

ECON526: Quantitative Economics with Data Science Applications

Regression Analysis

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Overview



Summary

- Previously we have discussed estimating treatment effects using the difference in means.
- In this lecture we will discuss how to estimate treatment effects using regression.
- I assume that you are familiar with the basics of linear regression, including the assumptions underlying OLS and the interpretation of regression coefficients.



Linear Regression



- Simple linear regression is a statistical method that allows us to model the relationship between two variables: a dependent variable and an independent variable.
- The goal of simple linear regression is to find the line of best fit that describes the relationship between the two variables.
- The line of best fit is defined as the line that minimizes the sum of the squared differences between the observed values and the predicted values.
- Mathematically, the linear regression equation is the least-norm solution to an overdetermined system of linear equations.

$$\rightarrow y = X\beta + \epsilon \longrightarrow \epsilon = y - X\beta$$

$$\rightarrow \min_{\beta} ||\epsilon||_2^2 = \min_{\beta} \epsilon' \epsilon = \min_{\beta} (y - X\beta)' (y - X\beta)$$

$$\rightarrow \hat{\beta} = \arg\min_{\beta} (y - X\beta)'(y - X\beta) = (X'X)^{-1}X'y$$



- Aside: Computational complexity of OLS
- We could solve for the OLS coefficient in a number of ways. However, some ways will be substantially faster than others.

```
from scipy.optimize import minimize
   import statsmodels.api as sm
   import timeit
   # Generate data
   np.random.seed(123)
   n = 10000
 8 X = np.random.normal(size=(n, 10))
   y = np.random.normal(size=n)
10
   # Define some different OLS functions
12
   # Most naive way - never do this!
   def ols lstsq(X, y):
       return minimize(lambda beta: np.linalg.norm(X @
```

```
Time to compute OLS coefficients numerically:
12 ms \pm 259 \mus per loop (mean \pm std. dev. of 7 runs, 100
loops each)
Time to compute OLS coefficients with a matrix inverse:
221 \mus \pm 10.1 \mus per loop (mean \pm std. dev. of 7 runs,
1,000 loops each)
Time to compute OLS coefficients by solving a linear
system:
120 \mus \pm 2.09 \mus per loop (mean \pm std. dev. of 7 runs,
10,000 loops each)
```



- Aside: Computational complexity of OLS
- Solving a matrix inverse involves computing the inverse of a matrix, which is a computationally expensive operation.
 - \rightarrow The most common way to find the inverse of a matrix is to solve n linear systems, where n is the number of columns in the matrix.
 - \rightarrow Each linear system is of the form $(X'X)\beta=e_i$, where X is a matrix and e_i is a vector of zeros with a 1 in the $i_{
 m th}$ position.
 - \rightarrow Solved this way, it would take $O(n^3)$ operations to find the full inverse.
 - \rightarrow The fastest known way to take a matrix inverse is $\sim O(n^{2.375})$.
- On the other hand, we don't need to know the full inverse of (X'X). We only need to know the solution to the linear system $(X'X)\beta = X'y$.
 - \rightarrow This can be solved in just one set of $O(n^2)$ operations.



- In this course, we will use the statsmodels package to estimate linear regression models.
 - \rightarrow It may not be the fastest, but it gets us a whole lot more information.

```
import statsmodels.api as sm
    import statsmodels.formula.api as smf
    sm_results = sm.OLS(y, X).fit()
    print(sm results.summary())
                                 OLS Regression Results
Dep. Variable:
                                         R-squared (uncentered):
                                                                                    0.001
Model:
                                        Adj. R-squared (uncentered):
                                  OLS
                                                                                    0.000
Method:
                        Least Squares F-statistic:
                                                                                    1.198
                     Sun, 29 Oct 2023 Prob (F-statistic):
                                                                                    0.286
Date:
Time:
                             23:27:59 Log-Likelihood:
                                                                                  -14130.
No. Observations:
                                10000
                                        AIC:
                                                                                2.828e+04
Df Residuals:
                                 9990
                                         BIC:
                                                                                2.835e+04
Df Model:
                                   10
Covariance Type:
                            nonrobust
                                                  P>|t|
                 coef
                         std err
                                           t
                                                             [0.025
                                                                          0.9751
```





