

Lecture 1

Chapter 2

1

PATTERN RECOGNITION

- ❖ Typical application areas
 - Machine vision
 - Character recognition (OCR)
 - Computer aided diagnosis
 - Speech recognition
 - Face recognition
 - Biometrics
 - Image Data Base retrieval
 - Data mining
 - Bionformatics
- ❖ The task: Assign unknown objects – patterns – into the correct class. This is known as classification.

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❖ **Features:** These are measurable quantities obtained from the patterns, and the classification task is based on their respective values.

❖ **Feature vectors:** A number of features

$$x_1, \dots, x_l,$$

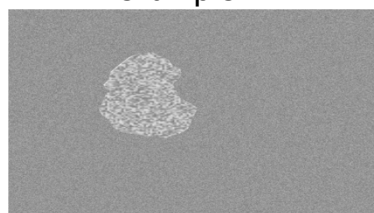
constitute the feature vector

$$\underline{x} = [x_1, \dots, x_l]^T \in R^l$$

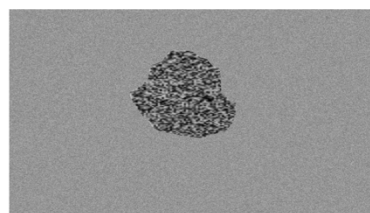
Feature vectors are treated as random vectors.

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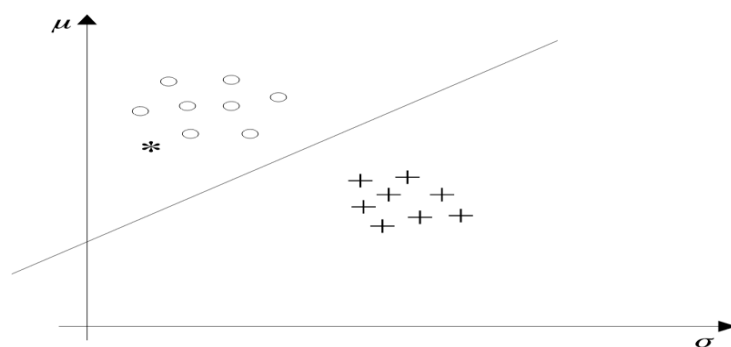
An example:



(a)



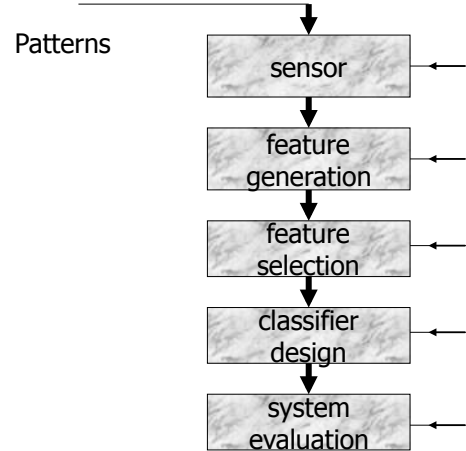
(b)



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- ❖ The classifier consists of a set of functions, whose values, computed at \underline{x} , determine the class to which the corresponding pattern belongs

- ❖ Classification system overview



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- ❖ **Supervised – unsupervised – semisupervised** pattern recognition:

The major directions of learning are:

- **Supervised:** Patterns whose class is known a-priori are used for training.
- **Unsupervised:** The number of classes/groups is (in general) unknown and no training patterns are available.
- **Semisupervised:** A mixed type of patterns is available. For some of them, their corresponding class is known and for the rest is not.

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CLASSIFIERS BASED ON BAYES DECISION THEORY

- ❖ Statistical nature of feature vectors

$$\underline{x} = [x_1, x_2, \dots, x_l]^T$$

- ❖ Assign the pattern represented by feature vector \underline{x} to the most probable of the available classes

$$\omega_1, \omega_2, \dots, \omega_M$$

That is $\underline{x} \rightarrow \omega_i : \underset{\text{maximum}}{P(\omega_i | \underline{x})}$

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- ❖ Computation of a-posteriori probabilities

