Chapter 6 in ISL: Regularization

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Regularization

- Another way to control bias and variance is through regularization or shrinkage.
- Rather than selecting a few predictors that seem reasonable, maybe trying a few combinations, use them all.
- I mean ALL.
- But, make your estimates of β "smaller"

Some optimization terms

- An optimization problem has 2 components:
 - 1. The "Objective function": e.g. $\frac{1}{n}\sum_i(y_i-x_i'\beta)^2$. 2. The "constraint": e.g. "fewer than 5 non-zero entries in β ".
- A constrained minimization problem is written

$$\min_{\beta} f(\beta)$$
 subject to $C(\beta)$

- $f(\beta)$ is the objective function
- $C(\beta)$ is the constraint

Regularization

One way to do this for regression is to solve (say):

$$\min_{\beta} \frac{1}{n} \sum_{i} (y_i - x_i' \beta)^2$$
s.t.
$$\sum_{j} \beta_j^2 < t$$

for some t > 0.

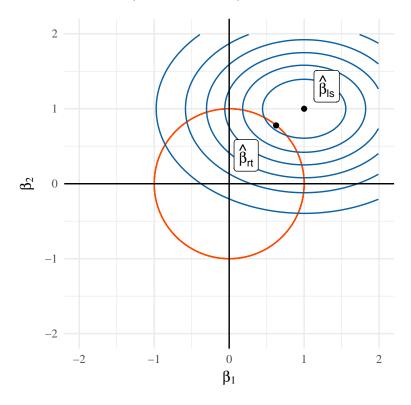
- This is called "ridge regression".
- The minimizer of this problem is called $\widehat{\beta}_{r,t}$

Compare this to least squares:

$$\min_{\beta} \frac{1}{n} \sum_{i} (y_i - x_i' \beta)^2$$

s.t. $\beta \in \mathbb{R}^p$

Geometry of ridge regression (2 dimensions)



Ridge regression

An equivalent way to write

$$\widehat{\beta}_{r,t} = \arg\min_{\|\beta\|_2^2 \le t} \frac{1}{n} \sum_i (y_i - x_i'\beta)^2$$

is in the $\frac{\text{Lagrangian}}{\text{Lagrangian}}$ form

$$\widehat{\beta}_{r,\lambda} = \arg\min_{\beta} \frac{1}{n} \sum_{i} (y_i - x_i'\beta)^2 + \lambda ||\beta||_2^2.$$

For every λ there is a unique t (and vice versa) that makes

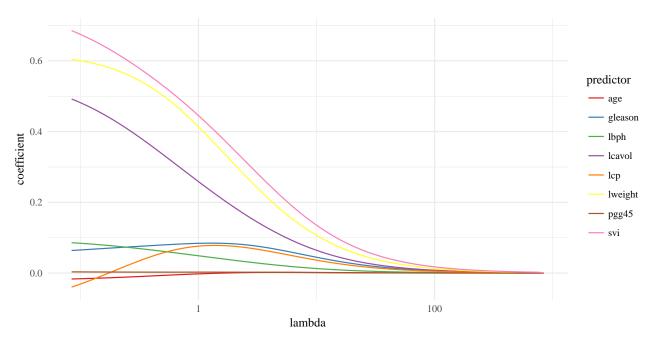
$$\widehat{\beta}_{r,t} = \widehat{\beta}_{r,\lambda}$$

Observe:

- $\lambda=0$ (or $t=\infty$) makes $\widehat{\beta}_{r,\lambda}=\widehat{\beta}_{ls}$ Any $\lambda>0$ (or $t<\infty$) penalizes larger values of β , effectively shrinking them.

Note: λ and t are known as tuning parameters

Ridge regression path



Regularization and rescaling

Least squares is invariant to rescaling

Let's multiply our design matrix by a factor of 10 to get $\widetilde{\mathbb{X}} = 10\mathbb{X}$.