• When estimating "treatment effects" with linear regression, we run into exactly the same problem as when we estimate using the difference in means.

$$\rightarrow Y_i = Y_{0i}(1-T_i) + Y_{1i}T_i$$

- → If we don't know both potential outcomes, we can't estimate the treatment effect.
- However, we can estimate the treatment effect if we make some assumptions about the data generating process.
 - → The best case scenario is that we have an RCT.
 - → Otherwise, we have to make sure that we close off all other channels through which the treatment could affect the outcome.



- For starters, let's assume that we have an RCT.
- We'll compute the treatment effect both using the difference in means we used before, and also using regression.

```
df = pd.read_csv("data/online_classroom.csv")
df_no_blended = df[df["format_blended"] == 0]

# Difference in means
te = df_no_blended.groupby("format_ol")["falsexam"].mean().diff()
print(f"Estimated treatment effect: {te[1]:.3f}\n")

# Regression
sm_results = smf.ols("falsexam ~ format_ol", data=df_no_blended).fit()
print(sm_results.summary().tables[1])
```

Estimated treatment effect: -4.912

| ======== | coef | std err | t | P> t | [0.025 | 0.975] |
|--------------------------------|---------|---------|--------|-------|--------|--------|
| <pre>Intercept format_ol</pre> | 78.5475 | 1.113 | 70.563 | 0.000 | 76.353 | 80.742 |
| | -4.9122 | 1.680 | -2.925 | 0.004 | -8.223 | -1.601 |



 Because we just have a binary treatment, the difference in means and the regression coefficient are the same.

$$egin{array}{l}
ightarrow E[Y_i \mid T_i] = eta_0 + eta_1 T_i + \epsilon_i \end{array}$$

 \longrightarrow

$$ar{ au} = E[Y_i \mid T_i = 1] - E[Y_i \mid T_i = 0] = (eta_0 + eta_1 * 1) - (eta_0 + eta_1 * 0) = eta_1$$

- With only one (binary) variable, it may not be clear why we (usually) prefer regression over the difference in means.
 - → However, regression makes it much easier to control for other variables.



- For example, suppose we want to control for the student's gender.
 - → In this example, gender is not a confounder because we have an RCT.
 - → However, we can still control for it in the regression, and it should improve the precision of our estimate.
- To "control for" a covariate using the difference in means, we have to create subsamples of the data that have no variation in that variable
 - → In other words, we look at *conditional* differences in means.
- With regression, controlling for a variable is as simple as including it on the right hand side.





- What if we don't think that there is a linear relationship between the treatment and the outcome?
- We can still use linear regression, but we need to be careful about how we interpret the coefficients.
- For example, if we think that the treatment effect is nonlinear, we can include a
 quadratic term for the treatment.

$$\rightarrow Y_i = eta_0 + eta_1 T_i + eta_2 T_i^2 + \epsilon_i$$

→ This works simply by creating a new column in the data that is the square of the treatment variable.



