# Identification of and correction for publication bias

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- Fundamental requirement of science: replicability
- Different researchers should reach same conclusions
- Methodological conventions should ensure this (e.g., randomized experiments)
- Replicability often appears to fail, e.g.
  - Experimental economics (Camerer et al., 2016)
  - Experimental psychology (Open Science Collaboration, 2015)
  - Medicine (Ionnidias, 2005)
  - Cell Biology (Begley et al, 2012)
  - Neuroscience (Button et al, 2013)

- Possible explanation: selective publication of results
- Due to:
  - Researcher decisions
  - Journal selectivity
- Possible selection criteria:
  - Statistically significant effects
  - Confirmation of prior beliefs
  - Novelty
- Consequences:
  - Conventional estimators are biased
  - Conventional inference does not control size

Literature

### Identification of publication bias:

- Good overview: Rothstein et al. (2006)
- Regression based:
   Egger et al. (1997)
- Symmetry of funnel plot ("trim and fill"):
   Duval and Tweedie (2000)
- Parametric selection models:
   Hedges (1992), Iyengar and Greenhouse (1988)
- Distribution of p-values, parametric distribution of true effects:
   Brodeur et al. (2016)

Literature

### **Corrected inference:**

McCrary et al. (2016)

### Replication- and meta-studies for empirical part:

- Replication of econ experiments: Camerer et al. (2016)
- Replication of psych experiments: Open Science Collaboration (2015)
- Minimum wage: Wolfson and Belman (2015)
- Deworming: Croke et al. (2016)

#### Our contributions

- Nonparametric identification of selectivity in the publication process, using
  - a) Replication studies: Absent selectivity, original and replication estimates should be symmetrically distributed
  - Meta-studies: Absent selectivity, distribution of estimates for small sample sizes should be noised-up version of distribution for larger sample sizes
- Corrected inference when selectivity is known
  - a) Median unbiased estimators
  - b) Confidence sets with correct coverage
  - c) Allow for nuisance parameters and multiple dimensions of selection
  - d) Bayesian inference accounting for selection
- Applications to
  - a) Experimental economics
  - b) Experimental psychology
  - c) Effects of minimum wages on employment
  - d) Effects of de-worming

# **Outline**

- Introduction
- Setup
- 3 Identification
- Bias-corrected inference
- 5 Applications
- 6 Conclusion

- Assume there is a population of latent studies indexed by i
- True parameter value in study i is Θ<sub>i</sub>\*
  - $\Theta_i^*$  drawn from some population  $\Rightarrow$  empirical Bayes perspective
  - Different studies may recover different parameters
- Each study reports findings X<sub>i</sub>\*
  - Distribution of  $X_i^*$  given  $\Theta_i^*$  known
- A given study may or may not be published
  - Determined by both researcher and journal: we don't try to disentangle
- Probability of publication  $P(D_i = 1 | X_i^*, \Theta_i^*) = p(X_i^*)$
- Published studies are indexed by j

### Definition (General sampling process)

Latent (unobserved) variables:  $(D_i, X_i^*, \Theta_i^*)$ , jointly i.i.d. across i

$$egin{aligned} \Theta_i^* &\sim \mu \ X_i^* | \Theta_i^* &\sim f_{X^* | \Theta^*}(x | \Theta_i^*) \ D_i | X_i^*, \Theta_i^* &\sim \mathit{Ber}(p(X_i^*)) \end{aligned}$$

