

Back to stochastic actor-oriented models

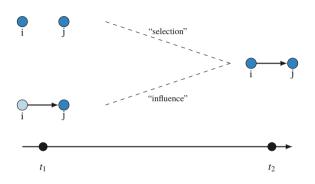
- Model for network panel data observation of a network at discrete time points
- Developed to explain the micro-mechanism that led to the network evolution
- Extended to explain the micro-mechanism that led to the co-evolution of networks and behaviours

The only model enabling to distinguish selection from influence mechanisms

In the following, we denote by X the network and by Z the behaviour



Competing explanatory stories



- ► Study of influence requires the consideration of selection and vice versa
- ▶ Only longitudinal data allows distinguishing between selection and influence



Model formulation: network dynamics

- Rate function $\tau_m^{[X]}$: waiting time between opportunities for a network change rate describing the average number of opportunities of change between $x(t_{m-1})$ and $x(t_m)$
- Evaluation function: choice of actor i

$$f_i^{[X]}(\beta, x', z) = \sum_{k=1}^K \beta_k^{[X]} s_{ik}^{[X]}(x', z)$$

Conditional on i

$$p_{(x',z)}^{[X]} = \frac{\exp(f_i^{[X]}(\beta, x', z))}{\sum_{x''} \exp(f_i^{[X]}(\beta, x'', z))}$$

with $x' = x(i \leadsto j)$



Model formulation: behavioural dynamics

- Rate function $\tau_m^{[Z]}$: waiting time between opportunities for a behavioural change rate describing the average number of opportunities of change between $z(t_{m-1})$ and $z(t_m)$
- Evaluation function: choice of actor i

$$f_i^{[Z]}(\beta, x, z') = \sum_{k=1}^K \beta_k^{[Z]} s_{ik}^{[Z]}(x, z')$$

Conditional on i

$$p_{(x,z')}^{[Z]} = \frac{\exp\left(f_i^{[Z]}(\beta, x, z')\right)}{\sum_{z''} \exp\left(f_i^{[Z]}(\beta, x, z'')\right)}$$

with
$$z' = \{z, z + 1, z - 1\}$$

For convenience, we omit v and w from the arguments of s_{ik}



Learning objectives of today

- ► Learn how to simulate from SAOMs
- ▶ Understand the Method of Moments (MoM) estimation
- ▶ Understand the Maximum likelihood estimation
- ▶ Learn about the stochastic approximation algorithm in SAOMs



Simulate the longitudinal panel data $x(t_1), \ldots, x(t_M)$

For simplicity, we assume M=2 (extension to more than two waves is straightforward)

Input

```
x(t_1) = network at time t_1
```

 τ = rate parameter

$$\beta = (\beta_1, \dots, \beta_k)$$
 = vector of evaluation function parameters

Output

```
x^{\text{sim}}(t_2) = network at time t_2
```



Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{\text{sim}}(t_2)
t \leftarrow 0
x \leftarrow x(t_1)
while condition = TRUE do
     dt \sim Exp(n\tau)
     i \sim \text{Uniform}(1, \dots, n)
     j \sim Multinomial(p_{i1}, \dots, p_{in})
     if i \neq j then
      x \leftarrow x(i \rightsquigarrow j)
     else
      x^{\text{sim}}(t_2) \leftarrow x
return x^{sim}(t_2)
```

t = time

dt = holding time between consecutive opportunities to change \sim = generated from



$$n = 4$$

$$\tau = 1.5$$

$$\beta = (\beta_{out}, \beta_{rec}, \beta_{trans})$$

$$= (-1, 0.5, -0.25)$$

ETH zürich

Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{\text{sim}}(t_2)
t \leftarrow 0
x \leftarrow x(t_1)
while condition = TRUE do
      dt \sim Exp(n\tau)
      i \sim \text{Uniform}(1, \dots, n)
      i \sim \text{Multinomial}(p_{i1}, \dots, p_{in})
      if i \neq j then
       x \leftarrow x(i \rightsquigarrow i)
      else
      \perp x \leftarrow x
      t \leftarrow t + dt
x^{\text{sim}}(t_2) \leftarrow x
return x^{sim}(t_2)
```