- Assume known

- a-priori probabilities

$$P(\omega_1), P(\omega_2), \dots, P(\omega_M)$$

- $p(\underline{x} | \omega_i), i = 1, 2, \dots, M$

This is also known as the likelihood of

$$\underline{x} \text{ w.r. to } \omega_i.$$

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➤ The Bayes rule ($M=2$)

$$p(\underline{x})P(\omega_i|\underline{x}) = p(\underline{x}|\omega_i)P(\omega_i) \Rightarrow$$

$$P(\omega_i|\underline{x}) = \frac{p(\underline{x}|\omega_i)P(\omega_i)}{p(\underline{x})}$$

where

$$p(\underline{x}) = \sum_{i=1}^2 p(\underline{x}|\omega_i)P(\omega_i)$$

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❖ The Bayes classification rule (for two classes $M=2$)

➤ Given \underline{x} classify it according to the rule

$$\begin{array}{l} \text{If } P(\omega_1|\underline{x}) > P(\omega_2|\underline{x}) \quad \underline{x} \rightarrow \omega_1 \\ \text{If } P(\omega_2|\underline{x}) > P(\omega_1|\underline{x}) \quad \underline{x} \rightarrow \omega_2 \end{array}$$

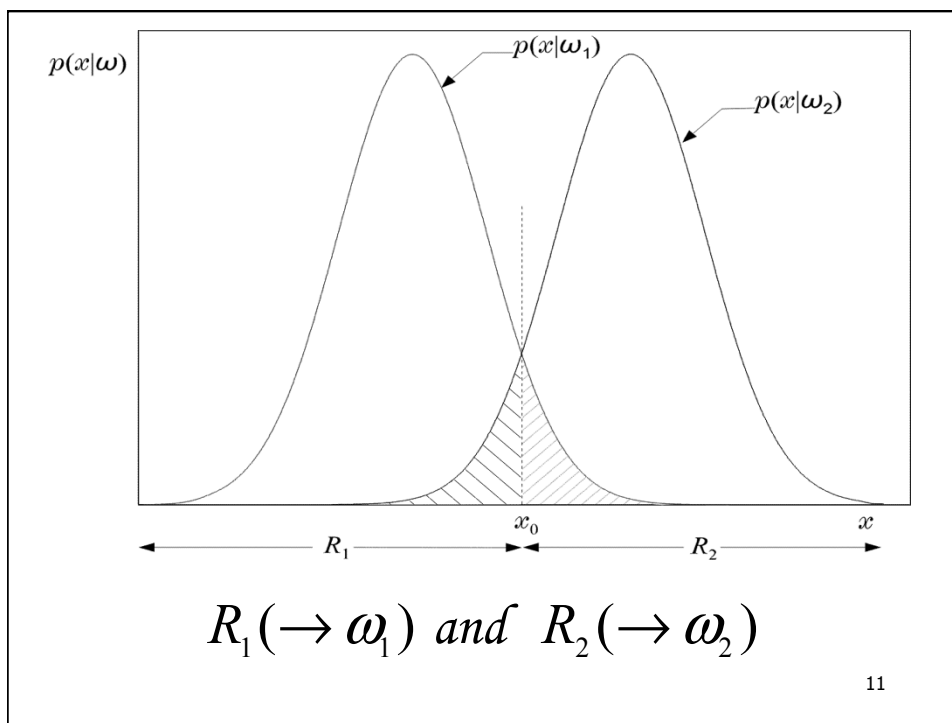
➤ Equivalently: classify \underline{x} according to the rule

$$p(\underline{x}|\omega_1)P(\omega_1) (><) p(\underline{x}|\omega_2)P(\omega_2)$$

➤ For equiprobable classes the test becomes

$$p(\underline{x}|\omega_1) (><) p(\underline{x}|\omega_2)$$

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❖ Equivalently in words: Divide space in two regions