Truncation: We observe i.i.d. draws of  $X_j$ , where

$$I_{j} = \min\{i: D_{i} = 1, i > I_{j-1}\}$$
  
 $\Theta_{j} = \Theta_{l_{j}}^{*}$   
 $X_{j} = X_{l_{j}}^{*}$ 

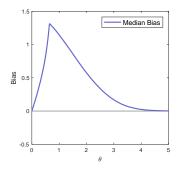
#### Example: treatment effects

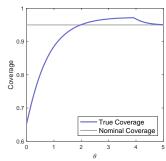
- Journal receives a stream of studies i = 1, 2, ...
- Each reporting experimental estimates  $X_i^*$  of treatment effects  $\Theta_i^*$
- Distribution of  $\Theta_i^*$ :  $\mu$
- Suppose that  $X_i^*|\Theta_i^* \sim N(\Theta_i^*, 1)$
- Publication probability: "significance testing,"

$$p(X) = \begin{cases} 0.1 & |X| < 1.96 \\ 1 & |X| \ge 1.96 \end{cases}$$

• Published studies: report estimate  $X_j$  of treatment effect  $\Theta_j$ 

### Example continued - Publication bias





- Left: median bias of  $\hat{ heta}_j = X_j$
- Right: true coverage of conventional 95% confidence interval

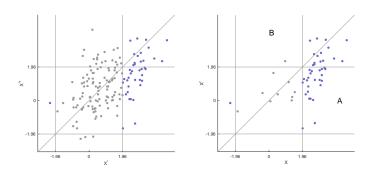
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Identification of the selection mechanism  $p(\cdot)$ 

- Key unknown object in model: publication probability  $p(\cdot)$
- We propose two approaches for identification:
  - Replication experiments:
    - ullet replication estimate  $X^r$  for the same parameter  $\Theta$
    - selectivity operates only on X, but not on X<sup>r</sup>
  - Meta-studies:
    - Variation in  $\sigma^*$ , where  $X^* \sim N(\Theta^*, \sigma^{*2})$
    - Assume  $\sigma^*$  is (conditionally) independent of  $\Theta^*$  across latent studies i
    - Standard assumption in the meta-studies literature; validated in our applications by comparison to replications
- Advantages:
  - Replications: Very credible
  - Meta-studies: Widely applicable

Intuition: identification using replication studies



- Left: no truncation
   ⇒ areas A and B have same probability
- Right:  $p(Z) = 0.1 + 0.9 \cdot \mathbf{1}(|Z| > 1.96)$  $\Rightarrow$  A more likely then B

Approach 1: Replication studies

## Definition (Replication sampling process)

• Latent variables: as before,

$$egin{aligned} \Theta_i^* &\sim \mu \ X_i^* | \Theta_i^* &\sim f_{X^* | \Theta^*}(x | \Theta_i^*) \ D_i | X_i^*, \Theta_i^* &\sim \mathit{Ber}(p(X_i^*)) \end{aligned}$$

Additionally: replication draws,

$$X_i^{*r}|X_i^*, D_i, \Theta_i^* \sim f_{X^*|\Theta^*}(x|\Theta_i^*)$$

Observability: as before,

$$I_{j} = \min\{i: D_{i} = 1, i > I_{j-1}\}$$
  
 $\Theta_{j} = \Theta_{I_{j}}$   
 $(X_{j}, X_{j}^{r}) = (X_{l_{j}}^{*}, X_{l_{j}}^{*r})$ 

# Theorem (Identification using replication experiments)

Assume that the support of  $f_{X_i^*,X_i^{*r}}$  is of the form  $A \times A$  for some set A. Then  $p(\cdot)$  is identified on A up to scale.

### Intuition of proof:

• Marginal density of  $(X, X^r)$  is

$$f_{X,X'}(x,x') = \frac{p(x)}{E[p(X_i^*)]} \int f_{X^*|\Theta^*}(x|\theta_i^*) f_{X^*|\Theta^*}(x'|\theta_i^*) d\mu(\theta_i^*)$$

