# Power and model assumptions

#### **Session 5**

MATH 80667A: Experimental Design and Statistical Methods for Quantitative Research in Management HEC Montréal

#### Outline

Power

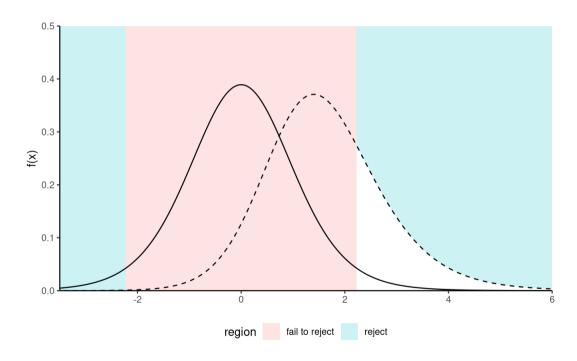
Effect and sample sizes

**Model assumptions** 

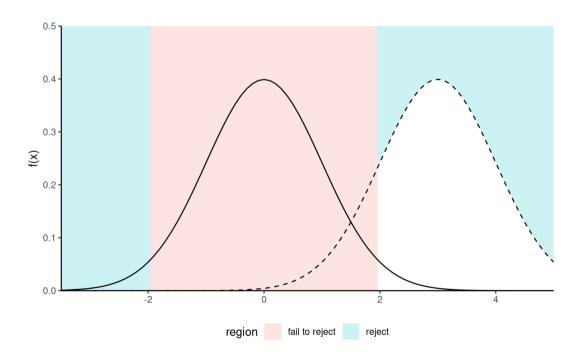
## Power

### I cried power!

The null alternative corresponds to a single value:  $\mu_1 = \cdots = \mu_K$ , whereas there are infinitely many alternatives...



Power is the ability to detect when the null is false, for a given alternative (dashed).



Power is the area in white under the dashed curved, beyond the cutoff.

## Parametrization of one-way ANOVA

#### group j has $n_j$ observations

population average of group j is  $\mu_j$ 

We can parametrize the model in terms of the overall sample average,

$$\mu = rac{1}{n} \sum_{j=1}^K \sum_{i=1}^{n_j} \mu_j = rac{1}{n} \sum_{j=1}^K n_j \mu_j,$$

where  $n=n_1+\cdots+n_K$  is the total sample size.

### What determines power?

Think in your head of potential factors.

- 1. The size of the effects,  $\delta_1=\mu_1-\mu$ ,  $\ldots$  ,  $\delta_K=\mu_K-\mu$
- 2. The background noise (intrinsic variability,  $\sigma^2$ )
- 3. The level of the test,  $\alpha$
- 4. The sample size in each group,  $n_j$
- 5. The choice of experimental design
- 6. The choice of test statistic

We focus on the interplay between

effect size | power | sample size

## Power and sample size calculations

Journals and grant agencies oftentimes require an estimate of the sample size needed for a study.

- large enough to pick-up effects of scientific interest (good signal-to-noise)
- efficient allocation of resources (don't waste time/money)

Same for study replication: how many participants needed?

## Living in an alternative world

Recall that with K treatments (groups) n observations, the F-statistic is

$$F = \frac{\text{between sum of squares}/(K-1)}{\text{within sum of squares}/(n-K)}$$

The null distribution is F(K-1, n-K).

The denominator is an estimator of  $\sigma^2$  under both the null and alternative.

So how does the *F*-test behaves under an alternative?

#### Numerator of the *F*-test

What happens to the numerator?

$$\mathsf{E}(\text{between sum of squares}) = \sigma^2\{(K-1) + \Delta\}.$$

where

$$\Delta = rac{\sum_{j=1}^K n_j (\mu_j - \mu)^2}{\sigma^2}.$$

Under the null hypothesis,  $\mu_j = \mu$  for  $j = 1, \ldots, K$  and  $\Delta = 0$ .

The greater  $\Delta$ , the further the mode (peak of the distribution) is from zero.

