CWRU DSCI351-451: Week12a-p Multiple Regression

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12.2.0.1 Reading, Homeworks, Projects, SemProjects

- Homework:
 - HW6 Due Thursday, November 14th
- Readings:
 - ISLR4 today
 - ISLR6 Thursday
- Projects: We will have four 2 week EDA projects
 - You have Proj 3
- 451 SemProjects:

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- Final Exam
 - Tuesday December 18th, 12 noon to 3pm, Olin 303

12.2.0.2 Textbooks

- Peng: R Programming for Data Science
- Peng: Exploratory Data Analysis with R
- Open Intro Stats, v3
- Wickham: R for Data Science
- Hastie: Intro to Statistical Learning with R

12.2.0.3 Syllabus

12.2.0.4 Multiple Regression Practicum

Day:Date	Foundation	Practicum	Reading	Due
w1a:Tu:8/29/17	ODS Tool Chain	R, Rstudio, Git		
w1b:Th:8/31/17	Setup ODS Tool Chain	Bash, Git, Twitter	PRP4-33	HW1
W2a:Tu:9/5/17	What is Data Sci- ence	OIS:Intro2R	PRP35-64	HW1 Due
w2b:Th:9/7/17	Data Analytic Style, Git	Teatime:Intro2R, For loops	PRP65-93	HW2
w3a:Tu:9/12/17	Struct. of Data Analysis, SemProj	ISLR:Intro2R	PRP94-116	HW2 Due
w3b:Th:9/14/17	OIS3 Intro to Data	GapMinder, Dplyr, Magrittr	OI1-1.9,	
w4a:Tu:9/19/17	OIS3, Intro2Data part 2, Data	EDA: PET Degr.	EDA1-31	Proj1
w4b:Th:9/21/17	Hypothesis Testing	GGPlot2 Tutorial	EDA32-58	HW3
w5a:Tu:9/26/17	Distributions	SemProj RepOut1	R4DS1-3	HW3 Due
w5b:Th:9/28/17*	Wickham DSCI in Tidyverse	SemProj RepOut1	R4DS4-6	SemProj1,
w6a:Tu:10/3/17*	OIS Found. of Infer- ence	Inference	R4DS7-8	Proj1 Due
w6b:Th:10/5/17	Summ. Stats & Vis.	Data Wrangling	R4DS9-16 Wrangle	
w7a:Tu:10/10/17		Midterm Review		
w7b:Th:10/12/17*	MIDTERM			HW4
,	EXAM			
w8a:Tu:10/17/17	Numerical Inference	Tidy Check Explore	OIS4	HW4 Due
w8b:Th:10/19/17	Algorithms, Models	Pairwise Corr. Plots	OIS5.1-4	Proj 2, HW5
Tu:10/24	CWRU FALL BREAK		R4DS17-21 Program	
w9b:Th:10/26/17	Categorical Infer	Predictive Analytics	OIS6.1,2	
w10a:Tu:10/31/17	SemProj	SemProj	OIS7	SemProj2 HW5 Du
w10b:Th:11/2/17	Lin. Regr.	Lin. Regr.	OIS8	Proj.2 due
w11a:Tu:11/7/17	Inf. for Regression	Curse of Dim.	OIS8	Proj 3
w11b:Th:11/9/17	Model Accuracy	Training Testing	ISLR3	HW6
w12a:Tu:11/14/17	Multiple Regr.	Mul. Regr. & Pred.	ISLR4	HW6 due
w12b:Th:11/16/17	Classification		ISLR4	
w13a:Tu:11/21/17	Classification	Clustering	ISLR6	Proj 3 due
Th:11/23/17	THANKSGIVING			Proj 4
w14a:Tu:11/28/17	Big Data	Hadoop		
w14b:Th:11/30/17	InfoSec	VerisDB		SemProj3
w15a:Tu:12/5/17*	SemProj Re- portOut3			
w15b:Th:12/7/17*	SemProj Re- portOut3			Proj4
	FINAL EXAM	Tues.12/18, 12:00-3:00pm	Olin 303	SemProj4 due