• Thus, for all a, b, if p(a) > 0,

$$\frac{p(b)}{p(a)} = \frac{f_{X,X^r}(b,a)}{f_{X,X^r}(a,b)}$$

#### Practical complication

- Replication experiments follow the same protocol
   ⇒ estimate same effect Θ
- But often different sample size
   ⇒ different variance ⇒ symmetry breaks down
- Additionally: replication sample size often determined based on power calculations given initial estimate
- $p(\cdot)$  is still identified (up to scale):
  - Assume X normally distributed
  - Intuition: Conditional on X, σ, (de-)convolve X<sup>r</sup> with normal noise to get symmetry back
  - μ is identified as well

#### Further complication

- What if selectivity is based not only on observed X, but also on unobserved W?
- Would imply general selectivity of the form

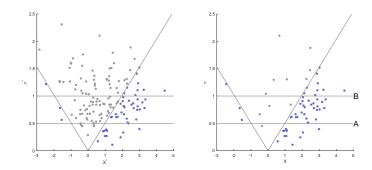
$$D_i|X_i^*,\Theta_i^*\sim Ber(p(X_i^*,\Theta_i^*))$$

Again assume normality,

$$X_i^{*r}|\sigma_i, D_i, X_i^*, \Theta_i^* \sim N(\Theta_i^*, \sigma_i^2)$$

- ⇒ Solution:
  - Identify  $\mu_{\Theta|X}$  from  $f_{X^r|X}$  by deconvolution
  - Recover f<sub>X|Θ</sub> by Bayes' rule (f<sub>X</sub> is observed)
  - This density is all we need for bias corrected inference
- We use this to construct specification tests for our baseline model

Intuition: identification using meta-studies



- Left: no truncation dist for higher  $\sigma$  noised up version of dist for lower  $\sigma$
- Right:  $p(Z) = 0.1 + 0.9 \cdot \mathbf{1}(|Z| > 1.96)$  $\Rightarrow$  "missing data" inside the cone

Approach 2: meta-studies

# Definition (Independent $\sigma$ sampling process)

$$egin{aligned} \sigma_i^* &\sim \mu_\sigma \ \Theta_i^* | \sigma_i^* &\sim \mu_\Theta \ X_i^* | \Theta_i^*, \sigma_i^* &\sim extstyle N(\Theta_i^*, \sigma_i^{*2}) \ D_i | X_i^*, \Theta_i^*, \sigma_i^* &\sim extstyle Ber(p(X_i^*/\sigma_i^*)) \end{aligned}$$

We observe i.i.d. draws of  $(X_j, \sigma_j)$ , where

$$I_j = \min\{i : D_i = 1, i > I_{j-1}\}\$$
  
 $(X_j, \sigma_j) = (X_{l_j}^*, \sigma_{l_j}^*)$ 

Define  $Z^* = \frac{X^*}{\sigma^*}$  and  $Z = \frac{X}{\sigma}$ 

## Theorem (Nonparametric identification using variation in $\sigma$ )

Suppose that the support of  $\sigma$  contains a neighborhood of some point  $\sigma_0$ . Then  $p(\cdot)$  is identified up to scale.

### Intuition of proof:

• Conditional density of Z given  $\sigma$  is

$$f_{Z|\sigma}(z|\sigma) = \frac{p(z)}{E[p(Z^*)|\sigma]} \int \varphi(z-\theta/\sigma) d\mu(\theta)$$

Thus

$$\frac{f_{Z|\sigma}(z|\sigma_2)}{f_{Z|\sigma}(z|\sigma_1)} = \frac{E[p(Z^*)|\sigma=\sigma_1]}{E[p(Z^*)|\sigma=\sigma_2]} \cdot \frac{\int \varphi(z-\theta/\sigma_2)d\mu(\theta)}{\int \varphi(z-\theta/\sigma_1)d\mu(\theta)}$$

 Recover μ from right hand side, then recover p(·) from first equation

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- Once we know  $p(\cdot)$ , can correct inference for selection
- For simplicity, here assume X,  $\Theta$  both 1-dimensional
- Density of published X given Θ:

$$f_{X|\Theta}(x|\theta) = \frac{p(x)}{E[p(X^*)|\Theta^* = \theta]} \cdot f_{X^*|\Theta^*}(x|\theta)$$

