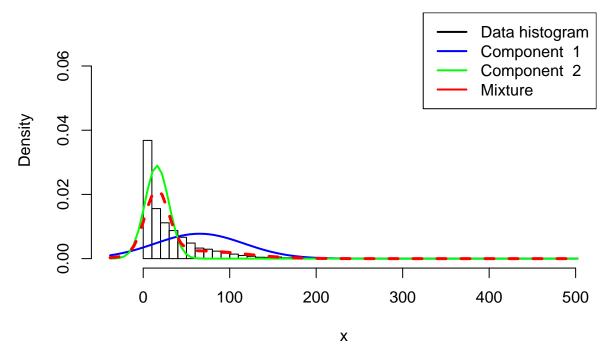
[1] 2206 4713

Iteration number: 7

Iteration number 6



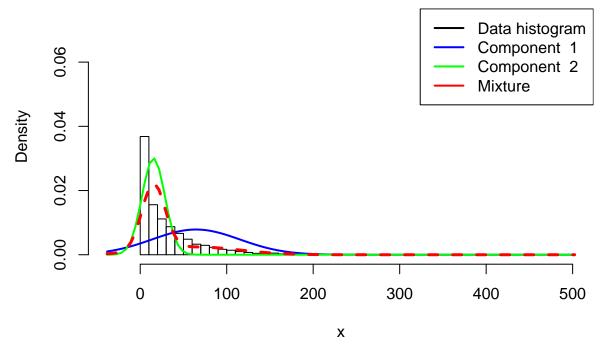
[1] 2129 4790



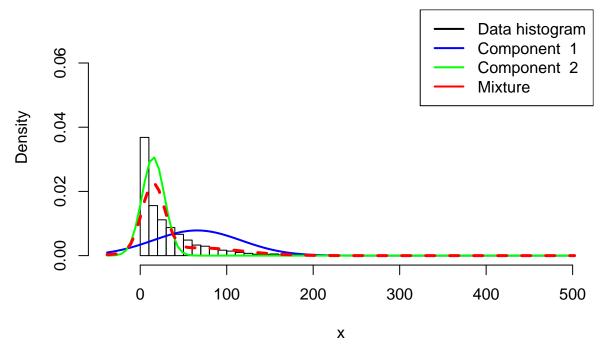
[1] 2144 4775

Iteration number: 9

Iteration number 8

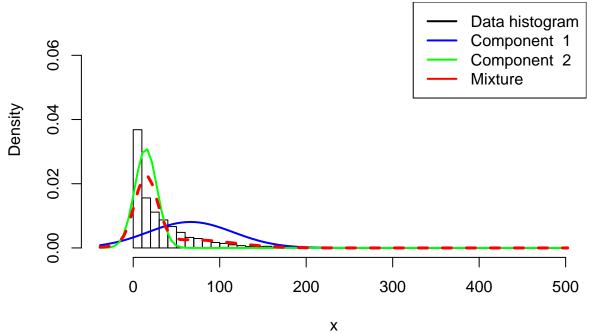


[1] 2195 4724



[1] 2158 4761

Iteration number 10



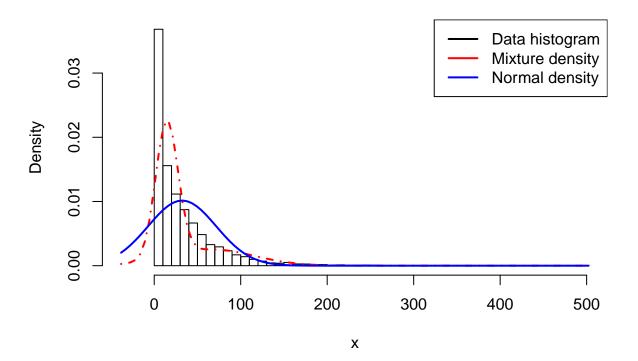
plots above shows the samples covnerging on the distribution of the data. The multinomial model fits better to the data then the model in part (a). It is always at least as good or better then the other model but instead risks to overfit on the given data.