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Generate the time elapsed between t_1 and the first opportunity to make a change

The more intuitive way to generate dt is:

generate the waiting time for each actor i

$$t_i \sim \operatorname{Exp}(\tau)$$

$$- dt = \min_{1 \le i \le n} \{t_i\}$$

But this requires the generation of n numbers...



Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{\text{sim}}(t_2)
t \leftarrow 0
```

 $x \leftarrow x(t_1)$

while condition = TRUE **do**

 $dt \sim Exp(n\tau)$

else $x \leftarrow x$

 $t \leftarrow t + dt$

 $x^{\text{sim}}(t_2) \leftarrow x$

return $x^{sim}(t_2)$

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Generate the time elapsed between t_1 and the first opportunity for a change

To avoid the generation of n numbers, we use the following result:

If

$$T_i \sim \operatorname{Exp}(\tau_i), \quad 1 \leq i \leq n$$

and T_1, \ldots, T_n are mutually independent, then

$$dt = \min\{T_1, \dots, T_n\} \sim \operatorname{Exp}(\sum_{i=1}^n \tau_i)$$

e.g. dt = 0.0027

ETH zürich

Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{\text{sim}}(t_2)
t \leftarrow 0
x \leftarrow x(t_1)
while condition = TRUE do
     dt \sim Exp(n\tau)
     i \sim \text{Uniform}(1, \dots, n)
     i \sim \text{Multinomial}(p_{i1}, \dots, p_{in})
     if i \neq j then
      x \leftarrow x(i \rightsquigarrow j)
     else
      x^{\text{sim}}(t_2) \leftarrow x
return x^{sim}(t_2)
```

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Select the actor i who has the opportunity to change

e.g. i = 1





```
Algorithm: Network evolution
```

```
Input: x(t_1), \tau, \beta, n
Output: x^{sim}(t_2)
t \leftarrow 0
x \leftarrow x(t_1)
while condition = TRUE do
     dt \sim Exp(n\tau)
     i \sim \text{Uniform}(1, \dots, n)
     i \sim \text{Multinomial}(p_{i1}, \dots, p_{in})
     if i \neq j then
      x \leftarrow x(i \rightsquigarrow j)
     else
      t \leftarrow t + dt
```

 $x^{\text{sim}}(t_2) \leftarrow x$

return $x^{\text{sim}}(t_2)$

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Select j

| $i \rightarrow j$ | f_i | p_{ij} |
|-------------------|-------|----------|
| $1 \rightarrow 1$ | -1.75 | 0.15 |
| $1 \rightarrow 2$ | -1.00 | 0.31 |
| $1 \rightarrow 3$ | -3.25 | 0.03 |
| $1 \rightarrow 4$ | -0.5 | 0.51 |
| | | |

ETH zürich

Algorithm: Network evolution

```
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     dt \sim Exp(n\tau)
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      x \leftarrow x(i \rightsquigarrow j)
     else
      x^{\text{sim}}(t_2) \leftarrow x
return x^{sim}(t_2)
```

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Select j

e.g. j = 4





Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{\text{sim}}(t_2)
t \leftarrow 0
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while condition = TRUE do
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     if i \neq j then
      x \leftarrow x(i \rightsquigarrow j)
     else
      x^{\text{sim}}(t_2) \leftarrow x
return x^{sim}(t_2)
```

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Select j e.g. j = 4





Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{\text{sim}}(t_2)
t \leftarrow 0
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     if i \neq j then
      x \leftarrow x(i \rightsquigarrow j)
     else
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return x^{sim}(t_2)
```

t = time

dt = holding time between consecutive opportunities to change \sim = generated from

