

# Multiple testing

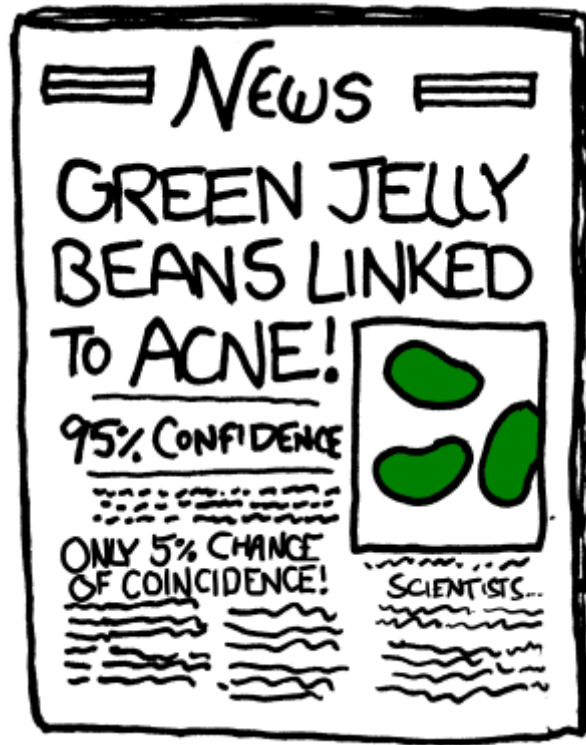
## Session 4

MATH 80667A: Experimental Design and Statistical Methods  
for Quantitative Research in Management  
HEC Montréal

# Multiple testing

# Scientifist, investigate!

- Consider the Cartoon *Significant* by Randall Munroe (<https://xkcd.com/882/>)



It highlights two problems: lack of accounting for multiple testing and selective reporting.

# How many tests

Consider a one-way ANOVA with  $K$  groups.

Having found a significant difference between group means (global null), you proceed to look at all pairwise differences:  $\binom{K}{2}$  tests for  $K$  groups.

- 3 tests if  $K = 3$  groups
- 10 tests if  $K = 5$  groups
- 45 tests if  $K = 10$

Many tests!

# Family-wise error rate

If you do a single hypothesis test and your testing procedure is well calibrated (model assumptions met), there is a probability  $\alpha$  of making a type I error if the null is true.

Why  $\alpha = 5\%$ ? Essentially **arbitrary**...

If one in twenty does not seem high enough odds, we may, if we prefer it, draw the line at one in fifty or one in a hundred. Personally, the writer prefers to set a low standard of significance at the 5 per cent point, and ignore entirely all results which fails to reach this level.

Fisher, R.A. (1926). The arrangement of field experiments. *Journal of the Ministry of Agriculture of Great Britain*, 33:503-513.

# How many tests?

Dr. Yoav Benjamini looked at the number of inference / tests performed in the Psychology replication project

Open Science Collaboration. (2015). Estimating the reproducibility of psychological science. *Science*, 349(6251), aac4716.

The number of tests performed ranged from 4 to 700, with an average of 72.

Most studies did not account for selection.

# Motivation

- If we do  $m$  **independent** comparisons, each one at the level  $\alpha$ , the probability of making at least one type I error, say  $\alpha^*$ , is

$$\begin{aligned}\alpha^* &= 1 - \text{probability of making no type I error} \\ &= 1 - (1 - \alpha)^m\end{aligned}$$

With  $\alpha = 5\%$

- $m = 4$  tests,  $\alpha^* \approx 0.185$ .
- $m = 72$  tests,  $\alpha^* \approx 0.975$ .

Tests need not be independent... but can show  $\alpha^* \leq m\alpha$ .

# Family of hypothesis

Consider a family of  $m$  null hypothesis  $\mathcal{H}_{01}, \dots, \mathcal{H}_{0m}$  tested.

- The family may depend on the context, but all hypothesis that are scientifically relevant and could be reported.

