

ST308 - Lent term

Bayesian Inference

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Statistical Decision Theory - Bayes Estimators

Summary of previous lecture

- Frequentist and subjective probability
- Setup of Frequentist and Bayesian Inference
- Examples: Conjugate models

Outline

Topics covered: Loss function, Frequentist, Posterior and Bayes Risk, Bayes (decision) Rules, Minimax criterion, Point Estimation, Bayes Estimators

1 Statistical Decision Theory

2 Point Estimation

Outline

1 Statistical Decision Theory

2 Point Estimation

Statistical Inference

- Collect or consider the **data** $x = (x_1, \dots, x_n)$ from an experiment.
- Assign model-likelihood to real world problem with **uncertainty**.
- (In Bayesian Inference) Assign a **prior** to the parameters.
- Based on the above **decide** on
 - ▶ A best guess for θ - Point Estimation
 - ▶ A range of values for θ - Interval Estimation
 - ▶ Choosing between H_0 and H_1 - Hypothesis Testing
 - ▶ A best guess/range for a future value of x - Prediction

Statistical Decision Theory

Given observations x and model $f(x|\theta)$, a statistical **decision problem** consists of

- 1 The **parameter space** Θ .
- 2 A set \mathcal{A} of all possible **actions** a .
- 3 A **loss function** $L(a, \theta) : \mathcal{A} \times \Theta \rightarrow \mathcal{R}$, reflecting the loss for action a and true parameter value θ .

Notes:

- The sets \mathcal{A} , Θ could be finite or infinite.
- The negative of $L(a, \theta)$ is called utility function.

Decision Rule and Frequentist Risk

Decision Rule

The function $\delta(x) : \mathcal{R} \rightarrow \mathcal{A}$ that indicates the action a after observing $X = x$.

Every decision rule is associated with a random risk.

Frequentist Risk

$$R[\delta(x), \theta] = E_{X|\theta} (L(\delta(x), \theta)) = \int L(\delta(x), \theta) f(x|\theta) dx$$

Posterior Risk

$$\rho[\delta(x), \pi(\theta)] = E_{\theta|x} (L(\delta(x), \theta)) = \int L(\delta(x), \theta) \pi(\theta|x) d\theta$$

Using Frequentist Risk

How to **choose** between two decision rules, $\delta_1(x)$ and $\delta_2(x)$?

- 1 Choose a **risk function**, e.g. frequentist risk.
- 2 If $R[\delta_1(x), \theta] < R[\delta_2(x), \theta]$ for all $\theta \in \Theta$, then $\delta_1(x)$ is **uniformly better** than $\delta_2(x)$.

If there exists $\delta^*(x)$ such that $R[\delta^*(x), \theta] \leq R[\delta(x), \theta]$ for all $\delta(x) \in \mathcal{D}$ and all $\theta \in \Theta$, then $\delta^*(x)$ is an **admissible** decision rule.

Issue: It is usually very difficult to minimise frequentist risk for **all** $\theta \in \Theta$.

The Minimax criterion

Minimax criterion

A **minimax** estimator $\delta^M(x)$ satisfies

$$\max_{\theta \in \Theta} R[\delta^M(x), \theta] = \min_{\delta \in \mathcal{D}} \left[\max_{\theta \in \Theta} R[\delta(x), \theta] \right]$$

Notes:

- Optimisation under worst case scenario - too conservative.
- Very difficult to find.
- Does not use prior information.
- inf and sup may also be used.

Bayes Risk criterion

Bayes Risk

Given the prior $\pi(\theta)$ and $\delta(x)$, **Bayes risk** is the function

$$r[\delta(x), \pi(\theta)] = E_{\theta} [R[\delta(x), \theta]] = \int R[\delta(x), \theta] \pi(\theta) d\theta$$

The decision rule $\delta^B(x)$ that minimises the Bayes risk is called **Bayes rule**.

Notes on Bayes Risk criterion

- Bayes risk is a **number**. Hence, given a loss function $L(a, \theta)$, $f(x|\theta)$ and $\pi(\theta)$ there exists an **optimal solution** to the statistical decision problem.
- Bayes risk can also be written as an expectation of the **posterior risk** wrt X

$$r[\delta(x), \pi(\theta)] = E_X [\rho[\delta(x), \pi(\theta)]] = \int \rho[\delta(x), \pi(\theta)] m(x) dx$$

where $m(x)$ is the **marginal likelihood**. It unifies the two approaches based on the frequentist and posterior risks.

Example: Vaccination problem

A public health organisation considers **vaccination** to prevent a disease. A test to determine immunity exists. Let X denote the test outcome (x_1 : positive, x_2 : negative) and θ whether the person is immune to the disease (θ_1 : immune, θ_2 : susceptible). The **likelihood** is given below

$f(x \theta)$	x_1	x_2
θ_1	0.65	0.35
θ_2	0.25	0.75

Example: Vaccination problem (cont'd)

Consider the **actions** regarding vaccination (a_1 : yes, a_2 : no). The **loss function** is

$L(a, \theta)$	a_1	a_2
θ_1	8	0
θ_2	0	20

Which of the 4 strategies (decision rules) is **better**?

