

SDS 383D Ex 04:
Hierarchical Models

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Math Tests

The data set in “mathtest.csv” shows the scores on a standardized math test from a sample of 10th-grade students at 100 different U.S. urban high schools, all having enrollment of at least 400 10th-grade students. (A lot of educational research involves “survey tests” of this sort, with tests administered to all students being the rare exception.)

Let θ_i be the underlying mean test score for school i , and let y_{ij} be the score for the j th student in school i . Starting with the “mathtest.R” script, you’ll notice that the extreme school-level averages \bar{y}_i (both high and low) tend to be at schools where fewer students were sampled.

Part 1

Briefly explain why this would be.

The extreme school-level averages occur in the schools with smaller sample sizes because we do not do a very good job of estimating the mean when sample size is small. These schools do not have min and max observation values that are more extreme than the other schools; they just have fewer observations to balance out the calculation of the mean. The smaller the sample size, the more influential an extreme observation is over the group mean.

Part 2

Fit a normal hierarchical model to these data via Gibbs sampling:

$$\begin{aligned} y_{ij} &\sim N(\theta_i, \sigma^2) \\ \theta_i &\sim N(\mu, \tau^2 \sigma^2) \end{aligned}$$

Decide upon sensible priors for the unknown model parameters (μ, σ^2, τ^2) .

The model is as follows.

$$\begin{aligned} (y_{ij} | \theta_i, \sigma^2) &\sim N(\theta_i, \sigma^2) \\ (\theta_i | \mu, \sigma^2, \tau^2) &\sim N(\mu, \sigma^2 \tau^2) \\ \mu &\sim I_{\mathbb{R}}(\mu), \text{ a flat prior on the real line} \\ \tau^2 &\sim I_{\mathbb{R}^+}(\tau^2), \text{ a flat prior on the positive real line} \\ \sigma^2 &\sim \left(\frac{1}{\sigma^2}\right) I_{\mathbb{R}^+}(\sigma^2), \text{ Jeffreys prior} \end{aligned}$$

where

- $i = 1, \dots, p$ indexes the p groups.
- n_i = sample size in each group.
- $j = 1, \dots, n_i$ indexes observations in a group.
- n = total number of observations.

The likelihood is

$$L(y | \theta_1, \dots, \theta_p, \sigma^2) \sim \prod_{i=1}^p \prod_{j=1}^{n_i} \left(\frac{1}{\sigma^2}\right)^{\frac{1}{2}} \exp\left[-\frac{1}{\sigma^2} (y_{ij} - \theta_i)^2\right] = (\sigma^2)^{-\frac{n}{2}} \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^p \sum_{j=1}^{n_i} (y_{ij} - \theta_i)^2\right]$$

The full conditionals are as follows.

$$(\theta_i | y, \mu, \sigma^2, \tau^2)$$

Note that \bar{y}_i is a sufficient statistic for the y 's, with $\bar{y}_i \sim N\left(\theta_i, \frac{\sigma^2}{n}\right)$.

$$(\theta_i | y, \mu, \sigma^2, \tau^2) \propto (\sigma^2)^{-\frac{1}{2}} \exp\left[-\frac{1}{2\sigma^2/n}(\bar{y}_i - \theta_i)^2\right] (\tau^2 \sigma^2)^{-\frac{1}{2}} \exp\left[-\frac{1}{2\sigma^2 \tau^2}(\theta_i - \mu)^2\right]$$

This is the normal-normal model, therefore

$$(\theta_i | y, \mu, \sigma^2, \tau^2) \sim N(m^*, v^*)$$

with

$$\begin{aligned} v^* &= \left[\frac{n_i}{\sigma^2} + \frac{1}{\sigma^2 \tau^2} \right]^{-1} = \left[\frac{n_i \tau^2 + 1}{\sigma^2 \tau^2} \right]^{-1} = \sigma^2 \left[\frac{\tau^2}{n_i \tau^2 + 1} \right] \\ m^* &= v^* \left[\left(\frac{n_i}{\sigma^2} \right) \bar{y}_i + \left(\frac{1}{\sigma^2 \tau^2} \right) \mu \right] \\ &= \sigma^2 \left[\frac{\tau^2}{n_i \tau^2 + 1} \right] \left[\left(\frac{n_i}{\sigma^2} \right) \bar{y}_i + \left(\frac{1}{\sigma^2 \tau^2} \right) \mu \right] \\ &= \left[\frac{n_i \tau^2}{n_i \tau^2 + 1} \right] \bar{y}_i + \left[\frac{1}{n_i \tau^2 + 1} \right] \mu \\ &= w \bar{y}_i + (1 - w) \mu \end{aligned}$$