Then:

$$\widetilde{\beta}_{ls} = (\widetilde{\mathbb{X}}^{\top}\widetilde{\mathbb{X}})^{-1}\widetilde{\mathbb{X}}^{\top}Y = \frac{1}{10}(\widetilde{\mathbb{X}}^{\top}\widetilde{\mathbb{X}})^{-1}\widetilde{\mathbb{X}}^{\top}Y = \frac{\widehat{\beta}_{ls}}{10}$$

So, multiplying our data by ten just results in our estimates being reduced by one tenth. Hence, any prediction is left unchanged:

$$\widetilde{\mathbb{X}}\widetilde{\beta}_{\mathrm{ls}} = \mathbb{X}\widehat{\beta}_{ls}$$

This means, for instance, if we have a covariate measured in miles, then we will get the "same" answer if we change it to kilometers

Least squares is invariant to rescaling: example

```
n = 20
set.seed(2018-04-10)
X = matrix(runif(2*n,0,1), ncol=2)
Y = X %*% c(.5,1.5) + rnorm(n,0,.25)
Xtilde = 2*X
Ytilde = Y - mean(Y)
summary(lm(Y~X))$coefficients
```

Estimate Std. Error t value Pr(>|t|) ## (Intercept) 0.08267477 0.1479918 0.5586444 5.836887e-01

```
## X1
           0.33116021 0.2329056 1.4218645 1.731531e-01
           1.43583489
                    0.2033167 7.0620598 1.908643e-06
## X2
summary(lm(Y~Xtilde))$coefficients
             Estimate Std. Error
                              t value
                                        Pr(>|t|)
## (Intercept) 0.08267477
                    0.1479918 0.5586444 5.836887e-01
## Xtilde1
           ## Xtilde2
summary(lm(Ytilde~Xtilde))$coefficients
##
             Estimate Std. Error
                              t value
                                        Pr(>|t|)
## (Intercept) -0.8103178  0.1479918 -5.475425  4.104055e-05
## Xtilde1
            ## Xtilde2
            0.7179174
                     0.1016584 7.062060 1.908643e-06
```

Ridge regression (and other regularized methods) is not

- ## 0.08864934 0.41708721
 - lm.ridge automatically scales every column of X to have mean zero and Euclidean norm 1.
 - It also centers Y.
 - Together, this means there is no intercept. (We don't penalize the intercept)
 - In R: scale(X) defaults to mean 0, SD 1. But you can change either.
 - Another version is in the package glmnet. More on this in a bit.

Solving the minimization

• One nice thing about ridge regression is that it has a closed-form solution (like OLS)

$$\widehat{\beta}_{r,\lambda} = (\mathbb{X}'\mathbb{X} + \lambda I)^{-1}\mathbb{X}'Y$$

- This is easy to calculate in R for any λ .
- However, computations and interpretation are simplified if we examine the Singular Value Decomposition of $\mathbb{X} = UDV'$.
- Then,

$$\widehat{\beta}_{r,\lambda} = (\mathbb{X}'\mathbb{X} + \lambda I)^{-1}\mathbb{X}'Y = (VD^2V' + \lambda I)^{-1}VDU'Y = V(D^2 + \lambda I)^{-1}DU'Y.$$

- For computations, now we only need to invert a diagonal matrix.
- For interpretations, we can compare this to OLS:

$$\widehat{\beta}_{ls} = (\mathbb{X}'\mathbb{X})^{-1}\mathbb{X}'Y = (VD^2V')^{-1}VDU'Y = VD^{-2}DU'Y = VD^{-1}U'Y$$

- Notice that $\widehat{\beta}_{ls}$ depends on d_j/d_i^2 while $\widehat{\beta}_{r,\lambda}$ depends on $d_j/(d_i^2 + \lambda)$.
- Ridge regression makes the coefficients smaller relative to OLS.
- But if X has small singular values, ridge regression compensates with λ in the denominator.

Ridge regression and multicollinearity

Multicollinearity is a phenomenon in which a combination of predictor variables is extremely similar to another predictor variable. Some comments:

- A better phrase that is sometimes used is "X is ill-conditioned"
- It means that one of its columns is nearly (or exactly) a linear combination of other columns. This is sometimes known as "(numerically) rank-deficient".
- If $\mathbb{X} = UDV'$ is ill-conditioned, then some elements of D are nearly zero
- If we form $\widehat{\beta}_{ls} = VD^{-1}U'Y$, then we see that the small entries of D are now huge (due to the inverse). This in turn creates a huge variance.
- Recall: $\mathbb{V}\widehat{\beta}_{ls} = \sigma^2(\mathbb{X}'\mathbb{X})^{-1} = \sigma^2 V D^{-2} V'$

Ridge Regression fixes this problem by preventing the division by a near zero number

Conclusion: $(\mathbb{X}^{\top}\mathbb{X})^{-1}$ can be really unstable, while $(\mathbb{X}^{\top}\mathbb{X} + \lambda I)^{-1}$ is not.

