

Exercise 07: Survival Analysis

2018 Spring

1 Concepts of Survival Analysis

A survival analysis is a method for analyzing time to events, where the events can be “death” or “failure”, etc.

The task of survival analysis is to

- (1) Estimate and interpret survivor and/or hazard functions;
- (2) Compare survivor and/or hazard functions
- (3) Assess the relationship between explanatory variables and survival time

1.1 Concepts

Censoring The survival time is not exactly known due to

- A subject does not experience the event until the study ends.
- A subject is lost-to-follow-up during the study period.
- A subject withdraws from the study due to some other reason.

Right-censored Unknown but $T > t$

Left-censored Unknown but $T < t$

Interval-censored Unknown but $t_1 < T < t_2$

Survival time T : the outcome variable (time to event)

Risk set The set of subjects with $T \geq t$

1.1.1 Survivor function

- $S(t) = P(T > t)$
- Probability that the survival time T exceeds a specified time t
- We often use the empirical survivor function $\hat{S}(t)$.

1.1.2 Hazard function: Instantaneous hazard

- **Force of mortality**
- $h(t) = \lim_{\Delta t \rightarrow 0} \frac{P(t \leq T \leq t + \Delta t)}{\Delta t}$
- $h(t) \geq 0$ and has no upper bounds
- Hazard function is also called **failure rate**.
- *Rate of events occurring per time unit*, e.g., 50 events per month = 600 events per year.

1.1.3 Relationship between $S(t)$ and $h(t)$

- $S(t) = \exp \left[- \int_0^t h(u) du \right]$
- $h(t) = - \frac{dS(t)/dt}{S(t)}$
- If $S(t) = e^{-\lambda t}$, $h(t) = \lambda$

1.1.4 Basic Descriptive Analysis

- Mean survival time \bar{T} (平均生存时间, ignoring the censorship)
- Median survival time (中位生存时间, $t|\hat{S}(t) = 0.5$)
- Average hazard rate (平均风险率) $\bar{h} = \#failures / \sum_{i=1}^n t_i$

Example 1 (Descriptive Analysis of Survival Time)

| <i>individual</i> | <i>t(weeks)</i> | <i>δ (failed=1;censored=0)</i> |
|-------------------|-----------------|--------------------------------|
| 1 | 3.5 | 0 |
| 2 | 3.5 | 1 |
| 3 | 5 | 1 |
| 4 | 6 | 0 |
| 5 | 8 | 0 |
| 6 | 12 | 0 |

- The mean survival time is $\bar{T} = \frac{3.5+5}{2} = 4.25$ weeks.
- The average hazard rate is $\bar{h} = 2/(3.5 + 3.5 + 5 + 6 + 8 + 12) = 0.0526$ failures per week.

2 Survival Analysis - Inference

The survival analysis can be classified into three main categories:

- Parametric methods: The survival times follow some parametric distribution
 - Lognormal distribution
 - Weibull distribution
 - Exponential distribution
 - Gamma distribution
- Nonparametric methods
 - Survival rate through Kaplan-Meier or life tables
 - Comparing n groups of survival rates through logrank test ($n=2$) or Breslow test ($n=3+$)
- Semi-parametric methods
 - Cox-proportional hazards model

2.1 Life tables

2.1.1 Assumptions

- There are no changes in survivorship over calendar time
- The experience of individuals who are lost to follow-up is the same as the experience of those who are followed.
- Withdrawal occurs uniformly within the interval.
- Event occurs uniformly within the interval.

| Ordered failure times ($t_{(i)}$) | failures (m_i) | censored (q_i) | Risk set $R(t_{(i)})$ |
|-------------------------------------|--------------------|--------------------|-----------------------|
| $t_{(0)} = 0$ | m_0 | q_0 | $R(t_{(0)})$ |
| $t_{(1)}$ | m_1 | q_1 | $R(t_{(1)})$ |
| \vdots | \vdots | \vdots | \vdots |
| $t_{(n)}$ | m_n | q_n | $R(t_{(n)})$ |

