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STATS 244

HOMEWORK 3

Rice Chapter 4, Question 78

Show that if a density is symmetric about zero, its skewness is zero.

SKewNESS can be determined using $E(X^3)$. For example if $E(X^3)$ is zero, then skewness is zero.

Let f_Y be a density symmetric around zero. Then, because of this symmetry $-f_Y = f_Y$. Which implies $E(Y^3) = E(-Y^3)$ Which means that $E(Y^3) = 0$.

Question 2

Consider the bivariate density of X and Y

$$f(x, y) = 4(x + y + xy)/5 \text{ for } 0 < x, y < 1, 0 \text{ otherwise}$$

1. Verify that this is a bivariate density
(the total volume of $\int \int f(x, y) dx dy = 1$).

$$\begin{aligned} \int \int f(x, y) dx dy &= \int_0^1 \int_0^1 \frac{4}{5} (x + y + xy) dx dy \\ &= \int_0^1 \left[\frac{4}{5} \left(\int_0^1 x + y + xy dx \right) \right] dy \\ &= \int_0^1 \left[\frac{4}{5} \left(\int_0^1 x dx + \int_0^1 y dx + \int_0^1 xy dx \right) \right] dy \\ &= \int_0^1 \left[\frac{4}{5} \left(\frac{x^2}{2} \Big|_0^1 + y + y \left(\frac{x^2}{2} \Big|_0^1 \right) \right) \right] dy \\ &= \int_0^1 \left[\frac{4}{5} \left(\frac{1}{2} + 1 \frac{y}{2} \right) \right] dy \\ &= \int_0^1 \frac{4}{5} \left(\frac{3y}{2} + \frac{1}{2} \right) dy \\ &= \int_0^1 \frac{4}{10} (3y + 1) dy \\ &= \frac{4}{10} \left[\int_0^1 3y dy + \int_0^1 1 dy \right] \\ &= \frac{4}{10} \left[3 \left(\frac{1}{2} \right) + 1 \right] \\ &= \frac{4}{10} \left[\frac{3}{2} + 1 \right] \\ &= 1 \end{aligned}$$

2. Find the marginal density of Y

We basically found this in a step from part 1

$$\begin{aligned}
 f_Y(y) &= \int_0^1 f_{XY}(x, y) dx \\
 &= \int_0^1 \frac{4}{5} (x + y + xy) dx \\
 &= \frac{4}{5} \int_0^1 (x + y + xy) dx \\
 &= \frac{4}{10} (3y + 1)
 \end{aligned}$$

3. Find the conditional density of X given $Y = 0.5$.

$$\begin{aligned}
 f_{X|Y}(x | Y = .5) &= \frac{f(x, .5)}{f_Y(.5)} \\
 &= \frac{\frac{4}{5} (x + .5 + .5x)}{\frac{4}{10} (2.5)} \\
 &= \frac{4}{5} (1.5x + .5)
 \end{aligned}$$

4. Find $E(X)$, $E(X^2)$, $\text{Var}(X)$, $E(XY)$, $\text{Cov}(X, Y)$.

$$\begin{aligned}
 E(X) &= \int_0^1 x f_X(x) dx \\
 &= \int_0^1 x \frac{4}{10} (3x + 1) dx \\
 &= \frac{4}{10} \left[\int_0^1 3x^2 + x dx \right] \\
 &= \frac{4}{10} \left[\int_0^1 3x^2 dx + \int_0^1 x dx \right] \\
 &= \frac{4}{10} \left[1 + \frac{1}{2} \right] \\
 &= \frac{3}{5}
 \end{aligned}$$

$$\begin{aligned}
E(X^2) &= \int_0^1 x^2 f_X(x) dx \\
&= \int_0^1 x^2 \frac{4}{10} (3x + 1) dx \\
&= \frac{4}{10} \int_0^1 3x^3 + x^2 dx \\
&= \frac{4}{10} \left[\frac{3}{4} x^4 \Big|_0^1 + \frac{x^3}{3} \Big|_0^1 \right] \\
&= \frac{4}{10} \left(\frac{3}{4} + \frac{1}{3} \right) \\
&= \frac{13}{30}
\end{aligned}$$

$$\begin{aligned}
\text{Var}(X) &= E(X^2) - E(X)^2 \\
&= \frac{13}{30} - \frac{3^2}{5} \\
&= \frac{11}{150}
\end{aligned}$$

$$\begin{aligned}
E(XY) &= \int_0^1 \int_0^1 xy f(x, y) dx dy \\
&= \int_0^1 \int_0^1 xy \frac{4}{5} (x + y + xy) dx dy \\
&= \int_0^1 \left[\frac{4}{5} \int_0^1 xy (x + y + xy) dx \right] dy \\
&= \int_0^1 \left[\frac{4}{5} y \left(\int_0^1 (x^2 + xy + x^2 y) dx \right) \right] dy \\
&= \int_0^1 \left[\frac{4}{5} y \left(\frac{1}{3} + y \left(\frac{1}{2} \right) + y \left(\frac{1}{3} \right) \right) \right] dy \\
&= \int_0^1 \frac{4}{15} y \left(\frac{5y}{2} + 1 \right) dy \\
&= \frac{4}{15} \int_0^1 \frac{5y^2}{2} + y dy \\
&= \frac{4}{15} \left[\frac{5}{6} + \frac{1}{2} \right] \\
&= \frac{16}{45}
\end{aligned}$$