Can we get the best of both worlds?

To recap:

- Deciding which predictors to include, adding quadratic terms, or interactions is model selection.
- Ridge regression provides regularization, which trades off bias and variance and also stabilizes multicollinearity.

Ridge regression: $\min ||\mathbb{Y} - \mathbb{X}\beta||_2^2$ subject to $||\beta||_2^2 \leq t$

Best linear regression model: $\min ||\mathbb{Y} - \mathbb{X}\beta||_2^2$ subject to $||\beta||_0 \le t$

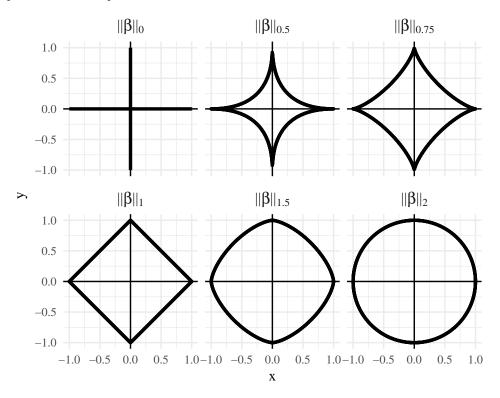
 $||\beta||_0$ is the number of nonzero elements in β

Finding the best linear model is a nonconvex optimization problem (In fact, it is NP-hard)

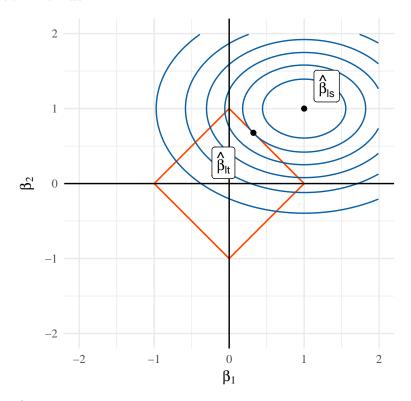
Ridge regression is convex (easy to solve), but doesn't do model selection

Can we somehow "interpolate" to get both?

Geometry of convexity



The best of both worlds



This regularization set. . .

- ... is convex (computationally efficient)
- ... has corners (performs model selection)

The lasso

ℓ_1 -regularized regression

Known as

- "lasso"
- "basis pursuit"

The estimator satisfies

$$\widehat{\beta}_{l,t} = \arg\min_{||\beta||_1 \le t} ||\mathbb{Y} - \mathbb{X}\beta||_2^2$$

In its corresponding Lagrangian dual form:

$$\widehat{\beta}_{l,\lambda} = \arg\min_{\beta} ||\mathbb{Y} - \mathbb{X}\beta||_{2}^{2} + \lambda ||\beta||_{1}$$

Lasso

While the ridge solution can be easily computed

$$\widehat{\beta}_{r,\lambda} = \arg\min_{\beta} ||\mathbb{Y} - \mathbb{X}\beta||_2^2 + \lambda ||\beta||_2^2 = (\mathbb{X}^\top \mathbb{X} + \lambda I)^{-1} \mathbb{X}^\top Y$$

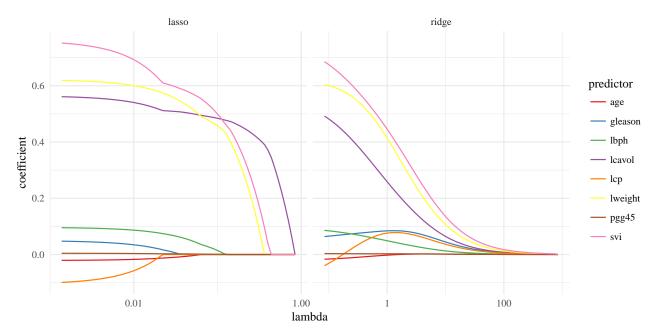
the lasso solution

$$\widehat{\beta}_{l,\lambda} = \arg\min_{\beta} ||\mathbb{Y} - \mathbb{X}\beta||_2^2 + \lambda ||\beta||_1 = ??$$

doesn't have a closed form solution.

