US-China differences in cognition and perception across 12 tasks: Replicability, robustness, and within-culture variation

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Abstract

Cultural differences between the US and China have been investigated using a broad array of psychological tasks measuring differences between cognition, language, perception, and reasoning. We examine the robustness of several classic experimental paradigms in cross-cultural psychology. Using online convenience samples of adults, we conducted two large-scale replications of 12 tasks previously reported to show cross-cultural differences. Our results showed a heterogeneous pattern of successes and failures: five tasks yielded robust cultural differences across both experiments, while six showed no difference between cultures, and one showed a small difference in the opposite direction. We observed moderate reliability in all of the multi-trial tasks, but there was little shared variation between tasks. Additionally, we did not see within-culture variation across a range of demographic factors in our samples. Finally, as in prior work, cross-cultural differences in cognition (in those tasks showing differences) were not strongly related to explicit measures of cultural identity and behavior. All of our tasks, data, and analyses are available openly online for reuse by future researchers, providing a foundation for future studies that seek to establish a robust and replicable science of cross-cultural difference.

*Keywords:* replication; cross-cultural differences; cognition; perception; US-China comparison

Word count: 10486

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## Introduction

Cross-cultural differences are a striking part of the broader landscape of human variation. Differences in values and behavior across cultures are obvious to even a casual observer, and researchers have attempted to quantify these differences via a wide range of measures. Comparisons between Western and East Asian cultures have been especially well-researched, with differences attested in a wide range of cognitive domains, including visual attention (Chua, Boland, & Nisbett, 2005; Ji, Peng, & Nisbett, 2000; Waxman et al., 2016), executive function (Sabbagh, Xu, Carlson, Moses, & Lee, 2006; B. Tan, 2020), language learning (Chan et al., 2011, 2011; Tardif, 1996; Waxman et al., 2016), relational reasoning (Carstensen et al., 2019; Cheng, 2020; Richland, Chan, Morrison, & Au, 2010; Su, 2020), similarity judgments (Ji, Zhang, & Nisbett, 2004), values (Ji, Nisbett, & Su, 2001; Kwan, Bond, & Singelis, 1997; Spencer-Rodgers, Williams, Hamilton, Peng, & Wang, 2007), preferences (Corriveau et al., 2017; DiYanni, Corriveau, Kurkul, Nasrini, & Nini, 2015; Liang & He, 2012) and self-concepts (Spencer-Rodgers, Boucher, Mori, Wang, & Peng, 2009; Spencer-Rodgers, Boucher, Peng, & Wang, 2009). As a result, Western and East Asian cultures are increasingly treated as cultural poles in efforts to measure cultural differences (Muthukrishna et al., 2020) and to correct for the pervasive bias in psychology research toward US and European samples (Arnett, 2016; Henrich, Heine, & Norenzayan, 2010; Nielsen, Haun, Kärtner, & Legare, 2017).

Despite a long empirical tradition of comparisons between these cultures and an abundance of psychological accounts for observed differences, estimates of differences are difficult to compare quantitatively because of the varying samples, measures, and methods used in different reports. Further, many of the most prominent reports of cross-cultural differences predate the field-wide discussion of methodological issues in psychology research

during the past 10 years (Open Science Collaboration, 2015). For example, much research in this tradition has been exploratory and hence has not followed current guidance regarding limiting analytic flexibility in order to decrease false positives (Simmons, Nelson, & Simonsohn, 2011). Given the importance of evidence about specific cross-cultural differences for constructing theories of culture more broadly (e.g., Markus & Kitayama, 1992, 2010), further investigation of many empirical findings is likely warranted.

Some empirical evidence points to issues in the robustness of cross-cultural measurements. Typically, measures used in this literature are not standardized and do not have published evidence about reliability and validity (Flake & Fried, 2020). The few extant direct comparisons between measures of cultural difference suggest that theoretically related tasks, such as implicit and explicit measures of the same construct, might not cohere (e.g., Kitayama, Park, Sevincer, Karasawa, & Uskul, 2009). Further, in a study with twenty cross-cultural measures used within a single US sample, Na et al. (2010) found a lack of coherence between tasks measuring social orientation and cognitive style, observing only 8 significant correlations between tasks across 90 statistical tests.[1] Finally, more recent work failed to replicate cultural differences on several related measures (Mercier, Yama, Kawasaki, Adachi, & Van der Henst, 2012; Mercier, Zhang, Qu, Lu, & Van der Henst, 2015; Zhou, Gotch, Zhou, & Liu, 2008). Thus, there is a need to explore the reliability of individual tasks as well as the intercorrelations between them.

Our goal in the current study was to collect a large dataset on a range of cross-cultural measures that had previously been used in comparisons of East Asian and Western cultures, enabling investigations of the robustness of these differences in new samples of Chinese and US participants. We decided to gather relatively large and heterogeneous convenience samples using online recruitment, rather than recruiting

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[1] These authors interpreted their findings as implying that the measures are orthogonal – indexing different constructs – and concluded that group-level differences between cultures are unlikely to relate to within-group individual differences. However, an alternative possibility is that the reliabilities of many individual tasks are low, a feature which would ensure low correlations between them.

smaller, more matched samples using in-lab recruitment. Our reasoning was that the larger samples that we could access using online recruitment would allow us to conduct highly-powered statistical tests, allowing us to make well-powered tests for cultural differences. Further, larger samples afford the analysis of individual and demographic differences within culture, a topic of considerable interest in this literature (e.g., Na et al., 2010, 2020). Finally, the development of browser-based online versions of prominent cross-cultural tasks would allow their inspection and reuse by other researchers, thus promoting a more cumulative approach to the measurement of cultural differences.

Our experiments were intended to be close replications of the original studies, but differences in format of administration introduced inevitable variation, in some cases more substantial than others. The interpretation of discrepant outcomes between an original study and a replication is complex, given that disparate outcomes can occur for many reasons (Machery, 2020; Nosek & Errington, 2020; Zwaan, Etz, Lucas, & Donnellan, 2018). In our case, interpretation is especially difficult and we explicitly avoid interpreting our results as bearing on the status of the original findings we investigate.

There were some significant differences between our experiments and the original studies. First, we recruited online convenience samples from the U.S. and China. Previous work varied in the participants' country of origin (in several cases, Japan for East Asian participants; Canada for Western participants), largely recruited either college students or community members, and was administered more than a decade ago. Within-culture variation and generational differences between our samples and previous samples make results difficult to compare directly. Furthermore, our strategy of constructing a battery of replication studies and administering them uniformly online altered the contexts in which participants engaged with the tasks and in some cases required alterations to the tasks themselves.

Accordingly, our replication studies should be viewed as an assessment of robustness:

specifically, we assess whether a set of previously-reported East-West cross-cultural

differences can be recovered in online convenience populations. These are not assessments

of the veracity of the original findings. Nevertheless, we believe that cross-cultural

psychology can be advanced via the identification of tasks that yield cross-cultural

differences robustly across a variety of samples and administration formats – we hope our

work contributes to this aim. We return to these interpretive issues in the General

Discussion.

Our task selection process was initially shaped by an interest in relational reasoning

and accounts explaining it with reference to cross-cultural differences in visual attention

and social cognition (Duffy, Toriyama, Itakura, & Kitayama, 2009; Kuwabara & Smith,

2012; Moriguchi, Evans, Hiraki, Itakura, & Lee, 2012). Additionally, in Experiment 1, we

selected tasks that could potentially be administered to young children as well as adults,

for use in future work addressing developmental questions about the relative time course of

cross-cultural differences across the visual, social, and cognitive domains. We balanced

three desiderata in our task selection, preferentially choosing tasks that (1) had been

theoretically or empirically implicated in relational reasoning, (2) were associated with

differential performance in US-China comparisons or related cultural contrasts (i.e., East

Asian vs. Western cultures), and (3) were relatively short, accessible tasks appropriate for

web administration. We also conducted an extensive set of pilot tests to ensure that

participants understood instructions and that the tasks yielded interpretable data.

In Experiment 2, we selected a second set of tasks to investigate based in part on the

results of Experiment 1. In particular, we repeated a handful of tasks from Experiment 1,

in some cases, varying task parameters. We then selected a further set of tasks that probed

both cross-cultural differences in higher-level cognition (e.g., language and reasoning) and

perception, again respecting the desideratum that the tasks should be relatively short and

amenable to administration in a web browser. The final set of tasks included in each

experiment is listed in Table 1.

139    In addition to the goal of replicating individual tasks, our hope was that the

140    relatively large dataset we collected could be used to explore the structure of within- and

141    between-culture variation in cognition and perception more broadly. Towards this goal, we

142    included a relatively extensive demographic questionnaire in both of our experiments, with

143    the aim of using these measures to explore variation within our samples. In the final

144    section of the paper, we report a series of exploratory analyses. The first of these assesses

145    the reliability of individual tasks to gauge whether these tasks are reliable enough from a

146    psychometric point of view to support further individual differences analyses. We then

147    report correlations across tasks, aiming to discover covariation between tasks that might

148    indicate that they load on the same construct. Finally, we turn to analyses of whether

149    within-culture demographic variables predict variation in task performance. Overall, a

150    number of tasks revealed acceptable levels of reliability, but tasks did not cluster together,

151    and we found relatively few demographic predictors of within-culture variation.

152    We make all code and data from our experiments available for further data collection

153    and analysis in hopes of promoting further cumulative work on measures and theories of

154    cross-cultural variation.

Table 1
*Tasks included in each experiment and the final sample size after exclusions. We conducted
simulations for power analysis and found that our sample sizes are well-powered to detect
medium and above effect sizes. We included the power analyses in Appendix A.*

| Experiment | Task | Citation | Task Description | CN | US |
|---|---|---|---|---|---|
| 1 | Ambiguous Relational Match-To-Sample (cRMTS) | Carstensen et al. (2019) | Infer whether an object or relation is causally relevant | 186 | 178 |
| | Picture Free Description | Imada, Carlson, & Itakura (2013) | Describe pictures from memory after a brief study period | 169 | 172 |
| | Ebbinghaus Illusion | Imada, Carlson, & Itakura (2013) | Judge the size of circles in a context designed to bias size judgments | 190 | 180 |

| | Horizon Collage | Senzaki, Masuda, & Nand (2014) | Make an image by dragging and dropping stickers onto a display | 187 | 175 |
|---|---|---|---|---|---|
| | Symbolic Self-Inflation (Family) | Kitayama et al. (2009) | Draw self and family members as circles | 150 | 114 |
| | Uniqueness Preference | Kim & Markus (1999) | Choose a sticker from five stickers, four of which are the same color | 191 | 180 |
| | Child Causal Attribution | Seiver, Gopnik, & Goodman (2013) | Watch short vignettes and explain the decisions of the characters | 177 | 170 |
| | Raven's Progressive Matrices | Su (2020) | Use analogical reasoning to complete visually-presented patterns | 191 | 180 |
| 2 | Ambiguous Relational Match-To-Sample (cRMTS) | Carstensen et al. (2019) | Infer whether an object or relation is causally relevant | 174 | 293 |
| | Picture Free Description | Imada, Carlson, & Itakura (2013) | Describe pictures from memory after a brief study period | 132 | 284 |
| | Change Detection | Mausda & Nisbett (2007) | Find differences in the foreground or background of two images | 160 | 253 |
| | Symbolic Self-Inflation (Friends) | Kitayama et al. (2009) | Draw a sociogram with self and friends as nodes, relationships as edges | 158 | 252 |
| | Adult Causal Attribution | Morris & Peng (1994) | Read a crime story and explain the criminal's motivations | 114 | 293 |
| | Taxonomic-Thematic Similarity | Ji, Zhang, & Nisbett (2004) | Match items based on taxonomic or thematic similarity (e.g., cow: chicken / grass) | 178 | 295 |

| Semantic Intuition | Li, Liu, Chalmers, & Snedeker (2018) | Decide whether a story refers to a named character (whose actions are mischaracterized) or the person who performed the actions (but had a different name) | 181 | 298 |
| Raven's Progressive Matrices | Su (2020) | Use analogical reasoning to complete visually-presented patterns | 181 | 298 |

## Experiment 1

In Experiment 1, our goal was to evaluate cross-cultural differences in a variety of constructs. We assembled a web-based battery of tasks and tested these on a snowball sample of US and Chinese participants.

### Methods

**Participants.** We recruited participants through snowball sampling seeded at large universities in the US and China, in which participants directly recruited by the researchers were encouraged to recruit their friends and family members through email forwarding and social media sharing. Participants in the US were compensated with $5 gift certificates (USD) and participants in China received ¥35 (CNY).

We recruited 203 and 201 participants each from the US and China, respectively. Since we did not have strong a priori expectations about specific effect sizes, our overall preregistered sample size was chosen to meet or exceed the sample sizes used in prior reports in the literature from which our tasks were drawn. Our sample size, methods, and main analyses were pre-registered and are available at https:// aspredicted.org/37y6a.pdf.

170     Our preregistered exclusion plan was to exclude people from the full dataset if they

171 failed quality checks[2] on any one task, unless this excluded 20% or more of our sample.

172 Due to a task demand associated with the Symbolic Self-Inflation task, this criterion would

173 have led to the exclusion of 107 people (US: 66, CN: 41) due to this task alone. This

174 triggered the less restrictive exclusion approach in our preregistration, using task-specific

175 quality checks to exclude participants only from the relevant individual task.

176     After exclusions, the US sample included 180 participants (49 Male, 120 Female, 9

177 Non-binary, 2 Declined to answer), with a mean age of 22.02 years old, all of whom were

178 native English speakers. The China sample included 191 participants (60 Male, 127

179 Female, 1 Non-binary, 3 Declined to answer), with a mean age of 22.42 years old, who were

180 all native speakers of Mandarin Chinese. This sample size is shared among all tasks except

181 for the Symbolic Self-Inflation task, which included 114 US participants and 150 CN

182 participants.

183     In addition to age, gender, and linguistic background, we collected a range of

184 demographic information including subjective socioeconomic status measured using the

185 MacArthur Ladder (Adler, Epel, Castellazzo, & Ickovics, 2000), level of maternal

186 education, the state or province the participant grew up in, residential mobility, and

187 number of international experiences.

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[2] We performed exclusion based on three aspects: side bias, missing data, and demographic exclusions.
(1) Side bias: If more than 90% of selections by a participant in a 2AFC task were a single response button (left or right), data from this participant in this task would be excluded.
(2) Missing data: If no data/more than 25% missing/data not codeable/side bias/participant did not follow instructions for any one task, then we would exclude all data from that participant, across tasks. If this led to 20% or more of participants being excluded, we would not apply this exclusion criteria to all data from the participant – only at the level of individual tasks.
(3) Demographic exclusions: We would exclude data from [CN/US] participants who reported living abroad for more than 2 years in regions with predominantly [European/Asian] populations (respectively); and [CN/US] participants who reported speaking or understanding [English/any Chinese language or dialect] with proficiency at or above 3 out of 10. If this led to 20% or more of participants being excluded from either test population (US/CN), we would drop this exclusion criterion for the relevant population and use exploratory regressions to examine how the factor relates to responding in our tasks.

