

## BAYESIAN METHODS

In many imaging situations—for instance, image recording by film—the observation model is nonlinear of the form

$$v = f(\mathcal{K}u) + \eta \quad (8.210)$$

where  $f(x)$  is a nonlinear function of  $x$ . The a posteriori conditional density given by Bayes' rule

$$p(u|v) = \frac{p(v|u)p(u)}{p(v)} \quad (8.211)$$

is useful in finding different types of estimates of the random vector  $u$  from the observation vector  $v$ . The minimum mean square estimate (MMSE) of  $u$  is the mean of this density. The *maximum a posteriori* (MAP) and the *maximum likelihood* (ML) estimates are the modes of  $p(u|v)$  and  $p(v|u)$ , respectively. When the

observation model is nonlinear, it is difficult to obtain the marginal density  $p(v)$  even when  $u$  and  $\eta$  are Gaussian. (In the linear case  $p(u|v)$  is easily obtained since it is Gaussian if  $u$  and  $\eta$  are). However, the MAP and ML estimates do not require  $p(v)$  and are therefore easier to obtain.

Under the assumption of Gaussian statistics for  $u$  and  $\eta$ , with covariances  $\mathcal{R}_u$  and  $\mathcal{R}_\eta$ , respectively, the ML and MAP estimates can be shown to be the solution of the following equations:

$$\text{ML estimate, } \hat{u}_{ML}: \mathcal{K}^T \mathcal{D} \mathcal{R}_\eta^{-1} [v - f(\mathcal{K}\hat{u}_{ML})] = 0 \quad (8.212)$$

where

$$\mathcal{D} \triangleq \text{Diag} \left\{ \left. \frac{\partial f(x)}{\partial x} \right|_{x=\hat{u}_i} \right\} \quad (8.213)$$

and  $\hat{u}_i$  are the elements of the vector  $\hat{u} \triangleq \mathcal{K}\hat{u}_{ML}$ .

$$\text{MAP estimate, } \hat{u}_{MAP}: \hat{u}_{MAP} = \mu_u + \mathcal{R}_u \mathcal{K}^T \mathcal{D} \mathcal{R}_\eta^{-1} [v - f(\mathcal{K}\hat{u}_{MAP})] \quad (8.214)$$

where  $\mu_u$  is the mean of  $u$  and  $\mathcal{D}$  is defined in (8.213) but now  $\hat{u} \triangleq \mathcal{K}\hat{u}_{MAP}$ .

Since these equations are nonlinear, an alternative is to maximize the appropriate log densities. For example, a gradient algorithm for  $\hat{u}_{MAP}$  is

$$\hat{u}_{j+1} = \hat{u}_j - \alpha_j \{ \mathcal{K}^T \mathcal{D}_j \mathcal{R}_\eta^{-1} [v - f(\mathcal{K}\hat{u}_j)] - \mathcal{R}_u^{-1} [\hat{u}_j - \mu_u] \} \quad (8.215)$$

where  $\alpha_j > 0$ , and  $\mathcal{D}_j$  is evaluated at  $\hat{u}_j \triangleq \mathcal{K}\hat{u}_j$ .

### Remarks

If the function  $f(x)$  is linear, say  $f(x) = x$ , and  $\mathcal{R}_\eta = \sigma_\eta^2 \mathbf{I}$ , then  $\hat{u}_{ML}$  reduces to the least squares solution

$$\mathcal{K}^T \mathcal{K} \hat{u}_{ML} = \mathcal{K}^T v \quad (8.216)$$

and the MAP estimate reduces to the Wiener filter output for zero mean noise [see (8.87)],

$$\hat{u}_{MAP} = \mu_u + \mathcal{G} (v - \mu_v) \quad (8.217)$$

where  $\mathcal{G} = (\mathcal{R}_u^{-1} + \mathcal{K}^T \mathcal{R}_\eta^{-1} \mathcal{K})^{-1} \mathcal{K}^T \mathcal{R}_\eta^{-1}$ .

In practice,  $\mu_u$  may be estimated as a local average of  $v$  and  $\mu_u \approx \mathcal{K}^+ f^{-1}(\mu_v)$ , where  $\mathcal{K}^+$  is the generalized inverse of  $\mathcal{K}$ .