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STATS 244

HOMEWORK 1

Rice Chapter 1, 20

A deck of 52 cards has been shuffled thoroughly. What is the probability that the four aces are next to each other?

Let A be the event that 4 aces are next to each other. Consider 4 aces as being one card, then there are $(48!4!)$ ways to shuffle the deck with the four aces together. Since the 4 aces can be anywhere in the deck, there are 49 locations they can appear in the deck. Thus there are $\frac{(49)48!4!}{52!}$

$$\frac{49!4!}{52!} = \frac{4!}{(52)(51)(50)} = \frac{1}{5525}$$

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A factory runs three shifts. In a given day, 1% of items produced by the first shift are defective, 2% of the second shift's items are defective, and 5% of the third shift's items are defective. If the shifts all have the same productivity, what percentage of items produced in a day are defective? If an item is defective, what is the probability that it was produced by the third shift?

SUPPOSE WLOG, the factory produces 300 items a day. Since all shifts have the same productivity, each shift manufactures 100 items. Then the percentage of defective item produced per day is $\frac{8}{300} \approx 2.7\%$. Conditional on a defective item being chosen, the probability that it was produced by the third shift is $P(B | D) = \frac{\frac{.05}{3}}{\frac{.08}{3}} = \frac{5}{8}$ where D is a defective item.

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Suppose that n components are connected in series. For each unit, there is a backup unit, and the system fails if and only if both a unit and its backup fail. Assuming all the units are independent and fail with probability p , what is the probability that the system works? For $n = 10$ and $p = .05$.

FOR ANY GIVEN component the probability of failure is

$$p^2 = .05^2 = .0025$$

Let F be probability of failure. Following example F, we'll find the probability of the complement of this event, $P(F^c)$ or probability that

all the components work.

$$P(F) = 1 - P(F^c) \quad (1)$$

$$= 1 - (1 - p)^n \quad (2)$$

$$= 1 - (1 - .0025)^{10} \quad (3)$$

$$= 1 - (.975) \quad (4)$$

$$= .025 \quad (5)$$

The probability of failure in example F was .40 so this is a considerable improvement.

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This problem introduces some aspects of a simple genetic model. Assume that genes in an organism occur in pairs and that each member of the pair can be either of the types a or A . The possible genotypes of an organism are then AA , Aa and aa (Aa and aA are equivalent). When two organisms mate, each independently contributes one of its two genes; either one of the pair is transmitted with probability .5.

1. Suppose that the genotypes of the parents are AA and Aa . Find the possible genotypes of their offspring and the corresponding probabilities.

CONSIDER THE TABLE

	A	A
A	AA	AA
a	Aa	Aa

The possible genotypes are $\{AA, Aa\}$ with both with .5 probability.

2. Suppose that the probabilities of the genotypes AA , Aa and aa are p , $2q$, and r , respectively, in the first generation. Find the probabilities in the second and third generations and show that these are the same. This result is called the Hardy-Weinberg Law.

THE PROBABILITIES TABLES for each genome type given parent types are as follows.

$P(AA)$

	AA	Aa	aa
AA	1	.5	0
Aa	.5	.25	0
aa	0	0	0

$P(Aa)$

	AA	Aa	aa
AA	0	.5	1
Aa	.5	.5	.5
aa	1	.25	0

$P(aa)$

	AA	Aa	aa
AA	0	0	0
Aa	0	.25	.5
aa	0	.5	1

$P_g(x)$ is the probability of a genotype for generation g .

$$\begin{aligned}
 P_2(AA) &= P_2(AA \mid AA, AA)P_1(AA, AA) \\
 &\quad + 2(P_2(AA \mid AA, Aa)P_1(AA, Aa)) \\
 &\quad + P_2(AA \mid Aa, Aa)P_1(Aa, Aa) \\
 &= (1)(p^2) + 2(.5)(2pq) + (.25)(4q^2) \\
 &= p^2 + 2pq + q^2 \\
 &= (p + q)^2
 \end{aligned}$$

$$\begin{aligned}
 P_2(Aa) &= P_2(Aa \mid Aa, Aa)P_1(Aa, Aa) \\
 &\quad + 2(P_2(Aa \mid AA, Aa)P_1(AA, Aa)) \\
 &\quad + 2(P_2(Aa \mid Aa, aa)P_1(Aa, aa)) \\
 &\quad + 2(P_2(Aa \mid AA, aa)P_1(AA, aa)) \\
 &= (.5)(4q^2) + 2(.5)(2pq) \\
 &\quad + 2(.25)(2qr) + 2(1)(pr) \\
 &= 2q^2 + 2pq + 2pqr + 2pr \\
 &= 2(q + p)(q + r)
 \end{aligned}$$

$$\begin{aligned}
 P_2(aa) &= P_2(aa \mid Aa, Aa)P_1(Aa, Aa) \\
 &\quad + 2(P_2(aa \mid Aa, aa)P_1(Aa, aa)) \\
 &\quad + P_2(aa \mid aa, aa)P_1(aa, aa) \\
 &= (.25)(4q^2) + 2(.5)(2pr) + (1)(r^2) \\
 &= q^2 + 2qr + r^2 \\
 &= (q + r)^2
 \end{aligned}$$