The

(c)

```
hist(x, breaks = 50, freq = FALSE, xlim = c(xGridMin,xGridMax), main = "Final fitted density")
lines(xGrid, mixDens, type = "l", lwd = 2, lty = 4, col = "red")
lines(xGrid, dnorm(xGrid, mean = mean(x), sd = apply(x,2,sd)), type = "l", lwd = 2, col = "blue")
legend("topright", box.lty = 1, legend = c("Data histogram", "Mixture density", "Normal density"), col=c(
```

Final fitted density



Task 2

(a)

```
rawData <- read.table("ebayNumberOfBidderData.dat",header=TRUE)</pre>
y <- as.vector(rawData[,1]); # Data from the read.table function is a data frame. Let's convert y and X
X <- rawData[,2:10];</pre>
model <- glm(y ~ PowerSeller+VerifyID+Sealed+Minblem+MajBlem+LargNeg+LogBook+MinBidShare, data=X, famil
print(model)
##
## Call: glm(formula = y ~ PowerSeller + VerifyID + Sealed + Minblem +
      MajBlem + LargNeg + LogBook + MinBidShare, family = poisson(),
##
       data = X)
##
## Coefficients:
## (Intercept) PowerSeller
                                VerifyID
                                               Sealed
                                                           Minblem
                                              0.44384
                   -0.02054
                                -0.39452
                                                          -0.05220
##
       1.07244
##
                   LargNeg
                                LogBook MinBidShare
      MajBlem
                    0.07067
##
     -0.22087
                                -0.12068
                                             -1.89410
##
## Degrees of Freedom: 999 Total (i.e. Null); 991 Residual
## Null Deviance:
                        2151
## Residual Deviance: 867.5
                                AIC: 3610
print(summary(model))
##
## Call:
## glm(formula = y ~ PowerSeller + VerifyID + Sealed + Minblem +
       MajBlem + LargNeg + LogBook + MinBidShare, family = poisson(),
##
       data = X)
## Deviance Residuals:
      Min
                 1Q
                     Median
                                   3Q
                                           Max
## -3.5800 -0.7222 -0.0441
                                        2.4605
                               0.5269
## Coefficients:
              Estimate Std. Error z value Pr(>|z|)
##
## (Intercept) 1.07244
                           0.03077 34.848 < 2e-16 ***
## PowerSeller -0.02054
                           0.03678 -0.558 0.5765
## VerifyID
              -0.39452
                           0.09243 -4.268 1.97e-05 ***
## Sealed
               0.44384
                           0.05056
                                    8.778 < 2e-16 ***
## Minblem
              -0.05220
                           0.06020 -0.867
                                             0.3859
## MajBlem
              -0.22087
                           0.09144 -2.416
                                            0.0157 *
## LargNeg
               0.07067
                           0.05633
                                    1.255
                                             0.2096
                           0.02896 -4.166 3.09e-05 ***
## LogBook
              -0.12068
## MinBidShare -1.89410
                          0.07124 -26.588 < 2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for poisson family taken to be 1)
```

```
##
## Null deviance: 2151.28 on 999 degrees of freedom
## Residual deviance: 867.47 on 991 degrees of freedom
## AIC: 3610.3
##
## Number of Fisher Scoring iterations: 5
```

The covariates that are significant is the variables: VerifyID, Sealed, MajBlem, LogBook and MinBidShare. I.e. those that have a Pr(abs(z)) > 2, also denothed with "*" in the output above.

(b)

```
xMatrix = as.matrix(X)
sigmaGPrior = 100*solve((t(xMatrix)%*%xMatrix))
tau = 10
chooseCov <- c(1:9)
covNames <- names(rawData)[2:length(names(rawData))];</pre>
xMatrix <- xMatrix[,chooseCov]; # Here we pick out the chosen covariates.
covNames <- covNames[chooseCov];</pre>
nPara <- dim(X)[2];</pre>
# Setting up the prior
mu <- as.vector(rep(0,nPara)) # Prior mean vector</pre>
Sigma <- tau^2*diag(nPara);</pre>
PoiPost <- function(theta,y,X, mu, SigmaGPrior) {</pre>
  nPara <- length(theta);</pre>
  linPred <- X%*%theta;</pre>
  logPoiLik <- sum( linPred*y -exp(linPred) - log(factorial(y)));</pre>
  if (abs(logPoiLik) == Inf) logPoiLik = -20000; # Likelihood is not finite, stear the optimizer away f
  logBetaPrior <- dmvnorm(theta, matrix(0,nPara,1), SigmaGPrior, log=TRUE);</pre>
  return(logPoiLik + logBetaPrior)
initVal <- as.vector(rep(0,dim(X)[2]));</pre>
logPost = PoiPost;
OptimResults<-optim(initVal,logPost,gr=NULL,y,xMatrix,mu,Sigma,method=c("BFGS"),control=list(fnscale=-1
approxPostStd <- sqrt(diag(-solve(OptimResults$hessian)))</pre>
names(approxPostStd) <- covNames # Naming the coefficient by covariates</pre>
betatilde = OptimResults$par
print("Betatilde: ")
## [1] "Betatilde: "
print(betatilde)
## [1] 1.07245146 -0.02054160 -0.39448259 0.44382930 -0.05219690 -0.22084701
## [7] 0.07066972 -0.12065535 -1.89401063
print("Jacobiany beta: ")
## [1] "Jacobiany beta: "
```

```
print(approxPostStd)
         Const PowerSeller
                              VerifvID
                                             Sealed
                                                        Minblem
                                                                    MajBlem
##
   0.03077417 0.03678186 0.09242044 0.05056213 0.06019999 0.09143186
##
       LargNeg
                   LogBook MinBidShare
## 0.05633101 0.02896411 0.07123677
(c)
RWMSampler <- function(logPostFunc, n, c, covar, ...) {</pre>
  currentTheta = as.vector(rep(0, dim(covar)[1]))
  draws = matrix(0, nrow = n, ncol = dim(covar)[1])
  oldProbability = logPostFunc(currentTheta, ...)
  for(i in 1:n) {
    currentDraw <- rmvt(1, mu = currentTheta, S = c*covar)</pre>
    newProbability <- logPostFunc(as.vector(currentDraw), ...)</pre>
    alpha = min(1, newProbability/oldProbability)
    uniformDraw = runif(1, 0, 1)
    if(uniformDraw >= alpha) {
      oldProbability = newProbability
      currentTheta = currentDraw
    }
    draws[i,] = currentDraw
```