- 1 Vaccinate **everyone** $\delta_1(x_1) = \delta_1(x_2) = a_1$
- 2 Vaccinate **positives** $\delta_2(x_1) = a_1, \delta_2(x_2) = a_2$
- 3 Vaccinate **negatives** $\delta_3(x_1) = a_2, \delta_3(x_2) = a_1$
- 4 **Don't vaccinate** anyone $\delta_4(x_1) = \delta_4(x_2) = a_2$

Example: Vaccination problem (cont'd)

$$\begin{aligned} R[\delta_1(x), \theta_1] &= E_{X|\theta}(L(\delta_1(x), \theta_1)) = \sum_{i=1}^2 L(\delta_1(x_i), \theta_1) f(x_i|\theta_1) \\ &= L(a_1, \theta_1) f(x_1|\theta_1) + L(a_1, \theta_1) f(x_2|\theta_1) = \dots = 8 \end{aligned}$$

Similarly we can get

$R(a, \theta)$	$\delta_1(x)$	$\delta_2(x)$	$\delta_3(x)$	$\delta_4(x)$
θ_1	8	5.2	2.8	0
θ_2	0	15	5	20

No admissible (optimal) strategy (decision rule).

Example: Vaccination problem (cont'd)

Suppose that the **prior** is $\pi(\theta_1) = 0.6$ and $\pi(\theta_2) = 0.4$.

$$\begin{aligned}r[\delta_1(x), \pi(\theta)] &= E_{\theta}(R[\delta_1(x), \theta]) = \sum_{i=1}^2 R[\delta_1(x), \theta_i] \pi(\theta_i) \\&= R[\delta_1(x), \theta_1] \pi(\theta_1) + R[\delta_1(x), \theta_2] \pi(\theta_2) = \dots = 4.8\end{aligned}$$

Similarly we can get $r[\delta_2(x), \pi(\theta)] = 9.12$, $r[\delta_3(x), \pi(\theta)] = 3.68$ and $r[\delta_4(x), \pi(\theta)] = 8$.

Hence, the **optimal strategy** (decision rule) is $\delta_3(x)$, to vaccinate those who are tested negative for immunity.

Outline

1 Statistical Decision Theory

2 Point Estimation

Point Estimation problem

- Collect or consider the **data** $x = (x_1, \dots, x_n)$ from an experiment.
- Assign model-likelihood to real world problem with **uncertainty**.
- (In Bayesian Inference) Assign a **prior** to the parameters.
- Based on the above decide on a **best guess** for θ - Point Estimation.

Decision theory elements

- **Action:** report a value for the θ , action set $\mathcal{A} = \Theta$.
- **Decision rule:** $\delta(x)$ is an estimator also denoted with $\hat{\theta}$, e.g. \bar{x} , $\frac{1}{n} \sum_i x_i^2$ etc.
- **Loss function:** absolute error, quadratic error, 0-1 loss etc.

Note: In the case of quadratic error loss function $L(\hat{\theta}, \theta) = (\hat{\theta} - \theta)^2$, frequentist risk is the **mean squared error (MSE)**

$$R(\hat{\theta}, \theta) = E_{X|\theta}(\hat{\theta} - \theta)^2 = \text{MSE}(\hat{\theta})$$

Example: Partial ordering with frequentist risk

Let $x = (x_1, \dots, x_n)$ be a random sample from $N(\theta, \sigma^2)$. Consider the **quadratic loss** function, i.e. the MSE as frequentist risk.

Consider $\delta_1(x) = \bar{x}$, the minimum variance unbiased estimator for this problem, and the naive estimator $\delta_2(x) = 100$.

We get $R[\delta_1(x), \theta] = \frac{1}{n}\sigma^2$, $R[\delta_2(x), \theta] = (100 - \theta)^2$.

Note that $\delta_1(x)$ is **not better** than $\delta_2(x)$ for all $\theta \in \mathcal{R}$ in terms of **MSE!**

Bayes Estimators

Theorem (Construction of Bayes estimators)

Bayes estimators minimise the Bayes risk $r[\delta(x), \pi(\theta)]$

This can be achieved if for every $x \in \mathcal{X}$ we select the value $\delta(x)$ that minimises the posterior risk $\rho[\delta(x), \pi(\theta)]$, since

$$r[\delta(x), \pi(\theta)] = \int \rho[\delta(x), \pi(\theta)] m(x) dx$$

Minimising the posterior risk is the **same** as minimising the Bayes risk.