Proof. It has been shown above that $P_2(AA) = (p + q)^2$, $P_2(Aa) = 2(q + p)(q + r)$ and $P_2(aa) = (q + r)^2$. To find the Hardy-Weinberg

Law result, we must show that $P_2(AA) = P_3(AA)$, $P_2(Aa) = P_3(Aa)$ and $P_2(aa) = P_3(aa)$.

First, it will be important to note that the three probabilities for a given generation's offspring will sum to 1. Ie...

$$(q + p)^2 + 2(q + p)(q + r) + (q + r)^2 = 1$$

Using the method to determine probabilities for generation 2 we can determine generation 3.

$$\begin{aligned} P_3(AA) &= 1(p + q)^4 + 2(.5)(p + q)^2 \\ &\quad + 2(q + p)(q + r) + (.25)4(p + q)(q + r)^2 \\ &= (p + q)^4 + 2(p + q)^3(q + r) + (p + q)^2(q + r)^2 \\ &= (p + q)^2[(p + q)^2 + 2(p + q)(q + r) + (q + r)^2] \\ &= (p + q)^2 \end{aligned}$$

$$\begin{aligned} P_3(Aa) &= (.25)(4)(q + p)^2(q + r)^2 + 2(.5)2(p + q)(q + r)(p + q)^2 \\ &= 2(.5)2(p + q)(q + r)(q + r)^2 + 2(p + q)^2(q + r)^2 \\ &= 2(p + q)^2(q + r)^2 + 2(p + q)^3(q + r) \\ &\quad + 2(p + q)(q + r)^3 + 2(p + q)^2(q + r)^2 \\ &= 2(q + p)(q + r)[(q + p)(q + r) + (p + q)^2 \\ &\quad + (q + r)^2(q + r)(q + r)] \\ &= 2(q + p)(q + r)[(p + q)^2 + 2(q + p)(q + r) + (q + r)^2] \\ &= 2(q + p)(q + r) \end{aligned}$$

$$\begin{aligned} P_3(aa) &= .25(4)(q + p)^2(q + r)^2 \\ &\quad + 2(.5)(q + r)^2 2(q + p)(q + r) + 1(q + r)^2 \\ &= (q + p)^2(q + r)^2 + 2(q + r)^3(q + p) + (q + r)^4 \\ &= (q + r)^2[(q + p)^2 + 2(q + r)(q + p) + (q + r)^2] \\ &= (q + r)^2 \end{aligned}$$

□

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Show that the binomial probabilities sum to 1.

Proof. The binomial distribution (n, θ) is given as

$$b(x; n, \theta) = \binom{n}{x} \theta (1 - \theta)^{n-x} \text{ for } x = 1, 2, \dots, n$$

Since $(a + b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$

$$\begin{aligned} \sum b(x; n, \theta) &= \sum_{x=0}^n \binom{n}{x} \theta (1 - \theta)^{n-x} \\ &= (\theta + 1 - \theta)^n \\ &= 1 \end{aligned}$$

□

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Let p_0, p_1, \dots, p_n denote the probability mass function of the binomial distribution with parameters n and p . Let $q = 1 - p$. Show that the binomial probabilities can be computed recursively by $p_0 = q^n$ and

$$p_k = \frac{(n - k + 1)p}{kq} p_{k-1}, \quad k = 1, 2, \dots, n$$

Use this relation to find $P(X \leq 4)$ for $n = 9000$ and $p = .0005$.

Proof. The binomial mass function for the parameters n and p and $q = 1 - p$ for $P(X = 0)$ is

$$\begin{aligned} p_0 &= \binom{n}{0} p^0 (1 - p)^{n-0} \\ &= \frac{n!}{0!n!} p^0 (q)^n \\ &= q^n \end{aligned}$$

Now consider

Binomial identity: $\binom{n}{k} = \frac{n+1-k}{k} \binom{n}{k-1}$

$$\begin{aligned} p_1 &= \binom{n}{1} p^1 q^{n-1} \\ &= \frac{n+1-1}{1} \binom{n}{1-1} p^1 q^{n-1} \\ &= n \binom{n}{0} p q^{n-1} \\ &= npq^{n-1} \end{aligned}$$

Taking the ratio of p_1 and p_0

$$\begin{aligned} \frac{p_1}{p_0} &= \frac{npq^{n-1}}{q^n} \\ &= \frac{np}{q} \end{aligned}$$