However, because the optimization problem is convex, there exist efficient algorithms for computing it

Coefficient path: ridge vs lasso



Packages

There are two main R implementations for finding lasso

- Using glmnet: lasso.out = glmnet(X, Y, alpha=1).
- Setting alpha=0 gives ridge regression (as does lm.ridge in the MASS package)
- Setting $alpha \in (0,1)$ gives a method called the "elastic net" which combines ridge regression and lasso.
- Alternatively, there is lars: lars.out = lars(X, Y)
- lars also other things called "Least angle", "forward stepwise", and "forward stagewise" regression

Two packages

- 1. lars (this is the first one)
- 2. glmnet (this one is faster)

These use different algorithms, but both compute the path for a range of λ .

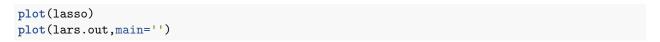
lars starts with an empty model and adds coefficients until saturated. The sequence of λ 's comes from the nature of the optimization problem.

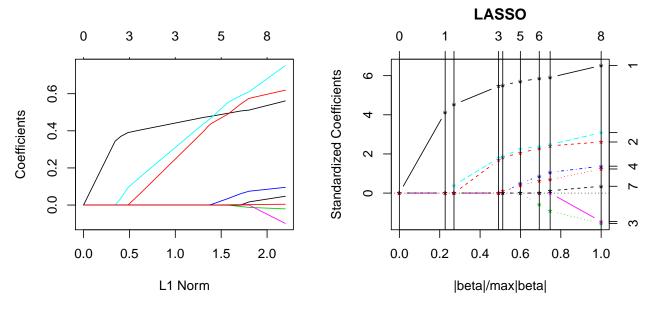
glmnet starts with an empty model and examines each value of λ using previous values as "warm starts". It is generally much faster than lars and uses lots of other tricks (as well as compiled code) for extra speed.

The path returned by lars as more useful than that returned by glmnet.

Lasso paths

```
lasso = glmnet(X,Y)
lars.out = lars(X,Y)
par(mfrow=c(1,2))
```





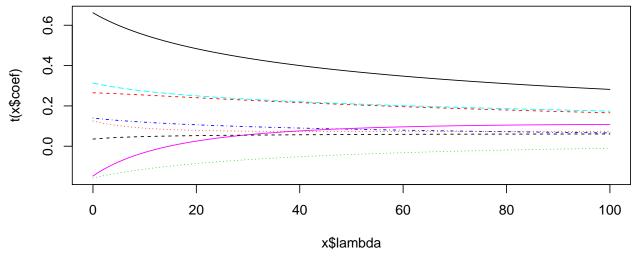
Model selection

Choosing the lambda

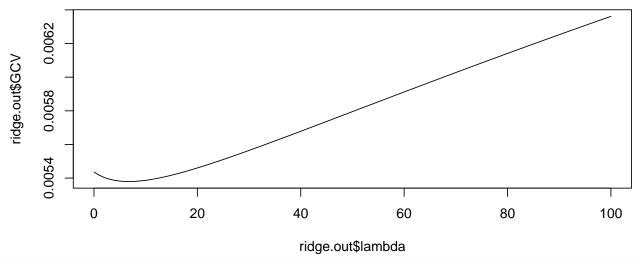
- You have to choose λ in lasso or in ridge regression
- lasso selects a model (by setting coefficients to zero), but the value of λ determines how many/which.
- All of these packages come with CV built in.
- However, the way to do it differs from package to package (whomp whomp)

Ridge regression, lm.ridge version

```
par(mfrow=c(1,1))
# 1. Estimate the model (note, this uses a formula, and you must supply lambda)
ridge.out = lm.ridge(lpsa~.-train, data=prostate, lambda = 0:100)
# 2. Plot it
plot(ridge.out)
```