Figure 1: DSCI351-451 Syllabus

12.2.0.4.1 First steps in the data analysis

```
library(psych)
if (!require("MASS")) install.packages("MASS")
## Loading required package: MASS
library(MASS)
?MASS
## No documentation for 'MASS' in specified packages and libraries:
## you could try '??MASS'
packageDescription('MASS')
## Package: MASS
## Priority: recommended
## Version: 7.3-47
## Date: 2017-02-26
## Revision: $Rev: 3474 $
## Depends: R (>= 3.1.0), grDevices, graphics, stats, utils
## Imports: methods
## Suggests: lattice, nlme, nnet, survival
## Authors@R: c(person("Brian", "Ripley", role = c("aut", "cre",
          "cph"), email = "ripley@stats.ox.ac.uk"), person("Bill",
##
##
          "Venables", role = "ctb"), person(c("Douglas", "M."),
##
          "Bates", role = "ctb"), person("Kurt", "Hornik", role =
##
          "trl", comment = "partial port ca 1998"),
          person("Albrecht", "Gebhardt", role = "trl", comment =
##
##
          "partial port ca 1998"), person("David", "Firth", role =
##
          "ctb"))
## Description: Functions and datasets to support Venables and
##
          Ripley, "Modern Applied Statistics with S" (4th edition,
          2002).
##
## Title: Support Functions and Datasets for Venables and Ripley's
          MASS
##
## LazyData: yes
## ByteCompile: yes
## License: GPL-2 | GPL-3
## URL: http://www.stats.ox.ac.uk/pub/MASS4/
## Contact: <MASS@stats.ox.ac.uk>
## NeedsCompilation: yes
## Packaged: 2017-02-26 08:41:42 UTC; ripley
## Author: Brian Ripley [aut, cre, cph], Bill Venables [ctb], Douglas
##
          M. Bates [ctb], Kurt Hornik [trl] (partial port ca 1998),
##
          Albrecht Gebhardt [trl] (partial port ca 1998), David Firth
##
          [ctb]
## Maintainer: Brian Ripley <ripley@stats.ox.ac.uk>
## Repository: CRAN
## Date/Publication: 2017-04-21 13:33:21 UTC
## Built: R 3.4.1; x86_64-pc-linux-gnu; 2017-08-27 18:42:12 UTC; unix
## -- File: /home/frenchrh/R/x86_64-pc-linux-gnu-library/3.4/MASS/Meta/package.rds
In what follows, we will use a dataset of 40 cases
```

• generated from a covariance matrix

- obtained from a subsample of real data we collected,
- which is about
 - burnout components,
 - work satisfaction,
 - work-family conflict, and
 - organizational commitment
- in hospitals.

There are six attributes in the dataset that we will analyze here;

- all are self-assessments made by nurses:
 - Commit: Commitment to their hospital (response here)
 - Exhaust: Emotional exhaustion (one of the three components of burnout)
 - Depers: Depersonalization (one of the three components of burnout)
 - Accompl: Accomplishment (one of the three components of burnout)
 - WorkSat: Work satisfaction
 - WFC: Work-family conflict

Our goal here is to understand

- how burnout dimensions and work satisfaction
- affect commitment of nurses to their hospital.

We start by

- generating the data and
- examining the correlation table
- and significance.