We should note that because the marginal density of X and Y are symmetric(?) $E(X) = E(Y) = \frac{3}{5}$. In anycase, we don't need to compute $E(Y)$ since the marginal densities look the same.

$$\begin{aligned}
\text{Cov}(X, Y) &= E(XY) - E(X)E(Y) \\
&= \frac{16}{45} - \frac{3^2}{5} \\
&= -\frac{1}{225}
\end{aligned}$$

5. Find $P(0.2 \leq X \leq .5 \text{ and } .4 \leq Y \leq .8)$

$$\begin{aligned}
P(0.2 \leq X \leq .5 \text{ and } .4 \leq Y \leq .8) &= \int_{.4}^{.8} \int_{.2}^{.5} f(x, y) dx dy \\
&= \int_{.4}^{.8} \int_{.2}^{.5} \frac{4}{5} (x + y + xy) dx dy \\
&= \int_{.4}^{.8} \left[\frac{4}{5} \int_{.2}^{.5} x + y + xy dx \right] dy \\
&= \int_{.4}^{.8} \left[\frac{4}{5} (.405y + .105) \right] dy \\
&= \frac{4}{5} (.0972 + .042) \\
&= .11136
\end{aligned}$$

I asked for a clarification on this question on the discussion board and no one answered! The CPR rule takes effect, which says if there is no response, the affirmative is implied. It's a gamble, but whatever. There are many cases where I would NOT employ this rule.

6. Find $P(X + Y \leq 1)$ Set $y = v - x$

$$\begin{aligned}
P(X + Y \leq 1) &= \int_0^y \int_0^x f(x, v) dx dv \\
&= \int_0^1 \int_0^{1-x} f(x, y) dy dx \\
&= \int_0^1 \int_0^{1-x} \frac{4}{5} (x + y + xy) dy dx \\
&= \int_0^1 \left[\frac{4}{5} \left(\int_0^{1-x} x dy + \int_0^{1-x} y dy + x \int_0^{1-x} y dy \right) \right] dx \\
&= \int_0^1 \frac{4}{5} \left[-(1-x)x + \frac{1}{2}(x-1)^2 + \frac{1}{2}(x-1)^2 x \right] dx \\
&= \int_0^1 \frac{4}{10} (x^3 - 3x^2 + x + 1) dx \\
&= \frac{4}{10} \left[\frac{1}{4} - 1 + \frac{1}{2} + 1 \right] \\
&= \frac{3}{10}
\end{aligned}$$

Question 3, Rice 4.81 and 4.82

- Find the moment-generating function of a Bernoulli random variable, and use it to find the mean, variance, and third moment.

A BERNOULLI random variable is one such that $f(x) = 1 - p$ where $f(0) = 1 - p$ and $f(1) = p$.

$$\begin{aligned} M(t) &= \sum e^{tx} f(x) \\ &= e^{t(0)} f(0) + e^{t(1)} f(1) \\ &= e^{t(0)} (1 - p) + e^{t(1)} (p) \\ &= 1 - p + e^t p \end{aligned}$$

To find the first, second and third moments, take the following derivative of $M(t)$.

$$\begin{aligned} M'(t) &= pe^t \\ M''(t) &= pe^t \\ M'''(t) &= pe^t \end{aligned}$$

Evaluating each of these at 0 gives us our moments, respectively.

$$\begin{aligned} E(X) &= p \\ E(X^2) &= p \\ E(X^3) &= p \end{aligned}$$

We also need find the variance.

$$\text{Var}(X) = E(X^2) - E(X)^2 = p - p^2$$

If we let $q = 1 - p$ then $\text{Var}(X) = p(1 - p) = pq$

2. Use the result of Problem 81 to find the mgf of a binomial random variable and its mean and variance.

LET $X_1, X_2, X_3, \dots, X_n$ be independently and identically distributed Bernoulli random variables with parameter p .

Let $Y = X_1 + X_2 + X_3, \dots, X_n$, so $Y = \sum_{i=1}^n X_i$. Then

$$\begin{aligned} M_Y(t) &= E(e^{ty}) \\ &= E(e^{(tx_1 + tx_2 + \dots + tx_n)}) \\ &= E(e^{tx_1}) E(e^{tx_2}) \dots E(e^{tx_n}) \\ &= M_{x_1}(t) M_{x_2}(t) \dots M_{x_n}(t) \\ &= (1 - p + pe^t)^n \end{aligned}$$

Now we take the first and second derivatives of $M_Y(t)$ in order to find the first and second moments.

$$\begin{aligned}M'_Y(t) &= \frac{d}{dt}(1 - p + e^t p)^n \\&= n(1 - p + e^t p)^{n-1} \frac{d}{dt}(1 - p + e^t p) \\&= n(1 - p + e^t p)^{n-1} e^t p \\&= npe^t(1 - p + pe^t)^{n-1}\end{aligned}$$