• Corresponding cumulative distribution function:  $F_{X|\Theta}(x|\theta)$ 

### Corrected frequentist estimators and confidence sets

- We are interested in bias, and the coverage of confidence sets
  - Condition on  $\theta$ : standard frequentist analysis
- Define  $\hat{\theta}_{\alpha}(x)$  via

$$F_{X|\Theta}\left(x|\hat{\theta}_{\alpha}\left(x\right)\right)=\alpha$$

Under mild conditions, can show that

$$P\left(\hat{\theta}_{\alpha}\left(X\right) \leq \theta | \theta\right) = \alpha \ \forall \theta$$

- Median-unbiased estimator:  $\hat{\theta}_{\frac{1}{2}}(X)$  for  $\theta$
- Equal-tailed level  $1 \alpha$  confidence interval:

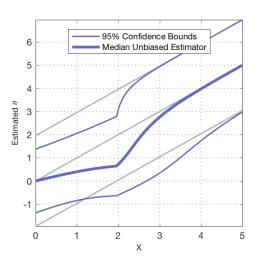
$$\left[\hat{\theta}_{\frac{\alpha}{2}}(X),\hat{\theta}_{1-\frac{\alpha}{2}}(X)\right]$$

Example: treatment effects

- Let us return to the treatment effect example discussed above
- Again assume  $X^*|\Theta^* \sim N(\Theta^*, 1)$  and

$$p(X) = 0.1 + 0.9 \cdot \mathbf{1}(|X| > 1.96)$$

Example continued – corrected confidence sets for  $\beta_p = 0.1$ 



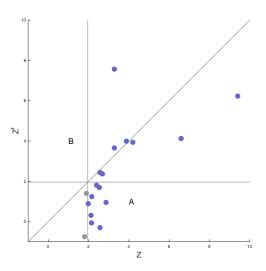
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Replications of Lab Experiments in Economics

- Camerer et al. (2016)
- Sample: all 18 between-subject laboratory experimental papers published in AER and QJE between 2011 and 2014
- Scatterplot next slide:
  - $Z = X/\sigma$ : normalized initial estimate
  - $Z^r = X^r/\sigma$ : replicate estimate
  - Initial estimates normalized to be positive

Economics Lab Experiments: Original and Replication Z Statistics



Economics Lab Experiments: Estimates of Selection model

Model:

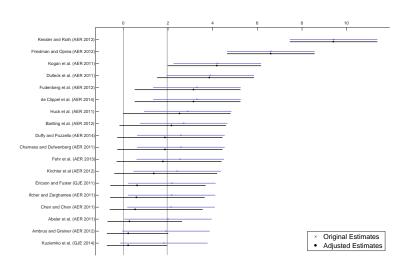
$$|\Theta^*| \sim \Gamma(\kappa, \lambda)$$
 $p(Z) \propto egin{cases} eta_p & |Z| < 1.96 \ 1 & |Z| \geq 1.96 \end{cases}$ 

Estimates:

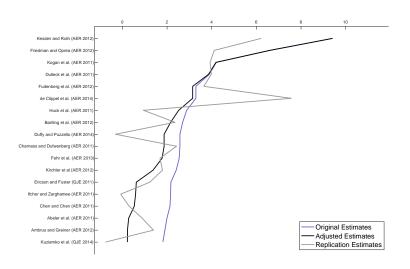
К	λ	$eta_{ ho}$
0.373	2.153	0.029
(0.266)	(1.024)	(0.027)

 Interpretation: insignificant (at the 5 % level) results about 3% as likely to be published as significant results

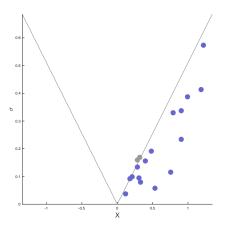
### Economics Lab Experiments: Adjusted Estimates