Select j

e.g. j = 1





Algorithm: Network evolution

```
Input: x(t_1), \tau, \beta, n
Output: x^{sim}(t_2)
t \leftarrow 0
x \leftarrow x(t_1)
while condition = TRUE do
      dt \sim Exp(n\tau)
      i \sim \text{Uniform}(1, \dots, n)
      i \sim \text{Multinomial}(p_{i1}, \dots, p_{in})
      if i \neq j then
      x \leftarrow x(i \rightsquigarrow j)
     else
      \perp x \leftarrow x
      t \leftarrow t + dt
x^{\text{sim}}(t_2) \leftarrow x
return x^{sim}(t_2)
t = time
```

e.g. t = 0 + 0.0027

dt = holding time between consecutive opportunities to change \sim = generated from

FTH zürich

Two different stopping rules:

1. *Unconditional* simulation:

the simulation of the network evolution carries on until a predetermined time length has elapsed (usually until t=1)



Two different stopping rules:

- 1. *Unconditional* simulation:
 - the simulation of the network evolution carries on until a predetermined time length has elapsed (usually until t = 1)
- 2. *Conditional* simulation on the observed number of changes: the simulation runs on until

$$\sum_{\substack{i,j=1\\ \text{B}\neq j}}^{n} \left| x_{ij}^{obs}(t_2) - x_{ij}(t_1) \right| = \sum_{\substack{i,j=1\\ \text{B}\neq j}}^{n} \left| x_{ij}^{\text{sim}}(t_2) - x_{ij}(t_1) \right|$$

Simulating the co-evolution of networks and behaviour

Simulate the longitudinal panel data $(x, z)_{t_1}, \dots, (x, z)_{t_M}$

For simplicity, we assume M=2 (extension to more than two waves is straightforward)

Input

```
n: number of actors (given) \tau^{[X]}, \tau^{[Z]}: network and behaviour rate parameters (given) \beta^{[X]} = (\beta_1^{[X]}, \dots, \beta_K^{[X]}): network evaluation function parameters (given) \beta^{[Z]} = (\beta_1^{[Z]}, \dots, \beta_K^{[Z]}): behavioural evaluation function parameters (given) (x, z)(t_1): network and behaviour at time t_1 (given)
```

Output

 $(x,z)^{\text{sim}}(t_2)$: network and behaviour at time t_2



Simulating the co-evolution of networks and behavior

Algorithm: Network-behavior co-evolution

```
Input: x(t_1), z(t_1), \tau^{[X]}, \tau^{[Z]}, \beta^{[X]}, \beta^{[Z]}, n
Output: x^{\text{sim}}(t_2), z^{\text{sim}}(t_2)
t \leftarrow 0; x \leftarrow x(t_1); z \leftarrow z(t_1)
while condition=TRUE do
       dt^{[X]} \sim Exp(n\tau^{[X]}); dt^{[Z]} \sim Exp(n\tau^{[Z]})
       if min{dt^{[X]}, dt^{[Z]}} = dt^{[X]} then
              i \sim \text{Uniform}(1, \ldots, n)
              j \sim Multinomial(p_{i1}, \ldots, p_{in})
             if i \neq i then
            x \leftarrow x(i \leadsto j)
           t \leftarrow t + dt^{[X]}
       else
              i \sim \text{Uniform}(1, \dots, n)
              z' \sim \text{Multinomial}(p_{z(z-1)}, p_{zz}, p_{z(z+1)})
              if z \neq z' then
x^{\text{sim}}(t_2) \leftarrow x; z^{\text{sim}}(t_2) \leftarrow z
return x^{\text{sim}}(t_2), z^{\text{sim}}(t_2)
```