188      **Procedure.**   Participants completed an online, browser-based sequence of eight

189   tasks (see Table 1) and a brief demographic questionnaire. All tasks were implemented in a

190   combination of jsPsych (De Leeuw, 2015) and custom HTML/JavaScript code. Tasks were

191   administered in English for the US sample and in Mandarin Chinese for the China sample.

192   To control for the impact of order-related inattention, task order was randomized across

193   participants with two exceptions: (1) the two drawing tasks (Symbolic Self-Inflation and

194   Horizon Collage) were always back-to-back in random order, and (2) Uniqueness Preference

195   was always the penultimate task (in keeping with the task cover story, which congratulated

196   participants on being nearly done with the experiment). In total, the experiment took

197   about 30 minutes to complete.

198      **Measures.**   Below, we give a short description of prior findings and methods for

199   each task.

200      ***Ambiguous cRMTS.***   Carstensen et al. (2019) observed cross-culturally distinct

201   developmental trajectories in a causal relational match-to-sample (cRMTS) task, and

202   different preferences in an ambiguous formulation of this task. Specifically, when

203   3-year-olds saw evidence consistent with both object-based (e.g., blue cubes make a

204   machine play music) and relational (pairs of different objects, AB, make a machine play

205   music) solutions, children in the US sample preferentially chose the object-based solution,

206   while those in China chose the relational solution.

207      We used this ambiguous version of the task (Carstensen et al., 2019, Experiment 3)

208   to explore whether adults in the US and China also show differing preferences for

209   object-based or relational solutions. Our participants saw two pairs of objects, AB and AC,

210   activate a machine, and were given a forced choice between an object-based solution (a

211   *same* pair of A objects, AA) and a relational solution (*different* pair BC).

212      ***Picture Free Description.***   Imada, Carlson, and Itakura (2013) found that

213   children around the age of 6 showed cultural differences in describing pictures to others.

214 Relative to US children, Japanese children were more likely to mention the objects in the

215 background first, as opposed to the focal objects in the picture. They also tended to

216 provide more descriptive accounts of the background objects than their US counterparts.

217 In our version of the task, we used a subset of seven images from the original study and

218 adapted the task for adult participants, who studied each image for 5 seconds and then

219 typed a description. We coded the first mentioned item (focal or background) and counted

220 descriptors for focal and background elements.

221     ***Ebbinghaus Illusion.***    Both Japanese adults and children have been found to be

222 more susceptible to the Ebbinghaus Illusion – in which context alters the perceived size of

223 a circle – than Western participants in the US and UK (Doherty, Tsuji, & Phillips, 2008;

224 Imada et al., 2013). We followed the Imada et al. (2013) implementation of the task, with

225 two testing blocks: the No Context block (10 trials) and Illusion block (24 trials). The No

226 Context block establishes baseline accuracy for discriminating which of two orange circles

227 is larger. In the Illusion trials, the two orange circles are flanked by a grid of 8 gray circles,

228 which are all smaller or larger than the center orange circle. The illusion occurs because

229 the orange circles appear larger when flanked by smaller gray circles, leading to distortions

230 in comparing the sizes of the two orange circles with differing contexts (i.e., small or large

231 flankers). Across the 24 Illusion trials, we measured accuracy of circle size judgments as a

232 function of the actual size difference and flanker context (helpful or misleading).

233     ***Horizon Collage.***    Senzaki, Masuda, and Nand (2014) found that school-age

234 children in Japan and Canada showed culture-specific patterns when creating a collage of

235 an outdoor scene. Japanese children drew the horizon higher, used more collage items, and

236 filled more space with collage items relative to Canadian children. We adapted the task

237 from Senzaki et al. (2014) Study 2, in which participants were prompted to make a collage

238 with stickers. Our participants could drag any of thirty images (line drawings of people,

239 animals, houses, etc.) onto a rectangular "canvas" in the middle of the screen. There was

240 also a sticker "horizon," a horizontal line that spanned the length of the canvas. All

stickers, including the horizon, could be clicked and dragged to the canvas to produce "a picture of the outside." Participants were asked to include a horizon and any number of other stickers to create their image. We measured the height of the horizon, the number of stickers used, and the total area occupied by stickers as in Senzaki et al. (2014).

***Symbolic Self-Inflation.*** Kitayama et al. (2009) found a difference between Western and East Asian cultures in the size of circles participants drew to represent themselves relative to other people in their social networks. Japanese participants drew circles of similar sizes to represent themselves and others, while those from Western countries (US, UK, Germany) tended to draw their "self" circles larger than those representing others, indicating a symbolic self-inflation in the three western cultures compared to Japan. We adapted this task, asking participants to draw themselves and the family members they grew up with as circles by clicking and dragging the mouse on a rectangular "canvas" to draw circles of varying sizes. They then labeled each circle for the person it represented. We measured the diameter of each circle and calculated a percent inflation score for each participant by dividing the diameter of the self circle by the average diameter of circles for all others.

***Uniqueness Preference.*** Kim and Markus (1999) tested East Asians' and Americans' preferences for harmony or uniqueness by asking them to pick one gift pen from five options. In the condition that we replicated, the options differed only in the barrel colors, with four that were the same and one that was unique. They found that European Americans were more likely to choose the unique colored pen than East Asian participants. We adapted our task to better fit the format of our online experiment by showing a virtual "sticker book" to measure progress through all tasks in our study. At the end of each task, participants received a virtual sticker. For the uniqueness preference task, we let them select one of five dinosaur stickers that were identical except for color: four blue and one yellow (with repeated and unique colors randomized between participants). Choice of the unique vs. repeated color was recorded.

268    ***Child Causal Attribution.***    Previous work has shown that participants from

269  South Korea and the U.S. attribute behaviors differently in situations where there is

270  evidence in favor of situational explanations (Choi, Nisbett, & Norenzayan, 1999).

271  Similarly, Chinese participants and media are more likely than their US counterparts to

272  attribute a person's behaviors to situational context as opposed to individual traits

273  (Morris, Nisbett, & Peng, 1995; Morris & Peng, 1994). We adapted the deterministic

274  situation condition in Seiver, Gopnik, and Goodman (2013), a task originally designed for

275  children. In this task, two children both engage in one activity and avoid another,

276  suggesting that situational constraints (e.g., the latter activity being dangerous) may be

277  guiding their decisions. Participants watched a series of four short, animated vignettes in

278  which two children both played in a pool and neither child played on a bicycle. We then

279  asked participants to explain in text why each child did not play on the bicycle, making for

280  two test trials per participant. We used the prompt question from Seiver et al. (2013),

281  which explicitly pits person attributions against situational ones: "Why didn't Sally play

282  on the bicycle? Is it because she's the kind of person who gets scared, or because the

283  bicycle is dangerous to play on?" We coded each response for per-trial count of (a)

284  personal and (b) situational attributions.

285    ***Raven's Standard Progressive Matrices.***    As an additional attention check as

286  well as an exploratory measure of relational reasoning assessing performance rather than

287  preference, we included the 12 questions from Set E of Raven's Standard Progressive

288  Matrices. Su (2020) found cross-cultural differences between adults in the US and China in

289  performance on this set. This set of questions was selected because it is the most difficult

290  subset and also the one most dependent on true analogical reasoning, without alternative

291  heuristic approaches like visual pattern completion.

292    **Analyitic approach.**    Data and analysis scripts are available at

293  https://github.com/anjiecao/CCRR_writeups

294    The papers that we drew on for our tasks used a heterogeneous set of analytic

295 methods. Rather than planning to replicate these specific analyses, we instead attempted

296 to follow current best practices by using linear mixed effects models with maximal random

297 effect structure as a unified analytic framework (Barr, Levy, Scheepers, & Tily, 2013). We

298 fit a separate model to each task. In case of convergence failure, we followed lab standard

299 operating procedures: pruning random slopes first and then random intercepts, always

300 maintaining random intercepts by participant. For linear models, we report p-values

301 derived from t-scores. For linear mixed models, we report p-values derived from z-scores,

302 which is appropriate for relatively large samples (Blouin & Riopelle, 2004). Our key tests

303 of interest were typically either the coefficient for a main effect of country (US/China) or

304 an interaction of country and condition.

## Results

306 **Ambiguous cRMTS.** To examine whether adults in the US and China show

307 differing preferences for object-based or relational solutions, we ran a mixed-effects logistic

308 regression predicting response choice (object or relation) with country (US or China) as a

309 fixed effect. There was no main effect of country on response choice (object or relation; US:

310 $M = 0.39$, $SD = 0.48$; CN: $M = 0.38$, $SD = 0.47$; $\beta = 0.04$, $SE = 0.84$, $z = 0.05$, $p =$

311 0.96). The preference for object-based solutions seen in US preschoolers in an ambiguous

312 context and the corresponding preference for relational solutions observed in China did not

313 extend to adults in our sample.

314 Our US results replicate findings by M. Goddu and Walker (2018), who reported that

315 US adults are at chance in this paradigm. It seems likely that adults in both groups of our

316 study are aware of the ambiguous evidence and their near-chance selections reflect

317 (reasonable) uncertainty.

318 **Picture Free Description.** Based on Imada et al (2013), we expected Chinese

319 participants would be more likely to mention background objects first and provide more

320 descriptive accounts for background objects relative to focal objects, in comparison with
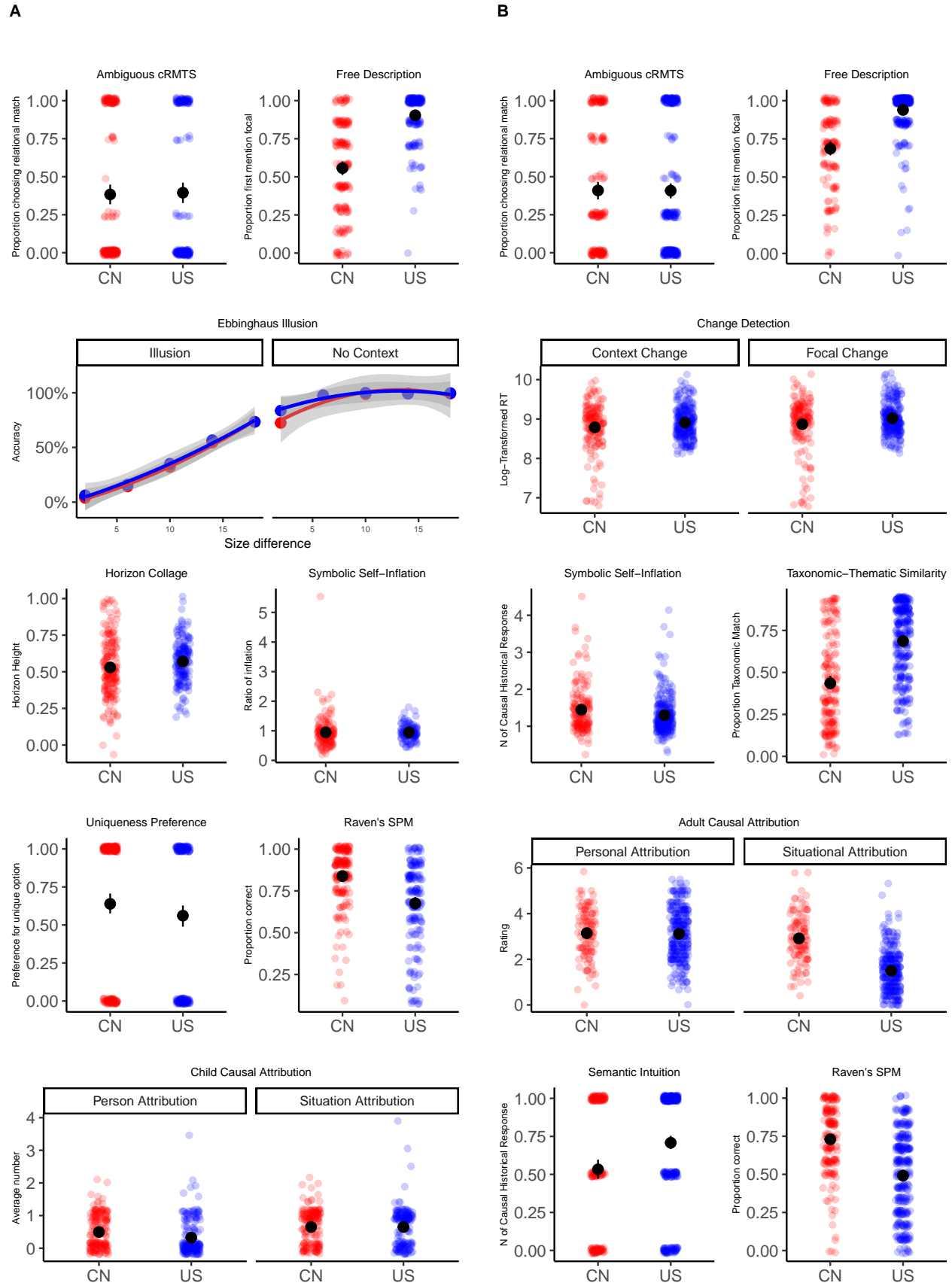
*Figure 1*. Results from each task. Results from the CN sample are plotted in red, and the US in blue. Panel A and B includes results from experiment 1 and 2 respectively. The results for other measures used in the Free Description and Horizon Collage were included in Appendix B.

321  US participants. Our results extend previous findings with the former metric (first

322  mention; US: $M = 0.90$, $SD = 0.17$; CN: $M = 0.56$, $SD = 0.30$) but not the latter (number

323  of descriptive accounts; for focal objects: US: $M = 1.06$, $SD = 0.51$; CN: $M = 0.87$, $SD =$

324  0.44; for background objects: US: $M = 1.30$, $SD = 0.93$; CN: $M = 0.94$, $SD = 0.72$).

325      For first mention, we ran a mixed-effects logistic regression predicting the type of first

326  mention (object or relation) with country (US or China) as a fixed effect. We found a main

327  effect of country ($\beta = 3.39$, $SE = 0.34$, $z = 9.98$, $p < 0.01$). For descriptive accounts, we

328  ran a mixed-effect Poisson regression model predicting the number of descriptive accounts,

329  with description type (focal or background), country (US or China), and their interaction

330  as fixed effects. There was a significant main effect of culture (with US participants

331  providing more descriptions overall: $\beta = 0.35$, $SE = 0.13$, $t = 2.67$, $p < 0.01$). The culture

332  effect interacted with the description types, but the effect was in the opposite direction,

333  with U.S participants providing more background descriptions than focal descriptions,

334  relative to Chinese participants ($\beta = $ -0.15, $SE = 0.07$, $t = $ -1.97, $p < 0.05$).

335      The mixed results between the first mention and descriptive accounts measures

336  suggest that there is some complexity in linking broader theoretical accounts to specific

337  measures; we interpret this result with caution and include the task in Experiment 2 to

338  follow up further.

