

Proof: $E[\epsilon|x] = 0 \rightarrow$ conditional form

$$E[\epsilon_i | x_i = x] = 0$$

$$E[y_i - \beta_0 - \beta_1 x_i | x_i = x] = 0$$

Population & Sample Mean

$$E[Y] = \int y f_y(y) dy$$

$$\bar{Y} = \frac{1}{n} \sum Y_i$$

Derivation: OLS Coefficients

$$\min_{b_0, b_1} \hat{E}[(Y_i - (b_0 + b_1 X_i))^2]$$

$$= \min_{b_0, b_1} \frac{1}{n} \sum_{i=1}^n [(Y_i^2 - 2Y_i(b_0 + b_1 X_i) + (b_0 + b_1 X_i)^2)]$$

$$\frac{\partial}{\partial b_0} \Rightarrow \left\{ \frac{1}{n} \sum_{i=1}^n [-2Y_i + 2(b_0 + b_1 X_i)] = 0 \right.$$

$$\hookrightarrow \hat{b}_0 = \frac{1}{n} \sum Y_i - \frac{1}{n} \sum X_i$$

$$\frac{\partial}{\partial b_1} \left\{ \Rightarrow \frac{1}{n} \sum_{i=1}^n [-2Y_i X_i + 2(\hat{b}_0 + \hat{b}_1 X_i) X_i] = 0 \right.$$

↓ plug in \hat{b}_0

$$\hat{b}_1 = \hat{Cov}(X, Y) / \hat{Var}(X)$$

$$\hat{b}_0 = \bar{Y} - \hat{b}_1 \bar{X}$$

Population & Sample Variance

$$Var(Y) = E[(Y - E[Y])^2]$$

$$\widehat{Var}(Y) = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$

Population & Sample Covariance

$$Cov(X, Y) = E[(Y - E[Y])(X - E[X])]$$

$$\widehat{Cov}(XY) = \frac{1}{n-1} \sum_{i=1}^n [(Y_i - \bar{Y})(X_i - \bar{X})]$$

Derivation: Omitted Variable Bias

$$Y_i = \beta_0 + \beta_1 x_i + \underbrace{\gamma s_i + \eta_i}_{\epsilon_i}$$

↳ s_i initially left out of regression

Equations & Proofs

$$\begin{aligned} \hat{b}_1 &= Cov(Y, X) / Var(X) \\ &= \frac{Cov(\beta_0, \bar{X}) + Cov(\beta_1, X_i, X_i) + Cov(\epsilon_i, X_i)}{Var(X)} \\ &= \beta_1 + \frac{Cov(\gamma s_i, X_i) + Cov(\eta_i, X_i)}{Var(X)} \end{aligned}$$

$$\boxed{\hat{b}_1 = \beta_1 + \gamma \frac{Cov(s_i, X_i)}{Var(X)}}$$

SSR, SST, SSE

SSR	SST	SSE
(Sum of Squared residuals)	(Sum of Squared Total)	(Sum of Squared Errors)

$$SSR = \sum \hat{Y}_i^2 \quad SST = \sum (Y_i - \bar{Y})^2 \quad SSE = \sum (Y_i - \hat{Y}_i)^2$$

