

Worksheet 16 (Solutions)

1. Let $X \sim \text{Bin}(1, p)$. Find the maximum likelihood estimator for p .

The MLE can be calculated by maximizing the log-likelihood

$$\begin{aligned} \frac{d}{dp} l(X; p) &\propto \frac{d}{dp} [x \log(p) + (n - x) \log(1 - p)] \\ &= \frac{x}{p} - \frac{n - x}{1 - p} \end{aligned}$$

Which is zero at:

$$\begin{aligned} \frac{x}{\hat{p}} &= \frac{1 - x}{1 - \hat{p}} \\ \hat{p}(1 - x) &= x \cdot (1 - \hat{p}) \\ \hat{p} &= x. \end{aligned}$$

2. Let $X \sim N(\mu, 1)$. Find the maximum likelihood estimator for p .

The MLE can be calculated by maximizing the log-likelihood

$$\begin{aligned} \frac{d}{d\mu} l(X; p) &\propto \frac{d}{d\mu} \left[\frac{-1}{2} (x - \mu)^2 \right] \\ &= (x - \mu) \end{aligned}$$

And therefore the MLE is at μ .

3. Let $X_1, \dots, X_n \sim U(a, b)$. Find the maximum likelihood estimator for a and b .

The likelihood is equal to $(b - a)^{-n}$ as long as all of the data is between a and b , and equal to zero otherwise. So we want to set:

$$\begin{aligned} \hat{a} &= \min_i (X_i) \\ \hat{b} &= \max_i (X_i) \end{aligned}$$

4. Let $X \sim \text{Gamma}(1, \beta)$. Find the maximum likelihood estimator for β .

The MLE can be calculated by maximizing the log-likelihood

$$\begin{aligned} \frac{d}{d\beta} l(X; p) &\propto \frac{d}{d\beta} [x/\beta - \log(\beta)] \\ &= -x/\beta^2 - \frac{1}{\beta} \end{aligned}$$

Which yields:

$$\hat{\beta} = x$$

5. Let $X_1, \dots, X_n \sim_{i.i.d.} N(\mu, \sigma^2)$. Find the maximum likelihood estimator for μ and σ^2 .

The MLE can be calculated by maximizing the log-likelihood for both μ and σ^2 :

$$\begin{aligned} \frac{d}{d\mu} l(X; p) &\propto \frac{d}{d\mu} \left[\frac{-1}{2\sigma^2} \sum_i (X_i - \mu)^2 \right] \\ &= \frac{-1}{\sigma^2} \cdot \sum_i (X_i - \mu) \end{aligned}$$

And therefore the MLE for the mean in terms of the variance is:

$$\hat{\mu} = \bar{X}$$

As before. The variance is:

$$\begin{aligned} \frac{d}{d\sigma^2} l(X; p) &\propto \frac{d}{d\sigma^2} \left[\frac{1}{2} \log(\sigma^2) \frac{-1}{2\sigma^2} \sum_i (X_i - \mu)^2 \right] \\ &= \frac{1}{\sigma^2} - \frac{1}{(\sigma^2)^2} \cdot \sum_i (X_i - \mu)^2 \end{aligned}$$

Which is zero at:

$$\hat{\sigma}^2 = \frac{1}{n} \cdot \sum_i (X_i - \mu)^2$$

And plugging in $\mu = \bar{X}$:

$$\begin{aligned} \hat{\mu} &= \bar{X} \\ \hat{\sigma}^2 &= \frac{1}{n} \cdot \sum_i (X_i - \bar{X})^2 \end{aligned}$$