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Abstract—This manual provides a simple introduction to the generation of random numbers

## 1 Uniform Random Numbers

Let U be a uniform random variable between 0 and 1.

1.1 Generate  $10^6$  samples of U using a C program and save into a file called uni.dat.

**Solution:** : Link to the code : C code

1.2 Load the uni.dat file into python and plot the empirical CDF of *U* using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr(U \le x) \tag{1.1}$$

**Solution:** The following code plots Fig: Python code

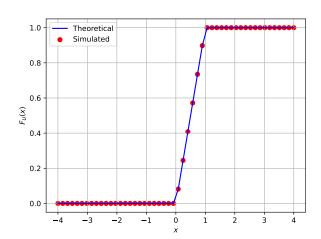


Fig. 1: The CDF of U

1.3 Find a theoretical expression for  $F_U(x)$ . **Solution:** As U is uniformly distributed random variable in the interval (0,1) and

$$p_U(x) = \begin{cases} 1, & x \in (0,1) \\ 0, & otherwise \end{cases}$$
 (1.2)

$$F_U(x) = \int_{-\infty}^x p_U(x) dx \tag{1.3}$$

$$F_U(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ x, & x \in (0, 1) \\ 1, & x \in (1, \infty) \end{cases}$$
 (1.4)

1.4 The mean of *U* is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^{N} U_i$$
 (1.5)

and its variance as

$$var[U] = E[U - E[U]]^2$$
 (1.6)

Write a C program to find the mean and variance of U.

**Solution:** Link to the code :C code

aniket@aniket-HP:~/Desktop/PR\$ gcc code/codel\_4.c aniket@aniket-HP:~/Desktop/PR\$ ./a.out Mean is 0.500007 Variance is 0.083301 aniket@aniket-HP:~/Desktop/PR\$

Fig. 1: Output

1.5 Verify your result theoretically given that

$$E\left[U^{k}\right] = \int_{-\infty}^{\infty} x^{k} dF_{U}(x) \tag{1.7}$$

Solution: we know that,

$$dF_U(x) = p_U(x)dx (1.8)$$

also mean  $(\mu)$  is E(U): Hence,

$$\mu = \int_{-\infty}^{\infty} x p_U(x) dx$$
$$= \int_{0}^{1} x dx$$
$$= \frac{x^2}{2} \Big|_{0}^{1}$$
$$= \frac{1}{2}$$

 $Variance(\sigma^2) = E(U^2) - E(U)^2$ 

$$E(U^2) = \int_{-\infty}^{\infty} x^2 p_U(x) dx$$
$$= \int_0^1 x^2 dx$$
$$= \frac{x^3}{3} \Big|_0^1$$
$$= \frac{1}{3}$$

$$\sigma^2 = \frac{1}{3} - \frac{1}{4}$$
$$= \frac{1}{12}$$

2 Central Limit Theorem

2.1 Generate 10<sup>6</sup> samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \tag{2.1}$$

using a C program, where  $U_i$ , i = 1, 2, ..., 12 are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

**Solution:** link to the code : C code

2.2 Load gau.dat in python and plot the empirical CDF of *X* using the samples in gau.dat. What properties does a CDF have?

Solution: Link to Python code: Python code

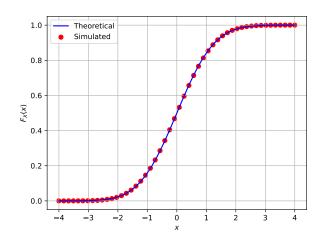


Fig. 2: The CDF of X

Properties:

- (1) Graph is symmetric about a single point
- (2) The  $F_X(x)$  is non-decreasing function
- (3)  $\lim_{x\to\infty} F_X(x)$
- 2.3 Load gau.dat in python and plot the empirical PDF of *X* using the samples in gau.dat. The PDF of *X* is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{2.2}$$

What properties does the PDF have? **Solution:** Link to the code :Python code

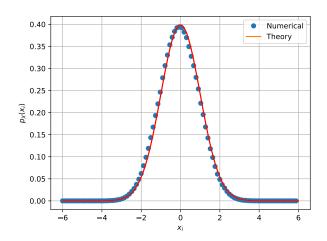


Fig. 2: The PDF of X

Properties:

- (1) Area under the curve is One.
- (2) Symmetric about line  $x = \mu$ .
- (3) Increasing in first half and decreasing in other half.
- 2.4 Find the mean and variance of *X* by writing a C program.