Derivation: \hat{b} matrix X

$$\min((Y - Xb)'(Y - Xb))$$

$$\min((Y'Y - Y'Xb) - (b'X'Y - b'X'Xb))$$

$$= \min(-2Y'Xb + b'X'Xb)$$

$$\downarrow \boxed{\frac{\partial}{\partial b}}$$

$$X'Xb = X'Y$$

$$\hat{b} = (X'X)^{-1} X'Y$$

Derivation: No Perfect collinearity

$$E[\hat{b}|x] = E[(X'X)^{-1} X'Y | x]$$

$$= E[(X'X)^{-1} X'(X\beta + \epsilon) | x]$$

$$= \beta + E[(X'X)^{-1} X' E[\epsilon | x]]$$

$$= \boxed{\beta}$$

Homoskedasticity of Error Term

$$V(\hat{b}|x) = (X'X)^{-1} X' V(Y|x)(X'X)^{-1}$$

$$= (X'X)^{-1} X' V(X'\beta + \epsilon | x) (X'X)^{-1}$$

$$= (X'X)^{-1} X' \sigma^2 I ((X'X)^{-1} X')$$

$$= \sigma^2 [(X'X)^{-1} X' X (X'X)^{-1}]$$

$$= \sigma^2 (X'X)^{-1}$$

Asymptotic Normal Distribution of \hat{b}

$$\sqrt{n}(\hat{b} - \beta)$$

$$= \sqrt{n}((X'X)^{-1} X' Y - \beta)$$

$$= \sqrt{n}((X'X)^{-1} X' (X\beta + \epsilon) - \beta)$$

$$= \sqrt{n}(X'X)^{-1} X' \epsilon$$

$$= \underbrace{\frac{1}{n} \sum X_i X_i'} * \underbrace{\frac{1}{n} \sum X_i \epsilon_i}_{Q_x \text{ by LLN}}$$

$N(0, A)$ by CLT

Simultaneity Problem

$$Cov(p_t, M_{t+1}) = Cov(Y + \delta a_t + z_t^d + \lambda^d + u_t, M_{t+1})$$

$$= \delta Cov(a_t, u_{t+1}) + Cov(u_t, M_{t+1})$$

$$= \delta Cov(\alpha + \beta p_t + z_t^d + \lambda^d + H_{t+1}, M_{t+1}) + "$$

$$= \delta (\beta Cov(p_t, M_{t+1}) + Var(M_{t+1})) + "$$

$$Cov(p_t, u_{t+1}) = \frac{\delta Var(M_{t+1}) + Cov(M_{t+1}, H_{t+1})}{1 - \delta \beta}$$

Derivation: IV estimator

$$E[Z_i \epsilon_i] = 0$$

$$E[Z_i(Y_i - x_i' \beta)] = 0$$

$$E[Z_i Y_i - Z_i x_i' \beta] = 0$$

$$E[Z_i Y_i] - E[Z_i x_i' \beta] = 0$$

$$\boxed{\beta = (Z'X)^{-1} Z' Y}$$

Equivalence of Simultaneous to Two Stage Least Squares

$$\hat{b} = (\hat{X}'X) \hat{X}' Y$$

$$= (X' P_2 X)^{-1} X' P_2 Y$$

$$= (X' Z (Z' Z)^{-1} Z' X)^{-1} X' Z (Z' Z)^{-1} Z' Y$$

$$= (Z' X)^{-1} (Z' Z) (Z' Z)^{-1} X' Z (Z' Z)^{-1} Z' Y$$

$$\boxed{(Z' X)^{-1} Z' Y}$$

Proving Random Effects

$$A3: \text{Cor}(\tilde{\epsilon}_{it}, \tilde{\epsilon}_{is}|x) = 0$$

$$\text{Cov}(\tilde{\epsilon}_{it}, \tilde{\epsilon}_{is}|x) = \text{Cov}(\epsilon_{it} - \lambda \bar{\epsilon}_i, \epsilon_{is} - \lambda \bar{\epsilon}_i|x)$$

$$= \text{Cov}(\epsilon_{it}, \epsilon_{is} - \lambda \bar{\epsilon}_i|x) - \lambda \text{Cov}(\bar{\epsilon}_i, \epsilon_{is} - \lambda \bar{\epsilon}_i)$$

$$= \boxed{\sigma_a^2 - 2\lambda \text{Cov}(\epsilon_{it}, \bar{\epsilon}_i) + \lambda^2 \text{Var}(\bar{\epsilon}_i|x)}$$

$$\text{Var}(\bar{\epsilon}_i|x) = \text{Var}\left(\frac{1}{T} \sum_{t=1}^T \epsilon_{it}|x\right)$$