### Noncentrality parameter and power

$$\Delta = rac{\sum_{j=1}^K n_j (\mu_j - \mu)^2}{\sigma^2}.$$

#### When does power increase?

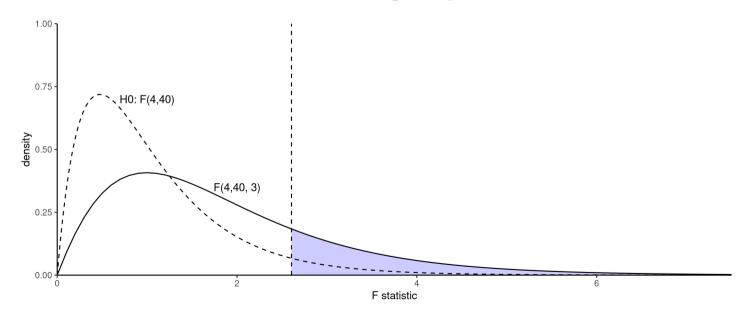
What is the effect of an increase of the

- group sample size  $n_1, \ldots, n_K$ .
- variability  $\sigma^2$ .
- true mean difference  $\mu_i \mu$ .

## Noncentrality parameter

The distribution is  $F(\nu_1, \nu_2, \Delta)$  distribution with degrees of freedom  $\nu_1$  and  $\nu_2$  and noncentrality parameter  $\Delta$ .

One-way ANOVA with n observations and K groups:  $u_1=K-1$  and  $u_2=n-K$ .



Note: the  $F(
u_1, 
u_2)$  distribution is indistinguishable from  $\chi^2(
u_1)$  for  $u_2$  large.

## Computing power

Given a value of  $\Delta$ , we can compute the tail probability as follows

1. Compute the cutoff point: the value under  $\mathscr{H}_0$  that leads to rejection at level  $\alpha$ .

```
cutoff <- qf(p = 1-alpha, df1 = df1, df2 = df2)</pre>
```

2. Compute probability below the alternative curve, from the cutoff onwards.

## Example from the OSC psychology replication

The key statistics provided in the paper to test the "depletion" hypothesis is the main effect of a one-way ANOVA with three experimental conditions and confirmatory information processing as the dependent variable (  $F(2,82)=4.05, p=0.02, \eta^2=0.09$ ). Considering the original effect size and an alpha of 0.05 the sample size needed to achieve 90% power is 132 subjects.

Replication report of Fischer, Greitemeyer, and Frey (2008, JPSP, Study 2) by E.M. Galliani

Q: What is the sample size for given power?

#### How do we compute the power

Assume that the design is balanced, meaning  $n_1=\dots=n_k=n/k$ . Then,

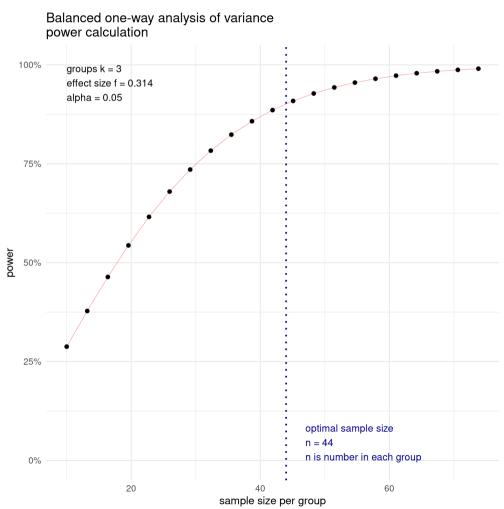
$$\Delta = rac{n}{k\sigma^2} \sum_{j=1}^k (\mu_j - \mu)^2.$$

Plug-in  $_{\rm df1}=k-1$ ,  $_{\rm df2}=n-k$  and  $_{\rm ncp}=\Delta$  for fixed mean difference, level and number of groups in the formulas of the previous slide.

#### Power curves

```
library(pwr)
power_curve <-
pwr.anova.test(
    f = sqrt(0.09/(1-0.09)),
    k = 3,
    power = 0.9,
    sig.level = 0.05)
plot(power_curve)</pre>
```

Convert  $\eta^2$  to Cohen's f (the effect size reported in pwr) via  $f^2=\eta^2/(1-\eta^2)$ . Explained later in the slides.