339      **Ebbinghaus Illusion.**   To test whether perception of the Ebbinghaus Illusion

340  varied across populations in our sample, we ran a mixed-effects logistic regression

341  predicting accuracy on each trial, with country (US or China), context (No Context or

342  Illusion context), and circle size difference (the percent difference in diameters) as fixed

343  effects, along with their interactions. We found main effects of context (with worse

344  performance in the Illusion context; $\beta = 5$, $SE = 0.27$, $z = 18.64$, $p < 0.01$) and circle size

345  difference (worse performance for smaller differences; $\beta = 0.34$, $SE = 0.01$, $z = 29.13$, $p < $

346  0.01). There was a marginally significant main effect of country in the opposite of the

347  predicted direction (US participants performed worse: $\beta = 0.52$, $SE = 0.25$, $z = 2.13$, $p < $

348  0.05) but no interactions with country (All $\beta < 0.01$; All $p > 0.05$).

349      In sum, we failed to replicate cultural differences found between Western and East
350  Asian participants in susceptibility to the Ebbinghaus Illusion.

351      **Horizon Collage.**  In the Horizon Collage task, three key measurements were
352  calculated from the "collage" participants created: the height of the horizon (in proportion
353  to the height of the frame), the number of stickers, and the total area that the stickers
354  covered (following the original analysis, we added up the area occupied by each individual
355  sticker). Senzaki et al. (2014) found that Japanese children tended to put the horizon
356  higher, include more stickers, and cover more area in their collage, compared with
357  Canadian children. We ran a fixed effect linear model with culture as the main predictor
358  for each of the measurements. Culture significantly predict the horizon height, but the
359  effect was in the opposite direction direction, with U.S. participants putting the horizon
360  sticker at higher than the Chinese participants (Sticker height: US: $M = 0.57$, $SD = 0.15$;
361  CN: $M = 0.53$, $SD = 0.20$; $\beta = 0.04$, $SE = 0.02$, $t = 2.16$, $p < 0.05$). Culture did not
362  significantly predict any of the other two measurements (Sticker number: US: $M = 11.41$,
363  $SD = 5.78$; CN: $M = 11.53$, $SD = 5.73$; Sticker area: US: $M = 16.83$, $SD = 8.28$; CN: $M$
364  $= 17.18$, $SD = 8.47$; Both $\beta < 0.03$; Both $p > 0.1$).

365      Our experiment contrasted Chinese and US adults, rather than Japanese and
366  Canadian children. Although Senzaki et al. (2014) found that the cultural differences were
367  more salient in older children than younger children, suggesting that cultural differences
368  might increase with development, interpretation of our failure to replicate is still qualified
369  by differences in culture and medium of administration.

370      **Symbolic Self-Inflation.**  To test whether US adults have a larger symbolic self
371  than Chinese adults, we ran a linear regression predicting percent inflation (the diameter of
372  the self circle divided by the average diameter of circles for others) with country (US or
373  China) as a fixed effect. No difference was found in the degree of symbolic self-inflation

374 between US and Chinese adults based on percent inflation (US: $M = 0.95$, $SD = 0.26$; CN:

375 $M = 0.94$, $SD = 0.53$; $\beta = < 0.01$, $SE = 0.05$, $t = 0.03$, $p = 0.98$).

376         As an exploratory analysis, we also considered whether the number of circles drawn

377 for others could be a potential confounding factor for the effect. Due to the limited space

378 on the canvas, people who drew more circles may draw each circle smaller, resulting in less

379 symbolic self-inflation. We found culture to be a significant predictor of the number of

380 others circles drawn, with US participants drawing more others circles than Chinese

381 participants (US: $M = 3.49$, $SD = 1.26$; CN: $M = 3.12$, $SD = 1.29$). However, when

382 controlling for the number of circles drawn, there was still no difference between US and

383 Chinese adults on the percent inflation ($\beta = -0.01$, $SE = 0.05$, $t = -0.22$, $p = 0.82$).

384         One possible explanation for our null result is that there are cultural differences

385 between Japan and China in self-concept; Japanese samples typically demonstrate

386 characteristics previously associated with East Asian cultures in general, with Chinese

387 samples deviating from these characteristics at times (Bailey, Chen, & Dou, 1997; Church

388 et al., 2012, 2014). In addition, the null results could also be attributed to differences

389 between our task design and that of Kitayama et al. (2009). Instead of asking participants

390 to draw their social network, our design asked participants to draw themselves and the

391 family members they grew up with. During the coding process, we noticed that people

392 from both cultures tended to draw older people, e.g., their parents, as larger circles, which

393 might have resulted in larger circles for others than for the self in both cultures, masking

394 any US-China difference in the degree of self-inflation. We follow up on this possibility by

395 changing the task prompt in Experiment 2.

396         **Uniqueness Preference.**   We examined cross-cultural preferences for uniqueness

397 by running a simple logistic regression predicting each participant's single choice (minority

398 or majority color) with country (US or China) as a fixed effect; we used logistic regression

399 rather than mixed effects logistic regression due to the absence of repeated observations.

400 There was no cross-cultural difference in the probability of choosing the unique sticker (US:

$M = 0.56$, $SD = 0.50$; CN: $M = 0.64$, $SD = 0.48$; $\beta = $ -0.32, $SE = 0.21$, $z = $ -1.52, $p = $ 0.13).

The difference between our result and that of the original study by Kim and Markus (1999) may be related to our use of an online format in our study. In the original study, participants were asked to pick a gift pen from five physical pens with different barrel colors. It could be that Asian American participants in the previous study chose the more common color because they wanted to leave a choice for the next participant in the face of resource scarcity, rather than because they were expressing values or identities influenced by East Asian cultural mandates favoring interpersonal harmony and similarity. Our finding is also consistent with previous work demonstrating that tendencies toward conformity in East Asian samples are linked to reputation management (Yamagishi, Hashimoto, & Schug, 2008); it may be that our online experiment did not establish a sufficient social context to motivate participants' concern about reputation, and accordingly failed to motivate reputation management in the form of a conformity preference.

**Child Causal Attribution.** To test whether Chinese participants tended to make more situational attributions than US adults, we ran a mixed-effects Poisson regression predicting the number of attributions included in each explanation, with attribution type (situational or personal), country (US or CN), and their interaction as fixed effects. We found a main effect of attribution type (situational attribution: US: $M = 0.65$, $SD = 0.61$; CN: $M = 0.65$, $SD = 0.52$; personal attribution: US: $M = 0.33$, $SD = 0.54$; CN: $M = 0.50$, $SD = 0.52$; $\beta = 0.27$, $SE = 0.10$, $z = 2.65$, $p < 0.01$). Neither the interaction nor the main effect of culture was significant (both $\beta < 0.3$; $p > 0.05$).

The failure to find cross-cultural differences in attribution could be related to the style of the tasks, which were relatively repetitive and originally designed for children; in Experiment 2, we follow up with a causal attribution task designed for adults.

426     **Raven's Standard Progressive Matrices.**    As an exploratory measure of

427 relational reasoning, we ran a mixed-effects logistic regression predicting per-trial accuracy,

428 with country as a fixed effect, random intercepts for each subject and question, and

429 by-question random slopes for country. We found a main effect of country, with Chinese

430 participants outperforming those from the US (US: $M = 0.68$, $SD = 0.25$; CN: $M = 0.84$,

431 $SD = 0.17$; $\beta$ = -1.36, $SE = 0.24$, $z$ = -5.77, $p < 0.01$).

432     Our findings replicate Su (2020) in finding an advantage for Chinese participants on

433 Raven's Matrices. In our context, we also interpret the relatively high scores we observed

434 as evidence that participants were engaging fully with our tasks.

435 **Discussion**

436     We did not observe cross-cultural differences in the majority of tasks in Experiment

437 1. The only exceptions were in Picture Free Description and our exploratory measure of

438 performance in relational reasoning (Raven's SPM). We also found cultural difference in

439 the opposite direction in one of the measures in the Horizon Collage task. Many of our

440 tasks did not have a manipulation check and could yield null results simply by virtue of

441 inattention. However, the results of Raven's SPM (and the Ebbinghaus Illusion) suggest

442 that participants were engaged in our tasks and performed at a high objective level. In

443 addition to minor methodological changes that we made, interpretation of our failure to

444 replicate individual tasks in many cases could be due to (1) differences in administration

445 (online vs. in-person), (2) differences in participant recruitment (e.g., university pool

446 vs. snowball sampling), (3) differences in target age (adults vs. children), and (4)

447 differences in sample (e.g. Japanese vs. Chinese adults in the East Asian group).

448     Our failure to find robust differences between Western and East Asian cultures in this

449 initial selection of tasks was dispiriting. We designed Experiment 2 to extend Experiment 1

450 by recruiting a different sample and identifying followup or replacement tasks that we

451 hoped would yield a broader set of cross-cultural differences.

## Experiment 2

Experiment 2 was designed to follow up on Experiment 1 and further evaluate cross-cultural differences across a battery of tasks. Because several of our tasks in Experiment 1 yielded no evidence for cross-cultural differences, we replaced these with alternative tasks selected to address similar or related constructs. We replaced the Ebbinghaus Illusion with a measure of Change Detection that has been argued to index context sensitivity (Masuda & Nisbett, 2006). We replaced the child-appropriate causal attribution task with a version designed for adults (Morris & Peng, 1994). We also included two tasks measuring linguistic or semantic intuitions more broadly (Taxonomic/Thematic Similarity and Semantic Intuition), following up on the detection of cross-cultural differences in the Picture Free Description task. Although our goal in Experiment 2 was to evaluate a further set of tasks, we also included the Ambiguous cRMTS, Picture Free Description, and Raven's Progressive Matrices tasks to replicate our results from Experiment 1, and we included a modified version of Symbolic Self-Inflation to address several issues with the earlier version of the task.

In Experiment 2, we made use of crowd-sourcing services – rather than snowball sampling – as our participant recruitment channel. We had two rationales. First, in Experiment 1 our samples were quite young (due to seeding our sampling with university students through email and social media). A younger sample may be less enculturated because they are less experienced or more exposed to international media and influences, and thus less likely to show distinct cross-cultural differences. Second, we were concerned that being recruited by friends and family (as in a snowball sample) might prime interdependent thinking among our participants, leading to decreased cross-cultural differences.

## Methods

[477]      **Participants.**    We recruited participants through online crowdsourcing websites.

[478] For the US, we used Prolific and applied the following screening criteria: a) US nationality,

[479] b) born in the US, and c) currently residing in the US; For China, we used Naodao

[480] (www.naodao.com), a platform designed for conducting online experiments in mainland

[481] China. Participants in US received \$12.25 in compensation and in China ¥35.

[482]      We recruited 304 participants from the U.S. and 185 participants from China. 10

[483] participants were excluded because they did not meet our demographic inclusion criteria[3].

[484] Following our preregistration (available at https://osf.io/u7mzg), we applied a task-based

[485] exclusion procedure in which we excluded a participant's responses in a particular task if

[486] they a) showed a response bias for a single response button or value[4], b) had missing data

[487] on more than 25% of trials[5], or c) failed to meet the inclusion criteria for that task as

[488] specified in the preregistration[6].

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[3] Demographic exclusions: we would exclude data from [CN/US] participants who reported living abroad for more than 2 years in regions with predominantly [European/Asian] populations (respectively); and [CN/US] participants who reported speaking or understanding [English/any Chinese language or dialect] with proficiency at or above 3 out of 10. If this led to 20% or more of participants being excluded from either test population (US/CN), we would drop this exclusion criterion for the relevant population and use exploratory regressions to examine how the factor relates to responding in our tasks.

[4] If more than 90% of selections by a participant in a 2AFC task were a single response button (left/right in RMTS or top/bottom in taxonomic/thematic similarity), data from this participant in this task would be excluded. If 100% of selections by a participant in a scalar (e.g., Likert scale used in causal attribution or in the demographics section) or multiple choice task (e.g. Raven's SPM) used a single response button/value, data from this participant in this task would be excluded.

[5] If no data/more than 25% missing/data not codeable/side bias/participant did not follow instructions for any one task, then we would exclude all data from that participant, across tasks. If this led to 20% or more of participants being excluded, we would not apply this exclusion criteria to all data from the participant – only at the level of individual tasks.

[6] Additional task and trial-level exclusions:
(1) Change detection: trials where the participant incorrectly identified the change would not be analyzed (but will not be counted as missing data unless a response was not attempted).
(2) Free description: trials where the response was not code-able would be discarded and considered missing data.
(3) Causal attribution: none.
(4) Symbolic self-inflation: data from participants who failed to draw exactly one "self" circle would be excluded, as well as data from participants who drew only a "self" circle.
(5) Taxonomic/thematic similarity task: There were two unambiguous catch trials intermixed with regular

Similar to Experiment 1, we collected demographic information from participants, including subjective socioeconomic status, the state or province the participant grew up in and the one they currently reside in, residential mobility, number of international experiences, education, and undergraduate area of study (STEM or non-STEM). We also administered scales to collect explicit measures of participants' cultural identities and behaviors (Cleveland & Laroche, 2007; Cleveland, Laroche, & Takahashi, 2015).

The sample size for each task after exclusions and the descriptive statistics for each demographic question are reported in Table 1.

**Procedure.** Similar to Experiment 1, participants completed eight tasks and a brief demographics questionnaire online. The experiment was administered in English for the US sample and in Mandarin Chinese for the Chinese sample, with the exception of the Adult Causal Attribution task. As in previous work, this task was administered in English, and only Chinese participants who self-identified as being able to read English participated in it. To control for the impact of order-related inattention, task order was randomized across participants with two exceptions: (1) the Free Description task always occurred before (not necessarily immediately) Change Detection (because Change Detection included a manipulation check that explicitly asked about focal objects, which could bias responding in Free Description), and (2) the two story-based tasks (Semantic Intuition and Adult Causal Attribution) always occurred together in a fixed order at the end of the study, with Semantic Intuition first and Adult Causal Attribution last. Adult Causal Attribution was always the last task (if run) because it was administered in English and we did not wish to prime CN participants with English stimuli before any of the other tasks, all of which were run in Mandarin.

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trials in this task (e.g., Choose cat: cat, dog). If a participant missed either unambiguous catch trial, all triad data from that participant would be excluded.
(6) Semantic intuition: Participants would be excluded for missing any of the 5 control questions.
(7) Ambiguous RMTS: none.
(8) Raven's SPM: none.

Measures.

**Tasks from Experiment 1.** We replicated three tasks from Experiment 1 using identical procedures: Ambiguous cRMTS, Picture Free Description, and Raven's Standard Progressive Matrices.

**Symbolic Self-Inflation.** Participants were asked to draw themselves and their friends as circles, as opposed to drawing themselves and their family members as circles in Experiment 1. They were also asked to draw lines between any two people who are friends, as in the original study by Kitayama et al. (2009). They then labeled each circle to indicate the person it represents. We calculated a percent inflation score for each participant by dividing the diameter of the self circle by the average diameter of circles for others.