So full conditional is

$$(\theta_i | y, \mu, \sigma^2, \tau^2) \sim N\left(\left[\frac{n_i \tau^2}{n_i \tau^2 + 1} \right] \bar{y}_i + \left[\frac{1}{n_i \tau^2 + 1} \right] \mu, \sigma^2 \left[\frac{\tau^2}{n_i \tau^2 + 1} \right]\right) \quad (1)$$

$$(\mu | \theta, y, \sigma^2, \tau^2)$$

$$\begin{aligned} (\mu | \theta, y, \sigma^2, \tau^2) &\propto \exp\left[-\frac{1}{2\sigma^2 \tau^2} \sum_{i=1}^p (\theta_i - \mu)^2\right] \cdot 1 \\ &= \exp\left[-\frac{1}{2\sigma^2 \tau^2} \{(\theta_1 - \mu)(\theta_1 - \mu) + \dots + (\theta_p - \mu)(\theta_p - \mu)\}\right] \\ &= \exp\left[-\frac{1}{2\sigma^2 \tau^2} \left\{ p\mu^2 - 2\mu \sum_{i=1}^p \theta_i + \sum_{i=1}^p \theta_i^2 \right\}\right] \\ &= \exp\left[-\frac{p}{2\sigma^2 \tau^2} \left\{ \mu^2 - 2\mu \left(\frac{\sum_{i=1}^p \theta_i}{p} \right) + \frac{\sum_{i=1}^p \theta_i^2}{p} \right\}\right] \\ &\propto \exp\left[-\frac{p}{2\sigma^2 \tau^2} \left\{ \mu^2 - 2\mu \bar{\theta}_i \right\}\right] \end{aligned}$$

We recognize this as a Normal kernel, therefore

$$(\mu | \theta, y, \sigma^2, \tau^2) \sim N\left(\bar{\theta}_i, \frac{\sigma^2 \tau^2}{p}\right) \quad (2)$$

$$(\sigma^2 | \theta, y, \mu, \tau^2)$$

$$\begin{aligned}
(\sigma^2 | \theta, y, \mu, \tau^2) &\propto (\sigma^2)^{-\frac{n}{2}} \exp \left[-\frac{1}{2\sigma^2} \sum_{i=1}^p \sum_{j=1}^{n_i} (y_{ij} - \theta_i)^2 \right] (\sigma^2)^{-\frac{p}{2}} \exp \left[-\frac{1}{2\sigma^2 \tau^2} \sum_{i=1}^p (\theta_i - \mu)^2 \right] \left(\frac{1}{\sigma^2} \right) \\
&= (\sigma^2)^{-\frac{(n+p)}{2}-1} \exp \left[-\left(\frac{1}{\sigma^2} \right) \cdot \left\{ \frac{1}{2} \sum_{i=1}^p \sum_{j=1}^{n_i} (y_{ij} - \theta_i)^2 + \frac{1}{2\tau^2} \sum_{i=1}^p (\theta_i - \mu)^2 \right\} \right]
\end{aligned}$$

We recognize this as an Inverse-Gamma kernel, therefore

$$(\sigma^2 | \theta, y, \mu, \tau^2) \sim IG \left(\frac{(n+p)}{2}, \left\{ \frac{1}{2} \sum_{i=1}^p \sum_{j=1}^{n_i} (y_{ij} - \theta_i)^2 + \frac{1}{2\tau^2} \sum_{i=1}^p (\theta_i - \mu)^2 \right\} \right) \quad (3)$$

$$(\tau^2 | \theta, y, \mu, \sigma^2)$$

$$(\tau^2 | \theta, y, \mu, \sigma^2) \propto (\tau^2)^{-\frac{p}{2}} \exp \left[-\frac{1}{2\sigma^2 \tau^2} \sum_{i=1}^p (\theta_i - \mu)^2 \right] \cdot 1$$

We recognize this as an Inverse Gamma kernel, therefore

$$(\tau^2 | \theta, y, \mu, \sigma^2) \sim IG \left(\frac{p}{2} - 1, \frac{1}{2\sigma^2} \sum_{i=1}^p (\theta_i - \mu)^2 \right) \quad (4)$$

Part 3

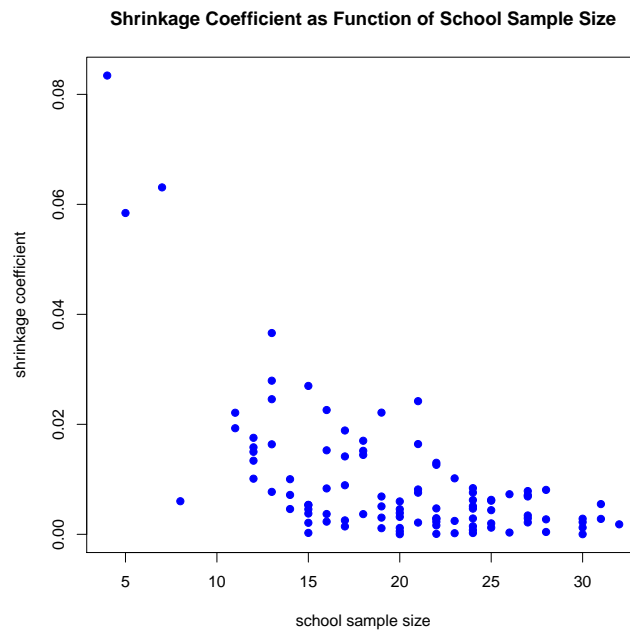


Figure 1: Shrinkage estimator by school sample size

Price Elasticity of Demand

A Hierarchical Probit Model via Data Augmentation

Gene Expression Over Time