Make sure the matcov.txt file is in your working directory before running this code:

```
matcov <- unlist(read.csv("./data/matcov.txt", header = F))</pre>
covs <- matrix(matcov, 6, 6)</pre>
means \leftarrow c(4.47, 14.95, 4.87, 36.08, 5, 1.88)
set.seed(987)
nurses <- data.frame(mvrnorm(n = 40, means, covs))</pre>
colnames(nurses) <- c("Commit", "Exhaus", "Depers", "Accompl",</pre>
                     "WorkSat", "WFC")
corr.test(nurses)
## Call:corr.test(x = nurses)
## Correlation matrix
##
           Commit Exhaus Depers Accompl WorkSat
                                                     WFC
## Commit
             1.00 -0.64
                           -0.27
                                     0.27
                                              0.76 - 0.52
                                     0.12
## Exhaus
             -0.64
                     1.00
                            0.20
                                             -0.50 0.68
## Depers
             -0.27
                     0.20
                             1.00
                                     0.04
                                             -0.51 - 0.02
## Accompl
             0.27
                     0.12
                            0.04
                                     1.00
                                              0.23 0.15
## WorkSat
             0.76
                    -0.50
                           -0.51
                                     0.23
                                              1.00 -0.39
## WFC
             -0.52
                     0.68 -0.02
                                     0.15
                                             -0.39 1.00
## Sample Size
## [1] 40
## Probability values (Entries above the diagonal are adjusted for multiple tests.)
           Commit Exhaus Depers Accompl WorkSat WFC
##
             0.00
                     0.00
                                     0.73
## Commit
                             0.73
                                              0.00 0.01
             0.00
                                     1.00
## Exhaus
                     0.00
                             1.00
                                              0.01 0.00
                                     1.00
## Depers
             0.10
                     0.23
                             0.00
                                              0.01 1.00
## Accompl
             0.09
                     0.45
                             0.79
                                     0.00
                                              0.96 1.00
## WorkSat
             0.00
                     0.00
                             0.00
                                     0.16
                                              0.00 0.13
```

WFC 0.00 0.00 0.92 0.35 0.01 0.00

##

 $\hbox{\tt \#\#}\quad \hbox{To see confidence intervals of the correlations, print with the short=FALSE option}$

The values with a probability value lower than 0.05

• are significant by common standards.

We can see, for instance, that, in this subsample,

- commitment is significantly correlated
 - with exhaustion, work satisfaction, and work-family conflict,
- but not with depersonalization and accomplishment.

We can also see that the predictors are intercorrelated—

• that is, they share part of their variance.

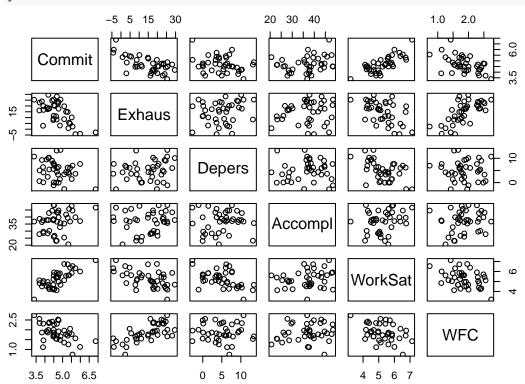
We will examine whether this constitutes a problem

• for a regression analysis later.

Let's plot the relationship

• to see if the relationships indeed seem linear:

plot(nurses)



Here, we will only comment on the scatterplots in which commitment is included.

We can see that there is visibly

- a negative linear association
 - between commitment and exhaustion and work-family conflict.
- There is visibly a positive linear relationship
 - between commitment and work satisfaction.
- Notice that there are also other relations visible on the plots,

- such as the visible relation between work- family conflict and exhaustion.
- From these scatterplots,
 - nothing in the data seems problematic for the relationships we are exploring.

12.2.0.5 Performing the multiple linear regression

We want to know if there is a relationship

• between our predictors and the response.

We first want to know

- whether the three burnout dimensions
- predict commitment to the hospital.

We create the model by

- using the formula syntax
 - as an argument in the lm() function.
- What is on the left of the tilde (~) sign
 - is the response,
- on the right are the predictors,
 - separated by a plus (+) sign:

Let's examine

- · the coefficients
- and their significance
- in the summary of the model:

```
model1 <- lm(Commit ~ Exhaus + Depers + Accompl, data = nurses)</pre>
```