**Adult Causal Attribution.** We speculated that the lack of cross-cultural differences in Causal Attribution in Experiment 1 might be due to the simplistic nature of our task, which was designed for use with young children. Therefore, in Experiment 2 we used a paradigm designed for adults, in which participants were asked to read a crime narrative from a news report that included substantial information on a criminal's background and the events leading up to their crime, and then rate the relevance of various situational and personal factors (Morris & Peng, 1994). In the original study, both Chinese participants and US participants read stories in English. We followed this procedure by selecting the subset of our Chinese participants who self-identified as comfortable reading short stories in English to participate. In the task, participants were told that they would read news stories and answer questions to help social scientists understand the factors that contribute to murders. Participants were randomly assigned to read one of two stories (Iowa shooting or Royal Oak shooting). After the stories, they were asked to write a short explanation for the murderer's behaviors. Then, they rated a list of statements about causes of the murder on a 7-point Likert scale. The statements included items that describe personal and situational factors, and we measured endorsement of these two factor types.

⁵³⁸        ***Change Detection.***    Masuda and Nisbett (2006) found differences in attention

⁵³⁹  allocation between Japanese and US participants in a change detection paradigm. They

⁵⁴⁰  found that Japanese participants were significantly faster than US participants in

⁵⁴¹  identifying changes in the background of images. We followed their original procedure and

⁵⁴²  used the same stimuli. In this task, participants were presented with 30 pairs of images.

⁵⁴³  On each trial, two pictures would alternate on the screen, each presented for 560ms with a

⁵⁴⁴  blank screen in between images for 80ms. The two pictures were almost identical with

⁵⁴⁵  subtle differences, either in the focal object (e.g., a tractor in daylight with its lights on or

⁵⁴⁶  off) or the background (e.g., a cloud with slightly different locations in the sky).

⁵⁴⁷  Participants were instructed to press a key when they spotted the difference, and then

⁵⁴⁸  describe the difference in a text box. If they did not detect a difference within 60 seconds,

⁵⁴⁹  the trial timed out. Only trials in which participants correctly identified the changes were

⁵⁵⁰  included in the analysis. After 30 trials, participants saw each pair of images again, this

⁵⁵¹  time side-by-side on the screen. They were asked to identify the focal object(s) in the

⁵⁵²  pictures by typing into a text box. These responses were used as a manipulation check to

⁵⁵³  ensure that participants in both cultures construed focal objects similarly.

⁵⁵⁴        We coded change descriptions to exclude trials in which participants did not identify

⁵⁵⁵  the change, and checked agreement on focal objects across cultures. We measured how

⁵⁵⁶  quickly participants identified the difference on trials in which they reported the difference

⁵⁵⁷  correctly.

⁵⁵⁸        ***Taxonomic-Thematic Similarity.***    Ji et al. (2004) showed that Chinese

⁵⁵⁹  participants are more likely to match items based on thematic similarity, whereas US

⁵⁶⁰  participants are more likely to match items based on taxonomic similarity. In this task,

⁵⁶¹  participants were presented with triads containing a cue word and two match options. In

⁵⁶²  each test set, one option was a taxonomic match (e.g., monkey - elephant) and the other a

⁵⁶³  thematic match (e.g., monkey - banana). In each filler set, the cue item and the options

⁵⁶⁴  were broadly similar, thematically and taxonomically, making for a more ambiguous

decision (e.g., monkey: elephant, tiger). Participants completed a two-alternative forced choice task in which they chose one match for each cue item.

The findings of Ji et al. (2004) were replicated in more recent work (Le, Frank, & Carstensen, 2021); we used a subset of testing materials from Le et al. (2021), with 15 test triads, 15 filler triads, and 2 attention check questions. The order of the triads was randomized between subjects. We measured taxonomic vs. thematic match selections on each of the test trials.

***Semantic Intuition.***    Li, Liu, Chalmers, and Snedeker (2018) found cultural differences in semantic intuitions about ambiguous referents in Chinese and US participants. Specifically, Chinese participants were more likely to determine the referent of a name based on the description of the speaker (the descriptivist view) whereas US participants were more likely to determine the referent based on the original usage (the causal-historical view). In the study, participants read five separate stories and judged the correctness of statements referring to a character after each story. Two comprehension check questions were included for each story. We followed the original procedure closely and used the same materials. We measured participants' semantic intuition as their judgment on the correctness of statements referring to the critical characters.

**Results**

**Ambiguous cRMTS.**    Our analysis was identical to that in Experiment 1. We did not observe a main effect of country on participants' preference for object vs relational matches (proportion relational match: US: $M = 0.41$, $SD = 0.44$; CN: $M = 0.41$, $SD = 0.42$; $\beta = -0.01$, $SE = 0.48$, $z = -0.03$, $p = 0.98$). As in Experiment 1, we did not find evidence that the differential preferences observed in preschoolers extend to adults. It seems likely that adults in both populations are aware of the mixed evidence for the relational

589  and object solution and that their responses reflect sensitivity to this ambiguous design.[7]

590      **Picture Free Description.**   US participants were more likely to initially mention

591  the focal objects than the background objects (first mention: US: $M = 0.94$, $SD = 0.14$;

592  CN: $M = 0.69$, $SD = 0.26$). We used the same regression analysis as in Experiment 1 and

593  found a main effect of country ($\beta = 3.09$, $SE = 0.32$, $z = 9.61$, $p < 0.01$). Our results

594  replicate the first mention finding in Experiment 1 with a comparable effect size

595  (standardized mean difference; Experiment 1: 1.50[1.26, 1.74]; Experiment 2: 1.57[1.34,

596  1.80]).

597      We also deviated from our pre-registered analysis plan and coded for the number of

598  descriptive accounts directed at the focal objects and background objects using the same

599  coding schemes as Experiment 1. We ran the same mixed-effect Poisson regression model

600  predicting the number of descriptive accounts, with the interactions between description

601  type (focal or background) and country (U.S. or China) as the predictor. Interestingly, we

602  did not replicate the results in Experiment 1, but found patterns similar to the results in

603  Imada et al. (2013). We found an interaction between country and type of description ($\beta$

604  $= 0.16$, $SE = 0.07$, $z = 2.24$, $p < 0.05$). Chinese participants, in contrast to U.S.

605  participants, provided more descriptive accounts of the background objects relative to the

606  focal objects (for focal objects: US: $M = 0.66$, $SD = 0.82$; CN: $M = 0.50$, $SD = 0.64$; for

607  background objects: US: $M = 0.69$, $SD = 1.07$; CN: $M = 0.61$, $SD = 0.80$).

608      In summary, these results extend Imada et al.'s (2013) findings to Chinese adults.

609      **Change Detection.**   We ran a linear mixed-effects model predicting the reaction

610  time to correctly identified changes in the pictures, with country (U.S. or China) and type

611  of change detected (focal or background) as main effects, as well as their interaction. We

612  did not find evidence for an interaction between culture and type of change detected ($\beta =$

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[7] Our reliability analysis shows that adults expressed this uncertainty only at the population level:
individuals tended to be consistent in choosing the same solution type across all four test trials, with
ambiguity expressed as disagreement between participants.

613   0.04, $SE = 0.03$, $z = 1.40$, $p = 0.16$). Participants in both countries identified changes to

614   the context faster than changes to focal objects (context changes: $M = 10,101.87$, $SD =$

615   4,257.15; focal object changes: $M = 10,646.54$, $SD = 4,816.10$; $\beta = 0.07$, $SE = 0.02$, $t =$

616   3.45, $p < 0.01$). Chinese participants identified both types of change more quickly than US

617   participants (US: $M = 10,689.49$, $SD = 4,406.73$; CN: $M = 9,875.67$, $SD = 4,733.57$; $\beta =$

618   0.12, $SE = 0.05$, $t = 2.27$, $p < 0.05$).

619         As an exploratory analysis, we also retroactively analyzed the coded accuracy of the

620   participants' responses. Interestingly, we found main effects of culture and the type of

621   change, as well as an interaction between culture and the type of change. Participants

622   across both countries were more accurate in identifying changes of the focal objects than

623   the context objects, and Chinese participants were more accurate in identifying the changes

624   than the U.S. participants on average (US focal: $M = 0.90$, $SD = 0.30$; US context: $M =$

625   0.85, $SD = 0.36$; CN focal: $M = 0.94$, $SD = 0.24$; CN context: $M = 0.89$, $SD = 0.32$).

626   However, the difference between Chinese participants and U.S. participants was larger in

627   focal changes than in background changes, with Chinese participants being more accurate

628   than U.S. participants in the focal changes than in the context changes ($\beta = $ -0.22, $SE =$

629   0.09, $z = $ -2.53, $p < 0.05$). This interaction is different from our predictions: If we

630   extrapolate from the original study, we should expect to see a difference in the background

631   context, with participants performing similarly in the background changes but Chinese

632   participants being more accurate in the background changes than U.S. participants.

633         In sum, we did not replicate the findings of Masuda and Nisbett (2006).

634         **Symbolic Self-Inflation.**   In Experiment 1, we did not find a significant difference

635   in the degree of symbolic self-inflation between adults in the US and China. Here, we

636   observed a pattern contrary to the prediction: US adults showed less self-inflation than

637   Chinese adults (US: $M = 1.30$, $SD = 0.51$; CN: $M = 1.45$, $SD = 0.65$; $\beta = $ -0.15, $SE =$

638   0.06, $t = $ -2.56, $p < 0.05$). We did not replicate the findings of Kitayama et al. (2009) with

639   Japanese participants in either of our experiments.

⁶⁴⁰       **Adult Causal Attribution.**     We ran a mixed-effects linear regression predicting

⁶⁴¹  endorsement of each potential cause with country (US or China) and attribution type

⁶⁴²  (personal or situational) as fixed effects, as well as their interaction. We found an

⁶⁴³  interaction in the predicted direction: Chinese participants endorsed situational

⁶⁴⁴  attributions to a greater extent than their counterparts in the US (situational ratings: US:

⁶⁴⁵  $M = 1.71$, $SD = 0.80$; CN: $M = 3.17$, $SD = 0.89$; personal ratings: US: $M = 3.12$, $SD = $

⁶⁴⁶  1.10; CN: $M = 3.14$, $SD = 1.07$; $\beta = $ -1.39, $SE = 0.14$, $t = $ -9.71, $p < 0.01$). This result

⁶⁴⁷  extends the original findings by Morris and Peng (1994), and suggests that the measure of

⁶⁴⁸  causal attribution in Experiment 1 (which was designed for use with child participants)

⁶⁴⁹  may not be appropriate for measuring cross-cultural differences in causal attribution

⁶⁵⁰  among adults.

⁶⁵¹       **Taxonomic-Thematic Similarity.**     We used a mixed-effects logistic regression

⁶⁵²  predicting response (taxonomic or thematic match) with country (US or China) as a fixed

⁶⁵³  effect. There was a significant effect in the predicted direction: participants in the US were

⁶⁵⁴  more likely to choose taxonomic matches than participants in China (proportion taxonomic

⁶⁵⁵  matches: US: $M = 0.69$; $SD = 0.46$; CN: $M = 0.44$; $SD = 0.50$; $\beta = 2.02$, $SE = 0.89$, $t = $

⁶⁵⁶  2.27, $p < 0.05$). This finding replicates the findings of Ji et al. (2004) and Le et al. (2021).

⁶⁵⁷       **Semantic Intuition.**     We ran a mixed-effects logistic regression predicting response

⁶⁵⁸  (descriptive or causal-historical) with country (US or China) as a fixed effect, and found

⁶⁵⁹  that US participants made significantly more causal-historical choices than Chinese

⁶⁶⁰  participants (proportion causal historical choice: US: $M = 0.71$; $SD = 0.46$; CN: $M = 0.53$;

⁶⁶¹  $SD = 0.50$; $\beta = 1.59$, $SE = 0.37$, $t = 4.37$, $p < 0.01$). We also replicated the item effect

⁶⁶²  identified by Li et al. (2018), though this was not among our preregistered analyses. In

⁶⁶³  sum, We replicated Li et al. (2018) with new samples of adults in the US and China.

⁶⁶⁴       **Raven's Standard Progressive Matrices.**     We replicated the findings from

⁶⁶⁵  Experiment 1. Chinese participants scored higher on Raven's Standard Progressive

⁶⁶⁶  Matrices than US participants (US: $M = 0.49$, $SD = 0.27$; CN: $M = 0.73$, $SD = 0.23$; $\beta = $

667  -1.82, $SE = 0.25$, $z = -7.39$, $p < 0.01$).

## Discussion

669  Overall, Experiment 2 was more successful than Experiment 1 in documenting
670  cross-cultural differences between participants in the US and China. This success can be
671  attributed to the inclusion of the successful tasks from Experiment 1 (e.g., Free Description
672  and Raven's Standard Progressive Matrices), and the exclusion of tasks designed for young
673  children (e.g., Child Causal Attribution, Horizon Collage).

<div align="center">

**Exploratory analyses**

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675  We conducted a set of exploratory analyses to consolidate results from the two
676  experiments. We first performed a miniature meta-analysis with the tasks from both
677  experiments. Then, we assessed the reliability of the tasks that included multiple trials, the
678  relationships between tasks, and finally, how explicit cultural identities and demographic
679  factors relate to task performance.

## Mini meta-analysis

681  As our first exploratory analysis, we identified the key effect of interest from our
682  pre-registration (usually a main effect of culture or an interaction of culture, depending on
683  task) and converted the coefficient into a standardized measure of effect size (standardized
684  mean difference; SMD) via the method described by Westfall, Kenny, and Judd (2014).
685  Because there is no "correct" direction for any task except Raven's SPM, we show the
686  absolute value of the effect sizes (Figure 2).