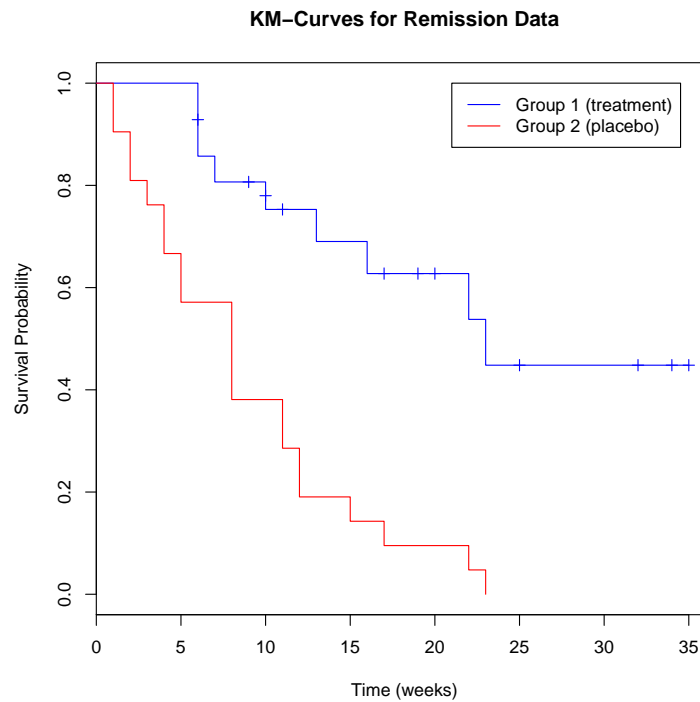
2.2 Kaplan-Meier Estimation

The Kaplan-Meier survival table is in the following form:

| Table 1: Kaplan-Meier Table Example | | | | | |
|-------------------------------------|----------|----------|-------------|----------|------|
| t_j | n_j | d_j | $P(t_j)$ | $S(t_j)$ | SE |
| 5 | 28 | 1 | 27/28=0.964 | 0.96 | 0.04 |
| 29 | 22 | 1 | 21/22=0.955 | 0.92 | 0.05 |
| 37 | 20 | 1 | 19/20=0.950 | 0.87 | 0.07 |
| \vdots | \vdots | \vdots | \vdots | \vdots | |

- The estimator $\hat{q}_i = d_j/n_j$ is the estimate of $h(t_j)$;
- The survival probability $P(t_j) = 1 - \hat{q}_j$
- The survival rate $S(t_j) = \prod_{t_i \leq t_j} P(t_i)$
- The standard error is $SE(S(t_j)) = S(t_j) \left[\sum_{i=1}^j \frac{d_j}{n_j(n_j - d_j)} \right]$

```
time1 <- c(6,6,6,7,10,13,16,22,23,6,9,10,11,17,19,20,25,32,32,34,35)
status1 <- c(1,1,1,1,1,1,1,1,1,0,0,0,0,0,0,0,0,0,0,0)
time2 <- c(1,1,2,2,3,4,4,5,5,8,8,8,8,11,11,12,12,15,17,22,23)
status2 <- c(1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1,1)
time <- c(time1, time2)
status <- c(status1, status2)
group <- factor(c(rep(0,21), rep(1,21)))
fit <- survfit(Surv(time,status) ~ group)
plot(fit, conf.int="none", col=c("blue", "red"),
     mark.time = TRUE,
     xlab="Time (weeks)", ylab="Survival Probability")
legend(21,1,c('Group 1 (treatment)', 'Group 2 (placebo)'),
      col = c('blue','red'), lty = 1)
title(main='KM-Curves for Remission Data')
```



2.3 Hypothesis testing for the survival data

There are three statistical test methods for comparing 2+ survival curves:

- **Log-rank test** using equal weights on different observed time points.
- **Breslow test** using the risk set size $|R_i|$ as the weight for each time point.
- **Tarone-Ware test** using the square root of the risk set size $\sqrt{|R_i|}$ as the weight for each time point.

2.3.1 Logrank test for comparing two survival curves/functions

- Sort the K unique times: $t_1 < t_2 < \dots < t_K$
- n_{ij} : number of persons in group i at risk at t_j
- $n_j = \sum_i n_{ij}$ the total number of subjects at risk at t_j
- o_{ij} : number of failures in group i at t_j
- $o_j = \sum_i o_{ij}$: total number of failures at t_j

Under the null hypothesis $H_0 : S_1(t) = S_2(t), 0 < t < \infty$, o_{1j} has the hypergeometric distribution conditional on the margins $\{n_{1j}, n_{2j}, o_j, n_j - o_j\}$:

$$\mathbb{P}(o_{1j}) = \frac{\binom{o_j}{o_{1j}} \binom{n_j - o_j}{n_{1j} - o_{1j}}}{\binom{n_j}{n_{1j}}}$$

Then we can get the conditional expectation and variance:

$$\begin{aligned}
 e_{1j} &= E(o_{1j} | \text{marginals}) \\
 &= \left(\frac{n_{1j}}{n_j} \right) o_j \\
 V_j &= \text{Var}(o_{1j} | \text{marginals}) \\
 &= \frac{n_j - n_{1j}}{n_j - 1} \times n_{1j} \left(\frac{o_j}{n_j} \right) \left(1 - \frac{o_j}{n_j} \right) \\
 &= \frac{n_{1j} n_{2j} o_j (n_j - o_j)}{n_j^2 (n_j - 1)}
 \end{aligned}$$

Since

$$z = \frac{\sum_{j=1}^K (o_{1j} - e_{1j})}{\sqrt{\sum_{j=1}^K V_j}} \sim N(0, 1) \text{ under } H_0$$

we can obtain the log-rank test statistic:

$$X^2 = \frac{\left(\sum_{j=1}^K (o_{1j} - e_{1j}) \right)^2}{\sum_{j=1}^K V_j} \sim \chi^2(df = 1)$$

with χ^2 we can get the p -value to decide whether to reject the null hypothesis.