$\begin{aligned} \text{If } \underline{x} \in R_1 &\Rightarrow \underline{x} \text{ in } \omega_1 \\ \text{If } \underline{x} \in R_2 &\Rightarrow \underline{x} \text{ in } \omega_2 \end{aligned}$
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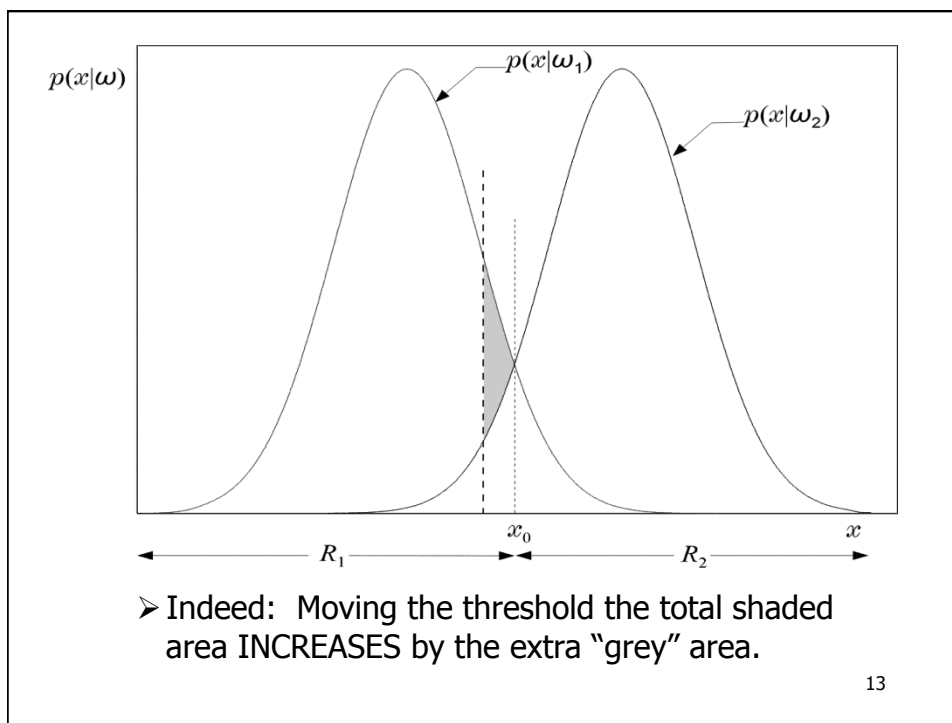
❖ Probability of error

➤ Total shaded area

$$\text{➤ } P_e = \frac{1}{2} \int_{-\infty}^{x_0} p(x|\omega_2) dx + \frac{1}{2} \int_{x_0}^{+\infty} p(x|\omega_1) dx$$

❖ Bayesian classifier is OPTIMAL with respect to minimising the classification error probability!!!!

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❖ The Bayes classification rule for many ($M > 2$) classes:

➤ Given \underline{x} classify it to ω_i if:

$$P(\omega_i|\underline{x}) > P(\omega_j|\underline{x}) \quad \forall j \neq i$$

➤ Such a choice also minimizes the classification error probability

❖ Minimizing the average risk

➤ For each wrong decision, a penalty term is assigned since some decisions are more sensitive than others

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➤ For $M=2$

- Define the loss matrix

$$L = \begin{pmatrix} \lambda_{11} & \lambda_{12} \\ \lambda_{21} & \lambda_{22} \end{pmatrix}$$

11 - cost of
deciding class 1
when the true
class is indeed 1.

- λ_{12} penalty term for deciding class ω_2 , although the pattern belongs to ω_1 , etc.

➤ Risk with respect to ω_1

$$r_1 = \lambda_{11} \int_{R_1} p(\underline{x}|\omega_1) d\underline{x} + \lambda_{12} \int_{R_2} p(\underline{x}|\omega_1) d\underline{x}$$

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➤ Risk with respect to ω_2

$$r_2 = \lambda_{21} \int_{R_1} p(\underline{x}|\omega_2) d\underline{x} + \lambda_{22} \int_{R_2} p(\underline{x}|\omega_2) d\underline{x}$$

➤ $\square \Rightarrow$ Probabilities of wrong decisions, weighted by the penalty terms

➤ Average risk

$$r = r_1 P(\omega_1) + r_2 P(\omega_2)$$

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❖ Choose R_1 and R_2 so that r is minimized