$$= \frac{1}{T^2} \text{Var}\left(\sum_{t=1}^T (\alpha_i + \beta_i t)|x\right)$$

$$= \frac{1}{T^2} \text{Var}(T\alpha_i + \sum_{t=1}^T \beta_i t|x)$$

$$= \frac{1}{T^2} (T^2 \sigma_a^2 + T \sigma_v^2)$$

$$= \sigma_a^2 + \frac{1}{T} \sigma_v^2$$

$$= \boxed{\frac{1}{T} (\sigma_v^2 + T \sigma_a^2)}$$

Now,

$$\text{Cov}(\epsilon_{it}, \bar{\epsilon}_i|x) = \text{Cov}(\epsilon_{it}, \frac{1}{T} \sum \epsilon_{is})$$

$$= \boxed{\frac{1}{T} (\sigma_v^2 + T \sigma_a^2)}$$

Final

$$\text{Cov}(\tilde{\epsilon}_{it}, \tilde{\epsilon}_{is}|x)$$

$$= \sigma_a^2 - 2\lambda \text{Cov}(\epsilon_{it}, \bar{\epsilon}_i) + \lambda^2 \text{Var}(\bar{\epsilon}_i|x)$$

$$= \boxed{0}$$

Proof ①

$$\begin{aligned} E[\epsilon_i | x_i = x] &= 0 \\ E[y_i - \beta_0 - \beta_1 x_i | x_i = x] &= 0 \\ E[y_i | x_i = x] - \beta_0 - \beta_1 x_i &= 0 \\ E[y_i | x_i = x] &= \beta_0 + \beta_1 x_i \end{aligned}$$

Population ② Sample Covariance

$$\begin{aligned} \text{Cov}(x, Y) &= E[(x - E[x])(Y - E[Y])] \\ \widehat{\text{Cov}}(x, Y) &= \frac{1}{n-1} \sum (y_i - \bar{Y})(x_i - \bar{x}) \end{aligned}$$

Proof that $\beta_1 = \hat{\beta}_1$, $\beta_0 = \hat{\beta}_0$ ④

$$\begin{aligned} \hat{\beta}_1 &= \frac{\text{Cov}(x, Y)}{\text{Var}(x)} \\ &= \frac{\text{cov}(x, \beta_0 + \beta_1 x + \epsilon)}{\text{Var}(x)} \\ &= \frac{0 + \beta_1 \text{Var}(x) + 0}{\text{Var}(x)} = \boxed{\beta_1} \end{aligned}$$

$$\begin{aligned} \hat{\beta}_0 &= E[Y_i] - \beta_1 E[x_i] \\ &= E[\beta_0 + \beta_1 x_i + \epsilon_i] - \beta_1 E[x_i] \\ &= \beta_0 + \beta_1 E[x_i] - \beta_1 E[x_i] \\ &= \boxed{\beta_0} \end{aligned}$$

Derivation of Omitted Variable Bias ⑤

$$Y_i = \beta_0 + \beta_1 x_i + \underbrace{\gamma s_i}_{\epsilon_i} + n_i$$

\hookrightarrow s_i initially left out of regression

$$\hat{\beta}_1 = \text{Cov}(Yx) / \text{Var}(x)$$

$$\begin{aligned} &= \frac{\text{Cov}(\beta_0, x) + \text{Cov}(\beta_1 x_i, x_i) + \text{Cov}(\epsilon_i, x)}{\text{Var}(x)} \\ &= \beta_1 \frac{\text{Var}(x)}{\text{Var}(x)} + \frac{\text{Cov}(\gamma s_i, x)}{\text{Var}(x)} + \frac{\text{Cov}(n_i, x)}{\text{Var}(x)} \\ &\hat{\beta}_1 = \beta_1 + \gamma \frac{\text{Cov}(s_i, x)}{\text{Var}(x)} \end{aligned}$$

SSR, SST, SSE

$$\begin{array}{|c|c|c|} \hline \text{SSR} & \text{SST} & \text{SSE} \\ \hline (\text{sum of squares of residuals}) & (\text{sum of squares total}) & (\text{sum of square errors}) \\ \hline \end{array}$$

$$\text{SSR} = \sum \hat{\beta}_0^2 \quad \text{SST} = \sum (y_i - \bar{Y})^2 \quad \text{SSE} = \sum (y_i - \hat{y}_i)^2$$