**Should be chosen a priori and pre-registered**

**Keep it small:** the number of planned comparisons for a one-way ANOVA should be less than the number of groups  $K$ .



# Notation

Define

$$R_i = \begin{cases} 1 & \text{if we reject } \mathcal{H}_{0i} \\ 0 & \text{if we fail to reject } \mathcal{H}_{0i} \end{cases}$$
$$V_i = \begin{cases} 1 & \text{type I error for } \mathcal{H}_{0i} \quad (R_i = 1 \text{ and } \mathcal{H}_{0i} \text{ is true}) \\ 0 & \text{otherwise} \end{cases}$$

with

- $R = R_1 + \cdots + R_m$  the total number of rejections ( $0 \leq R \leq m$ ).
- $V = V_1 + \cdots + V_m$  the number of null hypothesis rejected by mistake.

# Decision rule

Classify the decision on the  $m$  tests in a table based on whether the null hypothesis is true or false.

We reject the null hypothesis  $\mathcal{H}_0$  if the  $p$ -value is less than the level,  $p < \alpha$ .

Truth \ Decision	Reject null hypothesis	Fail to reject null
$\mathcal{H}_0$ is true	$R - V$ correct rejections	–
$\mathcal{H}_a$ is true	$V$ type I errors	–
Total	$R$ rejections	$m - R$ non-rejections

# Familywise error rate

The familywise error rate is the probability of making at least one type I error per family

$$\text{FWER} = \Pr(V \geq 1)$$

If we use a procedure that controls for the family-wise error rate, we talk about simultaneous inference (or simultaneous coverage for confidence intervals).

# Bonferroni's procedure

Consider a family of  $m$  hypothesis tests and perform each test at level  $\alpha/m$ .

- reject  $i$ th null  $\mathcal{H}_{i0}$  if the associated  $p$ -value  $p_i \leq \alpha/m$ .
- build confidence intervals similarly with  $1 - \alpha/m$  quantiles.

If the (raw)  $p$ -values are reported, reject  $\mathcal{H}_{0i}$  if  $m \times p_i \geq \alpha$  (i.e., multiply reported  $p$ -values by  $m$ )

# Holm's sequential method

Order the  $p$ -values of the family of  $m$  tests from smallest to largest

$$p_{(1)} \leq \cdots \leq p_{(m)}$$

associated to null hypothesis  $\mathcal{H}_{0(1)}, \dots, \mathcal{H}_{0(m)}$ .

**Idea** use a different level for each test, more stringent for smaller  $p$ -values.

Coupling Holm's method with Bonferroni's procedure: compare  $p_{(1)}$  to  $\alpha_{(1)} = \alpha/m$ ,  $p_{(2)}$  to  $\alpha_{(2)} = \alpha/(m-1)$ , etc.

**Holm-Bonferroni procedure is always more powerful than Bonferroni**

# Holm-Bonferroni procedure

## Sequential testing

- start with the smallest  $p$ -value
- check significance one test at a time
- stop when the first nonsignificant  $p$ -value is found or no more test in store.

## Conclusion

If  $p_{(j)} \geq \alpha_{(j)}$  but  $p_{(i)} \leq \alpha_{(i)}$  for  $i = 1, \dots, j-1$  (all smaller  $p$ -values)

- reject  $\mathcal{H}_{0(1)}, \dots, \mathcal{H}_{0(j-1)}$
- fail to reject  $\mathcal{H}_{0(j)}, \dots, \mathcal{H}_{0(m)}$

If  $p_{(i)} \leq \alpha_{(i)}$  for all test  $i = 1, \dots, m$

- reject  $\mathcal{H}_{0(1)}, \dots, \mathcal{H}_{0(m)}$

# Numerical example

Consider  $m = 3$  tests with raw  $p$ -values 0.01, 0.04, 0.02.

$i$	$p_{(i)}$	Bonferroni	Holm-Bonferroni
1	0.01	$3 \times 0.01 = 0.03$	$3 \times 0.01 = 0.03$
2	0.02	$3 \times 0.02 = 0.06$	$2 \times 0.02 = 0.04$
3	0.04	$3 \times 0.04 = 0.12$	$1 \times 0.04 = 0.04$

Reminder of Holm–Bonferroni: multiply by  $(m - i + 1)$  the  $i$ th smallest  $p$ -value  $p_{(i)}$ , compare the product to  $\alpha$ .

# Why choose Bonferroni's procedure?

- simple
- generally applicable (any design)
- but dominated by sequential procedures (Holm-Bonferroni uniformly more powerful)
- low power when the number of test  $m$  is large
- $m$  must be prespecified



# Alternative measures

The FWER does not make a distinction between one or multiple type I errors.

