

Section 2.2: Covariance, Correlation, and Least Squares

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Suggested reading: OpenIntro Statistics, Chapter 7.1, 7.2

A Deeper Look at Least Squares Estimates

Last time we saw that least squares estimates had some special properties:

- ▶ The fitted values \hat{Y} and x were **very** dependent
- ▶ The residuals $Y - \hat{Y}$ and x had no apparent relationship
- ▶ The residuals $Y - \hat{Y}$ had a sample mean of zero

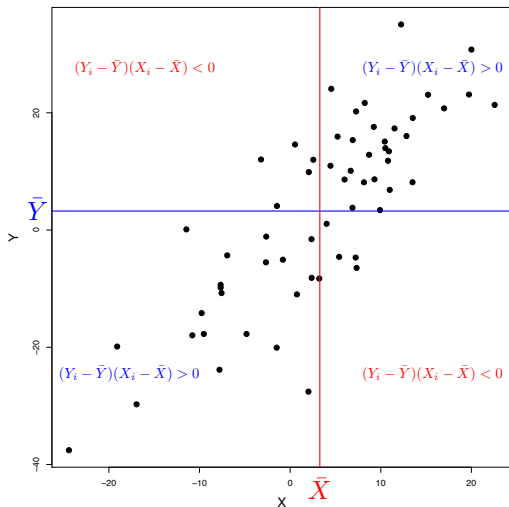
What's going on? And what exactly are the least squares estimates?

We need to review **sample covariance** and **correlation**

Covariance

Measure the *direction* and *strength* of the linear relationship between Y and X

$$\text{Cov}(Y, X) = \frac{\sum_{i=1}^n (Y_i - \bar{Y})(X_i - \bar{X})}{n - 1}$$



- ▶ $s_y = 15.98,$
 $s_x = 9.7$
- ▶ $\text{Cov}(X, Y) = 125.9$

How do we interpret that?

Correlation

Correlation is the standardized covariance:

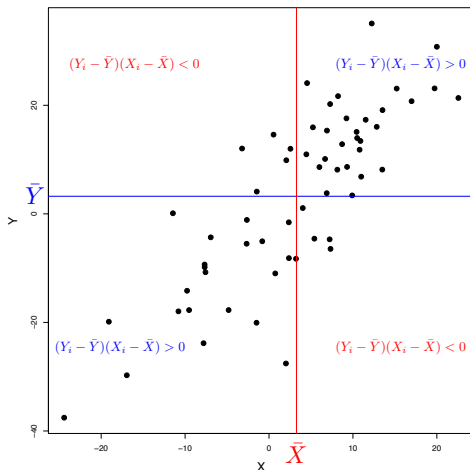
$$\text{corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{s_x^2 s_y^2}} = \frac{\text{cov}(X, Y)}{s_x s_y}$$

The correlation is scale invariant and the units of measurement don't matter: **It is always true that $-1 \leq \text{corr}(X, Y) \leq 1$.**

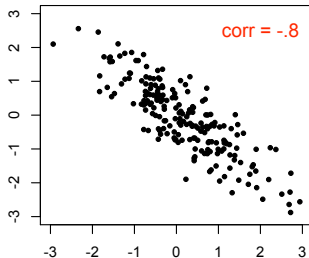
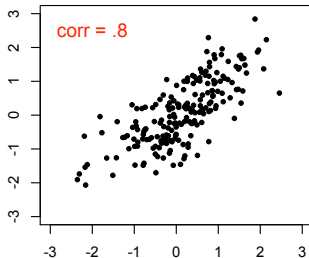
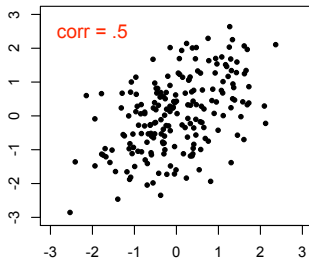
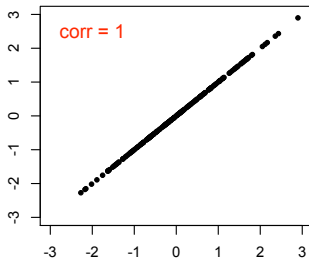
This gives the direction (- or +) and strength ($0 \rightarrow 1$) of the **linear** relationship between X and Y .

