

# Deriving the Least-Squares Estimates for Simple Linear Regression

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This document contains the mathematical details for deriving the least-squares estimates for slope ( $\beta_1$ ) and intercept ( $\beta_0$ ). We obtain the estimates,  $\hat{\beta}_1$  and  $\hat{\beta}_0$  by finding the values that minimize the sum of squared residuals (1).

$$SSR = \sum_{i=1}^n [y_i - \hat{y}_i]^2 = [y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)]^2 = [y_i - (\hat{\beta}_0 - \hat{\beta}_1 x_i)]^2 \quad (1)$$

Recall that we can find the values of  $\hat{\beta}_1$  and  $\hat{\beta}_0$  that minimize (1) by taking the partial derivatives of (1) and setting them to 0. Thus, the values of  $\hat{\beta}_1$  and  $\hat{\beta}_0$  that minimize the respective partial derivative also minimize the sum of squared residuals. The partial derivatives are

$$\begin{aligned} \frac{\partial SSR}{\partial \hat{\beta}_1} &= -2 \sum_{i=1}^n x_i (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) \\ \frac{\partial SSR}{\partial \hat{\beta}_0} &= -2 \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) \end{aligned} \quad (2)$$

Let's begin by deriving  $\hat{\beta}_0$ .

$$\begin{aligned} \frac{\partial SSR}{\partial \hat{\beta}_0} &= -2 \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \\ \Rightarrow - \sum_{i=1}^n (y_i + \hat{\beta}_0 + \hat{\beta}_1 x_i) &= 0 \\ \Rightarrow - \sum_{i=1}^n y_i + n\hat{\beta}_0 + \hat{\beta}_1 \sum_{i=1}^n x_i &= 0 \\ \Rightarrow n\hat{\beta}_0 &= \sum_{i=1}^n y_i - \hat{\beta}_1 \sum_{i=1}^n x_i \\ \Rightarrow \hat{\beta}_0 &= \frac{1}{n} \left( \sum_{i=1}^n y_i - \hat{\beta}_1 \sum_{i=1}^n x_i \right) \\ \Rightarrow \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x} \end{aligned} \quad (3)$$

Now, we can derive  $\hat{\beta}_1$  using the  $\hat{\beta}_0$  we just derived

$$\begin{aligned}
\frac{\partial \text{SSR}}{\partial \hat{\beta}_1} &= -2 \sum_{i=1}^n x_i (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \\
\Rightarrow - \sum_{i=1}^n x_i y_i + \hat{\beta}_0 \sum_{i=1}^n x_i + \hat{\beta}_1 \sum_{i=1}^n x_i^2 &= 0 \\
(\text{Fill in } \hat{\beta}_0) \Rightarrow - \sum_{i=1}^n x_i y_i + (\bar{y} - \hat{\beta}_1 \bar{x}) \sum_{i=1}^n x_i + \hat{\beta}_1 \sum_{i=1}^n x_i^2 &= 0 \\
\Rightarrow (\bar{y} - \hat{\beta}_1 \bar{x}) \sum_{i=1}^n x_i + \hat{\beta}_1 \sum_{i=1}^n x_i^2 &= \sum_{i=1}^n x_i y_i \\
\Rightarrow \bar{y} \sum_{i=1}^n x_i - \hat{\beta}_1 \bar{x} \sum_{i=1}^n x_i + \hat{\beta}_1 \sum_{i=1}^n x_i^2 &= \sum_{i=1}^n x_i y_i \\
\Rightarrow n \bar{y} \bar{x} - \hat{\beta}_1 n \bar{x}^2 + \hat{\beta}_1 \sum_{i=1}^n x_i^2 &= \sum_{i=1}^n x_i y_i \\
\Rightarrow \hat{\beta}_1 \sum_{i=1}^n x_i^2 - \hat{\beta}_1 n \bar{x}^2 &= \sum_{i=1}^n x_i y_i - n \bar{y} \bar{x} \\
\Rightarrow \hat{\beta}_1 \left( \sum_{i=1}^n x_i^2 - n \bar{x}^2 \right) &= \sum_{i=1}^n x_i y_i - n \bar{y} \bar{x} \\
\hat{\beta}_1 &= \frac{\sum_{i=1}^n x_i y_i - n \bar{y} \bar{x}}{\sum_{i=1}^n x_i^2 - n \bar{x}^2}
\end{aligned} \tag{4}$$

To write  $\hat{\beta}_1$  in a form that's more recognizable, we will use the following:

$$\sum x_i y_i - n \bar{y} \bar{x} = \sum (x - \bar{x})(y - \bar{y}) = (n-1) \text{Cov}(x, y) \tag{5}$$

$$\sum x_i^2 - n \bar{x}^2 = \sum (x - \bar{x})^2 = (n-1) s_x^2 \tag{6}$$

where  $\text{Cov}(x, y)$  is the covariance of  $x$  and  $y$ , and  $s_x^2$  is the sample variance of  $x$  ( $s_x$  is the sample standard deviation).

Thus, applying (5) and (6), we have

$$\begin{aligned}
\hat{\beta}_1 &= \frac{\sum_{i=1}^n x_i y_i - n \bar{y} \bar{x}}{\sum_{i=1}^n x_i^2 - n \bar{x}^2} \\
&= \frac{\sum_{i=1}^n (x - \bar{x})(y - \bar{y})}{\sum_{i=1}^n (x - \bar{x})^2} \\
&= \frac{(n-1) \text{Cov}(x, y)}{(n-1) s_x^2} \\
&= \frac{\text{Cov}(x, y)}{s_x^2}
\end{aligned} \tag{7}$$

The correlation between  $x$  and  $y$  is  $r = \frac{\text{Cov}(x, y)}{s_x s_y}$ . Thus,  $\text{Cov}(x, y) = r s_x s_y$ . Plugging this into (7), we have

$$\hat{\beta}_1 = \frac{\text{Cov}(x, y)}{s_x^2} = r \frac{s_y s_x}{s_x^2} = r \frac{s_y}{s_x} \tag{8}$$