We can also look at the more stringent criterion **per-family error rate**,  $\text{PFER} = E(V)$ , the expected (theoretical average) number of false positive.

One can show that

$$\text{FWER} = \Pr(V \geq 1) \leq E(V),$$

Any procedure that controls the per-family error rate thus also controls the familywise error rate: Bonferroni does.

# Dedicated methods for ANOVA

All methods valid with equal group variances and independent observations.

- **Tukey**'s honestly significant difference (HSD) method: to compare (all) pairwise differences between subgroups, based on the largest possible pairwise mean differences, with extensions for unbalanced samples.
- **Scheffé**'s method: applies to any contrast (properties depends on sample size  $n$  and number of groups  $K$ , not the number of test). Better than Bonferroni if  $m$  is large. Can be used for any design, but not powerful.
- **Dunnett**'s method: only for all pairwise contrasts relative to a specific baseline (control).

Described in Dean, Voss and Draguljić (2017) in more details.

# Adjustment for one-way ANOVA

Similar ideas but different **critical coefficients**. All derived using software.

Proceed only if there is a significant difference between groups, i.e. if we reject global null.

With  $K = 5$  groups and  $n = 9$  individuals per group (arithmetic example), critical value for two-sided test of zero difference with standardized  $t$ -test statistic and  $\alpha = 5\%$  are

- Scheffé's (all contrasts): 3.229
- Tukey's (all pairwise differences): 2.856
- Dunnett's (difference to baseline): 2.543
- unadjusted Student's  $t$ -distribution: 2.021

# False discovery rate

Suppose that  $m_0$  out of  $m$  hypothesis are true null (so  $\mathcal{H}_0$  holds  $m_0$  times).

The **false discovery rate** is the proportion of false discovery among rejected nulls,

$$\text{FDR} = \begin{cases} \frac{V}{R} & R > 0, \\ 0 & R = 0. \end{cases}$$

False discovery rate offers weak-FWER control

the property is only guaranteed under the scenario where all null hypotheses are true.

# False discovery rate vs FWER

A simultaneous procedure that controls family-wise error rate (FWER) ensure any selected test has type I error  $\alpha$ .

The false discovery rate (FDR) is less stringent: it's a guarantee for the proportion **among selected** discoveries.

But false discovery rate is scalable:

- 2 type I errors out of 4 tests is unacceptable.
- 2 type I errors out of 100 tests is probably okay.

# Controlling false discovery rate

The Benjamini-Hochberg (1995) procedure

1. Order the  $p$ -values from the  $m$  tests from smallest to largest:  $p_{(1)} \leq \dots \leq p_{(m)}$
2. For level  $\alpha$  (e.g.,  $\alpha = 5\%$ ), set

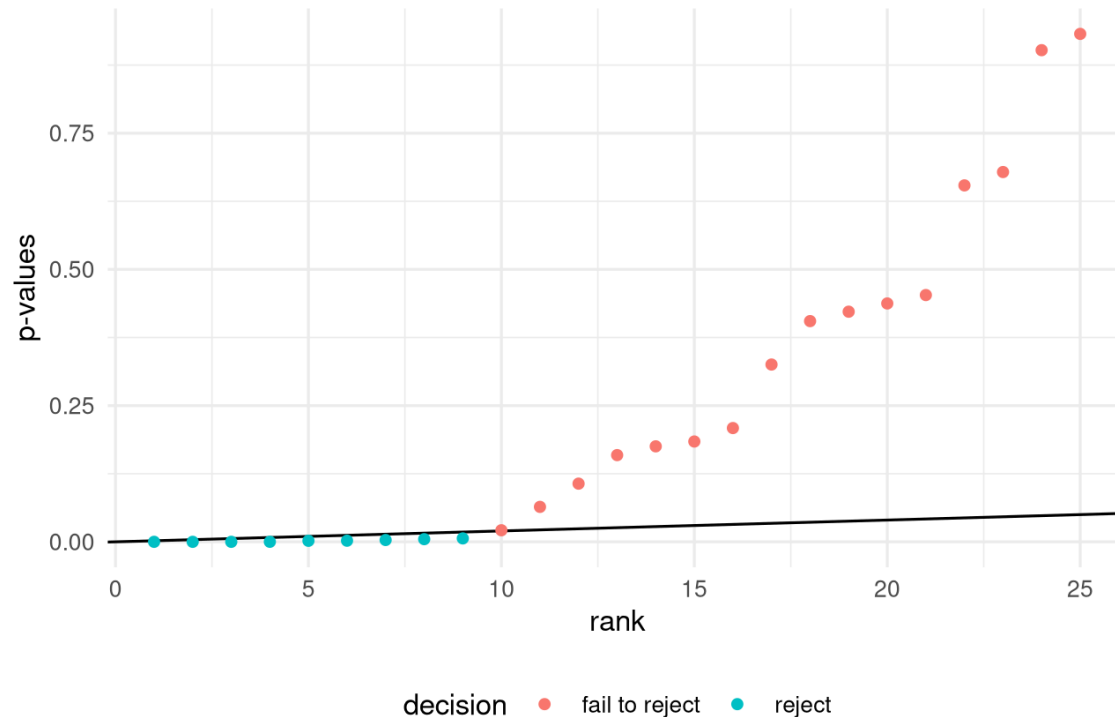
$$k = \max \left\{ i : p_{(i)} \leq \frac{i}{m} \alpha \right\}$$