The following output shows

- that exhaustion and accomplishment
 - are predictors of commitment to the hospital
 - (look at p-value under Pr(<|t|) or refer to *)
- exhaustion negatively
 - (more emotionally exhausted people are less committed)
- and accomplishment positively
 - (more accomplished people are more committed):

summary(model1)

```
##
## Call:
## lm(formula = Commit ~ Exhaus + Depers + Accompl, data = nurses)
##
## Residuals:
                       Median
                                    3Q
##
        Min
                  1Q
                                             Max
  -1.35915 -0.32590 0.02808 0.35635
##
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
               4.331261
                           0.398985 10.856 6.62e-13 ***
## (Intercept)
## Exhaus
               -0.048725
                           0.008625
                                     -5.649 2.05e-06 ***
## Depers
               -0.027053
                           0.019795
                                    -1.367 0.18021
## Accompl
                0.032923
                           0.010392
                                      3.168 0.00313 **
## ---
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4892 on 36 degrees of freedom
## Multiple R-squared: 0.55, Adjusted R-squared: 0.5125
## F-statistic: 14.67 on 3 and 36 DF, p-value: 2.116e-06
```

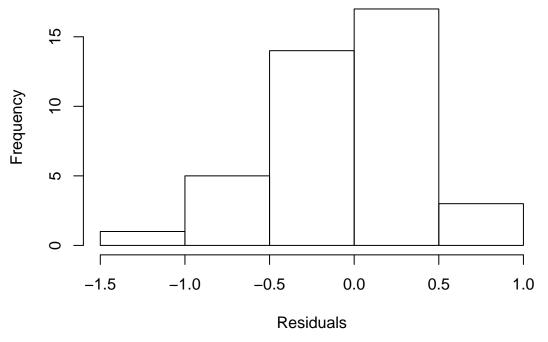
12.2.0.6 Checking for the normality of residuals

We have seen that it is important

• that residuals are normally distributed.

We can do this visually by plotting, as in the following line of code:

Histogram of residuals



From the preceding output, we might suspect

• a slight deviation from normality.

The Shapiro-Wilk test

• is a test of normality in frequentist statistics

```
shapiro.test(resid(model1))
```

```
##
## Shapiro-Wilk normality test
##
## data: resid(model1)
## W = 0.97757, p-value = 0.6001
```

We can also see that

- the p-value for F-statistic is significant (bottom of the output), and
- that 55 percent of variance (see Multiple R-squared) is predicted.

The adjusted R-squared

- considers the number of predictors
- in the calculation of its value.

It is recommended that you specify

- which value you use when reporting the results,
- or you can also report both values.

Here, we can see that

- Adjusted R-squared is just a bit lower than Multiple R-squared,
- meaning that the results are not much affected
 - by the number of predictors.

12.2.0.7 Checking for variance inflation

We also want to check whether

- there is a problem of variance inflation
- in our analysis
 - that is, whether the predictors are correlated a lot (multicollinear).
- For this purpose, we will rely on the vif() function of the HH package.
 - he function takes the lm formula as an argument:

```
# if (!require("HH")) install.packages("HH")
# install.packages("qmp")
# install.packages("Rmpfr")
# install.packages("HH")
library(HH)
## Loading required package: lattice
## Loading required package: grid
## Loading required package: latticeExtra
## Loading required package: RColorBrewer
## Loading required package: multcomp
## Loading required package: mvtnorm
## Loading required package: survival
## Loading required package: TH.data
##
## Attaching package: 'TH.data'
  The following object is masked from 'package:MASS':
##
##
       geyser
## Loading required package: gridExtra
##
## Attaching package: 'HH'
```

```
## The following object is masked from 'package:psych':
##
## logit
vif(Commit ~ Exhaus + Depers + Accompl, data = nurses)
## Exhaus Depers Accompl
## 1.054369 1.040318 1.015791
```

There are several rules-of-thumb to assess this.

- One is to consider vif values higher than 10 to be problematic,
- another is to consider a predictor as problematic
 - if the square root of the vif value is higher than 2.
- This is not the case here,
 - therefore, we consider our data to be non-multicollinear here.

12.2.0.8 Examining potential mediations and comparing models

Let's now examine whether

- including work-family conflict and work satisfaction
- permits to predict an additional part of variance.