Correlation

$$\text{corr}(Y, X) = \frac{\text{cov}(X, Y)}{\sqrt{s_x^2 s_y^2}} = \frac{\text{cov}(X, Y)}{s_x s_y} = \frac{125.9}{15.98 \times 9.7} = 0.812$$



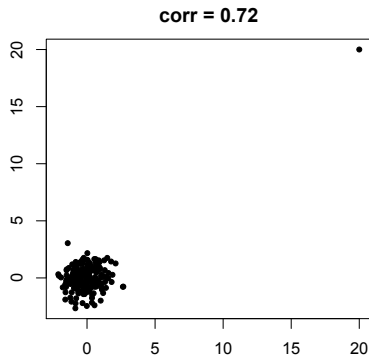
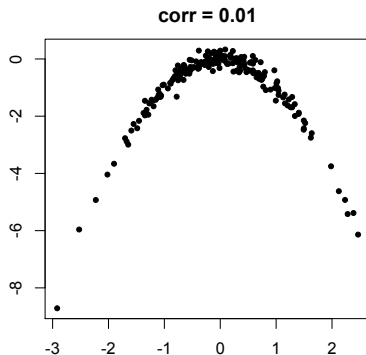
Correlation



Correlation

Only measures **linear** relationships:

$\text{corr}(X, Y) = 0$ does not mean the variables are not related!



Also be careful with influential observations...

The Least Squares Estimates

The values for b_0 and b_1 that minimize the least squares criterion are:

$$b_1 = r_{xy} \times \frac{s_y}{s_x} \quad b_0 = \bar{Y} - b_1\bar{X}$$

where,

- ▶ \bar{X} and \bar{Y} are the sample mean of X and Y
- ▶ $\text{corr}(x, y) = r_{xy}$ is the sample correlation
- ▶ s_x and s_y are the sample standard deviation of X and Y

These are the **least squares estimates** of β_0 and β_1 .

The Least Squares Estimates

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$$b_1 = r_{xy} \times \frac{s_y}{s_x} \quad b_0 = \bar{Y} - b_1 \bar{X}$$

How do we interpret these?

- ▶ b_0 ensures the line goes through (\bar{x}, \bar{y})
- ▶ b_1 scales the correlation to appropriate units by multiplying with s_y/s_x (what are the units of b_1 ?)

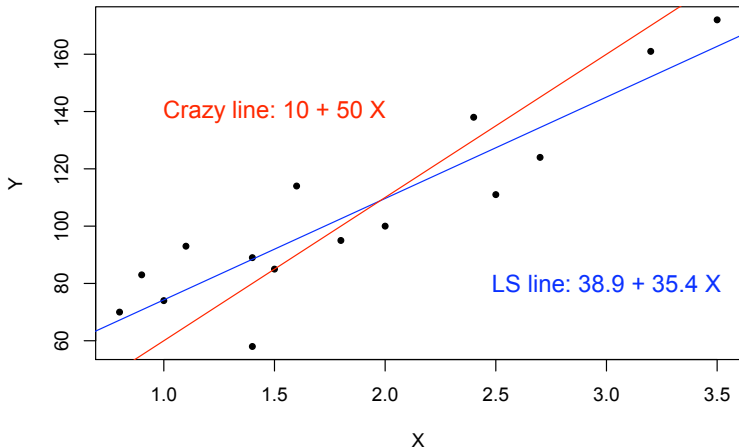
Properties of Least Squares Estimates

Remember from the housing data, we had:

- ▶ $\text{corr}(\hat{Y}, x) = 1$ (a perfect linear relationship)
- ▶ $\text{corr}(e, x) = 0$ (no linear relationship)
- ▶ $\text{mean}(e) = 0$ (sample average of residuals is zero)

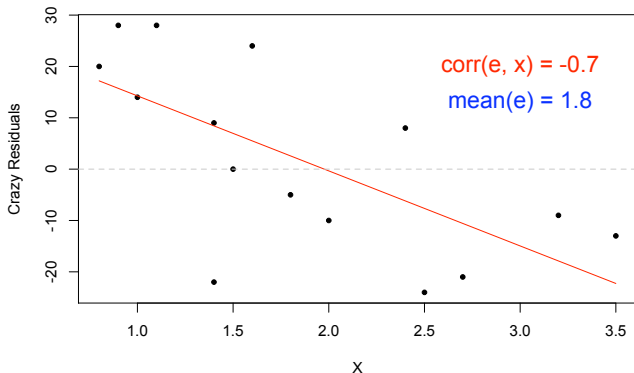
Why?