687  Across our two experiments, we saw consistent and generally large differences (SMD
688  $> 0.6$) in Free Description, Raven's SPM, Adult Causal Attribution, Semantic Intuition,
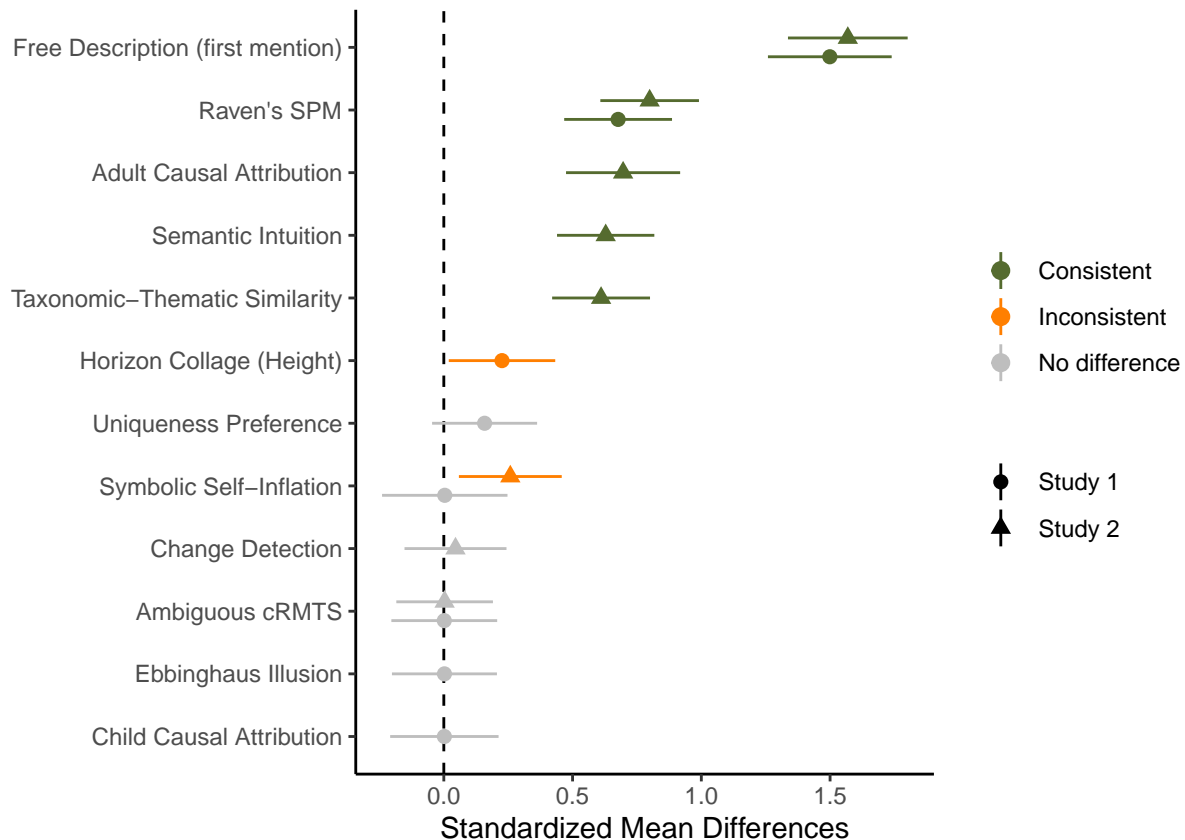689  and Taxonomic-Thematic Similarity tasks. Aside from Raven's SPM, all of these tasks

*Figure 2*. Forest plot of effect sizes (standardized mean difference) for each task across both experiments. Point shape shows experiment number and color indicates whether effects were consistent with prior literature.

690 have in common that they are deliberative linguistic tasks that tapped into relatively

691 high-level cognitive constructs. In contrast, we observed effect sizes close to zero for our

692 more aesthetic and perceptual tasks (Change Detection, Ebbinghaus Illusion, and Horizon

693 Collage). We also observed little consistent difference in four other tasks (Ambiguous

694 cRMTS, Symbolic Self-Inflation, Uniqueness Preference, and Child Causal Attribution),

695 perhaps for reasons idiosyncratic to each. We return to the broader question of

696 generalization across task types in the General Discussion.

697      We conducted three additional exploratory analyses to consolidate results from the

698 two experiments. First, we assessed the reliability of the tasks that included multiple trials.

699 Second, we examined whether there was shared variance between tasks. Finally, we

700 examined how explicit cultural identities and demographic factors relate to task

₇₀₁ performance.
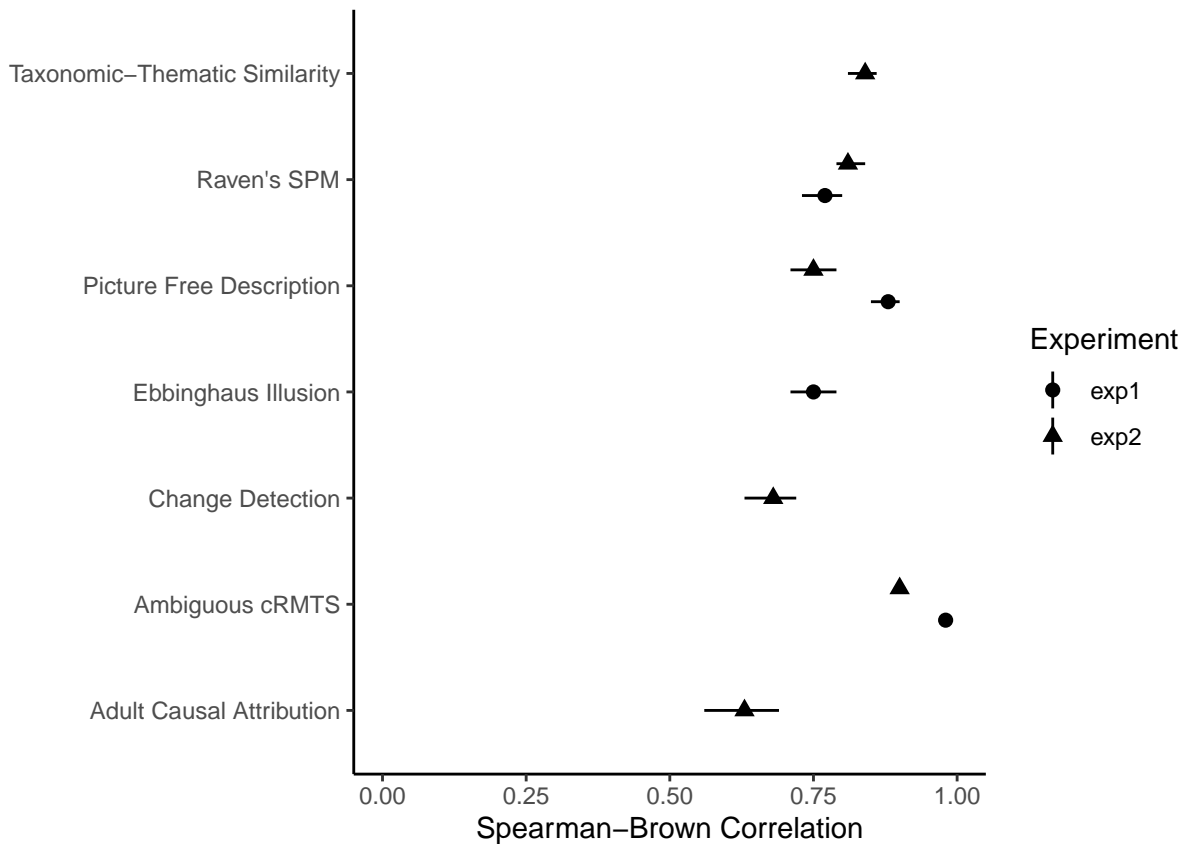
₇₀₂ **Reliability assessment**



*Figure 3*. Spearman-Brown adjusted reliabilities for tasks with four or more trials. Point shape shows experiment number. Error bars show 95% confidence intervals.

₇₀₃     One question motivating our work was whether the individual tasks we used were

₇₀₄ reliable enough – had low enough measurement error – to be used for further investigation

₇₀₅ of individual differences. The gold standard for evaluating whether a task yields stable

₇₀₆ within-person measurements is test-retest reliability (simply because test-retest gives a

₇₀₇ direct estimate of stability over time), but this method was outside the scope of our study.

₇₀₈ Instead, we used a split-half approach, asking whether participants' answers on individual

₇₀₉ questions relate to one another. Specifically, we used a permutation-based split half

₇₁₀ approach (Parsons, 2021) in which we made 5000 random splits of items into two simulated

<sub>711</sub> "halves" and then computed the within-person correlation between scores on these two

<sub>712</sub> halves, averaging across simulated runs. To estimate the reliability of the full-length

<sub>713</sub> instrument, we used the Spearman-Brown "prophecy" formula.

<sub>714</sub>     Since the split-half approach is only suitable for tasks with multiple trials, we

<sub>715</sub> removed tasks with fewer than four trials from the analysis. For tasks with more than one

<sub>716</sub> condition, we focused on the condition that was predicted to show cultural differences (i.e.,

<sub>717</sub> the Illusion context in the Ebbinghaus Illusion; situational factors in Adult Causal

<sub>718</sub> Attribution; background change scenes in Change Detection).

<sub>719</sub>     Figure 3 shows the corrected split-half reliabilities for all tasks in both of our

<sub>720</sub> experiments. Overall, the reliabilities were acceptable (all Spearman-Brown Correlations >

<sub>721</sub> 0.6). We further investigated whether there was cultural variation in the reliability of

<sub>722</sub> tasks. For most tasks, the reliabilities were relatively similar (within 0.1 of one another),

<sub>723</sub> but there were three tasks where reliability was lower for US participants than Chinese

<sub>724</sub> participants: Change Detection (US - CN = -0.19), Adult Causal Attribution (US - CN =

<sub>725</sub> -0.33), Free Description in Experiment 1 (US - CN = -0.23).

## Relations between individual tasks

<sub>726</sub>

<sub>727</sub>     One (perhaps simplistic) interpretation of the prior literature on cultural variation is

<sub>728</sub> that there is a general tendency toward holistic or analytic reasoning that varies across

<sub>729</sub> cultures and explains variation in tasks. This single dimension might correspond to broad

<sub>730</sub> (or focused) attention and contextualized, relational reasoning (or an emphasis on focal

<sub>731</sub> people or objects). As a first step towards investigating this interpretation, we explored

<sub>732</sub> whether there was a single dimension of individual variation in our data that corresponded

<sub>733</sub> to this general axis of cross-cultural difference. Because some data was missing, largely due

<sub>734</sub> to task-related exclusions, we treated the missing data using two approaches: listwise

<sub>735</sub> deletion and imputation with means. These approaches yielded comparable results, so here

736 we report correlations from listwise deletion.

737      Correlations between task scores were quite low on average, suggesting limited

738 support for a single factor explanation. Across both experiments, the largest absolute

739 magnitude of correlations observed were -0.29 (Taxonomic-Thematic Similarity and Adult

740 Causal Attribution in Experiment 2), -0.28 (Free Description and Raven's SPM in

741 Experiment 2), and -0.24 (Adult Causal Attribution and Free Description in Experiment

742 2). All other correlations were between -0.23 and 0.23. Hence, the amount of shared

743 variation between tasks was quite limited and our attempts at exploratory factor analysis

744 discovered structures with many distinct factors and very low loading on the first factor.

745      We also applied the same correlation analysis within each culture. Again, we found

746 limited correlations between tasks. This pattern replicated the findings from a previous

747 work showing negligible relationships among a battery of tasks that revealed cultural

748 differences (Na et al., 2010) (See Appendix C).


749 **Demographic variation and explicit measures of cultural identity**

750      As a final exploratory analysis, we asked whether demographic variation or variation

751 in cultural identity predicted responding in our tasks. Our approach to these questions was

752 to fit a set of exploratory regression models for each task, predicting task scores as a

753 function of an individual scale and its interaction with culture. This approach allowed us

754 to explore both within- and between-culture effects in a single model. Our predictors were

755 1) the summed score for our global/local cultural identity and consumption measures (with

756 local items reverse-scored, such that higher scores represent more global identities and

757 consumption patterns); 2) geographic information about where participants grew up

758 Markus & Conner (2014); and 3) a range of demographic factors, including age, gender

759 identity, residential mobility, number of international experiences, maternal education level,

760 and subjective socioeconomic status as measured by the MacArthur Ladder (Adler et al.,

761    2000).

762        **Task performance and global identity.**    We fit models predicting task

763    performance based on culture and its interaction with global (vs local) identity for tasks in

764    Experiment 2 (we did not collect these scales in Experiment 1). Two of these relationships

765    were statistically significant at .01 < p < .05 (Adult Causal Attribution: $p = 0.05$;

766    Taxonomic-Thematic Similarity: $p = 0.04$) but neither of these relationships survived

767    Bonferroni correction for multiple comparisons.

768        **Task performance and geographic origin.**    We next considered whether regions

769    within each country were meaningful predictors of task performance. We fit models

770    predicting task performance based on the regions that participants reported growing up in.

771    For China, provinces were categorized as rice-cultivating regions or wheat-cultivating

772    regions based on Talhelm et al. (2014). For the US, states were categorized based on either

773    the coastal locations (West Coast, East Coast, and Inland) or broad geographic locations

774    (West, South, Northeast, Midwest), following the categorization reported in Carstensen,

775    Saponaro, Frank, and Walker (2022).

776        5 out of the 48 models we ran showed statistically significant relationships between

777    regions and task performance. In Experiment 1, US coastal location was a significant

778    predictor for the Free Description task. Participants who grew up in Inland regions ($N =$

779    54) or on the East Coast ($N = 27$) were more likely to mention the focal object first when

780    describing the pictures than participants who grew up on the West Coast ($N = 84$; Inland:

781    $p = 0.02$; East Coast: $p = 0.05$). In Experiment 2, both coastal location and broad

782    geographic location were significant predictors for Raven's SPM, with participants from the

783    East Coast ($N = 89$) and Inland regions ($N = 159$) scoring higher than participants from

784    the West Coast ($N = 46$; Inland: $p < 0.01$; East Coast: $p = 0.05$), and participants from

785    the Midwest ($N = 78$) and South ($N = 94$) scoring higher than participants from the West

786    ($N = 59$; Midwest: $p < 0.01$; South: $p = 0.04$). In addition, both region categories

787    predicted performance in Change Detection. East Coast participants ($N = 75$) took longer

to respond than West Coast ($N = 42$) participants ($p = 0.02$), and Northeastern participants ($N = 52$) took longer to respond than participants who grew up in the West ($N = 52$; $p < 0.01$). However, none of these relationships survived Bonferroni correction.

**Basic demographic effects.** We fit 192 exploratory regression models to see if basic demographic factors could predict task performance. The demographic factors we explored were age, gender identity, residential mobility, number of international experiences, maternal education level, and subjective socioeconomic status as measured by the MacArthur Ladder (Adler et al., 2000). 26 were statistically significant, but only one model survived Bonferroni correction. Change detection was predicted by age in the US sample, with older participants taking longer to respond than younger participants (adjusted $p < 0.01$). Given some of the models could be considered as conceptual replications of previous work, we reported selected models with significant uncorrected results in Appendix D.

## General Discussion

The world's cultures are strikingly different, and psychologists have long sought to measure and characterize this variation, with differences between Western and East Asian cultures as a case study of particular interest. These efforts have given rise to a rich literature documenting cultural differences in a wide range of psychological tasks. Across two experiments, we selected a collection of tasks that had previously been shown to yield differences between Western and East Asian samples and replicated them with two relatively large online samples of participants from the US and China. In this discussion, we first consider the limitations of our study since these contextualize the remainder of our conclusions. Next, we consider the interpretation of our results within individual tasks. We end the discussion with a summary of the key findings of this work.

**General Limitations**

As discussed above and in the introduction, we did not design our experiments to replicate prior work directly, and hence one important limitation of our work is simply that it cannot be used as a test of the reliability of prior findings. Instead, our measures provide estimates of US-China differences on a range of constructs, specifically for online convenience samples. These estimates are likely biased downward – towards the null hypothesis of no difference between cultures – by several features of our experimental design.

Online experiments (especially grouped into a long battery as ours were) likely receive slightly less attention than in-person studies, though overall these effects have tended to be small in US samples (Buhrmester, Kwang, & Gosling, 2016). Contra this concern, however, participants did perform relatively accurately on those tasks that had correct answers (e.g., Raven's SPM, the Ebbinghaus Illusion), and in our exploratory analysis, we found relatively high reliabilities on all tasks. Further, our pre-registered exclusion criteria removed participants who performed poorly. Thus, we do not believe that participants were inattentive overall.