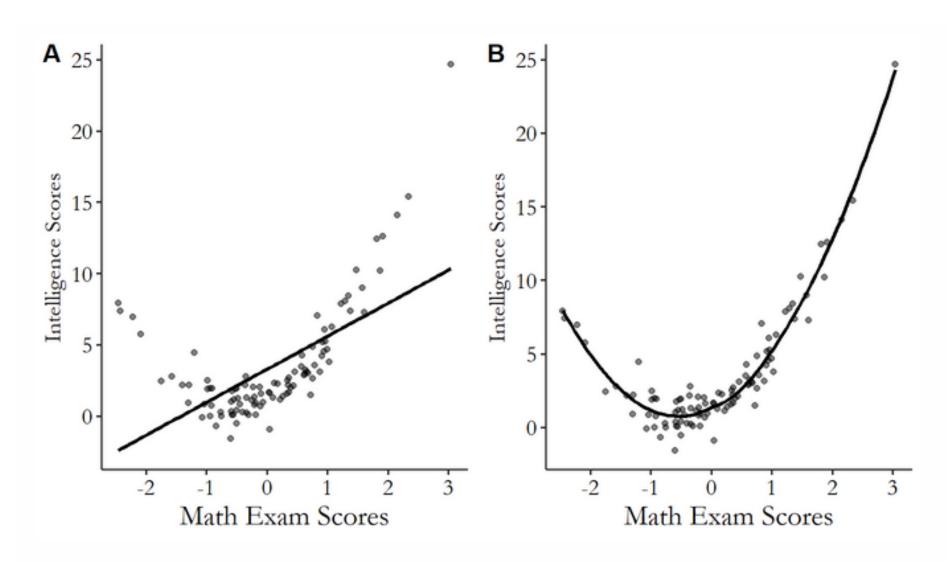
- We can also include other nonlinear transformations of the treatment variable.
 - $\rightarrow Y_i = eta_0 + eta_1 T_i + eta_2 \log(T_i) + \epsilon_i$

$$o Y_i = eta_0 + eta_1 T_i + eta_2 T_i^2 + eta_3 \log(T_i) + \epsilon_i$$

$$o Y_i = eta_0 + eta_1 T_i + eta_2 T_i^2 + eta_3 \log(T_i) + eta_4 \log(T_i)^2 + \epsilon_i$$

- But we should be careful doing stuff like this. It's easy to overfit the data.
 - ightharpoonup In fact, if we have k data points and specify a k-1-th order polynomial, we can always fit the data perfectly.







- We could also transform the dependent variable, rather than the treatment variable.
- For example, let's imagine we wanted to estimate a Cobb-Douglas function.
 - $o Y = z K^{lpha} L^{eta}$ where z is productivity, K is capital, and L is labor.
 - \rightarrow This equation is nonlinear in the parameters α and β .
 - → However, we can transform it to be linear using the log.
 - $\rightarrow \ln Y = \ln z + \alpha \ln K + \beta \ln L$
- This trick works whenever the dependent variable is an invertible function of an equation that is linear in parameters.

$$\rightarrow Y = f(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_k X_k)$$

$$ightarrow f^{-1}(Y)=eta_0+eta_1X_1+eta_2X_2+\cdots+eta_kX_k$$



- Sometimes the dependent variable is not continuous.
 - → For example, it might be binary (did the student pass the exam?)
 - → Or it might be a count (how many times did the student log in canvas?)
- In these cases, we shouldn't just apply OLS blindly
 - → Often we can work around this by transforming the dependent variable.
 - → For example, we can use a logistic regression to estimate the probability of passing the exam.
- Let $F(x) = rac{e^x}{1+e^x}$ be the logistic function and

$$\to P(Y_i = 1) = F(\beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \dots + \beta_k X_{ki})$$

 $o Y_i \sim \operatorname{Bernoulli}(P(Y_i = 1))$



```
1 from causaldata import restaurant_inspections
2 df = restaurant_inspections.load_pandas().data
3
4 df.head()
```

| | business_name | inspection_score | Year | NumberofLocations | Weekend |
|---|-----------------------------|------------------|------|-------------------|---------|
| 0 | MCGINLEYS PUB | 94 | 2017 | 9 | False |
| 1 | VILLAGE INN #1 | 86 | 2015 | 66 | False |
| 2 | RONNIE SUSHI 2 | 80 | 2016 | 79 | False |
| 3 | FRED MEYER - RETAIL FISH | 96 | 2003 | 86 | False |
| 4 | PHO GRILL | 83 | 2017 | 53 | False 2 |



```
# Statsmodels wants the dependent variable to be numeric

df["Weekend"] = 1*df["Weekend"]

# How is the frequency of weekend inspections changing over time?

m1 = smf.logit(formula = "Weekend ~ Year", data = df).fit()

print(m1.summary().tables[1])
```

Optimization terminated successfully.