Parameter estimation

Given the longitudinal data

$$(x,z)_{t_1},\ldots,(x,z)_{t_M}$$

we need to estimate 2(M-1)+K+W parameters

 \blacktriangleright M-1 rate parameters for the network rate function

$$au_1^{[X]}, \ldots, \, au_{M-1}^{[X]}$$

ightharpoonup M-1 rate parameters for the behaviour rate function

$$\tau_1^{[Z]}, \ldots, \tau_{M-1}^{[Z]}$$

► *K* and *W* parameters for the network and behavioural evaluation function, respectively

$$\begin{split} f_i^{[X]}(\beta^{[X]}, x', z) &= \sum_{k=1}^K \beta_k^{[X]} s_{ik}^{[X]}(x', z) \\ f_i^{[Z]}(\beta^{[Z]}, x, z') &= \sum_{w=1}^W \beta_w^{[Z]} s_{iw}^{[Z]}(x, z') \end{split}$$



Estimating the parameters of SAOMs

We would like to estimate

$$\theta = (\tau^{[X]}, \tau^{[Z]}, \beta^{[X]}, \beta^{[Z]}), \quad \theta \in \mathbb{R}^P, \quad P = 2(M-1) + K + W$$

- 1. Method of Moments (MoM) and its generalization (GMoM)
- 2. Maximum-likelihood estimation
- 3. Bayesian estimation

Estimating the parameters of SAOMs

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- 1. **Method of Moments (MoM)** and its generalization (GMoM)
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Estimating the parameters of SAOMs: MoM

$$\theta = (\tau^{[X]}, \tau^{[Z]}, \beta_1^{[X]}, \dots, \beta_K^{[X]}, \beta_1^{[Z]}, \dots, \beta_W^{[Z]}), \quad \theta \in \mathbb{R}^P, \quad P = 2(M-1) + K + W$$

- 1. Find *P* statistics $s(X) = (s_1(X, Z), \dots, s_p(X, Z))$ i.e. *P* variables that can be calculated from the network
- 2. Set the expected value of s(X, Z) equal to its sample counterpart s(x, z)

$$E_{\theta}[s(X,Z)] = s(x,z)$$

3. Solve the resulting system of equations with respect to θ .



1. Defining the statistics

The statistics s(X) must be sensitive to the parameter θ in the sense that

$$\frac{\partial E_{\theta}(s_p(x,z))}{\partial \theta_p} > 0$$

► Rate function:

- ightharpoonup models the frequency at which actors get opportunities for a change
- ▶ The higher τ_m , the higher the number of changes between t_{m-1} and t_m
- \triangleright Relevant statistics for τ are

$$s_{\tau}^{[X]}(X(t_m),X(t_{m-1})) = \sum_{ij} \left| X_{ij}(t_m) - X_{ij}(t_{m-1}) \right|$$

$$s_{\tau}^{[Z]}(Z(t_m),Z(t_{m-1})) = \sum_{\cdot} |Z_i(t_m) - Z_i(t_{m-1})|$$

Why do we sum over m?



1. Defining the statistics

► Evaluation function:

for the parameter β_k in

$$f_i(\beta^{[X]}, x, z)$$
 and $f_i(\beta^{[Z]}, x, z)$

- the higher β_k , the more strongly all actors strive after a high value of $s_{ik}(x,z)$
- this leads to the statistics

$$\begin{split} s_k^{[X]}(X(t_m), Z(t_{m-1})) &= \sum_i s_{ik}^{[X]} \left(X(t_m), Z(t_{m-1}) \right) \\ s_k^{[Z]}(X(t_{m-1}), Z(t_m)) &= \sum_i s_{ik}^{[Z]} \left(X(t_{m-1}), Z(t_m) \right) \end{split}$$

2. Setting the moment equations

▶ Rate function

$$E_{\theta}\left[s_{\tau}^{[X]}(X(t_m), X(t_{m-1})) \mid (X, Z)_{t_{m-1}}\right] = \sum_{ij} \left|x_{ij}(t_m) - x_{ij}(t_{m-1})\right| \quad m = 2, \dots, M$$

$$E_{\theta}\left[s_{\tau}^{[Z]}(Z(t_m), Z(t_{m-1})) \mid (X, Z)_{t_{m-1}}\right] = \sum_{i} \left|z_i(t_m) - z_i(t_{m-1})\right| \quad m = 2, \dots, M$$