One might also be concerned about the reaction time measures (e.g. Ebbinghaus Illusion and Change Detection) collected through the online experiment. Due to variations in individual participants' hardware and internet conditions, the reaction times collected could potentially be more variable than those from in-person experiment. However, the validity of online reaction time measure has been evaluated and compared against in-person data across many different classic cognitive psychology tasks such as the Stroop task and the Flanker task. Across many tasks that require precise reaction time measures, online replications were qualitatively identical to the in-person version of the task (Crump, McDonnell, & Gureckis, 2013). Moreover, some works have also evaluated the accuracy and reliabiltiy of jsPsych, the JavaScript library that we used (Anwyl-Irvine, Dalmaijer,

Hodges, & Evershed, 2021; Pinet et al., 2017). The consensus is that reaction time measures collected through jsPsych is reliable and accurate. Therefore, there is no reason to believe that the variability in our online reaction time measure deviated significantly from the in-person version of the reaction time task.

Another limitation of our estimates of US-China differences comes from variation in our sampling strategy between cultures. In Experiment 1, we used the same snowball sampling procedure, but this procedure may have yielded different samples due to differences in social networks or norms about sharing study information across cultures. In Experiment 2, because the platform we used to recruit US participants (Prolific) was not accessible in China, we used a different platform to recruit Chinese participants (Naodao). Prolific and Naodao have different levels of popularity and different participant pools, resulting in some asymmetry between the US and Chinese samples. Despite these differences between samples both across and within experiments, we do not see indications that our estimates were dramatically biased by our sampling decisions. First, our results were largely comparable in the tasks that were included in both experiments (e.g. Picture Free Description; Raven's SPM; and Ambiguous cRMTS). Second, in our exploratory analyses we did not find strong associations between participant demographics and cross-cultural effects (with some small exceptions discussed in that section). Finally, we reran all of our preregistered analyses with an age-matched subset of U.S. participants in Experiment 2 and found our results were qualitatively identical.

Thus, while our samples are certainly not representative samples of US or Chinese national populations – indeed to our knowledge, nearly all work to date has used convenience samples of one type or another – they appear to yield stable cross-sample estimates that do not reflect large biases due to sampling strategy or demographics.

One of the main ways in which our samples may not have been representative is that they are likely to be more globalized than the population on average simply by being young

864  (and thus less acculturated) and having access to a computer. Contra this concern,

865  variation in local cultural identity did not strongly relate to variation in any of our tasks,

866  but interestingly, we observed the strongest local identities (within our Chinese sample)

867  among the youngest participants.

868      Last but not least, another difference between our experiments and previous work

869  was the lack of an experimenter, and some of our tasks may be particularly sensitive to the

870  presence of an experimenter. In a web experiment, participants are often isolated in front

871  of their own computer. In contrast, in an in-person experiment, participants must interact

872  with and perform the task in front of experimenters who are often from the same social

873  group. Indeed, in the Uniqueness Preference pen choice task, cross-cultural differences are

874  dependent on the presence of an experimenter (Yamagishi et al., 2008). Our null results,

875  obtained in the absence of an experimenter, can be seen as a conceptual replication of this

876  work.

**Task-specific Limitations**

878      In addition to the general limitations discussed above, there are features of our

879  experimental adaptations that may have affected performance in specific tasks. In this

880  section, we highlight concerns about these issues and discuss their implications for

881  interpreting the results of these tasks. See Table 2 for a summary of this task-specific

882  discussion.

883      In the case of the Uniqueness Preference task, it is possible that adapting the task to

884  an online format in which resource scarcity was not strictly real and task choices had no

885  lasting effect (in the form of a new pen), may have trivialized the choice and undermined

886  the incentive for prosocial, harmonious behavior or expression. This possibility is

887  consistent with the chance responding we observed in both groups. Alternatively, our

888  results could be seen as a conceptual replication of Yamagishi et al. (2008), who argue that

differences in this task are moderated by the likelihood of evaluation, with no differences in pen choice observed in the absence of an experimenter.

The ambiguous developmental tasks, Ambiguous cRMTS and Child Causal Attribution, may have been too heavy-handed in their key manipulations; both were designed to highlight ambiguity for young children, but it may be that their explicit cues and repetitive instructions impressed this ambiguity too strongly for adult audiences, resulting in the adults' near-chance responding – a reasonable response to such marked ambiguity. Cultural differences in causal reasoning and attribution may only manifest when the task design is age-appropriate. Consistent with this view, we did replicate previously attested differences in the Adult Causal Attribution task in Experiment 2, and other recent work has shown cross-cultural differences in causal attribution among 4- to 9-year-olds in Germany, Japan, and Ecuador using a design similar to the Child Causal Attribution task (Jurkat, Iza Simba, Hernández Chacón, Itakura, & Kärtner, 2022).

Last but not least, variation within the broad cultural constructs of East Asia and the West could explain some of our findings, as a failure to extend previous work. Some of the tasks we included originally compared participants from other parts of East Asia and the West (e.g., Horizon Collage, Symbolic Self-Inflation, Change Detection; but c.f. Masuda, Ishii, and Kimura (2016) for an alternative account of mixed findings in change detection paradigms). For example, the Taxonomic-Thematic Similarity task replicated previously attested cross-cultural differences between the US and China both here and in other work (Le et al., 2021) but these differences failed to generalize to a US-Vietnam comparison, despite the cultural, historical, and geographic similarities between China and Vietnam. This variation suggests that similar psychological tendencies could be expressed differently under distinct sociocultural contexts and traditions, even across regions and countries that share many similarities. As another example, responding in the Horizon Collage could be modulated by variation between countries: Chinese and Japanese aesthetic traditions differ, so while Chinese and Japanese people may share a preference for highly contextualized

information, this preference may be typically expressed through distinct visual techniques.

**Conclusion**

We conducted two experiments to examine the robustness of several classic experimental paradigms in cross-cultural psychology. Our results showed a heterogeneous pattern of successes and failures: some tasks yielded robust cultural differences across both experiments, while others showed no difference between cultures. We estimated the reliability of the tasks to be moderate, with only minor variation in reliability across cultures. We also explored the effects of a range of demographic variables, including explicit identification with global identity, regional differences within cultures, and several demographic characteristics. All of these had minimal relation to task performance.

Our goal here was not to perform direct replications that would shed light on the replicability of specific findings. Instead, since our methods, administration medium, sample, and analytic approach differed from the prior literature, our hope was to examine the robustness of these paradigms as a method for measuring US-China differences in an online context. Our work has several strengths relative to the prior literature, including larger samples of participants from the US and China, two broad groups of tasks implemented openly online (and reusable by future researchers), and a preregistered analysis plan that allows for the unbiased estimation of cross-cultural effects. In sum, we hope that our work here provides a foundation for future studies that seek to establish a robust and replicable science of cross-cultural difference.

Table 2

*Differences between our tasks, the original refernce tasks, and speculation about how these differences may have impacted our results.*

| Task Name | Difference between the current and original version | Speculative reasons |
|---|---|---|
| Ambiguous causal Relational Match-To-Sample (cRMTS) | • Age: The original experiment was conducted with young children; the current experiment was conducted with adults (but cf. Goddu, Sullivan, & Walker (2021) for non-chance adult performance in a very similar paradigm). <br><br> • Test Format: The original experiment was conducted in person with physical stimuli and live experimenters; the current experiment was conducted through web-based interfaces with animation. | • The relation between objects might have been less salient when the stimuli were presented as schematic animation. <br><br> • Adults may have been overall less engaged with the tasks and paid less attention to the causal properties of the object pairs. |
| Ebbinghaus Illusion | • Age: The original study was conducted with young children; the current experiment was conducted with adults. <br><br> • Culture: The original experiment compared Japanese participants with Canadian participants; the current experiment compared Chinese participants with U.S. participants." | • Adults in both cultures reached ceiling performance. <br><br> • U.S. and Chinese participants may differ in visual context sensitivity compared to Canadian and Japanese participants. |

| | | |
|---|---|---|
| Horizon Collage | • Age: The original experiment was conducted with young children; the current experiment was conducted with adults. <br><br> • Test Format: The original experiment was conducted in person with paper and collage stickers; the current experiment was conducted through web-based interfaces. <br><br> • Culture: The original experiment compared Japanese participants with Canadian participants; the current experiment compared Chinese participants with U.S. participants. | • The task might have been too trivial for adults to engage with properly. <br><br> • The saliency of horizon height is diminished by the drag-and-drop online interface. |
| Symbolic Self-Inflation (Family version and Friends version) | • Test Format: The original experiment was conducted in person with pen and paper; the current experiment was conducted through web-based interfaces. <br><br> • Test Prompt (Family version): The original experiment asked the participants to draw their close social network; the current experiment asked the participants to draw their family members or friends growing up. | • The online interface might have altered participants' drawing process, making it more difficult to implicitly represent the symbolic meaning of the circle size. |

| | | |
|---|---|---|
| Uniqueness Preference | • Test Format: The original experiment was conducted in person; the current experiment was conducted through web-based interfaces. | • The choice between virtual stickers is less meaningful than between real pens. |
| | • Stimulus: The original experiment gave away physical pens to keep; the current experiment asked participants to select virtual stickers. | • The presence of a live experimenter might increase the social pressure on participants, causing them to consider the cultural perception of their choice. |
| | • Test Context: The original experiment prompted participants to make their pen selection after completing a questionnaire indicating their aesthetic preference for abstract art with unique or repeating shapes. | • This richer context may have prompted more careful consideration of the key prompt. |
| Child Causal Attribution | • Age: The original experiment was conducted with young children; the current experiment was conducted with adults. | • Adults may have failed to engage deeply in causal reasoning because the story is too simplistic. |
| Change Detection | • Test Format: The original experiment was conducted in person; the current experiment was conducted through web-based interfaces. | • The in-lab setting might facilitate participants' performance by making them pay more attention to the computer screen. |

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Appendix A

Prospective Power analysis

We conducted two power analyses: a low assumption t-test power analysis and a high assumption linear mixed effects model power analysis. Under both power analysis frameworks, we showed that our sample sizes are well-powered to detect medium and larger effect sizes.

**Low assumption t-test power analysis**

In this framework, we assumed the cultural differences are evaluated with a simple t-test between two samples. We varied the hypothetical effect sizes of the cultural differences. This analysis was conducted using the `pwr` package (Champely et al., 2018).

**High assumption mixed effects linear regression power analysis**

Conducting prospective power analysis for mixed-effects linear regression requires more assumptions. Here we specify a simplified model to run the power simulation on. The model predicts a response with culture as the fixed effect, and subjects as the random intercepts. We assumed the fixed effect's intercept and slope as 0.1 and 0.5, respectively. We varied the effect sizes of the model by manipulating different residual standard deviation. We conducted this analysis in the `simr` package (Green & MacLeod, 2016).
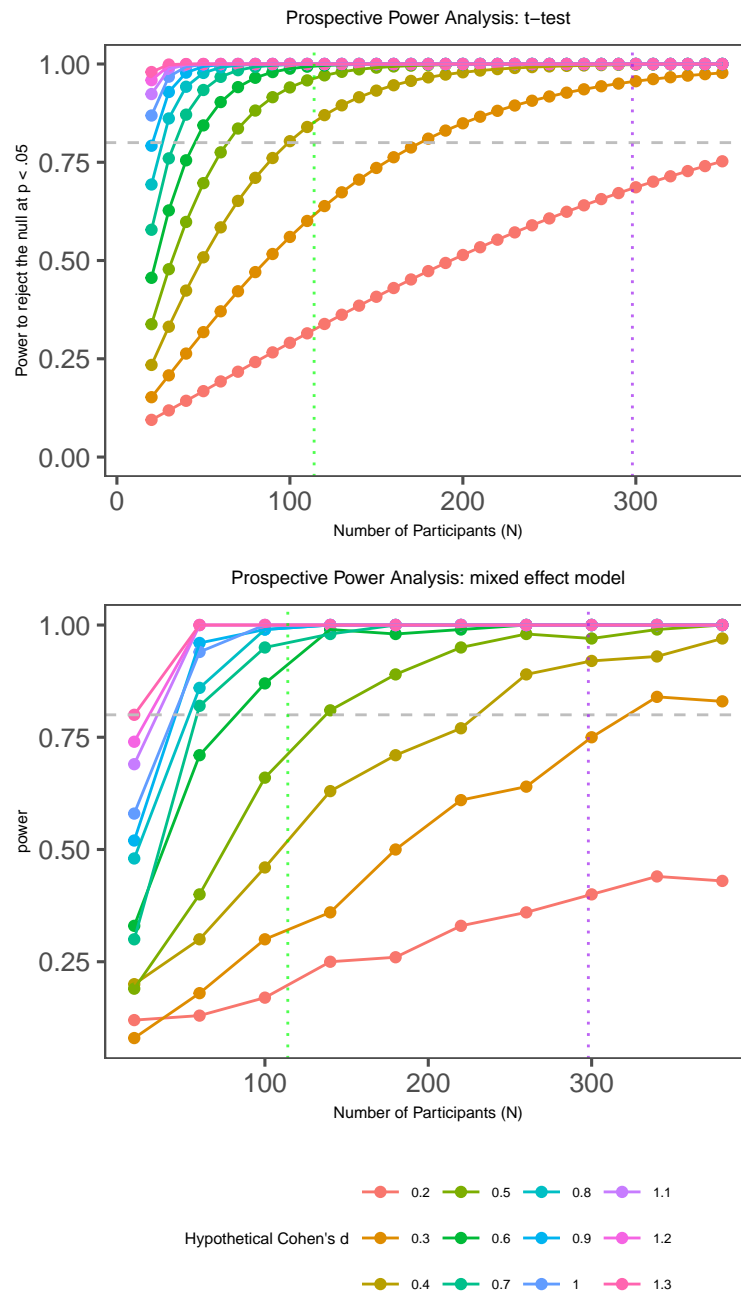
*Figure A1.* Simulation of power under different sample sizes and effect sizes. The top panel is under t-test assumptions, and the bottom panel is under linear mixed effect model assumptions. X-axis represents the number of particints in each condition. Y-axis represents the estimated power based on the effect sizes and the sample sizes. The horizontal gray dotted line on each plot represents 80% power. The two vertical lines, green and purple, represents the minimum sample size and the maximum sample size in our study, respectively.

