

STA380 Practice Problems for Quiz 7

These problems are not to be handed in, but they are for extra practice for students to be prepared for the quiz.

1. Use `optimize()` to maximise the $\text{Gamma}(\alpha = 2, \beta = 3)$ pdf with respect to x .
2. Consider the Anna-11 distribution with probability density function:

$$f_A(a) = \frac{1}{(1 - e^{-\rho})^{r-1}} \sum_{i=0}^{r-1} \binom{r-1}{i} (-1)^i e^{-\rho i} \frac{1}{\Gamma(r-1)} (a - \rho i)^{(r-1)-1} e^{-(a-\rho i)} \mathbf{1}_{\{a > \rho i\}}.$$

Use `optimize()` to maximise the above pdf with respect to a between the interval $[0, 10]$.

3. Consider the [Laplace distribution](#) with probability density function:

$$f(x | \mu, b) = \frac{1}{2b} \exp \left\{ -\frac{|x - \mu|}{b} \right\}, \quad x \in \mathbb{R}, \mu \in \mathbb{R}, b > 0.$$

Here, μ is called the location parameter and b is called the diversity or scale parameter.

- (a) Derive the log likelihood function.
- (b) With $n = 10^4$ replicates, use `VGAM::rlaplace(n, location = 4, scale = 2)` to generate samples from the Pareto distribution. You can read the documentation yourself (`??VGAM::rlaplace`) if you want to understand what the arguments `location` and `scale` are.
- (c) Use `optim()` to maximize the log likelihood. Use all 5 methods. Do any of them provide a good result?
- (d) The closed-form MLEs are as follows:

$$\hat{b}_{MLE} = \frac{\sum_{i=1}^n |X_i - \mu|}{n}, \quad \hat{\mu}_{MLE} = \text{median}(X_1, \dots, X_n)$$

You don't need to derive these. Compare the answers of the closed-form MLE to the results from `optim()`. What do you conclude?

4. Derive the log likelihood function of the $\text{Beta}(\alpha, \beta)$ distribution. With $n = 10^4$ replicates, use `rbeta(n, shape1 = 3, shape2 = 4)` to generate samples from the Beta distribution. Then, use the Newton-Raphson Method to approximate the MLE of the generated sample.
5. Consider $X_i \stackrel{i.i.d}{\sim} \text{Exponential}(\theta)$ where $i = 1, 2, \dots, n$ where $n = 10^4$. Suppose we had all of the realisations except for one of them, let's say X_1 . Construct the EM algorithm and code it in R. Use `rexp(n-1, rate = 1/3)` to generate an incomplete sample.

Below are the solutions.

```
1. op <- optimize(dgamma, interval = c(0, 100),
                 shape = 2, scale = 3, maximum = TRUE)
op$maximum
# graph; optional but nice to see
curve(dgamma(x, shape =2, scale = 3), 0, 10)
abline(h = dgamma(op$maximum, shape = 2, scale = 3),
       col = "red", lty = 2, lwd = 2)
```

You could also brute force and directly code up the pdf instead of using `dgamma()`.

```
2. anna_pdf <- Vectorize(function(a, rho, r){
  if(a > ((r-1)*rho)){return(0)}
  total <- 0
  denom <- (1 - exp(-rho))^(r-1)
  for(i in 0:(r-1)){
    part1 <- choose(r-1, i) * exp(-i*rho) * (-1)^(i)
    q <- a - i*rho
    part2 <- if((q > 0) && (a < (r-1)*rho)){
      dgamma(q, shape = r - 1)
    } else {0}
    total <- total + part1*part2
  }
  return(total/denom)
}, vectorize.args = "a")

op <- optimize(anna_pdf, interval = c(0, 10),
               rho = 2.7, r = 5, maximum = TRUE)
op$maximum

curve(anna_pdf(x, rho = 2.7, r = 5), 0, 10)
abline(h = anna_pdf(op$maximum, rho = 2.7, r = 5),
       col = "red", lty = 2, lwd = 2)
```