We first will ask R to fit a second model, and

• then will compare model1 and model2 using the anova() function:

The following output shows that

- indeed the second model predicts additional variance
 - in comparison to model1
- (see the significance of the F statistic for the comparison (under Pr(>F)):

anova(model1, model2)

```
## Analysis of Variance Table
##
## Model 1: Commit ~ Exhaus + Depers + Accompl
## Model 2: Commit ~ Exhaus + Depers + Accompl + WorkSat
## Res.Df RSS Df Sum of Sq F Pr(>F)
## 1 36 8.6161
## 2 35 5.7181 1 2.898 17.738 0.0001685 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

We will now examine the second model,

• as the additional variance predicted is significantly different from 0:

summary(model2)

```
##
## Call:
## lm(formula = Commit ~ Exhaus + Depers + Accompl + WorkSat, data = nurses)
##
## Residuals:
## Min 1Q Median 3Q Max
```

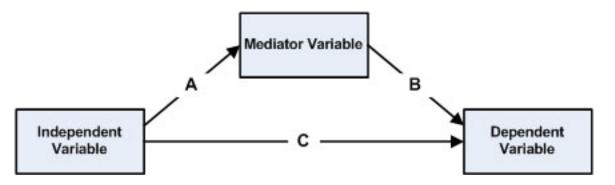


Figure 2: Mediation Model

```
-0.98119 -0.22736 -0.01279 0.26613 0.73625
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
                                      3.028 0.004598 **
## (Intercept)
               1.969672
                           0.650440
## Exhaus
               -0.029524
                           0.008460
                                     -3.490 0.001326 **
## Depers
                0.014686
                           0.019123
                                      0.768 0.447656
## Accompl
                                      1.861 0.071142 .
                0.017392
                           0.009345
## WorkSat
                0.463720
                           0.110103
                                      4.212 0.000168 ***
## ---
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4042 on 35 degrees of freedom
## Multiple R-squared: 0.7014, Adjusted R-squared: 0.6673
## F-statistic: 20.55 on 4 and 35 DF, p-value: 8.662e-09
```

This model predicts 70 percent of variance in commitment,

• which is pretty good.

We can see that work satisfaction

- is a significant predictor of commitment to the hospital,
- that the unique contribution of accomplishment is no longer significant
 - (there is therefore a potential mediation),
- and that the contribution of exhaustion
 - has been reduced when including work satisfaction in the model
 - (there is therefore a potential partial mediation).
- This might be because of a mediation of the relationship
 - between the two burnout components
 - and commitment by job satisfaction.

What is Mediation?

- In statistics, a mediation model is
 - one that seeks to identify and explain the mechanism or process
 - that underlies an observed relationship between
 - an independent variable and a dependent variable
 - via the inclusion of a third hypothetical variable,
 - known as a mediator variable (also a mediating variable, intermediary

Let's test this relationship:

```
model3 <- lm(WorkSat ~ Exhaus + Depers + Accompl, data = nurses)</pre>
summary(model3)
##
## Call:
## lm(formula = WorkSat ~ Exhaus + Depers + Accompl, data = nurses)
##
## Residuals:
##
       Min
                      Median
                  1Q
                                    30
                                            Max
## -1.57359 -0.26967 -0.06299 0.24855
                                        1.47504
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 5.09270
                           0.49899
                                   10.206 3.59e-12 ***
                                    -3.839 0.000482 ***
## Exhaus
              -0.04141
                           0.01079
## Depers
               -0.09001
                           0.02476 -3.636 0.000860 ***
## Accompl
               0.03349
                           0.01300
                                     2.577 0.014217 *
##
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.6118 on 36 degrees of freedom
## Multiple R-squared: 0.5162, Adjusted R-squared: 0.4758
## F-statistic: 12.8 on 3 and 36 DF, p-value: 7.59e-06
```

We can notice that

- 51 percent of the variance of job satisfaction
- is predicted by the burnout components.

All three burnout components

- are significantly related to work satisfaction (p < .05),
- negatively for emotional exhaustion and depersonalization
- and positively for personal accomplishment.

In order to ascertain mediation,

- we need to proceed to Sobel tests.
- The bda package contains the necessary function,
 - called mediation.test().

The Sobel Test

- is basically a specialized t test
 - that provides a method to determine whether the reduction
 - in the effect of the independent variable,
 - after including the mediator in the model,
 - is a significant reduction and
- therefore whether the mediation effect is statistically significant.

Let's try to see whether the effect of exhaustion on commitment

• is mediated by work satisfaction:

```
if (!require("bda")) install.packages("bda")

## Loading required package: bda

library(bda)
mediation.test(nurses$WorkSat,nurses$Exhaus,nurses$Commit)
```

$$z = \frac{a*b}{\sqrt{(b^2*s_a^2 + a^2*s_b^2)}}$$

Figure 3: Sobel

Sobel Aroian Goodman ## z.value -2.972270400 -2.936471185 -3.009411683 ## p.value 0.002956062 0.003319697 0.002617542

In the following output, under Sobel,

- we can see that p.value is significant,
- as the presence of work satisfaction in the model
 - decreases the effect of exhaustion,
- that work satisfaction is significant
 - even though exhaustion is present in the model,
- and that, because the Sobel test is significant,
 - we can confirm that there is indeed
 - a partial mediation of the effect of exhaustion
 - on commitment by work satisfaction.

In other words,

- exhaustion decreases work satisfaction,
- and in turn, work satisfaction increases commitment.

The value resulting from the Sobel test follows a z distribution.

In order to obtain this value,

- the slope coefficients of the predictor regressed on the mediator (a)
 - are multiplied by the slope coefficient of the mediator
 - regressed on the response (b).

This value is then divided by the square root of b squared

- multiplied by the squared standard error of a
- plus a squared multiplied by the squared standard error of b. The formula is as follows:

Showing this is important, as very often,

- analysts include dozens or hundreds of predictors in their models
- without taking into consideration that the included predictors
 - could themselves be related to each other.

Readers are therefore advised to check

- for meaningful relationships between the attributes
- they intend to include as predictors in regression analyses
- before drawing conclusions on the final model!

12.2.0.9 Predicting new data

A particularly interesting use of regression

• is to examine how well a model predicts new data.

This is easily achieved in R.

We will first build the dataset named nurses2

• in the same way we did for the first dataset:

The following output shows that the correlation

- between the predicted values
- and the real values is 0.5766194.

This value is significant and might seem pretty good at first sight:

```
cor.test(predicted, nurses2$Commit)
```

```
##
## Pearson's product-moment correlation
##
## data: predicted and nurses2$Commit
## t = 4.3506, df = 38, p-value = 9.848e-05
## alternative hypothesis: true correlation is not equal to 0
## 95 percent confidence interval:
## 0.3231561 0.7528925
## sample estimates:
## cor
## 0.5766194
```

Let's square this value

- to know how much of the variance in the commitment of the individuals
- of the second sample is predicted by the model:

The output is 33.24899.

This means only 33 percent of the variance in commitment

- is predicted by the model,
- compared to 55 percent in the training data!

Now, we can also compute the residuals:

```
residuals_test <- nurses2$Commit - predicted
```

We are now able to compute the F value for our model.

Statistical Tests

t-statistic

F-test

We have seen that the F value is used

• to assess the overall significance of the model.

In our case, the F value is obtained as follows:

- 1) First, we need to know the number of degrees of freedom for the model;
- this is equal to the number of predictors we have, which is 3.
- We also need the degrees of freedom for the error;
 - this is the number of observations
 - minus the degrees of freedom of the model, minus 1.
- 2) We then compute the sum of squares for the model
- as the sum of squared differences
 - between the predicted values
 - and the mean of the response.

The sum of squares for the error is obtained as

- the sum of the squared differences
- between the observed and the predicted values.
- 3) We then compute the mean squares for the model
- as the sum of squares for the model
- divided by the degrees of freedom for the model.

We compute the mean squares for the error

- as the sum of squares for the error
- divided by the degrees of freedom for the error.
- 4) Finally, we obtain the F-statistic
- by dividing the means squares for the model
- by the mean squares for the error.

The following function does just that:

```
ComputeF <- function(predicted, observed, npred) {
   DFModel <- npred # the number of predictors
   DFError <- length(observed) - DFModel - 1
   SSModel <- sum((predicted - mean(observed))^2)
   SSError <- sum((observed - predicted)^2)
   MSModel <- SSModel / DFModel
   MSError <- SSError / DFError
   F <- MSModel / MSError
   F
}

ComputeF(unlist(model1[5]), nurses$Commit, 3)</pre>
```

```
## [1] 14.66868
```

```
ComputeF(predicted, nurses2$Commit, 3)
```

```
## [1] 10.4842
```

The outputted F value is 10.4842.

We can test this value using the following line of code.

The output shows that the threshold F value

• at a ceiling of 0.05 on the F distribution for our model is 2.866266:

```
qf(.95, df1 = 3, df2 = 36)
```

```
## [1] 2.866266
```

We can therefore, trust that our model significantly predicts new data.

12.2.0.10 Robust regression

In the example datasets that we used in this section,

- we have seen that some observations might threaten
- the reliability of our results,
 - because of the deviations of their residuals from a normal distribution.

The Shapiro-Wilk test performed on the residuals of model1 (nurses dataset)

- has shown that the distribution of the residuals
- was not significantly different from a normal distribution.

However, let's be particularly cautious

• and analyze the same data using robust regression.

As we mentioned earlier, robust regression

- does not require the residuals to be normally distributed,
- and therefore, fits our purpose.

We will not explore the algorithm.

- For details about this, the reader can consult
 - Robust Regression in R by Fox and Weisberg, in readings.
- Here, we simply perform robust regression using the rlm() function
 - of the MASS package.

Let's first install and load it:

```
model1.rr <- rlm(Commit ~ Exhaus + Depers + Accompl, data = nurses)
summary(model1.rr)</pre>
```

```
## Call: rlm(formula = Commit ~ Exhaus + Depers + Accompl, data = nurses)
## Residuals:
##
          Min
                      1Q
                             Median
                                             3Q
                                                       Max
## -1.4052046 -0.3233886 -0.0003426 0.3734567
                                                1.0108386
##
## Coefficients:
##
                       Std. Error t value
               Value
## (Intercept)
               4.3602 0.3849
                                  11.3271
## Exhaus
               -0.0518
                        0.0083
                                  -6.2306
## Depers
               -0.0279
                        0.0191
                                  -1.4602
## Accompl
                0.0338
                        0.0100
                                    3.3676
##
## Residual standard error: 0.5536 on 36 degrees of freedom
```

You might notice that the output of rlm() is laconical

• in comparison to the output of lm().

• There are no p-values provided, no R-squared values, no F test.

This makes the use of rlm() quite unpractical,

• as the user will have to compute them by hand.

There is so much controversy on how to do it

• that the computations in other software packages are currently questioned!

The reader interested in computing the robust R-squared

- can read the paper
 - A robust coefficient of determination for regression
 - by Renaud and Victoria-Feser (2010), which is in readings.

For our example,

- it seems that the results using lm() and rlm() are pretty similar
- (see the output of the preceding summary of model1).

Therefore, relying on lm() is advised here.

However, if you want to be really sure,

• why not try bootstrapping.

12.2.0.11 Bootstrapping (Advanced topic)

The bootstrap is covered in ISLR Chapter 5 Resampling Methods, in Section 5.2.

The principle of (nonparametric) bootstrapping

- is to create a number of sample K of size N
 - drawn with replacement from the original sample,
- where N is the original sample size.

The parameters are estimated for each sample separately.

This allows computing their confidence intervals,

• a measure of the variability of the parameters.

Apart from making deviations from normal distributions less problematic,

- using bootstrapping is useful for samples
- that have a small number of observations
 - (less than 100), as with ours.

Bootstrapping is easily performed using several functions in R

• for instance, the boot() function in the boot package.

But let's have a little fun and perform bootstrapping ourselves, 2,000 times.

We will first generate the samples and obtain the estimates.

We then display the estimates for the first six samples

• (rounded to the third decimal place):