❖ Then assign \underline{x} to ω_i if

$$\ell_1 \equiv \lambda_{11}p(\underline{x}|\omega_1)P(\omega_1) + \lambda_{21}p(\underline{x}|\omega_2)P(\omega_2) <$$

$$\ell_2 \equiv \lambda_{12}p(\underline{x}|\omega_1)P(\omega_1) + \lambda_{22}p(\underline{x}|\omega_2)P(\omega_2)$$

❖ Equivalently:

assign \underline{x} in $\omega_1(\omega_2)$ if

$$\ell_{12} \equiv \frac{p(\underline{x}|\omega_1)}{p(\underline{x}|\omega_2)} > (<) \frac{P(\omega_2)}{P(\omega_1)} \frac{\lambda_{21} - \lambda_{22}}{\lambda_{12} - \lambda_{11}}$$

ℓ_{12} : likelihood ratio

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❖ If $P(\omega_1) = P(\omega_2) = \frac{1}{2}$ and $\lambda_{11} = \lambda_{22} = 0$

$$\begin{aligned} \underline{x} &\rightarrow \omega_1 \text{ if } P(\underline{x}|\omega_1) > P(\underline{x}|\omega_2) \frac{\lambda_{21}}{\lambda_{12}} \\ \underline{x} &\rightarrow \omega_2 \text{ if } P(\underline{x}|\omega_2) > P(\underline{x}|\omega_1) \frac{\lambda_{12}}{\lambda_{21}} \end{aligned}$$

if $\lambda_{21} = \lambda_{12} \Rightarrow$ Minimum classification
error probability

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❖ An example:

$$\begin{aligned}
 - \quad p(x|\omega_1) &= \frac{1}{\sqrt{\pi}} \exp(-x^2) \\
 - \quad p(x|\omega_2) &= \frac{1}{\sqrt{\pi}} \exp(-(x-1)^2) \\
 - \quad P(\omega_1) &= P(\omega_2) = \frac{1}{2} \\
 - \quad L &= \begin{pmatrix} 0 & 0.5 \\ 1.0 & 0 \end{pmatrix}
 \end{aligned}$$

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➤ Then the threshold value is:

x_0 for minimum P_e :

$$x_0 : \exp(-x^2) = \exp(-(x-1)^2) \Rightarrow$$

$$x_0 = \frac{1}{2}$$

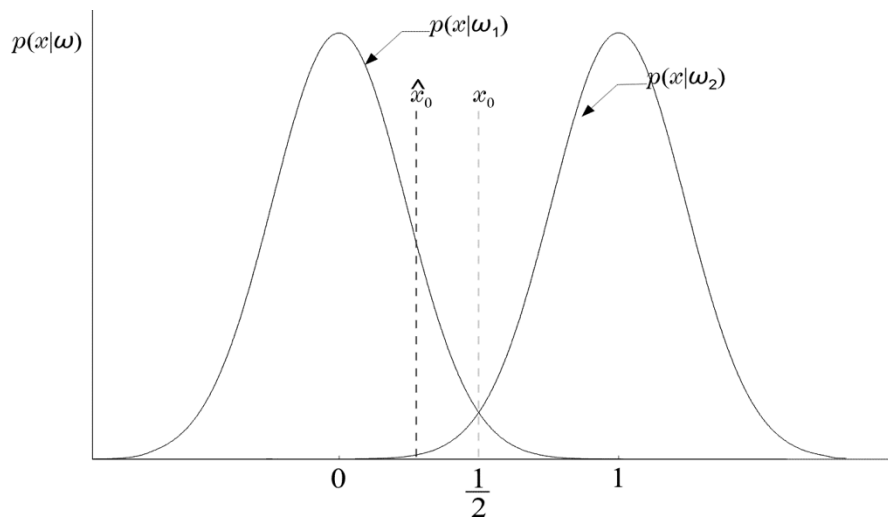
➤ Threshold \hat{x}_0 for minimum r

$$\hat{x}_0 : \exp(-x^2) = 2 \exp(-(x-1)^2) \Rightarrow$$

$$\hat{x}_0 = \frac{(1 - \ln 2)}{2} < \frac{1}{2}$$

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Thus \hat{x}_0 moves to the left of $\frac{1}{2} = x_0$
(WHY?)



