# CS 5824/ECE 5424: Advanced Machine Learning

## **Homework Assignment 1: Section A**

## **Ankit Parekh**

MS (Non-Thesis) Computer Engineering ECE Department ankitparekh@vt.edu

**Problem 1**: Question 1)

oblem 1: Question 1)
If 1,12, 1, 1 are Nobservations sampled from a Bernoulli
distribution with parameter P, the probability mass function for
each $\chi$ ; would be: $f(\chi_i:p) = p^{\chi_i} (1-p)^{1-\chi_i}$
$\pm (x_i:p) = p^{(1-p)}$
for n = 0 and 1 & 0 < p < 1.
n some de la company de la com
Likelihood Function $L(p) = \prod f(x_i; p) = p^{x_i}(1-p)^{1-x_i} p^{x_{in}}(1-p)^{1-x_{in}}$
CE (COLOR)
$L(p) = p = \chi_i (1-p)^n = \chi_i$
To find maximum liklihood, one can maximize ln L(P) instead of L(P),
since y = ln (x) is a naturally increasing function.
So me can find p at which loss L(P) is maximum.
or log L(P)
$\log L(p) = (\sum x_i) \log(p) + (n - \sum x_i) \log(1-p)$
Taking derivative and setting it to zero: $\frac{\partial \log(L(P))}{\partial P} = \frac{\sum x_i}{1 - \sum x_i} = 0$
Solving (1): multiply p(1-p) on both side:
$(\geq \chi_1)(1-1)-(\chi_1-2\chi_1)-1$
To indicate p as an estimate $\Rightarrow \hat{p} = \sum_{i=1}^{\infty} \hat{s}_i \cdot MLE \cdot \hat{q}_i \cdot \hat{p} = \sum_{i=1}^{\infty} \hat{z}_i$
n n
To show it is the maxima:
$\partial^2 \log (L(p)) = -\frac{\sum \chi_i}{i=1} - \frac{\sum (1-\chi_i)}{\sum (1-\chi_i)^2}$
2P P2 (1-P)2
This term in negative
Indicating p is a maxima.

Since  $y = \ln(x)$  is a naturally increasing function, we can try to find maxima of Log-liklihood to get the maxima of liklihood.

Log-liklihood function is  $l(1;x_1...x_N) = -N(1-\sum \ln(x_i!) + \ln(1)\sum x_i$ We need to maximize  $\log$ -liklihood and find the 1 for the derivative is 0.

 $\frac{dl(1;\chi_1...\chi_N) = -N + 1}{dl} \sum_{i=1}^{N} \chi_i = 0$ 

of  $\lambda = 1 > x_i$  MLE is  $\hat{\lambda} = 1 > x_i$  which is the mean  $N = 1 > x_i$  which is the mean  $N = 1 > x_i$  of the observations

By definition of expertation:  $E(Y) = \sum_{y \in Y} P(Y=y)$ For Poisson Distribution:  $E(Y) = \sum_{y > 0} y \cdot \lambda^{y} e^{-\lambda} \dots (I)$ Solving  $(I) : E(Y) = \lambda e^{-\lambda} \sum_{y > 1} \frac{1}{(y-1)!}$   $= \lambda e^{-\lambda} \sum_{z > 0} \frac{1}{z!}$   $= \lambda e^{-\lambda} e^{\lambda}$ Taylor Series Expansion for  $e^{\lambda}$ 

E(Y) = 1 : Expertation of Poisson Distribution

### **Problem 2:** Question 1) Used Python for calculations

import numpy as np

```
1. vehicle_speed_dataset = {'number_of_wheels':[4, 4, 2, 8, 4, 3], 'cost':[15000, 25000, 5000, 40000, 22000, 17000]}
2. def standardize_feature(arr):
3.    feature = np.array(arr)
4.    mean = np.mean(feature)
5.    print("Mean: ", mean)
6.    std = np.std(feature)
7.    print("Standard Deviation: ", std)
8.    arr = [(x - mean)/std for x in arr]
9.    print("Standardized Values:", arr)
10.    print("Number of Wheels : ", vehicle_speed_dataset['number_of_wheels'])
11.    standardize_feature(vehicle_speed_dataset['cost'])
12.    print("Cost : ", vehicle_speed_dataset['cost'])
13.    standardize_feature(vehicle_speed_dataset['cost'])
```