This can be executed in R:

```
fit <- survdiff(Surv(time, status) ~ group, data, rho=0)
summary(fit)
```

3 Exercises

1. (10 points) True or False

- (1) ___ The survival function $S(t)$ ranges between 0 and ∞ .
- (2) ___ A hazard rate of one per day is equivalent to seven per week.
- (3) ___ If you know the form of hazard function, then you can determine the corresponding survivor curve, and vice versa.
- (4) ___ If the survival curve for group 1 lies completely above the curve for group 2, the median survival time for group 2 is longer than that for group 1.
- (5) ___ The risk set at six weeks is the set of individuals whose survival time are less than or equal to six weeks.
- (6) ___ If the risk set at 6th week consists of 22 persons, and 4 persons failed and 3 are censored by the 7th week, then the risk set at 7th week consists of 18 persons.
- (7) ___ If a hazard ratio comparing group 1 relative to group 2 equals 10, then the potential for failure is 10 times higher in group 1 than in group 2.
- (8) ___ Survivor function is a proportion metric, while hazard function is a rate metric.
- (9) ___ Compared to standard log-rank test, Peto-Prentice test place more emphasis on the late-occurred failures.
- (10) ___ Compared to life table, the Kaplan-Meier table is more commonly used in actuary.

2. (5 points) The **mean residual life time (mrl)** can be defined as

$$\text{mrl}(t_0) = E[T - t_0 | T \geq t_0],$$

i.e. the *average remaining survival time given the population has survived beyond t_0* . Prove that

$$\text{mrl}(t_0) = \frac{\int_{t_0}^{\infty} S(t) dt}{S(t_0)}.$$

3. (10 points) The time (in days) to developing a tumor for rats exposed to a carcinogen follows a Weibull distribution with shape parameter $\lambda_0 = 0.5$ and scale parameter $\lambda_1 = 2$.
- (1) (2 points) Compute the probability that a random rat will be tumor-free at the 30-th day.
 - (2) (2 points) What is the average time to tumor development?
 - (3) (3 points) Find the hazard rate of time to tumor development at the 30-th day.
 - (4) (3 points) Find the median time to tumor development.
4. (5 points) Suppose we have a small data set with different kinds of censoring: 2+, 3, 4, 5-, 6, 7+, [5, 7], Suppose the distribution of the underlying survival time is an exponential distribution with a constant hazard λ . Write down the likelihood function of λ for this given data set.
5. (20 points) A survival analysis was conducted to compare the survival times (in years) for two groups each with 25 participants. **CHR** is used to indicate whether the group has history of chronic disease ($CHR = 1/0$).

| | |
|-----------------|--|
| Group 1 (CHR=0) | 12.3+, 5.4, 8.2, 12.2, 11.7, 10.0, 5.7, 9.8, 2.6, 11.0, 9.2, 12.1+, 6.6, 2.2, 1.8, 10.2, 10.7, 11.1, 5.3, 3.5, 9.2, 2.5, 8.7, 3.8, 3.0 |
| Group 2 (CHR=1) | 5.8, 2.9, 8.4, 8.3, 9.1, 4.2, 4.1, 1.8, 3.1, 11.4, 2.4, 1.4, 5.9, 1.6, 2.8, 4.9, 3.5, 6.5, 9.9, 3.6, 5.2, 8.8, 7.8, 4.7, 3.9 |

- (1) (10 points) Make a life table and Kaplan-Meier table for each group, respectively.
 - (2) (5 points) Compute the average survival times (\bar{T}) and average hazard rates \bar{h} for two groups. Which group has a better prognosis? Explain briefly.
 - (3) (5 points) How would a comparison of survivor curves provide additional information to what is provided in the table?
6. (20 points) Conduct the log-rank test procedure to compare these two survival data. You need to write down the details.

| | |
|---------|--|
| Group 1 | 6, 6, 6, 7, 10, 13, 16, 22, 23, 6+, 9+, 10+, 11+, 17+, 19+, 20+, 25+, 32+, 32+, 34+, 35+ |
| Group 2 | 1, 1, 2, 2, 3, 4, 4, 5, 5, 8, 8, 8, 8, 11, 11, 12, 12, 15, 17, 22, 23 |

7. (20 points) The dataset **veterans.dat** considers the survival times (days) for 137 patients from the Veterans Administration Lung Cancer Trial.