### Economics Lab Experiments: Adjusted Estimates



Economics Lab Experiments: Meta-study Approach



Economics Lab Experiments: Meta-study Results

Model:

$$|\Theta^*| \sim \Gamma(\tilde{\kappa}, \tilde{\lambda})$$
 $p(X/\sigma) \propto \begin{cases} \beta_p & |X/\sigma| < 1.96 \\ 1 & |X/\sigma| \ge 1.96 \end{cases}$ 

Recall replication-based estimates:

κ	λ	$eta_{p}$
0.373	2.153	0.029
(0.266)	(1.024)	(0.027)

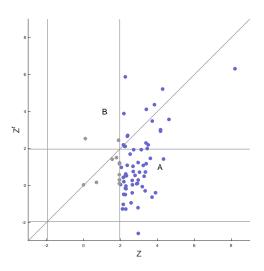
• Meta-study based estimates (only  $\beta_p$  comparable):

$ ilde{\kappa}$	$ ilde{\lambda}$	$\beta_{p}$
1.343	0.157	0.038
(1.310)	(0.076)	(0.051)

Replications of Lab Experiments in Psychology

- Open Science Collaboration (2015)
- 270 contributing authors
- Sample: 100 out of 488 articles published 2008 in
  - Psychological Science
  - Journal of Personality and Social Psychology
  - Journal of Experimental Psychology: Learning, Memory, and Cognition
- Some critiques by Gilbert et al. (2016):
  - statistical misinterpretation,
  - not all replication protocols endorsed by original authors
    - ⇒ we re-run estimators on subset of approved replications

Experiments in Psychology: Original and Replication Z Statistics



#### Experiments in Psychology: Estimates of Selection Model

Model:

$$|\Theta^*| \sim \Gamma(\kappa, \lambda)$$

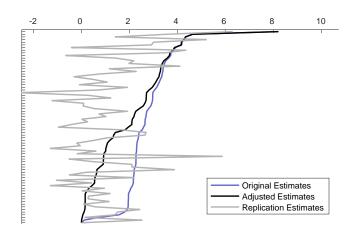
$$p(Z) \propto \begin{cases} \beta_{p1} & |Z| < 1.64 \\ \beta_{p2} & 1.64 \le |Z| < 1.96 \\ 1 & |Z| \ge 1.96 \end{cases}$$

Estimates:

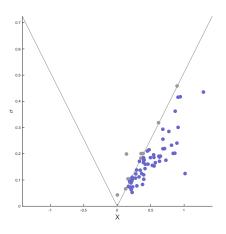
κ	λ	$eta_{p,1}$	$eta_{p,2}$
0.315	1.308	0.009	0.205
(0.143)	(0.334)	(0.005)	(0.088)

- Results insignificant at the 10% level 1% as likely to be published as results significant at 5% level
- Results significant at the 5% level five times as likely to be published as results significant at 10% level

Original and Replication Z Statistics: Psychology Lab Experiments



Psychology Lab Experiments: Meta-studies Approach



Psychology Lab Experiments: Estimates of Meta-studies Selection Model

Model:

$$|\Theta^*| \sim \Gamma(\tilde{\kappa}, \tilde{\lambda})$$

$$\rho(Z) \propto \begin{cases} \beta_{p1} & |Z| < 1.64 \\ \beta_{p2} & 1.64 \leq |Z| < 1.96 \\ 1 & |Z| \geq 1.96 \end{cases}$$

Recall replication-based estimates:

κ	λ	$eta_{p,1}$	$eta_{ ho,2}$
0.315	1.308	0.009	0.205
(0.143)	(0.334)	(0.005)	(0.088)

• Meta-study based estimates (only  $\beta_p$  comparable):