#### **Answer (Output):**

Number of Wheels: [4, 4, 2, 8, 4, 3]

Mean: 4.167

Standard Deviation: 1.863

Standardized Values: [-0.0894, -0.0894, -1.1628, 2.057, -0.0894, -0.626]

Cost: [15000, 25000, 5000, 40000, 22000, 17000]

Mean: 20666.667

Standard Deviation: 10687.479

Standardized Values: [-0.530, 0.405, -1.466, 1.809, 0.125, -0.343]

#### **Problem 2:** Question 2)

oblem 2: Question 2)		
Q.2) B	Linear Model hw (x) = No + N, x, + W2x2	
	Mean squared error is the sum of squared differences between the	
	Mean squared error is the sum of squared differences between the predicted and the true values.	
	Cost function J(w) for least equevres is given by: Features:	
	vine of whale	
	$J(\omega) = 1 \leq (\text{predicted}; -\text{true};)^2 \approx \cos t$	
	where n = number of samples	
X	$J(\omega) = \frac{1}{n} \sum_{i=1}^{n} \left( h_{\omega}(x_{i})^{2} - y_{i}^{\alpha} \right)^{2}$	
	$J(w) = \frac{1}{n} \left( \left( w_0 + w_1 x_1^{(i)} + w_2 x_2^{(i)} \right) - (y^{(i)})^2 \right)$	
S. B. B. B. A.	n i=1	

<u>obiem 2</u> : Questi	011 3)
Q.3) # W	le have used least Squares Cost function for this problem
	of find the parameters (w) of a hyperparameter plane
	which best fite the linear regression dataset.
* (	imputing the total squared error between the associated
1	hyperparameter plane and the true values gives an
	appropriate measure of how well does the line fits the
	detaset Right ve signi. It has been resultatives and the
	The order has a long number of known would be in bour
*	It also takes into account that the difference between
	medicted values and tome values might have different
	signs for different points and hence squaring the individual
-	count so that both positive and negative values are treated
	egually.
roblem 23 Q	2) 10 kg (c) = 10 hyperplane
	(hw(x))
((2)0)	o predicted outros & 14:- had
	sample point > ("true!" square
	C'true!
	Value) Yes
	The state of the s
* \	Ve get the area marked in blue Illi using Least squares framework
	and we want to minimize this area. We try to minimize this
	arrea using parameter (w) tuning.
	minimize $J(w) = minimize 1 \leq ((w_0 + w_1 x_1'' + w_2 x_2'') - y'')$
	minimize $J(w) = minimize $ $1 = \frac{\pi}{((w_0 + w_1 x_1^{(i)} + w_2 x_2^{(i)}) - y^{(i)})}$
	The state of the s

The residuals being normally distributed, L2 regularization performs better for linear regression.

I We want our approximation to pendize outliers more and capture The general trend. The Least Squares is also differentiable at all places and hence can be used as a cost function gradient descent

Problem 2: Question 4)

ucoción 1)
Q4) Partial differentiating the cost function $J(w)$ : xxisth sample
n.
$\frac{\partial}{\partial x} J(\omega) = \frac{\partial}{\partial x} \frac{1}{2} \sum_{i} \left( h_{i} \chi_{i}(x^{(i)}) - \chi_{i}(x^{(i)})^{2} \right)$
dwi dwi n ist
= 1.2 5 (hw(x(s)) - y(i)). 2 (hw(x(i)))
n = 1
= 1.2 = (hw (x(x)) - y(x)). 2 (wo+w1x1+w1x)
n = owi
+
°. ∂ J(w) = 5 (hw(x4)) - y(i)). (xi)
) dw; j=1
$\frac{n}{1-n}$
$\Im(\omega) = \Im((\hbar_{W}(\chi^{(Q')}) - \chi^{(q')})$
$\int \partial w_{o} \int_{j=1}^{j=1}$

Problem 2: Question 5)

Problem 2: Question 6)