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DISCRIMINANT FUNCTIONS DECISION SURFACES

❖ If R_i, R_j are contiguous: $g(\underline{x}) \equiv P(\omega_i|\underline{x}) - P(\omega_j|\underline{x}) = 0$

$$R_i : P(\omega_i|\underline{x}) > P(\omega_j|\underline{x})$$

$$\begin{array}{c} + \\ \hline - \end{array} g(\underline{x}) = 0$$

$$R_j : P(\omega_j|\underline{x}) > P(\omega_i|\underline{x})$$

is the surface separating the regions. On the one side is positive (+), on the other is negative (-). It is known as Decision Surface.

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- ❖ If $f(\cdot)$ monotonically increasing, the rule remains the same if we use:

$$\underline{x} \rightarrow \omega_i \text{ if: } f(P(\omega_i|\underline{x})) > f(P(\omega_j|\underline{x})) \quad \forall i \neq j$$

- ❖ $g_i(\underline{x}) \equiv f(P(\omega_i|\underline{x}))$ is a **discriminant function**.

- ❖ In general, discriminant functions may be defined independently of the Bayesian rule \rightarrow They lead to suboptimal solutions
- ❖ If chosen appropriately, they can be computationally more tractable.
- ❖ Moreover, in practice, they may also lead to better solutions.
- ❖ Sometimes the case when the underlying pdf's are unknown.

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THE GAUSSIAN DISTRIBUTION

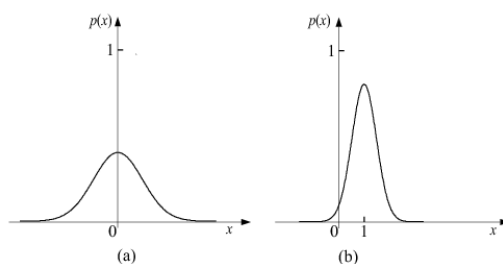
- ❖ The one-dimensional case

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

where

$$\mu \text{ is the mean value, i.e.: } \mu = E[x] = \int_{-\infty}^{+\infty} xp(x)dx$$

$$\sigma^2 \text{ is the variance, } \sigma^2 = E[(x - E[x])^2] = \int_{-\infty}^{+\infty} (x - \mu)^2 p(x)dx$$



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❖ The Multivariate (Multidimensional) case:

$$p(\underline{x}) = \frac{1}{(2\pi)^{\frac{\ell}{2}} |\underline{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (\underline{x} - \underline{\mu})^T \underline{\Sigma}^{-1} (\underline{x} - \underline{\mu})\right)$$

where $\underline{\mu}$ is the mean value, $\underline{\mu} = E[\underline{x}]$

and $\underline{\Sigma}$ is known as the covariance matrix and it is defined as:

$$\underline{\Sigma} = E[(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^T]$$

❖ **An example:** The two-dimensional case:

$$p(\underline{x}) = p(x_1, x_2) = \frac{1}{(2\pi) |\underline{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} [x_1 - \mu_1, x_2 - \mu_2] \underline{\Sigma}^{-1} \begin{bmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{bmatrix}\right)$$

$$\underline{\mu} = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} = \begin{bmatrix} E[x_1] \\ E[x_2] \end{bmatrix}, \quad \underline{\Sigma} = E\left[\begin{bmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{bmatrix} \begin{bmatrix} x_1 - \mu_1, x_2 - \mu_2 \end{bmatrix}\right] = \begin{bmatrix} \sigma_1^2 & \sigma \\ \sigma & \sigma_2^2 \end{bmatrix}$$

where $\sigma = E[(x_1 - \mu_1)(x_2 - \mu_2)]$

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BAYESIAN CLASSIFIER FOR NORMAL DISTRIBUTIONS

❖ Multivariate Gaussian pdf

$$p(\underline{x}|\omega_i) = \frac{1}{(2\pi)^{\frac{\ell}{2}} |\underline{\Sigma}_i|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (\underline{x} - \underline{\mu}_i)^T \underline{\Sigma}_i^{-1} (\underline{x} - \underline{\mu}_i)\right)$$

$\underline{\mu}_i = E[\underline{x}]$ is an $\ell \times 1$ vector, for $\underline{x} \in \omega_i$

$$\underline{\Sigma}_i = E[(\underline{x} - \underline{\mu}_i)(\underline{x} - \underline{\mu}_i)^T]$$

is the $\ell \times \ell$ covariance matrix.