► Evaluation function

$$\sum_{m=2}^{M} E_{\theta} \left[s_{k}^{[X]}(X(t_{m}), Z(t_{m-1}) \mid (X, Z)_{t_{m-1}} \right] = \sum_{m=2}^{M} s_{k}^{[X]}(x(t_{m}), z(t_{m-1})$$

$$\sum_{m=2}^{M} E_{\theta} \left[s_{k}^{[Z]}(X(t_{m-1}), Z(t_{m})) \mid (X, Z)_{t_{m-1}} \right] = \sum_{m=2}^{M} s_{k}^{[Z]}(x(t_{m-1}), z(t_{m}))$$



3. Solving the moment equation

The system of moment equation

$$E_{\theta}[s(X,Z)] = s(x,z)$$

cannot be solved by analytical or the usual numerical procedures, because

$$E_{\theta}[s(X,Z)]$$

cannot be calculated explicitly.

However, the solution can be approximated by the Robbins-Monro (1951) method for stochastic approximation

an iterative stochastic algorithm that attempts to find zeros of functions which cannot be analytically computed



► Iterative algorithm with step

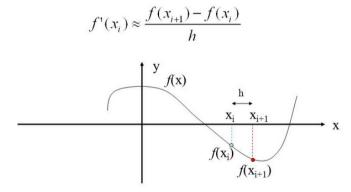
$$\hat{\theta}_{i+1} = \hat{\theta}_i - \alpha_i \ D^{-1} \left[\ s^{(i)}(x,z) - s(x,z) \ \right]$$

- ► The algorithm has three phases (Snijders (2001)):
 - 1. Phase 1: given an initial value θ_0 , n_1 simulations of the network evolution is used to approximate the matrix of the first order derivatives D and update θ
 - 2. Phase 2: divided into a few sub-phases characterized by a decreasing value of α_i
 - A step in each sub-phase consists in generating one network evolution and updating the value of $\hat{\theta_i}$ via the Robbins-Monro step
 - At the end of the sub-phase α_i is halved and the estimate for θ is the average of the values $\hat{\theta}_i$
 - This estimate is used as the initial value of θ for the next sub-phase
 - 3. Phase 3: is used to check the convergence of the algorithm and compute the standard errors of the estimates based on a large number of network trajectories simulated from the value of θ obtained at the end of phase 2



Phase 1: Approximation of $D = \frac{\partial E_{\theta}[S]}{\partial \theta}$

► Finite difference method:





Phase 1: Approximation of $D = \frac{\partial E_{\theta}[S]}{\partial \theta}$

- Finite difference method: at each step j, $j = 1, ..., n_1$,
 - \blacktriangleright simulate one network evolution from a SAOM with $\theta = \theta_0$ and compute the vector of statistics S_{i0}
 - define the vectors $\theta_{kj} = \theta_0 + \varepsilon_k e_k$ e_k the k-th unit vector in P dimensions, ε_k suitable constants
 - ightharpoonup simulate from θ_{kj} , compute the statistic S_{jk} and the ratio

$$d_{jk} = \frac{S_{jk} - S_{j0}}{\varepsilon_k}$$

for each element of θ

D is approximated by

$$\hat{D} = \frac{1}{n_1} \sum_{i=1}^{n_1} d_i$$

with
$$d_j = (d_{j1}, \ldots, d_{jp})$$



Phase 1: Approximation of $D = \frac{\partial E_{\theta}[S]}{\partial \theta}$

- Score function method
 - ► Score function: first order derivative of the log-likelihood