**Solution:** Link to the code : C code

```
aniket@aniket-HP:~/Desktop/PR$ gcc code/code2_4.c
aniket@aniket-HP:~/Desktop/PR$ ./a.out
Mean is 0.000294
Variance is 0.999560
aniket@aniket-HP:~/Desktop/PR$
```

Fig. 2: Output

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (2.3)$$

repeat the above exercise theoretically. **Solution:** 

(1) CDF is

$$F_X(x) = \int_{-\infty}^x p_X(x)dx$$
$$= 1$$

(2) Mean is

$$\mu = E(x) = \int_{-\infty}^{\infty} x p_X(x) dx$$

Due to symmetry clearly,  $\mu = 0$ 

(3) Variance is

$$var[X] = E(X^{2}) - E(X)^{2}$$
$$= \int_{-\infty}^{\infty} x^{2} p_{X}(x) dx - 0$$
$$= 1$$

(4) Q function is

$$Q(x) = Pr(x > X)$$
$$= 1 - F_X(x)$$

Then CDF is:

$$F_X(x) = 1 - Q(x)$$

- 3 From Uniform to Other
- 3.1 Generate samples of

$$V = -2\ln(1 - U) \tag{3.1}$$

and plot its CDF.

**Solution:** Link to the code :Python code

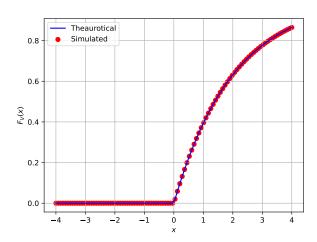


Fig. 3: CDF of V

3.2 Find a theoretical expression for  $F_V(x)$ . Solution:

$$F_V(X) = Pr(V \le x)$$

$$= Pr(-2\ln(1 - U) \le x)$$

$$= Pr((1 - U) \ge \exp(\frac{-x}{2}))$$

$$= Pr(U \le (1 - \exp(\frac{-x}{2})))$$

$$= F_U(1 - \exp(\frac{-x}{2}))$$

from equation (1.3),

$$F_U(x) = \begin{cases} 0, & x \in (-\infty, 0) \\ (1 - \exp\frac{-x}{2}), & x \in (0, \infty) \end{cases}$$

4 Triangular Distribution

4.1 Generate

$$T = U_1 + U_2 \tag{4.1}$$

**Solution:** Link to the code :C code

4.2 Find CDF of T

**Solution:** Link to the code :Pyhton code

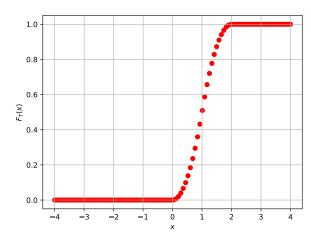


Fig. 4: CDF of T

4.3 Find the PDF of T.

**Solution:** Link to the code :Pyhton code

4.4 Find the theoretical expressions for the PDF and CDF of T.

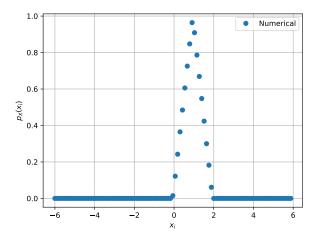


Fig. 4: PDF of T

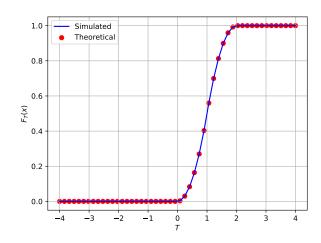


Fig. 4: Verification of CDF

## **Solution:**

$$T = U_1 + U_2 (4.2)$$

$$p_T(t) = p_{U_1} * p_{U_2} (4.3)$$

$$p_T(t) = \int_{-\infty}^{\infty} p_{U_1}(\tau) p_{U_2}(t-\tau) d\tau$$
 (4.4)

$$p_{T}(t) = \int_{-\infty}^{\infty} p_{U_{1}}(\tau) p_{U_{2}}(t-\tau) d\tau \qquad (4.4)$$

$$p_{T}(t) = \begin{cases} 0 & t \le 0 \\ \int_{0}^{t} d\tau & 0 \le t < 1 \end{cases}$$

$$p_{T}(t) = \begin{cases} \int_{t-1}^{1} d\tau & 1 \le t < 2 \\ 0 & t > 2 \end{cases} \qquad (4.5)$$

For CDF:

$$F_T(x) = \int_{-\infty}^x p_T(t)dt \tag{4.6}$$

From equation (4.5):

$$F_T(t) = \begin{cases} 0 & t \in (-\infty, 0) \\ \frac{x^2}{2} & t \in (0, 1) \\ \frac{-t^2}{2} + 2t - 1 & t \in (1, 2) \\ 1 & t \in (2, \infty) \end{cases}$$
(4.7)