3. (a) Let $\theta = (\mu, b)$. The likelihood function is:

$$L(\theta|\mathbf{x}) = (2b)^{-n} \exp \left\{ \sum_{i=1}^n \frac{-|x_i - \mu|}{b} \right\}$$

The log-likelihood function is:

$$l(\theta|\mathbf{x}) = -n \log(2b) - \frac{1}{b} \sum_{i=1}^n |x_i - \mu|$$

```
(b) n<-10^4
set.seed(1)
samp <- VGAM::rlaplace(n, location = 4, scale = 2)
```

```
(c) laplace_likelihood = function(theta, sample){
  mu <- theta[1]
  b <- theta[2]
  # need to add constraint for the optimizer...
  if (b < 0) {return(Inf)}
  n <- length(sample)
  # By default, optim() gives the MINIMUM. need to negate for the max!
  return(-(n*log(2*b) - (1/b)*sum(abs(sample-mu))))
}

start_val <- c(1, 1)

op_1 <- optim(start_val, sample = samp, laplace_likelihood, method = "Nelder-Mead")
```

```

op_2 <- optim(start_val, sample = samp, laplace_likelihood, method = "BFGS")
op_3 <- optim(start_val, sample = samp, laplace_likelihood, method = "CG")
op_4 <- optim(start_val, sample = samp, laplace_likelihood, method = "L-BFGS-B",
              lower = c(1e-6, 1e-6))
op_5 <- optim(start_val, sample = samp, laplace_likelihood, method = "SANN")

```

None of them actually provides a good answer. How disappointing. It seems like... you still need to learn how to take a derivative after all.

```

(d) > mu <- median(samp)
    > mu
    [1] 3.982892
    > sum(abs(samp - mu))/n
    [1] 2.030977

```

Conclusion: there are some expressions where the numerical approximation methods will fail. This likely happens when the function is not smooth.

(e) Let $\theta = (\alpha, \beta)$. The likelihood function is:

$$L(\theta|\mathbf{x}) = \left\{ \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \right\}^n \prod_{i=1}^n x_i^{(\alpha-1)} (1 - x_i)^{(\beta-1)}$$

The log-likelihood function is:

$$\begin{aligned} l(\theta|\mathbf{x}) &= n \log \left\{ \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \right\} + (\alpha - 1) \sum_{i=1}^n \ln(X_i) + (\beta - 1) \sum_{i=1}^n \ln(1 - X_i) \\ &= n \{ \log(\Gamma(\alpha + \beta)) - \log(\Gamma(\alpha)) - \log(\Gamma(\beta)) \} + (\alpha - 1) \sum_{i=1}^n \ln(X_i) + (\beta - 1) \sum_{i=1}^n \ln(1 - X_i) \end{aligned}$$

The code is:

```

n<-10^4
samp <- rbeta(n, shape1 = 3, shape2 = 4)

beta_likelihood = function(theta, sample){
  alpha <- theta[1]
  beta <- theta[2]
  # need to add constraint for the optimizer...
  if (alpha <= 0 || beta <= 0) {
    # we can properly maximize now
    return(-Inf)
  }
  gams <- lgamma(alpha + beta) - lgamma(alpha) - lgamma(beta)
  n * gams + (alpha-1)*sum(log(sample)) + (beta-1) * sum(log(1-sample))
}

newton_raphson_beta <- function(theta0, sample, tol = 0.01){
  theta <- theta0
  no_root <- TRUE
  while(no_root){
    L <- numDeriv::grad(beta_likelihood, theta, sample = sample)
    H <- numDeriv::hessian(beta_likelihood, theta, sample = sample)

    new_theta <- theta - solve(H, L)
    if(sqrt(sum((new_theta - theta)^(2)))) <= tol){
      no_root = FALSE
    }
    theta <- new_theta
  }
  return(theta)
}
newton_raphson_beta(theta0 = c(2, 1), sample = samp)