$$SF = \frac{\partial \ell(\theta)}{\partial \theta}$$

► We can prove that

$$\frac{\partial E_{\theta}[(S-s)]}{\partial \theta} = E_{\theta}[(S-s) \cdot SF]$$

and approximate D using the score function method: (Schweinberger and Snijders (2007))

$$\hat{D} = \frac{1}{n_1} \sum_{j=1}^{n_1} (S_j - s) SF_j$$

► The use of the score function method provide a better approximation



Phase 1: Update θ

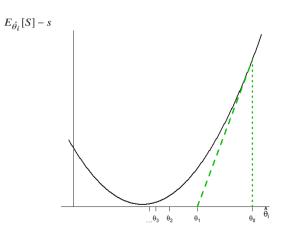
Newton Raphson step

$$\hat{\theta}_1 = \theta_0 - \hat{D}^{-1} \left(\overline{s} - s \right)$$

with

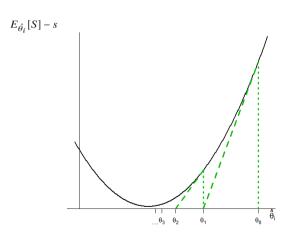
$$\bar{s} = \frac{1}{n_1} \sum_{j=1}^{n_1} S_{j0}$$

Phase 2: the Robbins-Monro step

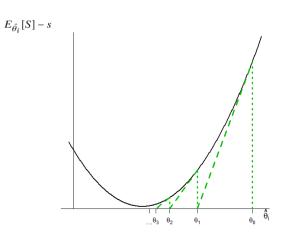




Phase 2: the Robbins-Monro step

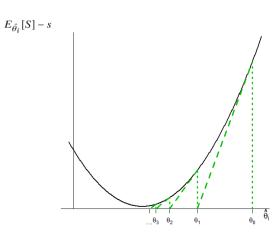


Phase 2: the Robbins-Monro step





Phase 2: the Robbins-Monro step





Phase 2: Sub-phase h

Each sub-phase consists of the following steps:

- Generate one network evolution from the current value of $\hat{\theta}$
- ▶ Update θ

$$\hat{\theta}_{i+1} = \hat{\theta}_i - \alpha_h \ \tilde{D}^{-1} \left[s^{(i)}(x, z) - s(x, z) \right], \quad \tilde{D} = \operatorname{diag}(\hat{D})$$

The steps are repeated until the maximum number of steps n_{2h}^* is reached or

$$\sum_{i} (s_k^{(i-1)}(x,z) - s_k(x,z))(s_k^{(i)}(x,z) - s_k(x,z)) < 0, \ \forall k$$

At the end of the sub-phase:

$$\hat{\theta} = \frac{1}{n_{2h}} \sum_{i}^{n_{2h}} \hat{\theta}_i \quad \alpha_{h+1} = \alpha_h/2$$

with n_{2h} the number of steps performed in sub-phase h



Phase 3: Checking convergence

- ▶ The value $\hat{\theta}$ returned by the last sub-phase in phase 2 is the estimate for θ
- ► A large number n_3 of network evolution is simulated given $\hat{\theta}$
- \blacktriangleright For each simulation, the vector of statistics S_i is computed

$$t\text{-conv}_k = \frac{\overline{s}_k - s_k}{\text{s.d}(s_k)}$$

maximum t-ratio =
$$(\overline{s} - s)'\Sigma^{-1}(\overline{s} - s)$$

with

$$\overline{s} = \frac{1}{n_3} \sum_{j=1}^{n_3} s_j, \quad j = 1, \dots, n_3$$

and

$$\Sigma = \text{Cov}[S] = \frac{1}{n_3} ss' - \overline{s}'\overline{s}, \quad j = 1, \dots, n_3$$