What is the intuition for the relationship between \hat{Y} and e and X ? Lets consider some “crazy” alternative line:



Fitted Values and Residuals

This is a bad fit! We are underestimating the value of small houses and overestimating the value of big houses.



Clearly, we have left some predictive ability on the table!

Summary: LS is the best we can do!!

As long as the correlation between e and X is non-zero, we could always adjust our prediction rule to do better.

We need to exploit all of the predictive power in the X values and put this into \hat{Y} , leaving no “ X ness” in the residuals.

In Summary: $Y = \hat{Y} + e$ where:

- ▶ \hat{Y} is “made from X ”; $\text{corr}(X, \hat{Y}) = 1$.
- ▶ e is unrelated to X ; $\text{corr}(X, e) = 0$.
- ▶ On average, our prediction error is zero: $\bar{e} = \sum_{i=1}^n e_i = 0$.

Decomposing the Variance

How well does the least squares line explain variation in Y ?

Remember that $Y = \hat{Y} + e$

Since \hat{Y} and e are uncorrelated, i.e. $\text{corr}(\hat{Y}, e) = 0$,

$$\begin{aligned}\text{var}(Y) &= \text{var}(\hat{Y} + e) = \text{var}(\hat{Y}) + \text{var}(e) \\ \frac{\sum_{i=1}^n (Y_i - \bar{Y})^2}{n-1} &= \frac{\sum_{i=1}^n (\hat{Y}_i - \bar{\hat{Y}})^2}{n-1} + \frac{\sum_{i=1}^n (e_i - \bar{e})^2}{n-1}\end{aligned}$$

Given that $\bar{e} = 0$, and the sample mean of the fitted values $\bar{\hat{Y}} = \bar{Y}$ (why?) we get to write:

$$\sum_{i=1}^n (Y_i - \bar{Y})^2 = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^n e_i^2$$

Decomposing the Variance

$$\underbrace{\sum_{i=1}^n (Y_i - \bar{Y})^2}_{\substack{\text{Total Sum of} \\ \text{Squares} \\ \text{SST}}} = \underbrace{\sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2}_{\substack{\text{Regression SS} \\ \text{SSR}}} + \underbrace{\sum_{i=1}^n e_i^2}_{\substack{\text{Error SS} \\ \text{SSE}}}$$

SSR: Variation in Y explained by the regression line.

SSE: Variation in Y that is left unexplained.

$\text{SSR} = \text{SST} \Rightarrow$ perfect fit.

Be careful of similar acronyms; e.g. SSR for “residual” SS.

The Coefficient of Determination R^2

The **coefficient of determination**, denoted by R^2 , measures how well the fitted values \hat{Y} follow Y :

$$R^2 = \frac{SSR}{SST} = 1 - \frac{SSE}{SST}$$

- ▶ R^2 is the proportion of variance in Y that is “explained” by the regression line (in the mathematical – not scientific – sense!): $R^2 = 1 - \text{Var}(e)/\text{Var}(Y)$
- ▶ $0 < R^2 < 1$
- ▶ For simple linear regression, $R^2 = r_{xy}^2$. Similar caveats to sample correlation apply!

R^2 for the Housing Data

```
summary(fit)

##
## Call:
## lm(formula = Price ~ Size, data = housing)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -30.425  -8.618   0.575  10.766  18.498
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   38.885      9.094   4.276 0.000903 ***
## Size          35.386      4.494   7.874 2.66e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 14.14 on 13 degrees of freedom
## Multiple R-squared:  0.8267, Adjusted R-squared:  0.8133
## F-statistic:    62 on 1 and 13 DF,  p-value: 2.66e-06
```

R^2 for the Housing Data

```
anova(fit)

## Analysis of Variance Table
##
## Response: Price
##              Df Sum Sq Mean Sq F value    Pr(>F)
## Size           1 12393.1  12393.1   61.998 2.66e-06 ***
## Residuals     13  2598.6    199.9
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0
```

$$R^2 = \frac{SSR}{SST} = \frac{12393.1}{2598.6 + 12393.1} = 0.8267$$

Back to Baseball

Three very similar, related ways to look at a simple linear regression... with only one X variable, life is easy!

	R^2	corr	SSE
OBP	0.88	0.94	0.79
SLG	0.76	0.87	1.64
AVG	0.63	0.79	2.49