```

4. Below is the write-up for the EM algorithm. First, we want to find our initial starting point $\hat{\theta}^{(0)}$ by computing the incomplete log likelihood.

$$\begin{aligned}
L(\theta|\mathbf{x}_{-1}) &= \frac{1}{\theta^{(n-1)}} e^{-\frac{1}{\theta} \sum_{i=2}^n x_i} \\
l(\theta|\mathbf{x}_{-1}) &= -(n-1) \ln(\theta) - \frac{1}{\theta} \sum_{i=2}^n x_i \\
\frac{dl(\theta|\mathbf{x}_{-1})}{d\theta} &= -\frac{n-1}{\theta} + \frac{1}{\theta^2} \sum_{i=2}^n x_i \\
0 &\stackrel{set}{=} -\frac{n-1}{\theta} + \frac{1}{\theta^2} \sum_{i=2}^n x_i \\
\hat{\theta}^{(0)} &= \frac{\sum_{i=2}^n x_i}{n-1}
\end{aligned}$$

Below is for the complete log likelihood:

$$\begin{aligned}
L(\theta|\mathbf{x}) &= \frac{1}{\theta^n} e^{-\frac{1}{\theta} \sum_{i=1}^n x_i} \\
l(\theta|\mathbf{x}) &= -n \ln(\theta) - \frac{1}{\theta} \sum_{i=1}^n x_i \\
&= -n \ln(\theta) - \frac{x_1}{\theta} - \frac{1}{\theta} \sum_{i=2}^n x_i
\end{aligned}$$

Now let's compute the expected value (expectation step):

$$\begin{aligned}
\mathbb{E}[l(\theta|\mathbf{x}) | \hat{\theta}^{(k)}, \mathbf{x}_{-1}] &= \mathbb{E} \left[-n \ln(\theta) - \frac{x_1}{\theta} - \frac{1}{\theta} \sum_{i=2}^n x_i \mid \hat{\theta}^{(k)}, \mathbf{x}_{-1} \right] \\
&= -n \ln(\theta) - \frac{\mathbb{E} \left[x_1 \mid \hat{\theta}^{(k)}, \mathbf{x}_{-1} \right]}{\theta} - \frac{1}{\theta} \sum_{i=2}^n x_i \\
&= -n \ln(\theta) - \frac{\hat{\theta}^{(k)}}{\theta} - \frac{1}{\theta} \sum_{i=2}^n x_i
\end{aligned}$$

Thus we claim that:

$$Q(\theta, \hat{\theta}^{(k)}) = -n \ln(\theta) - \frac{\hat{\theta}^{(k)}}{\theta} - \frac{1}{\theta} \sum_{i=2}^n x_i$$

Maximization step:

$$\begin{aligned}
\frac{dQ(\theta, \hat{\theta}^{(k)})}{d\theta} &= \frac{-n}{\theta} - \frac{\hat{\theta}^{(k)}}{\theta^2} + \sum_{i=2}^n \frac{x_i}{\theta^2} \\
0 &\stackrel{set}{=} \frac{-n}{\theta} - \frac{\hat{\theta}^{(k)}}{\theta^2} + \sum_{i=2}^n \frac{x_i}{\theta^2} \\
\hat{\theta}^{(k+1)} &= \frac{\sum_{i=2}^n x_i + \hat{\theta}^{(k)}}{n}
\end{aligned}$$

```

n <- 10^4
incomp_samp <- rexp(n-1, rate = 1/3)

# Assume the sample we put in here is incomplete.
em_exp <- function(sample, tol = 0.0001){
  # Step 1: get the starting point
  theta <- sum(sample) / length(sample)
  no_sol <- TRUE
  while(no_sol){
    new_theta <- (sum(sample) - theta) / (length(sample)+1)
    if(sqrt(sum((new_theta - theta)^2))) <= tol){
      no_sol = FALSE
    }
  }
}

```

```
    }  
    theta <- new_theta  
  }  
  return(theta)  
}  
  
em_exp(incomp_samp)
```