# Bayesian estimation of binary regression models

Applied Bayesian Statistics
Winter Term 2018

Bayesian estimation of binary regression models

Susumu Shikano

Introduction

Bayesian Regression

Conjugacy analysis

Wrap up

Susumu Shikano GSDS

- In Bayesian inference the (posterior) probability of individual models is at stake.
- You can obtain the posterior even after observing one single case.
- The posterior also depends on your prior belief.
- For a larger number of observations frequentists and Bayesian have similar results.
- Conjugacy of the beta distribution with the likelihood based on the binomial distribution.

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# **Example**

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### **Example**

Model: capacity of face-based inference of candidate ideology

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### **Example**

- Model: capacity of face-based inference of candidate ideology
- Parameter of interest:  $\pi$

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#### **Example**

- Model: capacity of face-based inference of candidate ideology
- Parameter of interest: π
- → Only one parameter is at stake.

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#### **Example**

- Model: capacity of face-based inference of candidate ideology
- Parameter of interest: π
- ullet Only one parameter is at stake.
- In praxis there are multiple parameters.

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#### **Example**

- Model: capacity of face-based inference of candidate ideology
- Parameter of interest: π
- ullet Only one parameter is at stake.
- In praxis there are multiple parameters.
- e.g. (bivariate) regression models

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# **Linear Regression Model**

#### Two approaches to obtain posterior

- Conjugacy analysis
  - Conjugacy: The property that the prior and posterior have the same probability form depending on the form of the distribution used to calculate the likelihood.
  - · Posterior can be obtained analytically.

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## **Linear Regression Model**

#### Two approaches to obtain posterior

- Conjugacy analysis
  - Conjugacy: The property that the prior and posterior have the same probability form depending on the form of the distribution used to calculate the likelihood.
  - Posterior can be obtained analytically.
- Deriving posterior per Gibbs Sampling
  - Conjugacy is not must.
  - Use of MCMC.

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## (Bivariate) linear regression model

$$y = \beta_0 + \beta_1 x + \epsilon$$
$$\epsilon \sim N(0, \sigma^2)$$

$$(y|\beta_0, \beta_1, \sigma^2, x) \sim N(\beta_0 + \beta_1 x, \sigma^2)$$

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The unknown parameters to be estimated

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The unknown parameters to be estimated

- β<sub>0</sub>
- β<sub>1</sub>
- σ<sup>2</sup>

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#### **How to proceed**

- Likelihood?
- Prior?
- Posterior?

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#### **How to proceed**

- Likelihood? → Normal distribution
- Prior?
- Posterior?

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#### **How to proceed**

- Likelihood? → Normal distribution
- Prior? → Normal-Inverse-Gamma
- Posterior?

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#### How to proceed

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Conjugacy!!

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#### How to proceed

- Likelihood? → Normal distribution
- Prior? → Normal-Inverse-Gamma
- Posterior? → Normal-Inverse-Gamma

Conjugacy!!

#### **Likelihood Function**

$$f_N(\boldsymbol{y}|\boldsymbol{\beta}, \sigma^2) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(y_i - \boldsymbol{X}\boldsymbol{\beta})'(y_i - \boldsymbol{X}\boldsymbol{\beta})}{2\sigma^2}\right\}$$

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#### Normal-Inverse-Gamma?

It looks like...

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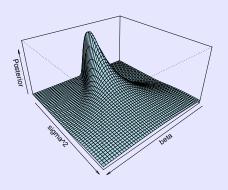
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#### **Normal-Inverse-Gamma?**

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## Inverse-Gamma IG(a, d) with a as rate and d as shape

If  $X \sim G(a, d)$ , then  $1/X \sim IG(a, d)$ :

$$f_{\Gamma}(\theta) = \frac{a^d}{\Gamma(d)} \theta^{d-1} \exp\left(-a\theta\right); \ f_{\Gamma^{-1}}(\theta) = \frac{a^d}{\Gamma(d)} \theta^{-d-1} \exp\left(-\frac{a}{\theta}\right)$$

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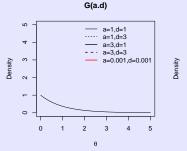
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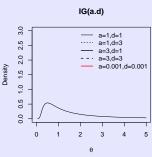
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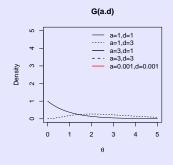
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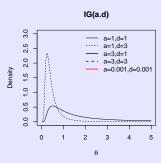
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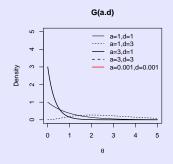
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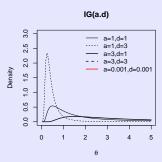
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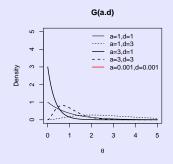
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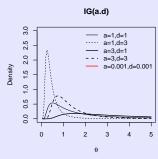
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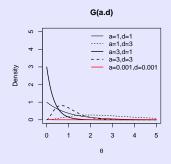
Bayesian Regression

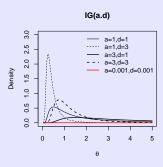
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#### **Prior: Normal-Inverse-Gamma**

$$\begin{split} f_{N-\Gamma-1}(\beta,\sigma^2|\mu,\lambda,a,d) &=& f_N(\beta|\mu,\sigma^2/\lambda)f_{\Gamma-1}(\sigma^2|a,d) \\ &=& \frac{\sqrt{\lambda}}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{\lambda(\beta-\mu)^2}{2\sigma^2}\right\}\underbrace{\frac{a^d}{\Gamma(d)}\sigma^{2(-d-1)}\exp\left(-\frac{a}{\sigma^2}\right)}_{\text{Inverse-Gamma-Dist.}} \end{split}$$