Population ② Sample Mean ②

$$E[Y] = \int y f_y(y) dy$$

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$$

Population ② Sample Variance

$$\text{Var}(Y) = E[(Y - E[Y])^2]$$

$$\widehat{\text{Var}}(Y) = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$

OLS COEFFICIENT Derivation ③

$$\begin{aligned} &\min_{b_0, b_1} \hat{E}[(Y_i - (b_0 + b_1 x_i))^2] \\ &= \min_{b_0, b_1} \frac{1}{n} \sum_{i=1}^n [Y_i^2 - 2Y_i(b_0 + b_1 x_i) + (b_0 + b_1 x_i)^2] \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial b_0} \left\{ \frac{1}{n} \sum_{i=1}^n [-2Y_i + 2(b_0 + b_1 x_i)] \right\} &= 0 \\ \hat{b}_0 &= \frac{1}{n} \sum Y_i - \frac{1}{n} \sum x_i \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial b_1} \left\{ \frac{1}{n} \sum_{i=1}^n [-2Y_i x_i + 2(\hat{b}_0 + \hat{b}_1 x_i) x_i] \right\} &= 0 \end{aligned}$$

$$\begin{aligned} \hat{b}_1 &= \frac{\sum Y_i x_i - \frac{1}{n} \sum Y_i \sum x_i}{\sum x_i^2 - \frac{1}{n} (\sum x_i)^2} \\ &\Downarrow \\ \hat{b}_1 &= \widehat{\text{Cov}}(x, Y) / \widehat{\text{Var}}(x) \end{aligned}$$

Derivation of $\hat{\beta}$ Matrix (Multiple) ⑥

$$\begin{aligned} &\min((Y - X\beta)'(Y - X\beta)) \\ &\min((Y'Y - Y'X\beta) - (b'X'Y - b'X'X\beta)) \\ &\min(-2Y'X\beta + b'X'X\beta) \\ &\quad \frac{\partial}{\partial \beta} \end{aligned}$$

$$X'X\beta = X'Y$$

$$\hat{\beta} = (X'X)^{-1} X'Y$$

P-value for t-test

$$P = F_{(n-(k+1))}(-|t_{\hat{\beta}_1}|) +$$

$$1 - F_{(n-(k+1))}(|t_{\hat{\beta}_1}|)$$

Asymptotic Normal Distribution of $\hat{\beta}$ ⑨

$$\begin{aligned} &\sqrt{n}(\hat{\beta} - \beta) \\ &= \sqrt{n}((X'X)^{-1} X'Y - \beta) \end{aligned}$$

$$= \sqrt{n}((X'X)^{-1} X'(X\beta + \epsilon) - \beta)$$

$$= \sqrt{n}(X'X)^{-1} X' \epsilon$$

$$= \frac{1}{n} \sum_{i=1}^n x_i x_i' * \frac{1}{\sqrt{n}} \sum_{i=1}^n x_i \epsilon_i$$

\Rightarrow by LLN

$N(0, A)$ by CLT

No Perfect Collinearity ⑦

$$\begin{aligned} E[\hat{\beta} | X] &= E[(X'X)^{-1} X'Y | X] \\ &= E[(X'X)^{-1} X'(X\beta + \epsilon) | X] \\ &= E[(X'X)^{-1} X'X\beta + (X'X)^{-1} X'\epsilon | X] \\ &= \beta + E[(X'X)^{-1} X' \epsilon | X] \\ &= \boxed{\beta} \end{aligned}$$

Homoskedasticity of error term ⑧

$$\begin{aligned} V(\hat{\beta} | X) &= (X'X)^{-1} X' V(Y | X) (X'X)^{-1} \\ &= (X'X)^{-1} X' V(X\beta + \epsilon | X) (X'X)^{-1} \\ &= (X'X)^{-1} X' \sigma^2 I (X'X)^{-1} \\ &= \sigma^2 (X'X)^{-1} X' X (X'X)^{-1} \\ &= \boxed{\sigma^2 (X'X)^{-1}} \end{aligned}$$