#### Power calculations with unknowns

Need to specify the following parameters:

Level  $\alpha$  (usually 5%)

adjust if you plan to correct for multiple testing!

Variance  $\sigma^2$ 

run a pilot study or look up plausible values from the literature

Mean difference size  $\mu_i - \mu$ 

## Estimating mean differences

If you care for a difference of D units between groups, the smallest value of  $\Delta$  (worst case analysis) is when

- ullet one group has mean difference  $\mu_1-\mu=-D/2$  ,
- ullet another has  $\mu_2-\mu=D/2$  and
- the rest with zero difference.

With a balanced design, we thus get

$$\Delta = rac{n}{k\sigma^2} \sum_{j=1}^k (\mu_j - \mu)^2 = rac{n}{k\sigma^2} igg(rac{D^2}{4} + rac{D^2}{4}igg) = rac{nD^2}{2k\sigma^2}.$$

#### Effect size estimates

#### **WARNING!**

Most effects reported in the literature are severely inflated.

#### Publication bias & the file drawer problem

Estimates reported in meta-analysis, etc. are not reliable. Use scientific knowledge

Replication reveal serious need for shrinkage.

The estimated effects size also have uncertainty: thus report confidence intervals.

#### Beware of small samples

Better to do a large replication than multiple small studies.

Otherwise, you risk being in this situation:



## Observed (post-hoc) power

Sometimes, the estimated values of the effect size, etc. are used as plug-in.

The (estimated) effect size in studies are noisy!

The post-hoc power estimate is also noisy and typically overoptimistic.

#### **Statistical fallacy**

Because we reject a null doesn't mean the alternative is true.

When is this relevant? If the observed difference seem important (large), but there isn't enough evidence (too low signal-to-noise).

# Effect size

#### Measures of effects

#### Multiple alternatives

- standardized differences/measures
- percentage of variability explained (ratio of sum of squares or coefficients of determination).

They are all interrelated, but different people have different preferences.

Popularized in the handbook

Cohen, Jacob. Statistical Power Analysis for the Behavioral Sciences, 2nd ed., Routhledge, 1988.

#### Cohen's d

Popular standardized measure of effect (dimensionless=no units): e.g., for a two-sample t-test between groups (comparing two means):

$$d=rac{\widehat{\mu}_1-\widehat{\mu}_2}{\widehat{\sigma}}$$

where  $\widehat{\sigma}$  is the pooled variance. Note: a finite sample correction (Hedge) can be used. Cohen's d is sometimes reported in terms of effect size

• small ( d=0.2 ), medium ( d=0.5 ) or large ( d=0.8 ) effect: these values change from one test to the next (above is for the two-sample t-test).

### Eta-squared

Can break down the variability for sum of squares as 'total = within + between'. Define the percentage of variability explained by effect.

$$\eta^2 = rac{\mathsf{SS}_{ ext{effect}}}{\mathsf{SS}_{ ext{total}}},$$

Problem: this depends on the variability of the measurement. Often, in more complex design, you see instead the following reported

$$\eta^2 = rac{\mathsf{SS}_{ ext{effect}}}{\mathsf{SS}_{ ext{error}} + \mathsf{SS}_{ ext{effect}}}.$$

For one-way ANOVA (no repeated measurements), the two are equivalent.

### Eta-squared

For the balanced one-way ANOVA with exhaustive factors=all possible choices in the population, write

$$\eta^2=rac{F
u_1}{F
u_1+
u_2}$$

For the replication example,

$$\eta^2 = rac{4.05 imes 2}{4.05 imes 2 + 84} = 0.088 pprox 0.09$$

#### Omega square

Yet another measure of effect size.

For one-way ANOVA, a transformation of the F-statistic gives

$$\omega^2=rac{
u_1(F-1)}{
u_1(F-1)+n}$$

Partial versions for more complex designs, etc.

#### **Uncertainty matters!**

Every effect size is random (because inputs are)

Report confidence intervals with estimates.

# Model assumptions