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❖ $\ln(\cdot)$ is monotonic. Define:

$$\begin{aligned} \text{➤ } g_i(\underline{x}) &= \ln(p(\underline{x}|\omega_i)P(\omega_i)) = \\ &\quad \ln p(\underline{x}|\omega_i) + \ln P(\omega_i) \end{aligned}$$

$$\text{➤ } g_i(\underline{x}) = -\frac{1}{2}(\underline{x} - \underline{\mu}_i)^T \Sigma_i^{-1}(\underline{x} - \underline{\mu}_i) + \ln P(\omega_i) + C_i$$

$$C_i = -\left(\frac{\ell}{2}\right) \ln 2\pi - \left(\frac{1}{2}\right) \ln |\Sigma_i|$$

$$\text{➤ Example: } \Sigma_i = \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix}$$

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$$\begin{aligned} \text{➤ } g_i(\underline{x}) &= -\frac{1}{2\sigma^2}(x_1^2 + x_2^2) + \frac{1}{\sigma^2}(\mu_{i1}x_1 + \mu_{i2}x_2) \\ &\quad -\frac{1}{2\sigma^2}(\mu_{i1}^2 + \mu_{i2}^2) + \ln(P\omega_i) + C_i \end{aligned}$$

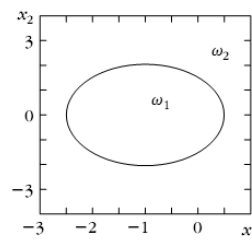
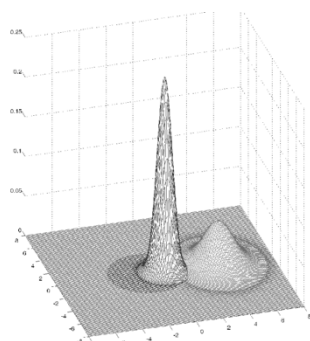
That is, $g_i(\underline{x})$ is quadratic and the surfaces

$$g_i(\underline{x}) - g_j(\underline{x}) = 0$$

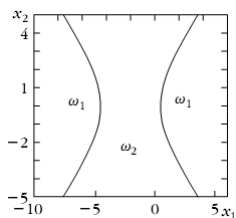
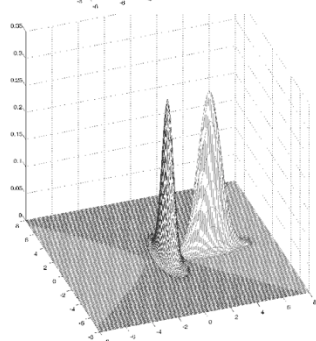
quadrics, ellipsoids, parabolas, hyperbolas, pairs of lines.

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❖ Example 1:



❖ Example 2:



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❖ Decision Hyperplanes

➤ Quadratic terms: $\underline{x}^T \Sigma_i^{-1} \underline{x}$

If ALL $\Sigma_i = \Sigma$ (the same) the quadratic terms are not of interest. They are not involved in comparisons. Then, equivalently, we can write:

$$\begin{aligned} g_i(\underline{x}) &= \underline{w}_i^T \underline{x} + w_{i0} \\ \underline{w}_i &= \Sigma^{-1} \underline{\mu}_i \\ w_{i0} &= \ln P(\omega_i) - \frac{1}{2} \underline{\mu}_i^T \Sigma^{-1} \underline{\mu}_i \end{aligned}$$

Discriminant functions are LINEAR.