$ ilde{\kappa}$	$ ilde{\lambda}$	$eta_{p,1}$	$eta_{p,2}$
0.974	0.153	0.017	0.306
(0.549)	(0.053)	(0.009)	(0.135)

Psychology Lab Experiments: Approved Replications

- 67 studies
- Replication-based estimates:

κ	λ	$eta_{p,1}$	$eta_{p,2}$
0.490	1.159	0.017	0.365
(0.268)	(0.402)	(0.011)	(0.165)

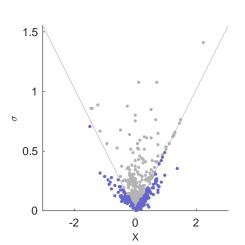
Meta-study based estimates:

$ ilde{\kappa}$	$\tilde{\lambda}$	$eta_{p,1}$	$eta_{ ho,2}$
0.634	0.198	0.022	0.440
(0.502)	(0.078)	(0.014)	(0.217)

ullet  $eta_{
ho}$  estimates larger than those in full dataset

Meta-study of the Effect of Minimum Wages on Employment

- Wolfson and Belman (2015)
- Elasticity of employment w.r.t. the minimum wage
   X > 0 ⇔ negative employment effect
- 1000 estimates from 37 studies using U.S. data that were circulated after 2000, either as articles in journals or as working papers
- For some: more than 1 estimate per study



### Estimates of selection model

Model:

$$\Theta^* \sim \bar{\theta} + t(v) \cdot \tilde{\tau}$$
 
$$p(X/\sigma) \propto \begin{cases} \beta_{p1} & X/\sigma < -1.96 \\ \beta_{p2} & -1.96 \le X/\sigma < 0 \\ \beta_{p3} & 0 \le X/\sigma < 1.96 \\ 1 & X/\sigma \ge 1.96 \end{cases}$$

- Recall X > 0 ⇔ negative employment effect.
- Estimates:

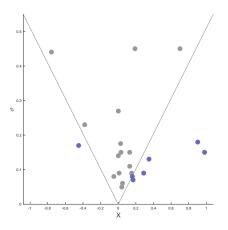
$\theta$	$ ilde{ au}$	$\tilde{v}$	$eta_{p,1}$	$eta_{ ho,2}$	$eta_{ ho,3}$	
0.018	0.019			0.270	0.323	
(0.009)	(0.011)	(0.279)	(0.350)	(0.111)	(0.094)	

Selection in favor of significant effects, negative effects.

Meta-Study of the Effects of Deworming

- Croke et al. (2016)
- Follow procedures outlined in the "Cochrane Handbook for Systematic Reviews of Interventions"
- Randomized controlled trials of deworming that include child body weight as an outcome
- 22 estimates from 20 studies

Meta-Study of the Effects of Deworming



Deworming: Estimates of selection model

Model:

$$\Theta^* \sim N(\bar{\theta}, \tau^2)$$

$$p(X) \propto \begin{cases} \beta_p & |X/\sigma| < 1.96 \\ 1 & |X/\sigma| \ge 1.96 \end{cases}$$

Estimates:

$ar{ heta}$	$ ilde{ au}$	$eta_{ ho}$
0.190	0.343	2.514
(0.120)	(0.128)	(1.869)

#### Conclusion

- Selectivity in the publication process is a potentially serious problem for statistical inference.
- We non-parametrically identify the form of selectivity:
  - Using replication studies:
     Original and replication estimates would be symmetrically distributed, absent selectivity
  - Using meta-studies:
     Under an independence assumption, higher-variance estimate distribution would be noised-up version of lower-variance estimate distribution, absent selectivity

#### Conclusion

- Easy correction for selectivity, if form is known:
  - Median unbiased estimators
  - Equal-tailed confidence sets with correct coverage
- Empirical findings:
  - Selectivity on significance in experimental economics, experimental psychology
  - Selectivity towards (negative) significant employment effects in minimum wage literature
  - Noisy estimates in meta-study for de-worming