Phase 3: Approximate the covariance matrix of the MoM estimator

We approximate the covariance matrix of the MoM estimator using the delta method

$$Cov[\hat{\theta}] \approx \hat{D}^{-1} \Sigma \hat{D}^{-1}$$

with

- \hat{D} the matrix of first order derivatives approximated as in phase 1

Estimating the parameters of SAOMs

We would like to estimate

$$\theta = (\tau^{[X]}, \tau^{[Z]}, \beta^{[X]}, \beta^{[Z]}), \quad \theta \in \mathbb{R}^P, \quad P = 2(M-1) + K + W$$

- 1. Method of Moments (MoM) and its generalization (GMoM)
- 2. Maximum-likelihood estimation
- 3. Bayesian estimation

Computing the (log-)likelihood of the evolution process

(Snijders et al., 2010)

The model assumptions allow to decompose the process in a series of micro-steps:

$$\{(T_r, i_r, j_r), r = 1, \ldots, R\}$$

- $ightharpoonup T_r$: time point for an opportunity for change,
- \triangleright i_r : actor who has the opportunity to change
- \triangleright j_r : actor towards whom the tie is changed

Given the sequence $\{(T_r, i_r, j_r), r = 1, \dots, R\}$, the likelihood of the evolution process

$$\log[p(x(t_2)|x(t_1);\theta)] = \log\left[\prod_{r=1}^{R} P_{\theta}((T_r, i_r, j_r))\right] \propto \log\left[\underbrace{\frac{(N\tau)^R}{R!}}_{\text{Prob. R steps}} \prod_{r=1}^{R} \underbrace{\frac{1}{N} p_{i_r, j_r}(\beta, x(T_r))}_{\text{Prob. } (i_r, j_r)}\right]$$

Maximizing the (log-)likelihood

Assuming M=2 for simplicitly, we look for the value $\hat{\theta}$ such that

$$\hat{\theta} = \max_{\theta \in \Theta} \log[p(x(t_2)|x(t_1);\theta)]$$

or equivalently

$$\hat{\theta} : \underbrace{\frac{\partial}{\partial \theta} \log[p(x(t_2)|x(t_1);\theta)]}_{\text{Score function}} = 0$$

- We cannot observe the complete data, i.e., the complete series of mini-steps that lead from $x(t_1)$ to $x(t_2)$, and thus we cannot compute the likelihood of the observed data
- ► A stochastic approximation method must be applied



Maximizing the (log-)likelihood

Stochastic approximation method: approximation

Augmented data method

The *augmented data* (or *sample path*) consists of the sequence of tie changes that brings the network from $x(t_1)$ to $x(t_2)$

$$(i_1,j_1),\ldots,(i_R,j_R)$$

Formally:

$$\underline{v} = \{(i_1, j_1), \dots, (i_R, j_R)\} \in \mathcal{V}$$

where \mathcal{V} is the set of all sample paths connecting $x(t_1)$ and $x(t_2)$.



Maximizing the (log-)likelihood

Stochastic approximation method: approximation

Augmented data method

The *augmented data* (or *sample path*) consists of the sequence of tie changes that brings the network from $x(t_1)$ to $x(t_2)$

$$(i_1,j_1),\ldots,(i_R,j_R)$$

Formally:

$$v = \{(i_1, j_1), \dots, (i_R, j_R)\} \in \mathcal{V}$$

where \mathcal{V} is the set of all sample paths connecting $x(t_1)$ and $x(t_2)$.

We can approximate the (log-)likelihood function of the observed data using the probability of \underline{v}

$$\log[p(\underline{v}|x(t_1),x(t_2))] \propto \log\left[\frac{(N\tau)^R}{R!}e^{-N\tau}\prod_{r=1}^R\frac{1}{N}p_{i_rj_r}(\beta,x(T_r))\right]$$



Questions you should be able to answer now

- ► How could you simulate from SAOMs?
- ► How are the parameters of SAOMs estimated?
- ▶ Why we need a stochastic approximation algorithm to compute the MoM estimate?
- ▶ Why MLE cannot be computed in SAOMs?



References

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