Then

$$p_1 = \frac{np}{q} p_0$$

More generally, since

$$\begin{aligned} p_k &= \binom{n}{k} p^k q^{n-k} \\ &= \frac{n+1-k}{k} \binom{n}{k-1} p^k q^{n-k} \end{aligned}$$

and

$$p_{k-1} = \binom{n}{k-1} p^{k-1} q^{n-(k-1)}$$

then

$$\begin{aligned} \frac{p_k}{p_{k-1}} &= \frac{\frac{n+1-k}{k} \binom{n}{k-1} p^k q^{n-k}}{\binom{n}{k-1} p^{k-1} q^{n-(k-1)}} \\ &= \frac{(n+1-k)}{k} p q^{-1} \\ &= \frac{(n+1-k)p}{kq} \end{aligned}$$

Therefore

$$p_k = \frac{(n+1-k)p}{kq} p_{k-1}$$

□

Using this relation we can find $P(X \leq 4)$ for $n = 9000$ and $p = .0005$.

$$\begin{aligned} P(X \leq 4) &= \sum_{k=0}^4 p_k \\ &= (.011) + (.05) + (.11) + (.16) + (.18) \\ &= .511 \end{aligned}$$

Question 7

In heads up Texas hold 'em (two players, each dealt two cards), find the probability that neither is dealt a pair (two cards of the same rank). If there are three players, what is the probability that none have a pair?

Let A be the event that the first player gets a pair and let B be the event that the second player gets a pair. The probability that no player gets a pair is

$$1 - (P(A) + P(B) - P(A \cap B))$$

$$1 - 2(.0588) - (.0048) = .8822$$

Now let C be the probability that the third player gets a pair. We can calculate the probability that none of the players get a pair using

$$1 - P(A \cup B \cup C) = 1 - (P(A) + P(B) + P(C)$$

$$- P(A \cap B) - P(C \cap A) - P(B \cap C)$$

$$+ P(A \cap B \cap C))$$

Where $P(A) = P(B) = P(C) = .0588$ and $P(A \cap C) \approx P(C \cap A) \approx P(B \cap C) \approx .0048$.

$$P(A \cap B \cap C) = P(A)P(B | A)P(C | A, B)$$

$$= (.0588)(.0816)(.0178)$$

The probability that no player has a pair is .838.

Question 8

For a Poisson process $N(,)$ with parameter λ , find the probabilities of the following events:

1. $N((1, 5]) > 1$ For a Poisson process the probability of k success over an interval of length t is

$$p(k) = \frac{\lambda t^k e^{-\lambda t}}{k!}$$

Since

$$P(N((1, 5]) > 1) = 1 - P(N(1, 5] = 0) - P(N(1, 5] = 1)$$

Compute and take the sum of the probabilities of each number of outcomes.

$$P(N((1, 5]) > 1) = 1 - \frac{\lambda(4)^0 e^{-\lambda(4)}}{0!} - \frac{\lambda(4)^1 e^{-\lambda(4)}}{1!}$$

$$= 1 - e^{-4\lambda} - 4\lambda e^{-4\lambda}$$

$$= 1 - e^{-4\lambda}(1 + 4\lambda)$$

2. $N((0,1]) = N((0,2])$

Take $P(N(0,1] = k) \cap P(N(0,2] = k)$

$$\frac{\lambda^k e^{-\lambda}}{k!} * \frac{\lambda 2^k e^{-\lambda 2}}{k!} = \frac{2\lambda^{2k} e^{-3\lambda}}{(k!)^2}$$

3. $N((1,2]) + N((3,4]) = 6$

Let I be the index set of all pairs x, y for which $x + y = 6$ where $x, y \geq 0$. Denote $A = N(1,2]$ and $B = N(2,3]$ Take

$$\sum_{i \in I} P(A = x_i) P(B = y_i) \text{ for } x, y \in I.$$

Which computes to $\frac{26.5}{180} (\lambda^{2k} e^{-2\lambda})$.

4. $N((0,1]) = N((1,2]) + m$ for m a nonnegative integer. Express your answer as an infinite series and then write your result in terms of a Bessel function (look it up). This result in terms of Bessel functions can be used to, for example, give a simple accurate approximation to this probability when m is large.

PROVED SOME integer m . Find $P(N(0,1]) = k + m \cap N(1,2] = k$, denote these $P(A)$ and $P(B)$, respectively.

$$p(A) = \frac{\lambda t^{(k+m)} e^{-\lambda t}}{(k+m)!}$$

$$p(B) = p(k) = \frac{\lambda t^k e^{-\lambda t}}{k!}$$

Multiplying the two yields

$$\frac{1}{k!(k+m)!} \lambda^{(2k+m)} e^{-2\lambda}.$$

As m gets larger the probability of m outcomes will get closer and closer to 0.