3. Reject  $\mathcal{H}_{0(1)}, \dots, \mathcal{H}_{0(k)}$ .

# Picture of Benjamini-Hochberg

Plot the  $m$   $p$ -values against their rank.

To ensure  $\text{FDR} \leq q$ , reject null hypotheses corresponding to  $p$ -value that fall below the line of slope  $\alpha/m$ .



# Exercise

**Table S3**

*Planned Comparisons in Study 2*

	Other (immersed & distanced) vs. Self-immersed	Self-distanced vs. Self-immersed	Other-distanced vs. Other-immersed	Other (immersed & distanced) vs. Self-distanced
Variables	<i>t</i> ( <i>p</i> -value)	<i>t</i> ( <i>p</i> -value)	<i>t</i> ( <i>p</i> -value)	<i>t</i> ( <i>p</i> -value)
LIMITS	1.74 (.09)	2.16 (.03)	0.06 (.96)	0.81 (.42)
COMPR	2.02 (.046)	1.95 (.05)	0.05 (.96)	0.31 (.76)
PERSP	4.82 (< .001)	2.83 (.005)	0.74 (.46)	1.28 (.20)
CHANGE	1.80 (.08)	0.06 (.96)	0.15 (.88)	1.63 (.11)

*Note.* LIMITS - Recognition of limits of knowledge; COMPR - Search for a compromise; PERSP - Consideration of others' perspectives; CHANGE - Recognition of change; Planned comparisons include information from all four cells in the denominator.

Grossman, I. and E. Kross (2014). Exploring "Solomon's paradox": Self-distancing eliminates the self-other asymmetry in wise reasoning about close relations in younger and older adults, *Psychological Science*, 25(8) 1571-1580



# Summary (1/2)

- Researchers often carry lots of hypothesis testing tests
  - the more you look, the more you find!
- Thus want to control probability of making a judicial mistake among all  $m$  tests performed
  - (family-wise error rate, FWER)
- Less stringent criterion: control for the **proportion** of condemned (rejections) that were innocent
  - (false discovery rate, FDR)
  - useful if you don't care about making some mistakes, but perform loads of test (potentially millions)

# Summary (2/2)

- ANOVA specific solutions: assumes normal data, equal variance, balanced samples...
  - Tukey's HSD (all pairwise differences),
  - Dunnett's method (only differences relative to a reference category)
  - Scheffé's method (all contrasts)
- General methods
  - FWER: Bonferroni (suboptimal), Bonferroni-Holm (more powerful)
  - FDR: Benjamini-Hochberg

Downside of adjustment is loss of power (but more robust findings).

# Rant about $p$ -values

The American Statistical Association (ASA) published a list of principles guiding (mis)interpretation of  $p$ -values.

- (2)  $P$ -values do not measure the probability that the studied hypothesis is true
- (3) Scientific conclusions and business or policy decisions should not be based only on whether a  $p$ -value passes a specific threshold.
- (4)  $P$ -values and related analyses should not be reported selectively
- (5)  $p$ -value, or statistical significance, does not measure the size of an effect or the importance of a result