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➤ Let in addition:

- $\Sigma = \sigma^2 I$. Then

$$g_i(\underline{x}) = \frac{1}{\sigma^2} \underline{\mu}_i^T \underline{x} + w_{i0}$$

- $g_{ij}(\underline{x}) = g_i(\underline{x}) - g_j(\underline{x}) = 0$
 $= \underline{w}^T (\underline{x} - \underline{x}_o)$

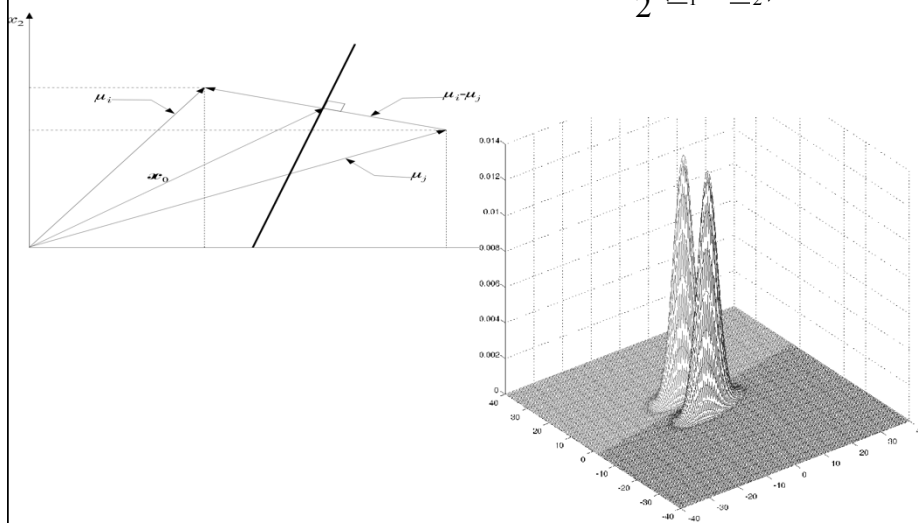
- $\underline{w} = \underline{\mu}_i - \underline{\mu}_j,$

- $\underline{x}_o = \frac{1}{2}(\underline{\mu}_i + \underline{\mu}_j) - \sigma^2 \ln \frac{P(\omega_i)}{P(\omega_j)} \frac{\underline{\mu}_i - \underline{\mu}_j}{\|\underline{\mu}_i - \underline{\mu}_j\|^2}$

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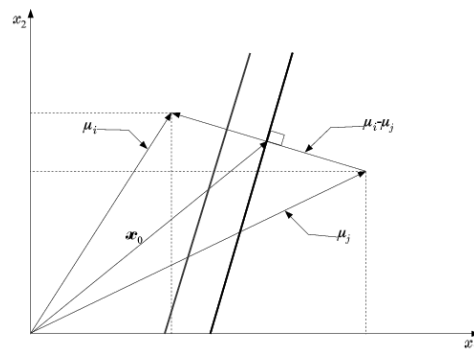
➤ Remark :

- If $p(\omega_1) = p(\omega_2)$, then $\underline{x}_o = \frac{1}{2}(\underline{\mu}_1 + \underline{\mu}_2)$



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- If $p(\omega_1) \neq p(\omega_2)$, the linear classifier moves towards the class with the smaller probability



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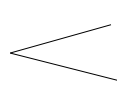
➤ Nondiagonal: $\Sigma \neq \sigma^2 I$

- $g_{ij}(\underline{x}) = \underline{w}^T (\underline{x} - \underline{x}_0) = 0$
- $\underline{w} = \Sigma^{-1}(\underline{\mu}_i - \underline{\mu}_j)$
- $\underline{x}_0 = \frac{1}{2}(\underline{\mu}_i + \underline{\mu}_j) - \ln\left(\frac{P(\omega_i)}{P(\omega_j)}\right) \frac{\underline{\mu}_i - \underline{\mu}_j}{\|\underline{\mu}_i - \underline{\mu}_j\|_{\Sigma^{-1}}^2}$

where

$$\|\underline{x}\|_{\Sigma^{-1}} \equiv (\underline{x}^T \Sigma^{-1} \underline{x})^{\frac{1}{2}}$$

➤ Decision hyperplane



not normal to $\underline{\mu}_i - \underline{\mu}_j$

normal to $\Sigma^{-1}(\underline{\mu}_i - \underline{\mu}_j)$

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❖ Minimum Distance Classifiers