4.5 Verify results through plot

## **Solution:**

Link to the code for CDF:Pyhton code Link to the code for PDF:Pyhton code

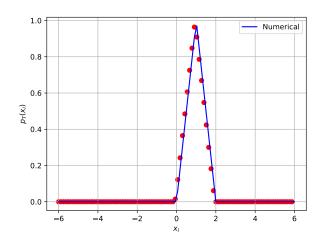


Fig. 4: verification of PDF

## 5 Maximum Likelihood

- 5.1 Generate equiprobable  $X \in \{1, -1\}$ **Solution:** Link to the code : C code
- 5.2 Generate

$$Y = AX + N, (5.1)$$

where A = 5 dB, and  $N \sim \mathcal{N}(0, 1)$ . **Solution:** Link to the code : C code

5.3 Plot Y using a scatter plot.

**Solution:** Link to the code :Python code

5.4 Guess how to estimate X from Y.

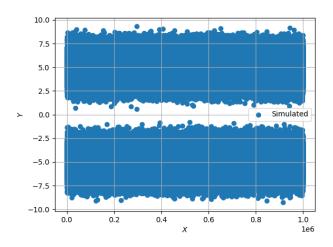


Fig. 5

**Solution:** From Data and Plot generally:

$$X = \begin{cases} 1 & Y > 0 \\ -1 & Y < 0 \end{cases}$$
 (5.2)

5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.3)

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.4)

**Solution:** Link to the code :C code

aniket@aniket-HP:~/Desktop/PR/code\$ gcc code5\_5.c
aniket@aniket-HP:~/Desktop/PR/code\$ ./a.out
\$P\_{e|0}\$ = 0.000000
\$P\_{e|1}\$ = 0.000000
aniket@aniket-HP:~/Desktop/PR/code\$

Fig. 5

From the Figure above:

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) = 0$$
  
 $P_{e|1} = \Pr(\hat{X} = 1|X = -1) = 0$ 

5.6 Find  $P_e$  assuming that X has equiprobable symbols.

Solution: Here.

$$Pr(X = 1) = Pr(X = -1) = \frac{1}{2}$$
 (5.5)

Hence,

$$P_e = \Pr(X = 1)P_{e|1} + \Pr(X = -1)P_{e|0}$$
  
=  $\frac{1}{2}(P_{e|1} + P_{e|0})$   
= 0

5.7 Verify by plotting the theoretical  $P_e$  with respect to A from 0 10dB.

**Solution:** Link to the code :Python code

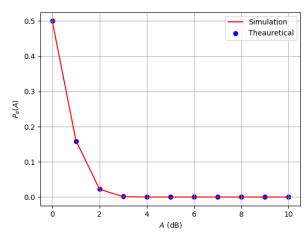


Fig. 5

5.8 Now consider a threshold  $\delta$  while estimating X from Y. Find the value of  $\delta$  that maximizes the theoretical  $P_e$ 

**Solution:** From equation (5.2) Hence,

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$

$$= \Pr(AX + N < \delta | X = 1)$$

$$= \Pr(N < \delta - A)$$

$$= O_N(A - \delta)$$

Where  $Q_N$  is the Q-function of the normal distribution. Similarly,

$$P_{e|1} = Q_N(A + \delta) \tag{5.6}$$

Therefore,

$$P_e = P_{e|0} \Pr(X = 1) + P_{e|1} \Pr(X = -1)$$
 (5.7)

$$= \frac{1}{2}(Q_N(A - \delta) + Q_N(A + \delta))$$
 (5.8)

To find minima by differentiate wrt  $\delta$ 

$$\begin{split} \frac{d(P_e)}{d\delta} &= \frac{1}{2} \frac{d}{d\delta} (Q_N(A - \delta) + Q_N(A + \delta)) \\ &= \frac{1}{2\sqrt{2\pi}} (e^{-\frac{(\delta - A)^2}{2}} - e^{-\frac{(\delta + A)^2}{2}}) = 0 \end{split}$$

From above equation

$$(\delta - A)^2 = (\delta + A)^2$$

Hence,

$$\delta = 0$$

5.9 Repeat the above exercise when

$$p_X(0) = p \tag{5.10}$$

**Solution:** From equation 5.8

$$P_e = P_{e|0}p + P_{e|1}(1-p)$$
  
=  $Q_N(A - \delta)p + Q_N(A + \delta)(1-p)$ 

Again by differentiating,

$$p\frac{1}{\sqrt{2\pi}}e^{-\frac{(\delta-A)^2}{2}} - (1-p)\frac{1}{\sqrt{2\pi}}e^{-\frac{(\delta+A)^2}{2}} = 0$$

By taking log,

$$\ln p - \frac{(\delta - A)^2}{2} = \ln(1 - p) + \frac{(\delta + A)^2}{2}$$
$$\delta = \frac{1}{2A} \ln \frac{1 - p}{p}$$

5.10 Repeat the above exercise using the MAP criterion.