Appendix B

Additional Visualizations for Task Results

**Free Description**

One measure from the Picture Free Description task is the number of descriptive accounts directed at the focal objects versus the background objects. In Experiment 1, we found a main effect of culture, with U.S. participants providing more description overall. We also found an interaction effect between culture and description types, with U.S. participants providing more background descriptions than focal descriptions. This pattern was inconsistent with the results reported in Imada et al. (2013). However, we found a pattern consistent with the original findings in experiment 2: Chinese participants provided more descriptive accounts directed at the background objects than the focal objects relative to U.S. participants. We reported the descriptive and inferential statistics in the main text.
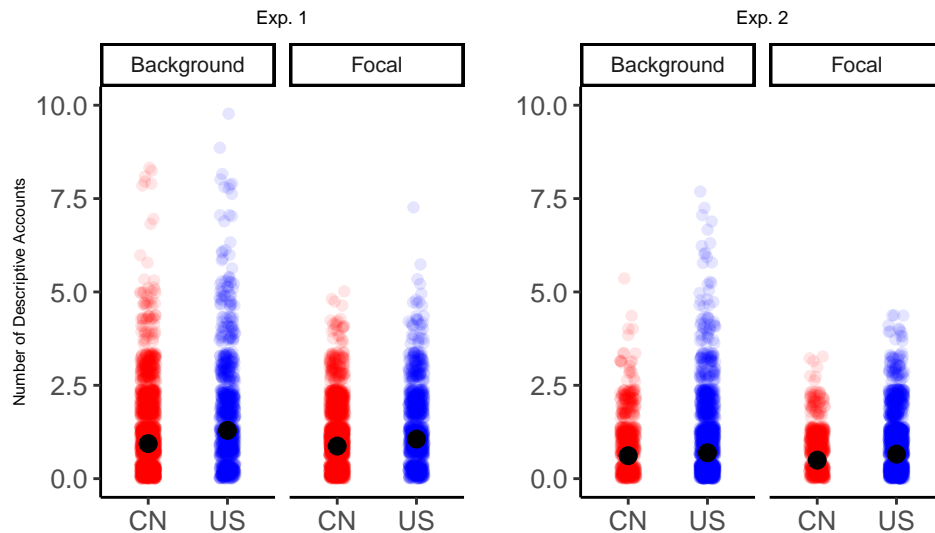


*Figure B1*. Results from the Picture Free Description task's descriptive accounts measures for experiment 1 and experiment 2 respectively. Each dot represents the number of descriptive accounts directed at either the focal object or the background object in each trial. Red indicates Chinese participants' responses, and blue indicates U.S. participants' responses. The black dots are the mean values for each group, with error bars showing 95% confidence intervals

1176 **Horizon Collage**

1177    In Horizon Collage task, we included two other measures in addition to the

1178 height of horizon: the number of stickers and the total area that the stickers

1179 covered. Here we included the visualization for the latter two measures. Unlike

1180 Senzaki et al. (2014), we did not see any cultural difference in either measure. We

1181 reported the descriptive and inferential statistics in the main text.
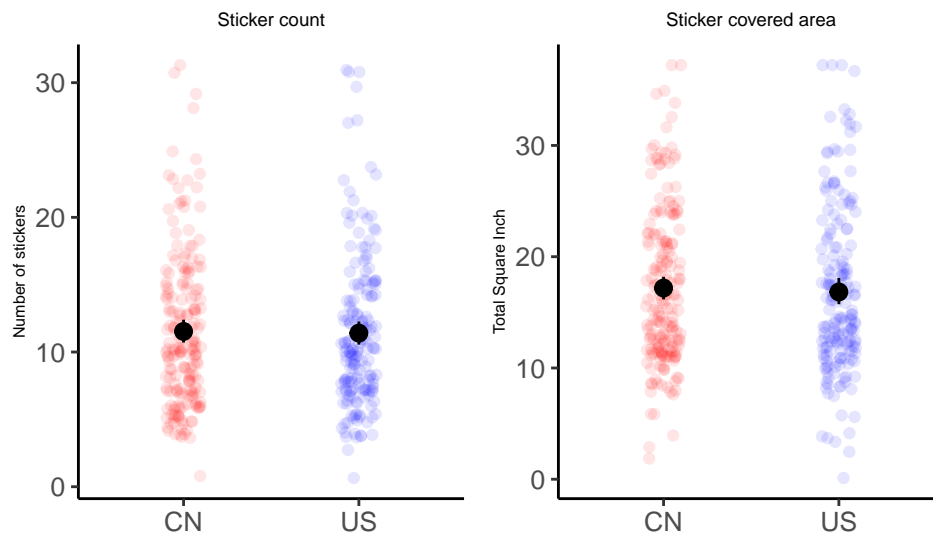


*Figure B2*. Results from the Horizon Collage task's sticker number and sticker covered area measures. Each dot represents one participant's response. Red indicates Chinese participants' responses, and blue indicates U.S. participants' responses. The black dots are the mean values for each group, with error bars showing 95% confidence intervals

Appendix C

Within-culture analysis on relationships between tasks.

1182    Na et al. (2010) found neglible correlations among a battery of tasks that revealed

1183    cultural differences within each culture. We found similar patterns in our data.

1184    Here we report four correlation matrices: US participants' correlation matrix is on

1185    the left, and Chinese participants' on the right. Each cell has a number indicating

1186    the correlation between the task marked on the column head and the task marked

1187    on the row head (Experiment 1: CA = Child Causal Attribution; EBB: Ebbinghaus

1188    Illusion; FD = Free Description; HZ = Horizon Collage; RMTS = Ambiguous

1189    cRMTS; RV = Raven's Progressive Matrices; SSI = Symbolic Self-Inflation; UP =

1190    Uniqueness Preference; Experiment 2: RMTS = Ambiguous cRMTS; RV = Raven's

1191    Progressive Matrices; FD = Free Description; CD = Change Detection; SSI =

1192    Symbolic Self-Inflation; CA = Causal Attribution; TTS = Taxonomic-Thematic

1193    Similarity; SeI = Semantic Intuition).

1194    **Experiment 1**



1195

1196 **Experiment 2**

### US

| | RMTS | RV | FD | CD | SSI | CA | TTS | Sel |
|------|------|------|------|------|------|------|------|------|
| Sel | 0 | −0.06 | −0.03 | −0.12 | −0.13 | −0.01 | 0.16 | 1 |
| TTS | 0.03 | 0.11 | 0.04 | 0.08 | 0.07 | −0.06 | 1 | 0.16 |
| CA | −0.06 | −0.16 | −0.02 | 0.14 | 0.09 | 1 | −0.06 | −0.01 |
| SSI | 0.07 | −0.09 | 0.08 | 0.11 | 1 | 0.09 | 0.07 | −0.13 |
| CD | −0.06 | −0.1 | 0.02 | 1 | 0.11 | 0.14 | 0.08 | −0.12 |
| FD | −0.06 | −0.05 | 1 | 0.02 | 0.08 | −0.02 | 0.04 | −0.03 |
| RV | −0.02 | 1 | −0.05 | −0.1 | −0.09 | −0.16 | 0.11 | −0.06 |
| RMTS | 1 | −0.02 | −0.06 | −0.06 | 0.07 | −0.06 | 0.03 | 0 |

### CN

| | RMTS | RV | FD | CD | SSI | CA | TTS | Sel |
|------|------|------|------|------|------|------|------|------|
| Sel | 0.11 | 0.21 | −0.13 | 0.28 | 0.12 | −0.08 | 0.09 | 1 |
| TTS | 0.08 | 0.19 | −0.16 | 0.28 | −0.27 | −0.12 | 1 | 0.09 |
| CA | −0.18 | −0.38 | 0.2 | −0.35 | −0.08 | 1 | −0.12 | −0.08 |
| SSI | −0.2 | 0.01 | −0.15 | −0.03 | 1 | −0.08 | −0.27 | 0.12 |
| CD | 0.23 | 0.31 | −0.27 | 1 | −0.03 | −0.35 | 0.28 | 0.28 |
| FD | −0.16 | −0.15 | 1 | −0.27 | −0.15 | 0.2 | −0.16 | −0.13 |
| RV | 0.26 | 1 | −0.15 | 0.31 | 0.01 | −0.38 | 0.19 | 0.21 |
| RMTS | 1 | 0.26 | −0.16 | 0.23 | −0.2 | −0.18 | 0.08 | 0.11 |

1197

Appendix D

Demographic variation in responses

1198       In the exploratory analysis section, we explored how basic demographic variables

1199 could moderate the task performances within each culture. To do so, we fit an fixed

1200 effect model predicting each task's response in each culture, with each demographic

1201 variable as the predictor. We fit 192 exploratory regression models in total. 26

1202 models were statistically significant, but only 1 survived Bonferroni correction.

1203 However, given that some of the demographic variables have been previously found

1204 to moderate task performance, we can treat this post-hoc exploratory analysis as a

1205 proxy for a conceptual replication. Therefore, we report selected uncorrected model

1206 results that are significant for key demographic variables in the tables below. These

1207 results should be interpreted with great cautions.

### Residential Mobility

1209       Higher residential mobility has been found to be related to individual's

1210 stronger emphasis on the independent self over the inter-dependent self (Ishii,

1211 Komiya, & Oishi, 2020; Oishi & Tsang, 2022). In our exploratory analysis's

1212 uncorrected models, residential mobility is a significant predictor to three task

1213 responses: U.S. participants' accuracy of Ebbinghaus Illusion task for U.S.

1214 participants, Chinese participants' proportion of first mention being focal objects in

1215 the Free Description task in Experiment 2, and U.S. participants' proportion of the

1216 taxonomic match selection in the Taxonomic-Thematic Similarity task. We reported

1217 the model results below.

| Study | Culture | Task Name | Beta | SE | t | p |
|---|---|---|---|---|---|---|
| Experiment 1 | US | Ebbinghaus Illusion | -0.01 | 0.01 | -2.01 | 0.05 |
| Experiment 2 | CN | Free Description | 0.03 | 0.01 | 2.45 | 0.02 |
| Experiment 2 | US | Taxonomic-Thematic Similarity | 0.01 | 0.00 | 2.59 | 0.01 |

### Socioeconomic Status

**MacArthur Ladder.** The MacAuthur ladder scale is a subjective measure of socioeconomic status (Adler et al., 2000). This measurement has been found to show moderate relationship with the objective socioeconomic measurements, and it has even more robust relationship to individual's subjective well beings than objective socioeconomic measurements (J. J. Tan, Kraus, Carpenter, & Adler, 2020). Past work has used this measure to investigate the relationship between socioeconomic status and individual's independent versus interdependent orientation (e.g., Miyamoto et al., 2018). In our current study, this measure was a significant predictor for five different tasks across two studies. In experiment 1, MacArthur Ladder scale response was negatively correlated with the proportion of participants choosing the unique pen in the U.S. participants. In experiment 2, this measure is negatively associated with the relational match Chinese participants selected in the Causal RMTS task, the score U.S. participants received in the Raven's Progressive Matrices, Chinese participants' reaction time in the Change Detection task, and the proportion of taxonomic-match Chinese participants selected in the Taxonomic-Thematic Similarity task. Theses results were summarized in the table below.

| Study | Culture | Task Name | Beta | SE | t | p |
|-------|---------|-----------|------|-----|-----|-----|
| Experiment 1 | US | Uniqueness Preference | -0.05 | 0.02 | -2.36 | 0.02 |
| Experiment 2 | CN | Causal RMTS | -0.04 | 0.02 | -2.05 | 0.04 |
| Experiment 2 | US | Raven's Progressive Matrices | -0.02 | 0.01 | -2.14 | 0.03 |
| Experiment 2 | CN | Change Detection | -735.86 | 221.15 | -3.33 | $< .001$ |
| Experiment 2 | CN | Taxonomic-Thematic Similarity | -0.03 | 0.01 | -2.05 | 0.04 |

**Maternal Education.** Maternal education is often regarded as a proxy for socioeconomic status (Hauser, 1994; Hoff & Tian, 2005). In the exploratory analysis, we first treated maternal education as a continuous variable. There was only one significant result: Chinese participants reported having higher maternal education made less situational attribution in the Causal Attribution task.

1241     Following common practices in the literature, we also binned the maternal education

1242     into a binary variable – whether participants have received college education or not.

1243     The binary variable was not significantly related to any task performance.

| Study | Culture | Task Name | Beta | SE | t | p |
|---|---|---|---|---|---|---|
| Experiment 1 | CN | Causal Attribution | -0.08 | 0.04 | -2.22 | 0.03 |

Appendix E

Examples of Tasks

**Experiment 1**

**Ambiguous Relational Match-To-Sample (cRMTS).** In the ambiguous cRMTS task (Carstensen et al., 2019, Experiment 3), participants were shown two pairs of objects, AB (A: blue square in the example setup shown in Figure E1; B: yellow circle) and AC (C: red triangle), activating a music machine during the training phase. In the test phase, they were given a forced choice between an object-based solution (a *same* pair of object A, i.e., a pair of blue squared in the example) and a relational solution (a *different* pair BC, i.e., yellow circle and red triangle in the example).
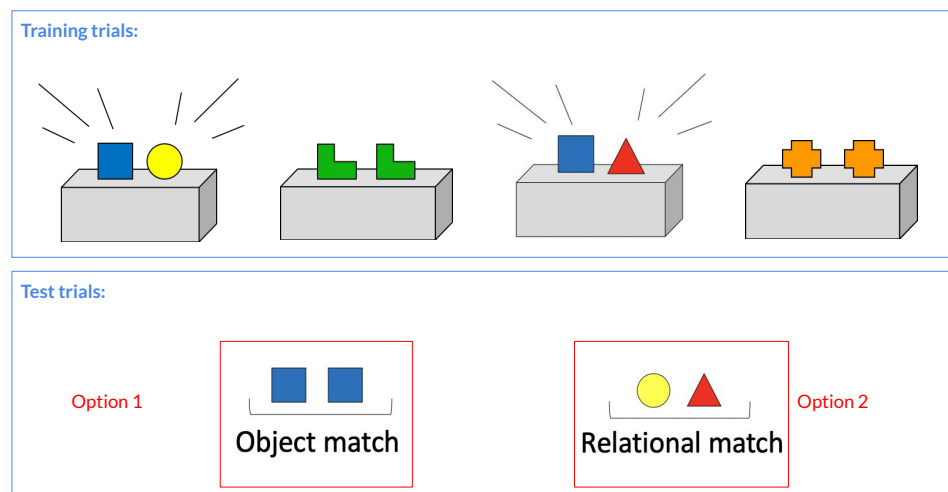


*Figure E1*. Ambiguous relational match-to-sample (cRMTS) task. Participants were given a forced choice between object-based solution (Option 1) and relational solution (Option 2) during the test phase.

**Picture Free Description.** In the picture free description task (Imada et al., 2013), Participants were shown a series of images with focal and background objects. For each image, they were asked to study it for 5 seconds, and then typed a description for it from memory.

... 5s ...

"Imagine you are talking to someone who did not see this picture. How would you tell them what you saw? "

*Figure E2*. A practice trial in picture free description task. In this image, the cat is the focal object, and the rest are background objects.

