# Section 2.2: Covariance, Correlation, and Least Squares

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Suggested reading: OpenIntro Statistics, Chapter 7.1, 7.2

# A Deeper Look at Least Squares Estimates

Last time we saw that least squares estimates had some special properties:

- ▶ The fitted values  $\hat{Y}$  and x were **very** dependent
- ▶ The residuals  $Y \hat{Y}$  and x had no apparent relationship
- ▶ The residuals  $Y \hat{Y}$  had a sample mean of zero

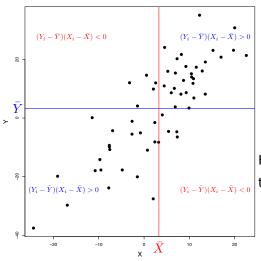
What's going on? And what exactly are the least squares estimates?

We need to review sample covariance and correlation

#### Covariance

Measure the *direction* and *strength* of the linear relationship between Y and X

$$Cov(Y,X) = \frac{\sum_{i=1}^{n} (Y_i - \overline{Y})(X_i - \overline{X})}{n-1}$$



- $s_y = 15.98,$   $s_x = 9.7$
- $\triangleright$  *Cov*(*X*, *Y*) = 125.9

How do we interpret that?

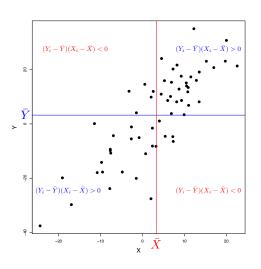
Correlation is the standardized covariance:

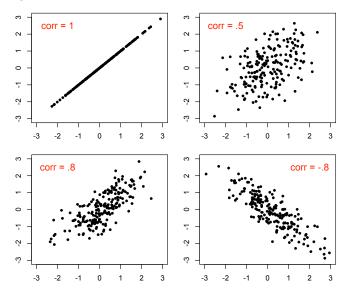
$$\operatorname{corr}(X,Y) = \frac{\operatorname{cov}(X,Y)}{\sqrt{s_X^2 s_y^2}} = \frac{\operatorname{cov}(X,Y)}{s_X s_y}$$

The correlation is scale invariant and the units of measurement don't matter: It is always true that  $-1 \le \operatorname{corr}(X, Y) \le 1$ .

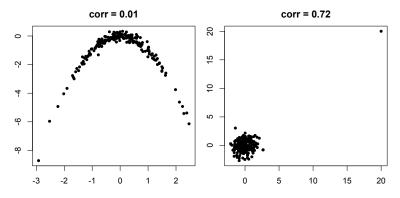
This gives the direction (- or +) and strength  $(0 \rightarrow 1)$  of the linear relationship between X and Y.

$$corr(Y,X) = \frac{cov(X,Y)}{\sqrt{s_X^2 s_y^2}} = \frac{cov(X,Y)}{s_X s_y} = \frac{125.9}{15.98 \times 9.7} = 0.812$$





Only measures linear relationships: corr(X, Y) = 0 does not mean the variables are not related!



Also be careful with influential observations...

# The Least Squares Estimates

The values for  $b_0$  and  $b_1$  that minimize the least squares criterion are:

$$b_1 = r_{xy} \times \frac{s_y}{s_x}$$
  $b_0 = \bar{Y} - b_1 \bar{X}$ 

where,

- $ightharpoonup \bar{X}$  and  $\bar{Y}$  are the sample mean of X and Y
- ightharpoonup correlation
- $ightharpoonup s_x$  and  $s_y$  are the sample standard deviation of X and Y

These are the **least squares estimates** of  $\beta_0$  and  $\beta_1$ ; they minimize the sum of squared residuals.

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How do we interpret these?

- ▶  $b_0$  ensures the line goes through  $(\bar{x}, \bar{y})$
- b<sub>1</sub> scales the correlation to appropriate units by multiplying with s<sub>y</sub>/s<sub>x</sub> (what are the units of b<sub>1</sub>?)

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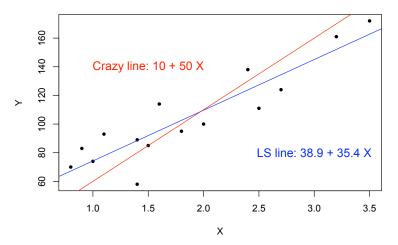
## Properties of Least Squares Estimates

### Remember from the housing data, we had:

- $corr(\hat{Y}, x) = 1$  (a perfect linear relationship)
- ightharpoonup corr(e,x)=0 (no linear relationship)
- ightharpoonup mean(e) = 0 (sample average of residuals is zero)

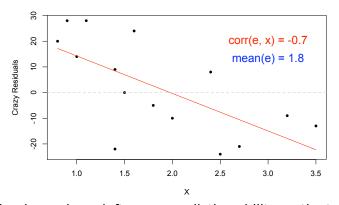
# Why?