myDraws <- RWMSampler(PoiPost, 10000, 5, diag(diag(-solve(OptimResults\$hessian))),

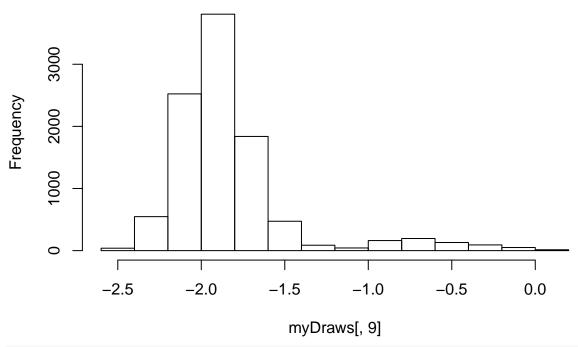
y, xMatrix, mu, Sigma)

return(draws)

hist(myDraws[,9])

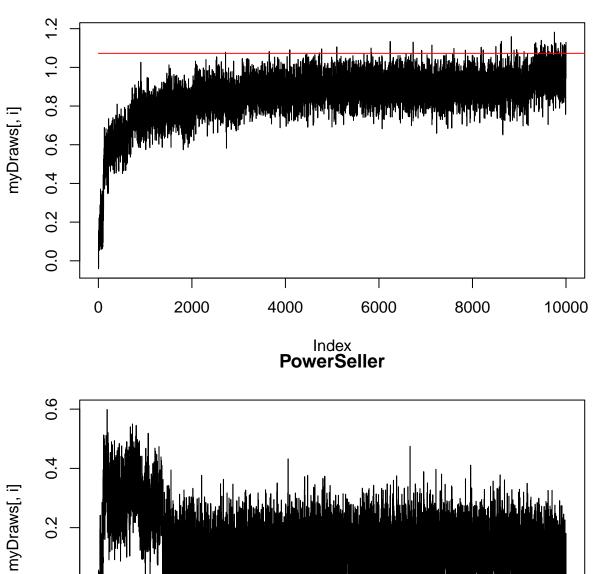
}

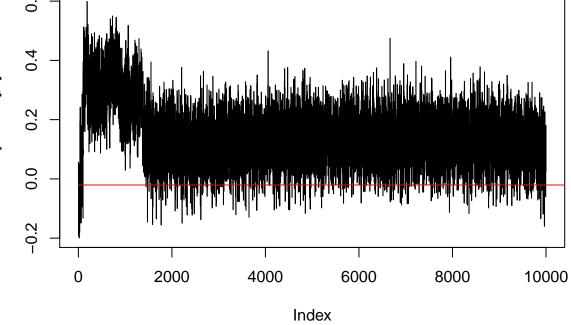
Histogram of myDraws[, 9]