```
ret <- data.frame(matrix(nrow = 0, ncol = 6))
set.seed(567)
for (i in 1:2000) {
   data <- nurses[sample(nrow(nurses), 40, replace = T),]
   model_i <- lm(Commit ~ Exhaus + Depers + Accompl,</pre>
```

```
data = data)
  ret <- rbind(ret,c(coef(model_i),summary(model_i)$r.square,</pre>
                    summary(model_i)$fstatistic[1]))
}
names(ret) <- c("Intercept", "Exhaus", "Depers",</pre>
               "Accomp", "R2", "F")
round(head(ret), 3)
     Intercept Exhaus Depers Accomp
##
         4.080 -0.037 -0.055 0.041 0.585 16.928
## 1
## 2
         4.196 -0.052 -0.048 0.040 0.694 27.273
## 3
         5.054 -0.052 -0.047 0.022 0.736 33.416
         4.103 -0.041 -0.042 0.037 0.545 14.373
         4.663 -0.041 -0.022 0.022 0.454 9.980
## 5
         4.525 -0.049 -0.035 0.029 0.497 11.874
## 6
set.seed(567)
sample(nrow(nurses), 40, replace = T)
## [1] 30 36 26 20 11 10 3 21 24 22 14 11 15 24 1 3 21 30 2 12 6 21 9
qnorm(0.975)
## [1] 1.959964
CIs <- data.frame(matrix(nrow = ncol(ret), ncol = 2))</pre>
for (j in 1:ncol(ret)) {
 M <- mean(ret[,j])</pre>
  SD <- sd(ret[,j])</pre>
  lowerb <- M - (1.96 * (SD / sqrt(2000)))
  upperb <- M + (1.96 * (SD / sqrt(2000)))
  CIs[j,1] <- round(lowerb,3)</pre>
 CIs[j,2] <- round(upperb,3)</pre>
names(CIs) <- c("95% C.I.lower bound", "95% C.I.upper bound")</pre>
rownames(CIs) <- colnames(ret)</pre>
CIs
##
             95% C.I.lower bound 95% C.I.upper bound
## Intercept
                                               4.325
                           4.297
                          -0.048
## Exhaus
                                              -0.048
## Depers
                          -0.029
                                              -0.027
## Accomp
                           0.033
                                               0.033
## R2
                           0.558
                                               0.570
## F
                          18.179
                                              19.139
```

The confidence intervals

- encompass all the values
- between the lower and upper bounds.

We can see that no confidence interval contains 0,

- meaning that, with a 95 percent threshold,
- values reported are statistically different from 0
 - (more correctly put, there is only a 5 percent chance
 - of observing values inside these bounds

- if the true value of the parameters in the population is 0).

So we conclude that

- bootstrapped coefficients are different from 0,
- as is the multiple R-squared value.

As you might have noticed,

- the value to which to compare the confidence intervals for F is not 0,
- but a value that depends upon the degrees of freedom.

We computed this value earlier and it was 2.866266.

As the confidence interval for F does not include this value,

- we can be assured that the bootstrapped model
- predicts a significant part of variance.

12.2.0.12 Summary

We examined how to develop functions

- that perform simple regression analyses,
- and how to multiply regression in R
- using a real life example.

We have examined the importance of significance tests for regression,

- and have briefly discussed
 - robust regression
 - and bootstrapping.

Note that, when data about the predictors and the response

- are collected simultaneously,
- causation cannot be established.

In order to ascertain causation,

- data must be collected longitudinally
- that is, the predictors before the response.

12.2.0.13 Links

12.2.0.13.1 Learning Predictive Analytics with R, Eric Mayor, Packtpub 2015