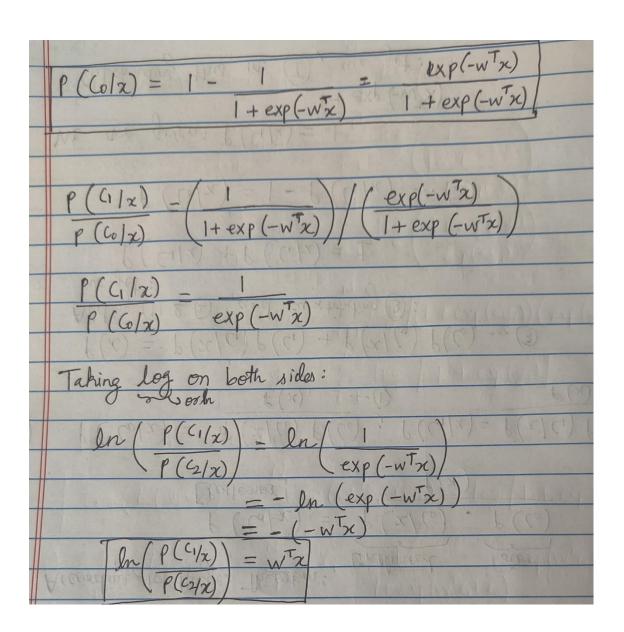
Problem 2:	Q6)
	1. We can incoease the number of features in the model
	to reduce bias. If we decrease the regularization it will
	increase variance, and the model might fit better to the
	distribution. Model complexity will increase.
	<u> </u>
*	2. We can decrease the model complexity by decreasing the number of features and increasing regularization. This will
	number of features and increasing regularization. This will
	introduce more bias and reduce variance respectively which
	should fix the over-fitting issue.

Problem 3: Question 1)

obicin 5.	Question 1)
Q1)	P(Co): Probability of patient having infectious (class Co) disease.
	P(Colx): Probability of patient having infectious (class Co) disease
	P(Co/x): Probability of patient having infectious (class Co) disease given he she exhibits observed symptoms in data vector
	Y. The state of th
	P(x/6): Probability of patient Behibiting observed symptoms (in data vector x) given that he/she has the infections disease (class Co).
	verter x) given that he/she has the infections disease
	(clas co).
	According to Baye's Theorem:
	According to Baye's Theorem: Likelihood Prior
	$P(\zeta_0 \chi) = P(\chi/\zeta_0) \cdot P(\zeta_0)$
	[Postenor] P(x)

**Problem 3:** Question 2)

02)	P(Colx) = P(x/co) P(Co); P(C1/x) = P(x/ci) P(Ci)
42/	
	$P(x) \cdots D$ $P(x) \cdots (2)$
	Taking Sey-of both jillsin - 10 1 1 1 1
	P(x) = P(x/c1) P(c1) + P(x/6) P(6) (3)
	Adding (1) & 2 and substituting (3): We get:
	P(C1/2) + P(C0/2) = 1
	1 (01-1) (1+ xx (-M3) ) (1+ xx (-M3))
	6 (1) P (Colx) = 1 - P (Colx) (I)
	THE RESTRICT OF THE PROPERTY OF THE PARTY OF
	We are given: P(G/x) = 1
	+ sxb (-w/x)
	Substituting this in (I), we get:



**Problem 3:** Question 3) Sigmoid function is mapping range of y; values into a [o,t] interval Class Mapping:  $\sigma(y_i) > 0.5 \Rightarrow (lass 1 ) Ci \in \{0,1\}$   $\sigma(y_i) < 0.5 \Rightarrow (lass 2)$   $y_i = w_{k_i} + w_0$  3 Legistic Regression for Binary ClassificationP.M.F. for o (yi) → f(x) (1 - f(x)) Likelihood Function = 1 +(x) (1-+(x)) Likelihood Function = L(w, wo) = To(yi) [ Taking log of likelihood function: log(L(w, w)) = log To(yi) (1- o(yi) log (o (y;)) + 1/1+exp(-yi) 1/1+exp(=yi) exp(-yi) Using (I) an (I): log(L(w, wo)) = > ciy; + > log(1-o(yi))