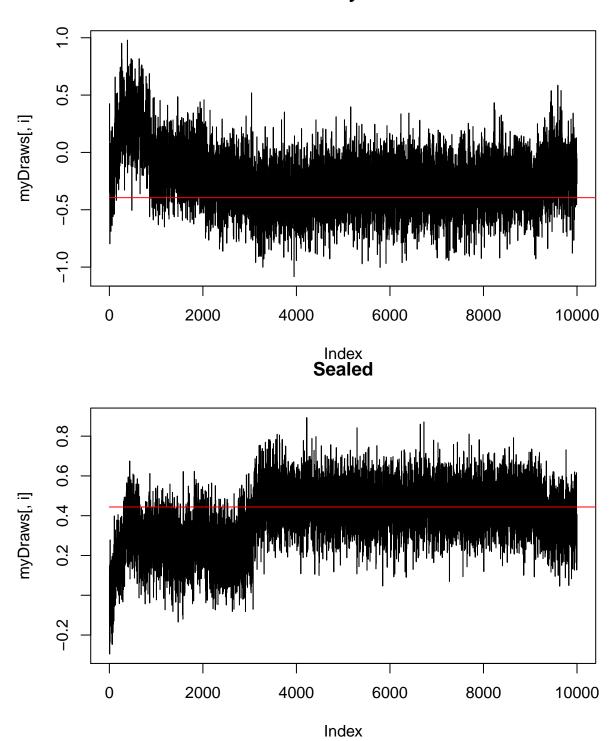
```
for(i in 1:9){
  plot(myDraws[,i], type='s')
  a = c(rep(betatilde[i],length(myDraws)))
  lines(a, col='red')
  title(covNames[i])
}
```

Const

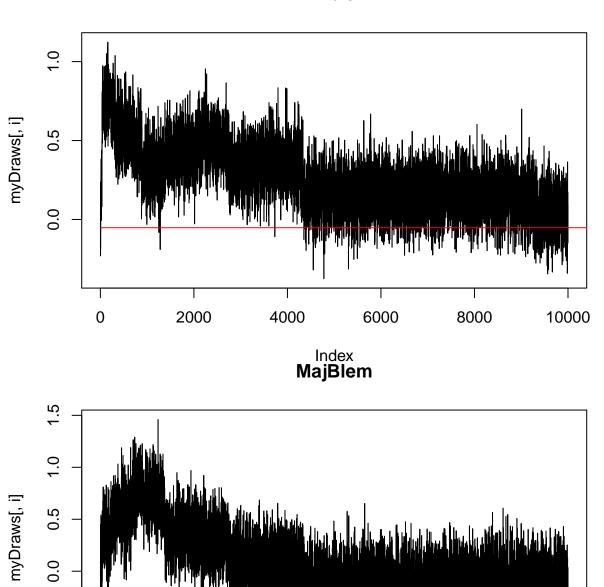




VerifyID



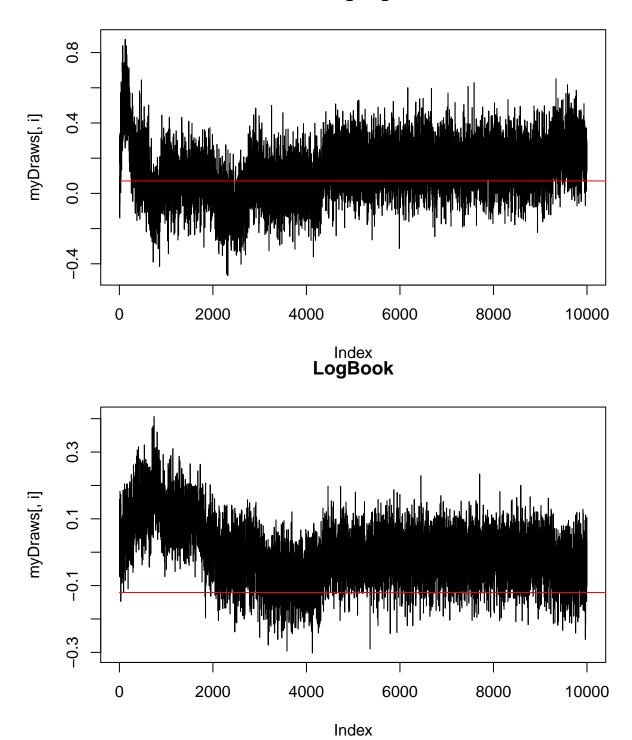
Minblem



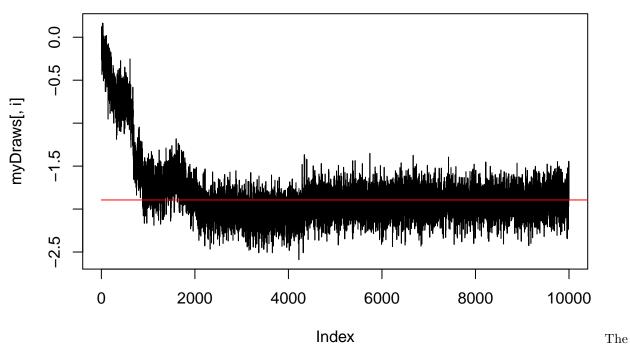
Index

-0.5

LargNeg



MinBidShare



draws seem to move in the right direction but not all variables are close to the approximated corresponding betatilde(red line) as in the earlier models in (a) nad (b), might be due to local optima or possible dependency on other vairables.