**Ebbinghaus Illusion.** In the Ebbinghaus illusion task (Imada et al., 2013), participants were asked to indicate which of the two orange circles (see Figure E3) was larger. They completed the task in two blocks: (1) no context block (upper right panel), where the two orange circles appeared independently; (2) illusional context block (bottom right panel), where each of the two orange circles were flanked by a grid of 8 gray circles which are all smaller or larger than the center orange circle.
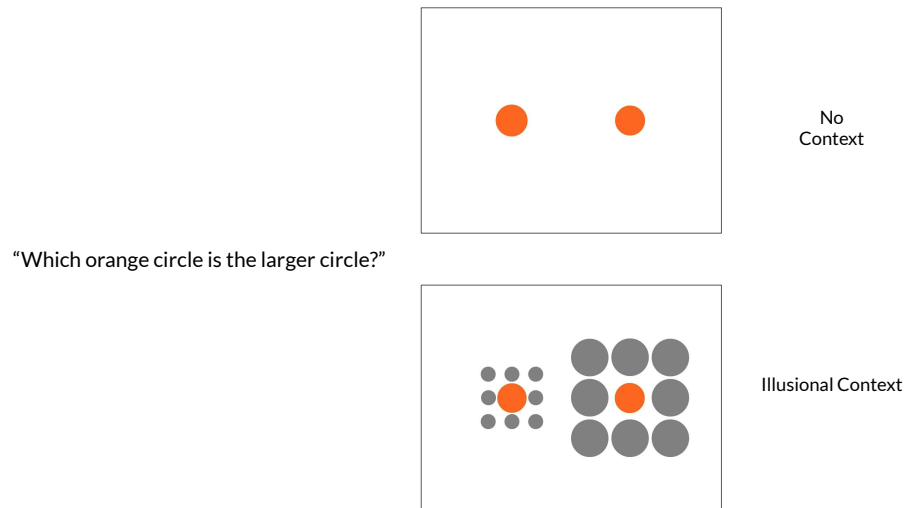
"Which orange circle is the larger circle?"

No Context

Illusional Context

*Figure E3*. Ebbinghaus illusion task. Upper right panel: no context block. Bottom right panel: illusional context block.

**Horizon Collage.** In the horizon collage task (Senzaki et al., 2014), participants were asked to create "a picture of the outside" by dragging-and-dropping sticker objects onto a rectangular "canvas." They were told to include a horizon line and could use any number of other stickers to create their image.
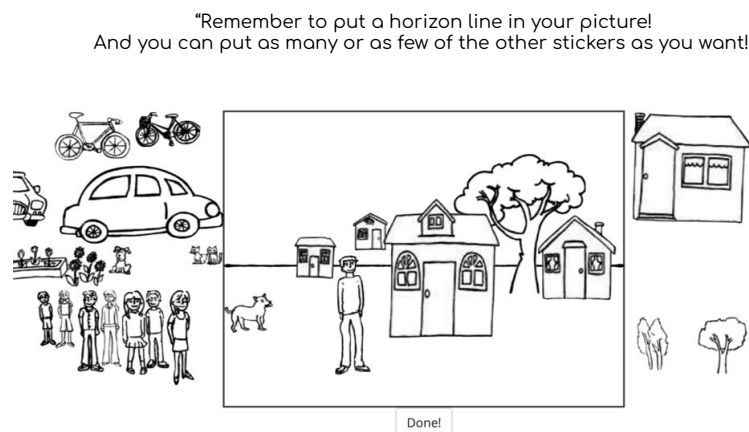


"Remember to put a horizon line in your picture!
And you can put as many or as few of the other stickers as you want!"

Done!

*Figure E4*. Horizon collage task. Participants created an image of the outside by dragging-and-dropping sticker objects onto a rectangular "canvas."

1268 **Symbolic Self-Inflation (Family).** In the adapted symbolic self-inflation

1269 task (Kitayama et al., 2009), we asked participants to draw themselves and their

1270 family members as circles by clicking-and-dragging their mouse on a rectangular
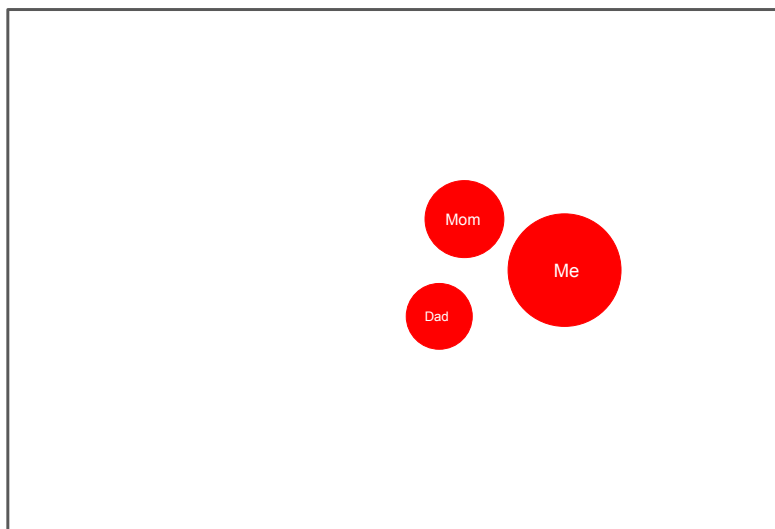
1271 "canvas."



*Figure E5*. Symbolic self-inflation task, family version. In this example, three circles were drawn representing the participant, their mother, and their father, respectively.

1272 **Uniqueness Preference.** In this adapted version of the uniqueness

1273 preference task (Kim & Markus, 1999), participants were asked to select one of the

1274 five dinosaur "stickers" that were identical except for color (four blue and one

1275 yellow).

"You can pick only one sticker. Which one do you want to pick?"



*Figure E6*. Uniqueness preference task. Participants were asked to select one of the five dinosaur "stickers."

**Child Causal Attribution.**   In the child causal attribution task (Seiver et al., 2013), participants watched a series of four short, animated vignettes in which two children both played in a pool and neither child played on a bicycle. They were then asked to explain in text why each child did not play on the bicycle.
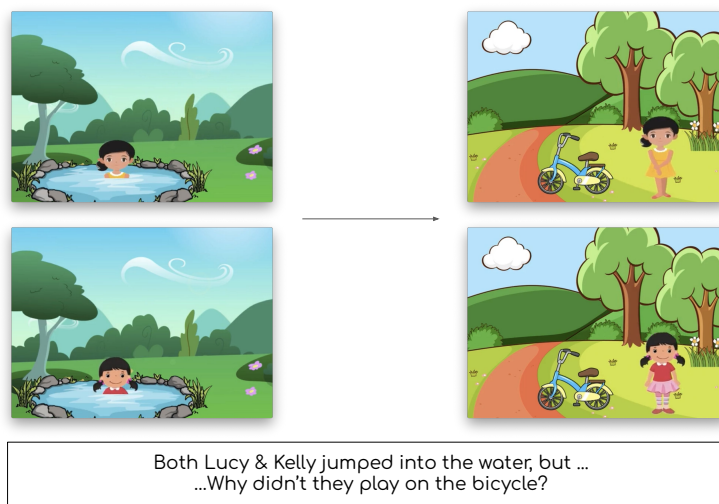


Both Lucy & Kelly jumped into the water, but ...
...Why didn't they play on the bicycle?

*Figure E7*. Child causal attribution task. Participants were asked to explain why each child did not play on the bicycle

1280 **Raven's Standard Progressive Matrices.** In Raven's standard
1281 progressive matrices task, participants were asked to select from a series of given
1282 options one that best completed the patterns demonstrated in the prompts.
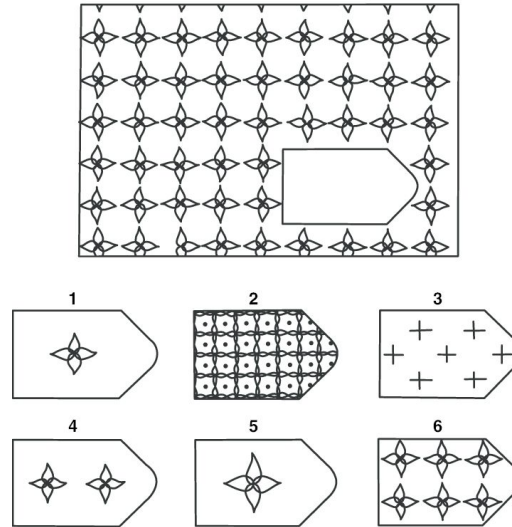


*Figure E8*. Raven's standard progressive matrices task. Participants were asked to select one option that best completed the patterns demonstrated in the prompts (upper panel).

1283 **Experiment 2**

1284 In Experiment 2, the following tasks followed the same procedures as in
1285 Experiment 1: ambiguous relational match-to-sample (ambiguous cRMTS), picture
1286 free description, and Raven's progressive matrices. The symbolic self-inflation task
1287 also followed similar procedures as in Experiment 1, except that participants were
1288 told to draw themselves and their *friends* into circles this time, as opposed to
1289 drawing draw themselves and their *family members* into circles in Experiment 1. We
1290 show examples for the remaining four tasks in Experiment 2 below.

1291 **Adult Causal Attribution.** In the adult version causal attribution task
1292 (Morris & Peng, 1994), participants were randomly assigned to read one of two
1293 crime stories (Royal Oak shooting or Iowa shooting). They were then asked to write
1294 a short explanation for the murderer's behaviors and rate a list of statements about

<sup>1295</sup> the causes of the murder (attributing to personal or situational factors) on a 7-point
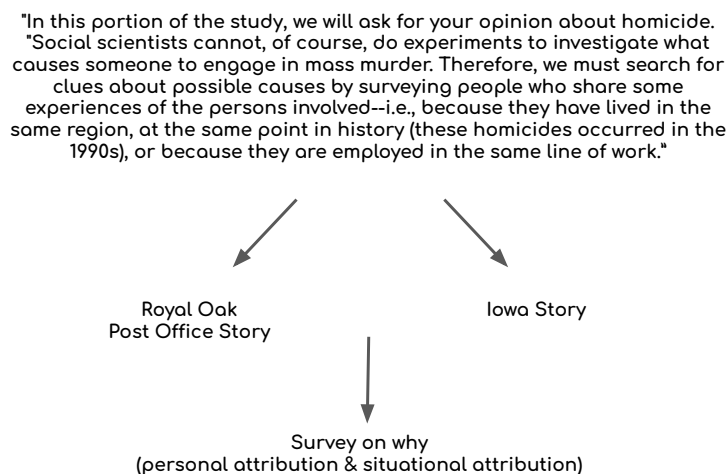
<sup>1296</sup> Likert scale.

"In this portion of the study, we will ask for your opinion about homicide. "Social scientists cannot, of course, do experiments to investigate what causes someone to engage in mass murder. Therefore, we must search for clues about possible causes by surveying people who share some experiences of the persons involved--i.e., because they have lived in the same region, at the same point in history (these homicides occurred in the 1990s), or because they are employed in the same line of work."

Royal Oak
Post Office Story

Iowa Story

Survey on why
(personal attribution & situational attribution)

*Figure E9*. Adult causal attribution task. Participants were randomly assiged to read one of two crime stories (Royal Oak shooting or Iowa shooting) and make attributions for the murder.

<sup>1297</sup> **Change Detection.** In each trial of the change detection task (Masuda &

<sup>1298</sup> Nisbett, 2006), a pair of images would alternate on the screen, each presented for

<sup>1299</sup> 560ms with a blank screen in between images for 80ms. The two pictures were

<sup>1300</sup> almost identical with subtle differences, either in the focal object (e.g., a tractor

<sup>1301</sup> with or without a wheel, as shown in the left panel of Figure E10) or the

<sup>1302</sup> background (e.g., a scaffold in the distance with varying size, as shown in the right

<sup>1303</sup> panel of Figure E10). Participants were instructed to press a key when they spotted

<sup>1304</sup> the difference, and then describe the difference in a text box. If they did not detect

<sup>1305</sup> a difference within 60 seconds, the trial timed out.

Change in <u>focal</u> objects    Change in <u>background</u> objects

*Figure E10*. Change detection task. Participants were instructed to detect the difference in each pair of images as quickly as they could. Left panel: change happened in a focal object. Right panel: change happened in the background.

**Taxonomic-Thematic Similarity.** In the taxonomic-thematic similarity task (Ji et al., 2004; Le et al., 2021), participants were presented with triads containing a cue word and two match options. In each test triad (as shown in Figure E11), one option was a taxonomic match (e.g., ladybug and bee are both insects) and the other a thematic match (e.g., bees usually appear in gardens). In each filler triads, the cue item and the options were broadly similar, thematically and taxonomically, making for a more ambiguous decision (e.g., monkey: elephant, tiger). Participants completed a two-alternative forced choice task in which they chose one match for each cue item.

In this game, you will see an item in bold, and two choices below it,
like this:

**bee**

ladybug  ← Taxonomic categorization

garden  ← Thematic categorization
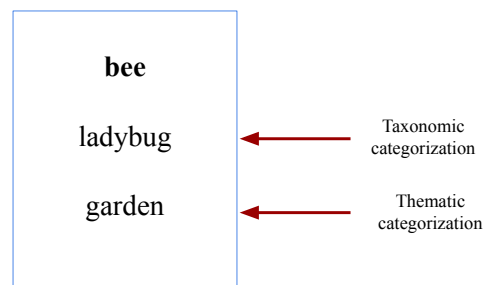
Choose the option that is most closely related to the bolded item.

*Figure E11*. Taxonomic-thematic similarity task. Participants completed a two-alternative forced choice task in which they chose one match for each cue item.

**Semantic Intuition.** In the semantic intuition task (Li et al., 2018), participants read five stories and judged the correctness of statements referring to a character after each story. The correctness judgement reflected their semantic intuition of either the descriptivist view (i.e., determine the referent of a name based on the description of the speaker) or the causal-historical view (i.e., determine the referent based on the original usage).

**Super Dog Race**

Long ago, there was a race called the Super Dog Race. Max, Pickles and Blaze participated in the race. Max crossed the finish line first, winning the race, but he got too excited and ran all the way to the North Pole. Pickles crossed the finish line second. He stopped and watched Max run away. The race announcer mistakenly thought that Pickles won the race. He told every newspaper in the world that Pickles won. He also told them that another dog, Blaze, ran very fast despite his short legs. Since then, everyone learned that Pickles won the race. They don't know anything else about Pickles.

Tom and Emily learned at school that Pickles won the Super Dog Race. This is the only thing they know about the dog race and Pickles. They don't know anything about Max. That night, their dad asked: Do you know who won the Super Dog Race?
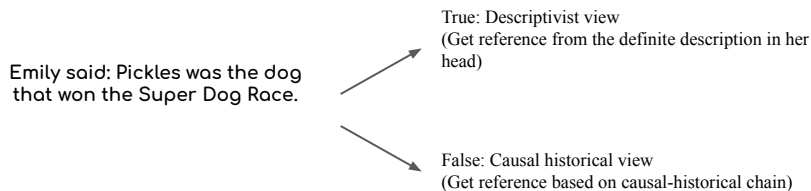
Emily said: Pickles was the dog that won the Super Dog Race.

True: Descriptivist view
(Get reference from the definite description in her head)

False: Causal historical view
(Get reference based on causal-historical chain)

*Figure E12*. Semantic intuition task. In this specific example, judging the statement as "true" indicates a descriptivist view, and judging it as "false" indicates a causal-historical view.