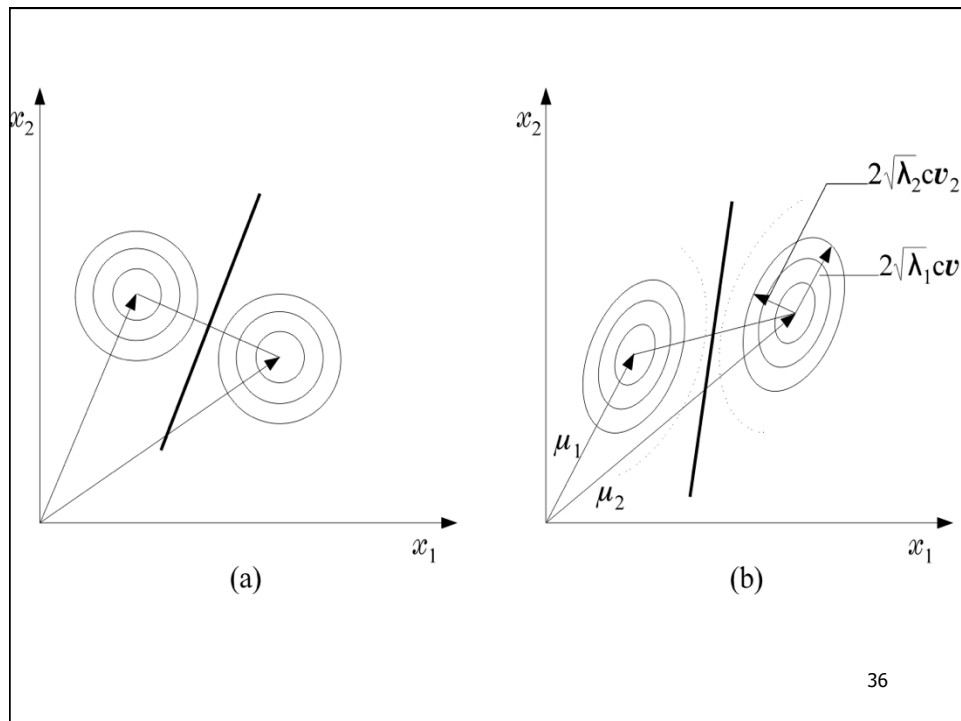
- $P(\omega_i) = \frac{1}{M}$ equiprobable
- $g_i(\underline{x}) = -\frac{1}{2}(\underline{x} - \underline{\mu}_i)^T \Sigma^{-1}(\underline{x} - \underline{\mu}_i)$
- $\Sigma = \sigma^2 I$: Assign $\underline{x} \rightarrow \omega_i$:

Euclidean Distance: $d_E \equiv \|\underline{x} - \underline{\mu}_i\|$
smaller

- $\Sigma \neq \sigma^2 I$: Assign $\underline{x} \rightarrow \omega_i$:

Mahalanobis Distance: $d_m = ((\underline{x} - \underline{\mu}_i)^T \Sigma^{-1}(\underline{x} - \underline{\mu}_i))^{\frac{1}{2}}$
smaller

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❖ Example:

Given $\omega_1, \omega_2 : P(\omega_1) = P(\omega_2)$ and $p(\underline{x}|\omega_1) = N(\underline{\mu}_1, \Sigma)$,

$$p(\underline{x}|\omega_2) = N(\underline{\mu}_2, \Sigma), \underline{\mu}_1 = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \underline{\mu}_2 = \begin{bmatrix} 3 \\ 3 \end{bmatrix}, \Sigma = \begin{bmatrix} 1.1 & 0.3 \\ 0.3 & 1.9 \end{bmatrix}$$

classify the vector $\underline{x} = \begin{bmatrix} 1.0 \\ 2.2 \end{bmatrix}$ using Bayesian classification :

$$\bullet \Sigma^{-1} = \begin{bmatrix} 0.95 & -0.15 \\ -0.15 & 0.55 \end{bmatrix}$$

- Compute Mahalanobis d_m from $\mu_1, \mu_2 : d_{m,1}^2 = [1.0, 2.2]$

$$\Sigma^{-1} \begin{bmatrix} 1.0 \\ 2.2 \end{bmatrix} = 2.952, d_{m,2}^2 = [-2.0, -0.8] \Sigma^{-1} \begin{bmatrix} -2.0 \\ -0.8 \end{bmatrix} = 3.672$$

- Classify $\underline{x} \rightarrow \omega_1$. Observe that $d_{E,2} < d_{E,1}$

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