Current function value: 0.045192

Iterations 10

| | coef | std err | Z | P> z | [0.025 | 0.975] |
|-----------|---------|---------|--------|-------|--------|--------|
| Intercept | 44.2360 | 23.502 | 1.882 | 0.060 | -1.828 | 90.300 |
| Year | -0.0244 | 0.012 | -2.088 | 0.037 | -0.047 | -0.002 |



- The coefficient on Year is negative, which indicates that the frequency of weekend inspections is decreasing over time.
- Notice that we have a new message: "Optimization terminated successfully"
 - → Since this GLM is nonlinear, we can't use the formula for OLS



 In order to actually interpret this coefficient, we prefer to know the marginal effect of Year rather than the coefficient value itself.

```
# Compute the marginal effect of Year
  2 print(m1.get_margeff(at="mean").summary())
        Logit Marginal Effects
Dep. Variable:
                              Weekend
Method:
                                 dydx
At:
                                 mean
                dy/dx
                         std err
                                                 P>|z|
                                                             [0.025
                                                                         0.9751
                                   -2.110
              -0.0002
                        8.79e-05
                                                 0.035
```





- What if the treatment effect itself is heterogeneous?
 - → For example, maybe the treatment effect is larger for students who are more engaged in the class.
- When comparing means, we would have to create subsamples of the data and estimate the treatment effect separately for each subsample.
- In a regression equation, we can incorporate heterogeneous treatment effects using **interaction terms**.

$$\rightarrow Y_i = \beta_0 + \beta_1 T_i + \beta_2 X_i + \beta_3 T_i X_i + \epsilon_i$$

$$egin{aligned} igsplus E[Y_i \mid T_i, X_i] &= eta_0 + eta_1 T_i + eta_2 X_i + eta_3 T_i X_i \end{aligned}$$



```
# Use * to include two variables independently
# plus their interaction
# (: is interaction-only, we rarely use it)
# m1 = smf.ols(formula = "inspection_score ~ NumberofLocations*Weekend + Year", data = df).fit()

print(m1.summary().tables[1])
# m1.t_test("NumberofLocations + NumberofLocations:Weekend = 0")
```

| | coef | std err | t | P> t | [0.025 | 0.975] |
|---------------------------|----------|---------|---------|-------|---------|---------|
| Intercept | 225.1260 | 12.415 | 18.134 | 0.000 | 200.793 | 249.460 |
| NumberofLocations | -0.0191 | 0.000 | -43.759 | 0.000 | -0.020 | -0.018 |
| Weekend | 1.7592 | 0.488 | 3.606 | 0.000 | 0.803 | 2.715 |
| NumberofLocations:Weekend | -0.0098 | 0.008 | -1.307 | 0.191 | -0.025 | 0.005 |
| Year | -0.0648 | 0.006 | -10.494 | 0.000 | -0.077 | -0.053 |

<class 'statsmodels.stats.contrast.ContrastResults'>

Test for Constraints

| | coef | std err | t | P> t | [0.025 | 0.975] |
|----|---------|---------|--------|-------|--------|--------|
| c0 | -0.0289 | 0.008 | -3.851 | 0.000 | -0.044 | -0.014 |



- In this example, it looks like the effect of the number of locations on the inspection score is strengthened on the weekend.
 - → (Assuming we have some causal model to justify that this is unbiased)
- Interactions are easiest to interpret when they are categorical; each group has their own treatment effect.
- However, they also work with continuous variables.
 - → In this case, the treatment effect is a function of the other variable.
 - → We could say that a certain variable "moderates" the treatment effect.



- Why not just make everything interact with everything else?
 - → We could, but it would be hard to interpret the coefficients.
 - → It would also be easy to overfit the data.
 - → In order to identify the interaction terms, we want lots of variation across both variables.
- Interactions are most useful when we have a *theory* about how the treatment effect varies with other variables.
 - → For example, we might think that the treatment effect of online class is larger for students who are more engaged in the class, or vice versa
- Generally, it's good to include interactions if they are a primary focus of the analysis.



Translating Causal Diagrams to Regression



Translating Causal Diagrams to Regression