What is the intuition for the relationship between  $\hat{Y}$  and e and X? Lets consider some "crazy" alternative line:



### Fitted Values and Residuals

This is a bad fit! We are underestimating the value of small houses and overestimating the value of big houses.



Clearly, we have left some predictive ability on the table!

# Summary: LS is the best we can do!!

As long as the correlation between e and X is non-zero, we could always adjust our prediction rule to do better.

We need to exploit all of the predictive power in the X values and put this into  $\hat{Y}$ , leaving no "Xness" in the residuals.

# In Summary: $Y = \hat{Y} + e$ where:

- $ightharpoonup \hat{Y}$  is "made from X";  $\operatorname{corr}(X, \hat{Y}) = 1$ .
- e is unrelated to X; corr(X, e) = 0.
- ▶ On average, our prediction error is zero:  $\bar{\mathbf{e}} = \sum_{i=1}^{n} \mathbf{e}_i = \mathbf{0}$ .

# Decomposing the Variance

How well does the least squares line explain variation in Y?

Remember that  $Y = \hat{Y} + e$ 

Since  $\hat{Y}$  and e are uncorrelated, i.e.  $corr(\hat{Y}, e) = 0$ ,

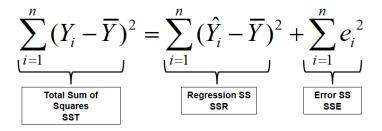
$$var(Y) = var(\hat{Y} + e) = var(\hat{Y}) + var(e)$$

$$\frac{\sum_{i=1}^{n} (Y_i - \bar{Y})^2}{n-1} = \frac{\sum_{i=1}^{n} (\hat{Y}_i - \bar{\hat{Y}})^2}{n-1} + \frac{\sum_{i=1}^{n} (e_i - \bar{e})^2}{n-1}$$

Given that  $\bar{e}=0$ , and the sample mean of the fitted values  $\bar{\hat{Y}}=\bar{Y}$  (why?) we get to write:

$$\sum_{i=1}^{n} (Y_i - \bar{Y})^2 = \sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^{n} e_i^2$$

# Decomposing the Variance



SSR: Variation in *Y* explained by the regression line.

SSE: Variation in *Y* that is left unexplained.

$$SSR = SST \Rightarrow perfect fit.$$

Be careful of similar acronyms; e.g. SSR for "residual" SS.

## The Coefficient of Determination $R^2$

The coefficient of determination, denoted by  $R^2$ , measures how well the fitted values  $\hat{Y}$  follow Y:

$$R^2 = \frac{\text{SSR}}{\text{SST}} = 1 - \frac{\text{SSE}}{\text{SST}}$$

- ▶  $R^2$  is the proportion of variance in Y that is "explained" by the regression line (in the mathematical not scientific sense!):  $R^2 = 1 Var(e)/Var(Y)$
- $ightharpoonup 0 < R^2 < 1$
- For simple linear regression,  $R^2 = r_{xy}^2$ . Similar caveats to sample correlation apply!

# R<sup>2</sup> for the Housing Data

```
summary(fit)
##
## Call:
## lm(formula = Price ~ Size, data = housing)
##
## Residuals:
##
      Min 10 Median 30 Max
## -30.425 -8.618 0.575 10.766 18.498
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 38.885 9.094 4.276 0.000903 ***
## Size 35.386 4.494 7.874 2.66e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 14.14 on 13 degrees of freedom
## Multiple R-squared: 0.8267, Adjusted R-squared: 0.8133
## F-statistic: 62 on 1 and 13 DF, p-value: 2.66e-06
```

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# R<sup>2</sup> for the Housing Data

```
anova(fit)
## Analysis of Variance Table
##
## Response: Price
##
            Df Sum Sq Mean Sq F value Pr(>F)
## Size 1 12393.1 12393.1 61.998 2.66e-06 ***
## Residuals 13 2598.6 199.9
## - - -
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0
```

$$R^2 = \frac{SSR}{SST} = \frac{12393.1}{2598.6 + 12393.1} = 0.8267$$

## Back to Baseball

Three very similar, related ways to look at a simple linear regression... with only one *X* variable, life is easy!

	R <sup>2</sup>	corr	SSE
OBP	0.88	0.94	0.79
SLG	0.76	0.87	1.64
AVG	0.63	0.79	2.49