Autoregressive models and simulation

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Topics

- semi-simple autoregressive models
- tips for statistical computing
- parametric bootstrapping
- the geometry of optimization
- the geometry of statistical inference

Semi-simple AR models

Brian Dennis and Taper (1994; B. Dennis et al. 2006)

```
library("ggplot2")
theme_set(theme_bw())
library("reshape2")
```

The basic Ricker model of population dynamics:

$$N_{t+1} = N_t \exp(a + bN_t + \sigma Z_t)$$
$$Z_t \sim N(0, 1)$$

Makes ecological sense (usually expressed in ecology as $N_{t+1} = rN_t \exp(-cN_t)$. Z_t incorporates *process error*. Log-transforming makes it into a *linear autoregressive* model (linearity in the *parameters a, b*).

$$X_{t+1} = X_t + a + be^{X_t} + \sigma Z_t$$

```
a=0, b=0 zero-drift Brownian motion a \neq 0, b=0 Brownian motion with drift (exp growth/decay) a \neq 0, b \neq 0 density-dependent growth a \neq 0, b < 0 regulation
```

Data from Dennis and Taper (Yellowstone grizzlies):

```
grizzly <- data.frame(n = c(33, 36, 34, 39, 35, 34, 38, 36, 37, 41, 39, 51, 47, 57, 48, 59, 64), t = 1973:1989)
```

Lagging variables in R is a little bit clunky; we have to drop the first and last elements respectively. head(x, -1) is an alternative to x[-n]. lag() doesn't work the way you think it does (!)

```
r <- log(grizzly$n[-1]/grizzly$n[-length(grizzly$n)])</pre>
  Look at the data:
par(las = 1, bty = "l", mfrow = c(1, 2))
plot(n ~ t, data = grizzly, type = "b", xlab = "time",
    ylab = "pop size")
plot(grizzly$n[-length(grizzly$n)], r, xlab = "pop size",
    ylab = "logarithmic growth rate")
   65
   60
                                        0.2
   55
                                     logarithmic growth rate
                                        0.1
   50
   45
                                        0.0
   40
                                       -0.1
         1975
                 1980
                        1985
                                                   40
                                                        45
                                                            50
```

Transform data and fit:

time

```
## R tip: use with() to simplify extraction
## from data frames -- **never** attach()
dd <- with(grizzly, data.frame(xlag1 = log(n[-length(n)]),</pre>
    x = log(n[-1]))
m2 \leftarrow lm(x \sim offset(xlag1) + exp(xlag1), data = dd)
## equivalently: lm(x-xlag1~exp(xlag1),data=dd)
## (offset may be easier to read)
s2 <- sigma(m2)^2
var.unbiased <- function(m) sigma(m)^2 * df.residual(m)/nobs(m)</pre>
## D&T use MLE rather than unbiased estimate of
## variance parameters
c(coef(m2), var = var.unbiased(m2))
    (Intercept)
                  exp(xlag1)
    0.141462270 -0.002411176 0.014524403
## test statistic
coef(summary(m2))[2, "t value"]
```

pop size

```
## [1] -0.6041122
## R tip: use update() when possible to
## simplify code
m1 <- update(m2, . ~ offset(xlag1))</pre>
c(coef(m1), var.unbiased(m1))
## (Intercept)
## 0.04139847 0.01490303
```

Parametric bootstrapping:

- for confidence intervals: simulate from the full model and record values. Compute quantiles.
- for *hypothesis tests*: simulate from the null model; fit the null model and the full model

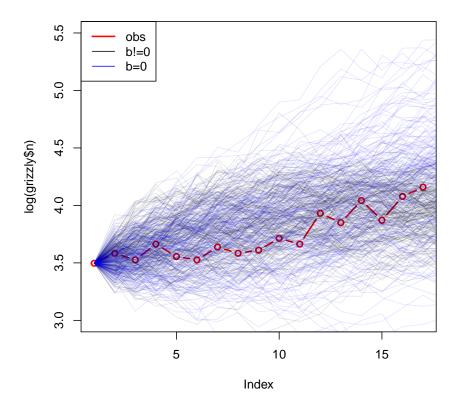
Simulation

lty = 1)

I thought I could use the simulate() method in R, but as it turns out I can't, so I have to define my own:

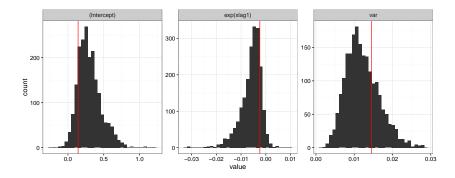
```
## simulate new data, add starting value
simfun \leftarrow function(fit, x0 = log(grizzly$n[1])) {
    cc <- coef(fit)
    a < - cc[1]
    ## make the function work for either
    b \leftarrow if (length(cc) == 2)
        coef(fit)[2] else 0
    sigma <- sqrt(var.unbiased(fit))</pre>
    res <- c(x0, numeric(nobs(fit) + 1))
    for (i in 2:length(res)) {
        res[i] \leftarrow res[i - 1] + a + b * exp(res[i - 1])
             1]) + rnorm(1, sd = sigma)
    }
    res
}
  Test simulation function:
sim2 <- replicate(200, simfun(m2))</pre>
sim1 <- replicate(200, simfun(m1))</pre>
plot(log(grizzly$n), type = "b", ylim = c(3, 5.5),
    lwd = 2, col = 2)
matlines(sim2, col = adjustcolor("black", alpha = 0.1),
```

```
matlines(sim1, col = adjustcolor("blue", alpha = 0.1),
    lty = 1)
legend("topleft", lty = 1, col = c("red", "black",
    "blue"), lwd = c(2, 1, 1), c("obs", "b!=0",
    "b=0"))
```



```
## re-fit model with new data
fitfun <- function(x, orig_fit) {</pre>
    dd <- data.frame(xlag1 = x[-length(x)], x = x[-1])
    update(orig_fit, data = dd)
}
## test:
all.equal(fitfun(log(grizzly$n), m2), m2)
## [1] TRUE
## summarize output
sumfun <- function(fit) {</pre>
    c(coef(fit), var = var.unbiased(fit))
}
## combine all of that stuff
bootfun <- function(fit) {</pre>
```

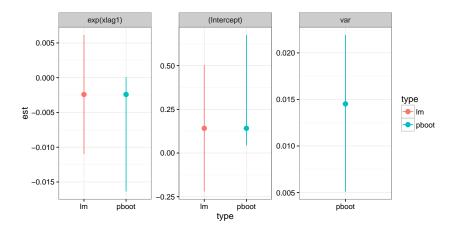
```
sumfun(fitfun(simfun(fit), fit))
}
set.seed(101)
v <- t(replicate(2000, bootfun(m2)))</pre>
## using ggplot, a fancy but powerful plotting
## tool
mv <- melt(as.data.frame(v)) ## restructure data in 'long' format</pre>
obsvals <- data.frame(variable = c("(Intercept)",</pre>
    "exp(xlag1)", "var"), value = sumfun(m2))
ggplot(mv, aes(value)) + facet_wrap(~variable,
    scale = "free") + geom_histogram(bins = 30) +
    geom_vline(data = obsvals, colour = "red",
        aes(xintercept = value))
```



```
## compare estimates
tfun <- function(x, type, est) {
    colnames(x) <- c("lwr", "upr")</pre>
    x <- data.frame(type, variable = rownames(x),</pre>
        est, x)
    rownames(x) <- NULL</pre>
    return(x)
}
bootconf <- t(apply(v, 2, quantile, c(0.025, 0.975)))
allCI <- rbind(tfun(confint(m2), "lm", coef(m2)),</pre>
    tfun(bootconf, "pboot", sumfun(m2)))
ggplot(allCI, aes(type, est, ymin = lwr, ymax = upr,
    colour = type)) + geom_pointrange() + facet_wrap(~variable,
    scale = "free")
```

Geometry of optimization (Press et al. 1994)

• need to find point estimate (best parameters), or get close



- frequentist MLE, Bayesian maximum a posteriori (mode of posterior distribution)
- geometric analogies; hill-climbing
- simplest cases: use derivatives
 - Newton-Raphson (good in theory, but fragile)
 - conjugate gradients (method="CG" in R optim)
 - quasi-Newton (method="BFGS" in R optim)
- derivatives can be expensive, fragile
 - Nelder-Mead (method="Nelder-Mead")
 - BOBYQA (nloptr package)
- complications: noisy surfaces
 - stochastic global optimizers: simulated annealing, genetic algorithms, differential evolution
- *complications*: multiple peaks
 - repeated starts
 - stochastic global optimization
- similar geometric issues apply to MCMC

Geometry of inference (Bolker 2008, ch. 6-7)

- Wald intervals: quadratic
- profile confidence intervals
- marginal intervals

References

Bolker, Benjamin M. 2008. Ecological Models and Data in R. Princeton, NJ: Princeton University Press.

Dennis, B., J. M Ponciano, S. R Lele, M. L Taper, and D. F Staples. 2006. "Estimating Density Dependence, Process Noise, and Observation Error." Ecological Monographs 76 (3): 323-41.

Dennis, Brian, and Mark L. Taper. 1994. "Density Dependence in Time Series Observations of Natural Populations: Estimation and Testing." Ecological Monographs 64 (2): 205-24. doi:10.2307/2937041.

Press, William H., Saul A. Teukolsky, William T. Vetterling, and Brian P. Flannery. 1994. Numerical Recipes in C: The Art of Scientific Computing. Cambridge University Press. http://www.nr.com/ oldverswitcher.html.