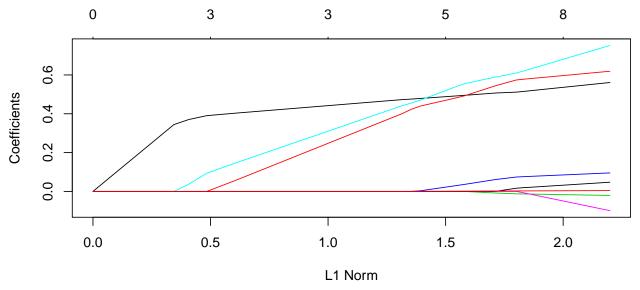
```
# (2a). If you chose lambda poorly, this will look bad, try again
# 3. Choose lambda using GCV
plot(ridge.out$lambda,ridge.out$GCV,ty='1')
```



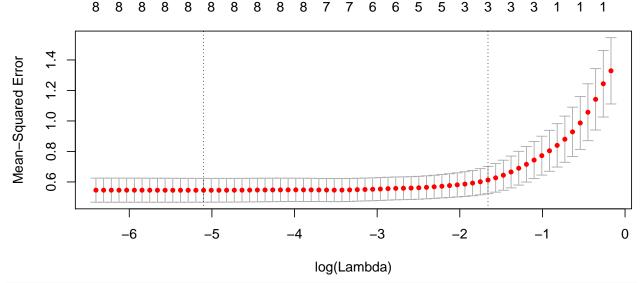
```
# 4. If there's a minimum, FIND IT, else try again
best.lam = which.min(ridge.out$GCV)
# 5. Return the coefs/predictions for the best model
coefs = coefficients(ridge.out)[best.lam,]
preds = as.matrix(dplyr::select(prostate,-lpsa,-train)) %*% coefs[-1] + coefs[1]
```

glmnet version (lasso or ridge)

```
# 1. Estimate cv and model at once, no formula version
lasso.glmnet = cv.glmnet(X,Y)
# 2. Plot the coefficient path
plot(lasso.glmnet$glmnet.fit) # the glmnet.fit == glmnet(X,Y)
```



3. Choose lambda using CV
plot(lasso.glmnet) #a different plot method for the cv fit



4. If the dashed lines are at the boundaries, redo it with a better set of lambda
best.lam = lasso.glmnet\$lambda.min # the value, not the location (or lasso\$lambda.1se)
5. Return the coefs/predictions for the best model
coefs.glmnet = coefficients(lasso.glmnet, s = best.lam)
preds.glmnet = predict(lasso.glmnet, newx = X, s = best.lam) # must supply `newx`

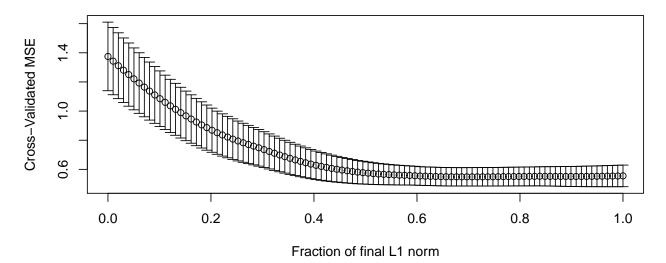
lars version

This is incredibly difficult to cross-validate.

The path changes from fold to fold, so things can get hairy.

In principle, the following should work.

```
# 1. Estimate cv, no formula version
lasso.lars.cv = cv.lars(X,Y) # also plots it
```



```
# 2. Choose lambda using CV
best.lam.lars = lasso.lars.cv$index[which.min(lasso.lars.cv$cv)] # the location, not the value
# 3. Estimate the lasso and plot
lasso.lars = lars(X,Y) # still the whole path
# 5. Return the coefs/predictions for the best model
coefs.lars = coefficients(lasso.lars, s = best.lam.lars, mode='fraction') # annoying mode argument is r
preds.lars = predict(lasso.lars, newx=X, s = best.lam.lars, mode='fraction') # must supply `newx`
```

Paths with chosen lambda (lasso and ridge)

```
ridge.glmnet = cv.glmnet(X,Y,alpha=0,lambda.min.ratio=1e-10) # added to get a minimum
par(mfrow=c(1,2))
plot(ridge.glmnet)
plot(lasso.glmnet)
```