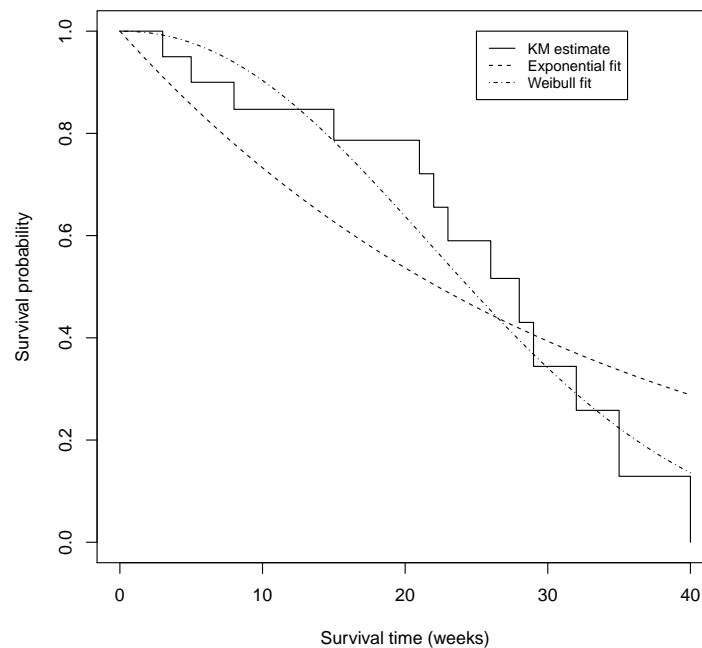
```
data <- read.table("http://cbb.sjtu.edu.cn/course/bi476/data/
  veterans.dat", header=F)
```

- (1) (5 points) Obtain the Kaplan-Meier plots for the two cell type (1=large, 0=other). Comment on how the two curves compare with each other. Moreover, draw a conclusion based on log-rank test.
 - (2) (5 points) Obtain Kaplan-Meier plots for the four cell types (large, adeno, small, and squamous). Note that you will need to recode the data to define a single variable which numerically distinguishes the four categories.
 - (3) (10 points) Compare the curves and use log-rank test and weighted log-rank test to draw the final conclusions.
8. (20 points) The dataset `tempsurv.dat` contains a series of survival times. We can use nonparametric Kaplan-Meier method and also the parametric models (e.g., exponential model, Weibull model, etc.) in R.

```
library(survival)
example <- read.table("data/tempsurv.dat", header=TRUE)
# fit a Kaplan-Meier model
fit1 <- survfit(Surv(survtime, status)~1, data=example, conf.type="plain")
# plot the Kaplan-Meier curve
plot(0,0, type="n", xlim=c(0,40), ylim=c(0,1),
     xlab="Survival time (weeks)", ylab="Survival probability")
lines(fit1, conf.int="none", lty=1)

x <- seq(0, 40, by=0.5)
# Fit an exponential model
fit2 <- survreg(Surv(survtime, status)~1, data=example, dist="exponential")
lambda <- exp(-fit2$coef)
sx <- exp(-lambda * x)
lines(x, sx, lty=2)

# Fit a Weibull model
fit3 <- survreg(Surv(survtime, status)~1, data=example, dist="weibull")
lambda <- exp(-fit3$coef/fit3$scale)
alpha <- 1/fit3$scale
sx <- exp(-lambda * x^alpha)
lines(x, sx, lty=4)
legend(25,1, c("KM estimate", "Exponential fit", "Weibull fit"),
     lty=c(1,2,4), cex=0.8)
```



- (1) (5 points) From the figure above, which model fits the data better? Exponential or Weibull? You can explain from both the theoretical and the observational perspective.
- (2) (5 points) Here are the outputs for the two model fitting, which model is better? Why? Hint: use log likelihood-ratio test to check.

```
## Call:
## survreg(formula = Surv(survtime, status) ~ 1, data = example,
##         dist = "exponential")
##
## Coefficients:
## (Intercept)
##          3.47
##
## Scale fixed at 1
##
## Loglik(model)= -58.1   Loglik(intercept only)= -58.1
## n= 20
```

```
## Call:
## survreg(formula = Surv(survtime, status) ~ 1, data = example,
##         dist = "weibull")
##
## Coefficients:
## (Intercept)
##          3.37
##
## Scale= 0.465
##
## Loglik(model)= -54.1   Loglik(intercept only)= -54.1
## n= 20
```

- (3) (5 points) You can also conduct the **Wald test** to check whether the data are from an exponential distribution.

- (4) (5 points) Use **score test** to test whether or not the survival times are from an exponential distribution.