 $= \quad \frac{\sqrt{\lambda}}{\sqrt{2\pi\sigma^2}} \frac{a^d}{\Gamma(d)} \left(\frac{1}{\sigma^2}\right)^{d+1} \exp\left\{-\frac{\lambda(\beta-\mu)^2+2a}{2\sigma^2}\right\}$ 

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#### **Prior: Normal-Inverse-Gamma**

 $f_{N-r-1}(\beta, \sigma^2 | \mu, \lambda, a, d) = f_N(\beta | \mu, \sigma^2 / \lambda) f_{r-1}(\sigma^2 | a, d)$ 

$$= \frac{\sqrt{\lambda}}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{\lambda(\beta-\mu)^2}{2\sigma^2}\right\} \underbrace{\frac{a^d}{\Gamma(d)}\sigma^{2(-d-1)} \exp\left(-\frac{a}{\sigma^2}\right)}_{\text{Inverse-Gamma-Dist.}}$$

$$= \frac{\sqrt{\lambda}}{\sqrt{2\sigma^2}} \underbrace{\frac{a^d}{\Gamma(d)} \left(\frac{1}{\sigma^2}\right)^{d+1} \exp\left\{-\frac{\lambda(\beta-\mu)^2 + 2a}{2\sigma^2}\right\}}_{2\sigma^2}$$

The prior is more precisely a multivariate normal-inverse gamma distribution:

$$\begin{split} f_{N-\Gamma-1}(\beta,\sigma^2|\mu,\Sigma,a,d) &= f_N(\beta|\mu,\sigma^2\cdot\Sigma)f_{\Gamma-1}(\sigma^2|a,d) \\ &= \frac{1}{\sqrt{2\pi\sigma^k|\Sigma|}}\frac{a^d}{\Gamma(d)}\left(\frac{1}{\sigma^2}\right)^{d+1} \\ &\times \exp\left\{-\frac{(\beta-\mu)'\Sigma^{-1}(\beta-\mu)+2a}{2\sigma^2}\right\} \end{split}$$

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### **Deriving posterior**

$$= \frac{1}{\sqrt{2\pi\sigma^k|\boldsymbol{\Sigma_0}|}} \frac{a_0^{d_0}}{\Gamma(d_0)} \left(\frac{1}{\sigma_0^2}\right)^{d_0+1}$$

$$\times \exp\left\{-\frac{(\beta_0-\mu_0)'\boldsymbol{\Sigma_0}^{-1}(\beta_0-\mu_0)+2a_0}{2\sigma^2}\right\}$$
Likelihood  $f(\boldsymbol{y}|\boldsymbol{\beta},\sigma^2) = \prod_i^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(y_i-\boldsymbol{X}\boldsymbol{\beta})'(y_i-\boldsymbol{X}\boldsymbol{\beta})}{2\sigma^2}\right\}$ 
Posterior  $f(\boldsymbol{\beta},\sigma^2|\boldsymbol{y}) = \frac{f(\boldsymbol{y}|\boldsymbol{\beta},\sigma^2)f(\boldsymbol{\beta},\sigma^2)}{f(\boldsymbol{y})}$ 

$$f(\boldsymbol{\beta},\sigma^2|\boldsymbol{y}) = f_{N-\Gamma^{-1}}(\boldsymbol{\mu}^*,\boldsymbol{\Sigma}^*,a^*,b^*)$$

Prior  $f(\beta, \sigma^2) = f_{N-\Gamma^{-1}}(\mu_0, \Sigma_0, a_0, d_0)$ 

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#### **Posterior**

Posterior 
$$f(\beta, \sigma^2 | \mathbf{y}) = \frac{f(\mathbf{y} | \beta, \sigma^2) f(\beta, \sigma^2)}{f(\mathbf{y})}$$
  
 $= f_{N-\Gamma^{-1}}(\mu^*, \Sigma^*, a^*, d^*)$   
with  
 $\mu^* = (\Sigma_0^{-1} + \mathbf{X}'\mathbf{X})^{-1}(\Sigma_0^{-1}\mu_0 + \mathbf{X}'\mathbf{y})$   
 $\Sigma^* = (\Sigma_0^{-1} + \mathbf{X}'\mathbf{X})^{-1}$   
 $a^* = a_0 + \frac{1}{2}(\mu_0'\Sigma_0^{-1}\mu_0 + \mathbf{y}'\mathbf{y} - \mu^{*'}\Sigma^{*-1}\mu^*)$   
 $d^* = d_0 + \frac{n}{2}$ 

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#### **Posterior**

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Wrap up

#### Relationship to ML/OLS

If n and/or  $\Sigma_0$  approach to infinity:

#### **Posterior**

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Wrap up

### **Relationship to ML/OLS**

If n and/or  $\Sigma_0$  approach to infinity:

$$\lim_{n\to\infty}\boldsymbol{\mu}^*=(\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{y}$$

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#### Posterior is not posterior...

- Joint posterior
- Conditional posterior
- Marginal posterior

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## Posterior is not posterior...

- Joint posterior
- Conditional posterior
- Marginal posterior

# Marginal posterior of $\beta$

By integrating out  $\sigma^2$ , we can obtain a multivariate student t-distribution with a degree of freedom of  $\nu = n - k$ :

$$f(\beta|\mathbf{y}) = \int_0^\infty f(\beta, \sigma^2|\mathbf{y}) d\sigma^2$$
$$= f_t(\nu, \mu^*, \Sigma^*)$$

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## Wrap up

# **Bayesian regression models**

- In conjugacy analysis posterior can be derived analytically.
- The larger n/the smaller the dispersion of prior, the more similar results with the maximum likelihood.

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