Quadratic error loss function

Suppose that we have $L(a, \theta) = (a - \theta)^2$. The **Bayes estimator** minimises the posterior risk

$$\rho[\alpha, \pi(\theta)] = \int (a - \theta)^2 \pi(\theta|x) d\theta.$$

We can write

$$\begin{aligned} \frac{\partial \rho[\alpha, \pi(\theta)]}{\partial a} &= \int \frac{\partial}{\partial a} (a - \theta)^2 \pi(\theta|x) d\theta = \int 2(a - \theta) \pi(\theta|x) d\theta \\ &= 2 \left[a \int \pi(\theta|x) d\theta - \int \theta \pi(\theta|x) d\theta \right] = 2[a - E(\theta|x)] \end{aligned}$$

Setting $\frac{\partial \rho[\alpha, \pi(\theta)]}{\partial a} = 0$ gives $a = E(\theta|x)$. Also $\frac{\partial^2}{\partial a^2} \rho[\alpha, \pi(\theta)] = 2 > 0$.

Quadratic error loss \longrightarrow Bayes estimator is the **posterior mean**.

Linear error loss function

Theorem: Assume that for positive k_0, k_1 the **loss function** is

$$L(a, \theta) = \begin{cases} k_0(a - \theta) & \text{if } a > \theta \\ k_1(\theta - a) & \text{if } a \leq \theta \end{cases}$$

The **Bayes estimator** is the $\frac{k_1}{k_0+k_1}$ -th **percentile** of $\pi(\theta|x)$ denoted by q .

$$\frac{k_1}{k_0 + k_1} = \pi(\theta \leq q|x) = \int_{-\infty}^q \pi(\theta|x) d\theta$$

Proof: We will show that $E_{\theta|x}(L(q, \theta)) \leq E_{\theta|x}(L(a, \theta))$. Assume $q < a$.

Linear error loss function (cont'd)

If $\theta \leq q < a$:

$$L(q, \theta) - L(a, \theta) = k_0(q - \theta) - k_0(a - \theta) = k_0(q - a)$$

If $q < a < \theta$:

$$L(q, \theta) - L(a, \theta) = k_1(\theta - q) - k_1(\theta - a) = k_1(a - q)$$

If $q < \theta < a$:

$$\begin{aligned} L(q, \theta) - L(a, \theta) &= k_1(\theta - q) - k_0(a - \theta) = k_1(\theta - q) + k_0(\theta - a) \\ &< k_1(\theta - q) < k_1(a - q) \end{aligned}$$

So putting everything **together** we get

$$L(q, \theta) - L(a, \theta) \leq \begin{cases} k_0(q - a) & \text{if } \theta \leq q \\ k_1(a - q) & \text{if } q < \theta \end{cases}$$

Linear error loss function (cont'd)

Taking **expectation** wrt to the posterior yields

$$\begin{aligned} E_{\theta|x}(L(q, \theta) - L(a, \theta)) &\leq k_0(q - a)\pi(\theta \leq q|x) + k_1(a - q)\pi(\theta > q|x) \\ &= k_0(q - a)\frac{k_1}{k_0 + k_1} + k_1(a - q)\left(1 - \frac{k_1}{k_0 + k_1}\right) = \dots = 0 \end{aligned}$$

So $E_{\theta|x}(L(q, \theta)) \leq E_{\theta|x}(L(a, \theta))$, i.e. **q minimises the posterior risk.** \square

Special case: For $k_0 = k_1 = 1$ we get the absolute error loss function

$$L(a, \theta) = |a - \theta|$$

Hence the $1/(1 + 1) = 0.5$ -percentile, or else the **posterior median** minimises the posterior risk and therefore is the **Bayes estimator**.

0 – 1 loss function

Finally consider the **0 – 1 loss** function

$$L(a, \theta) = \begin{cases} 0 & \text{if } |a - \theta| \leq \epsilon \\ 1 & \text{if } |a - \theta| > \epsilon \end{cases}$$

The posterior risk is the probability

$$\pi(|a - \theta| > \epsilon | x)$$

and is **minimised** when the following probability is **maximised**

$$\pi(|a - \theta| \leq \epsilon | x) = \pi(a - \epsilon < \theta < a + \epsilon | x)$$

This occurs at the **posterior mode** of $\pi(\theta | x)$ (draw a graph to check it).

Facts about Bayes Estimators

- Bayes estimators are also **minimax** estimators. But their risk (Bayes risk) is **smaller**.
- Bayes estimators are typically **admissible** estimators.
- For improper priors, Bayes estimators may not exist. If they do, they are called **generalised** Bayes estimators.
- Bayes estimators are **biased**.
- Like maximum likelihood estimators, Bayes estimators are **asymptotically unbiased and efficient and normally distributed**.
- **Famous examples:** **Lasso** and **Ridge** Regression estimators.
More in week 5

Reading

J.O. Berger:

Sections 1.3 1.5 2.4.1 2.4.2 4.3.1 4.4.1 and 4.4.2