## Unit 1: Generating Random Variables

Chapter 3 in "Statistical Computing with R"

## Anna Ly

Department of Mathematical and Computational Sciences
University of Toronto Mississauga

January 5, 2026

## Overview

- 1. The Inverse Transform Method
- 2. The Acceptance-Rejection Method
- 3. Transformation and Convolution Methods
- 4. Mixtures Methods

Strictly speaking, it is impossible to get random numbers from a computer; but programs can produce pseudo-random numbers.

In this course we assume that a suitable uniform pseudo-random number generator is available. (The methods for creating pseudo-randomness is not the focus of this course, and shall be omitted.)

Refer to 'help(.Random.seed)' for details about the default random number generator in R.

## Theorem 3.1 (Probability Integral Transform)

If X is a continuous random variable with cdf  $F_X(x)$  then  $U = F_X(x) \sim \textit{Uniform}(0,1)$ .

Proof.

## Corollary

Let  $U \sim Uniform(0,1)$ . Define  $X = F^{-1}(U)$ , where F is a cdf. Then, F is the cdf of X.

Proof.

The inverse transform method (continuous case) can be summarized as follows:

- 1. Find the cdf,  $F_X(x)$ .
- 2. Define the inverse function  $F_X^{-1}(u)$ .
- 3. For each random variate required:
  - a Generate a random u from Uniform(0,1).
  - b Deliver  $x = F_X^{-1}(u)$ .

**Warning:** we should only use this method if a closed form of  $F_X^{-1}(x)$  exists!

## Example

Use the inverse transform method to simulate a random sample from the distribution with density:

$$f_X(x) = 3x^2, \quad 0 < x < 1.$$

Additionally, write the R code.

## Example

Use the inverse transform method to simulate a random sample from the exponential distribution, which has the density:

$$f_X(x) = \frac{1}{\theta}e^{-\frac{1}{\theta}x}, \quad x > 0.$$

Additionally, write the R code.

## Example

Use the inverse transform method to simulate a random sample from the Weibull distribution, which has the density:

$$f_X(x) = egin{cases} rac{lpha}{eta} \left(rac{x}{eta}
ight)^{lpha-1} \mathrm{e}^{-\left(rac{x}{eta}
ight)^{lpha}} & x \geq 0, lpha > 0, eta > 0, \ 0 & ext{otherwise}. \end{cases}$$

To make life easier, we'll consider the case where  $\beta = 1$ .

- 1. Additionally, write the R code.
- 2. Find  $\mathbb{E}[X]$ , the theoretical mean, and check that the simulated mean is close to the theoretical mean.

If we just want to generate a random variable X with pmf:

$$\mathbb{P}(X=x_i)=p_i,\ i\in\mathbb{N},\ \sum_i p_i=1,$$

then inverse transform method (discrete case) can be summarized as follows:

- 1. Generate a random u from Uniform(0, 1).
- 2. Transform u into X as follows:

$$X = x_j$$
 if  $F_X(x_{j-1}) < u \le F_X(x_j)$ 

3. It follows that,

$$X = \begin{cases} x_1 & u \leq F_X(x_1), \\ x_2 & F_X(x_1) < u \leq F_X(x_2), \\ \dots & \\ x_j & F_X(x_{j-1}) < u \leq F_X(x_j), \\ \dots & \end{cases}$$

In other words, discrete random variables can be generated by slicing up the interval (0,1) into subintervals which define a partition on (0,1):

$$(0, F_X(x_1)], (F_X(x_1), F_X(x_2)], (F_X(x_2), F_X(x_3)], \dots, (F_X(x_{k-1}), 1].$$

We can also define:

$$p_1 = F_X(x_1), \quad p_2 = F_X(x_2) - p_1, \quad \dots, \quad p_j = F_X(x_j) - \sum_{i=1}^{j-1} p_i$$

## Proof of Previous Algorithm

Given X that was defined in the algorithm of the previous slide, Prove that  $\mathbb{P}(X = x_i) = p_i$ .

#### Example

Use the inverse transform method to simulate a random sample from the Bernoulli distribution with p=0.4. Additionally, write the R code.

#### Example

Let X be a discrete random variable with the following pmf:

Find  $F_X(x)$ . Additionally, write the R code.

#### Generalized Inverse Function

For a discrete random variable X with cdf  $F_X(x)$ , the inverse CDF  $F_X^{-1}(p)$ , also called the quantile function, is defined as:

$$F_X^{-1}(x) := \inf\{x : F_X(x) \ge p\}, \quad 0$$

Note that there isn't a unique x where F(x) = p for every  $p \in (0,1)$ .

## Example

From the previous example,

- Find  $F_X^{-1}(0.1)$ .
- Find  $F_X^{-1}(0.5)$ .
- Find  $F_X^{-1}(0.85)$ .

## Example

Use the inverse transform method to simulate a random sample from the Geometric distribution with pmf:

$$\mathbb{P}(X=i)=pq^i,\ i\in\mathbb{N},\ q=1-p.$$

Additionally, write the R code.

Remember that now we are slicing up the interval (0,1) into subintervals which define a partition on (0,1):

$$(0, F_X(x_1)], (F_X(x_1), F_X(x_2)], (F_X(x_2), F_X(x_3)], \dots, (F_X(x_{k-1}), 1].$$

Which basically means we are trying to evaluate  $F_X(x_i)$  for  $i \in \mathbb{N}$ .

- However, sometimes it's hard to obtain a closed form for  $F_X(x_i)$ .
- Thus, it is more useful to find a recursive formula.

## Example

Use the inverse transform method to simulate a random sample from the Binomial distribution with pmf:

$$p_i := \mathbb{P}(X=i) = \binom{n}{i} p^i (1-p)^{n-i}, \quad i=0,1,\ldots,n.$$

1. First, derive the following recursive formula:

$$p_{i+1}=\left(\frac{n-i}{i+1}\right)\left(\frac{p}{1-p}\right)p_i, \ i=0,1,\ldots,n.$$

2. Write the R code to simulate the algorithm.

## Example

Use the inverse transform method to simulate a random sample from the logarithmic distribution with pmf:

$$p_i := \mathbb{P}(X = i) = \frac{a\theta^i}{i}, \ i \in \mathbb{N}.$$

Where  $0 < \theta < 1$  and  $a = (-\ln(1 - \theta))^{-1}$ .

1. First, derive the following recursive formula:

$$p_{i+1} = \left(\frac{\theta i}{i+1}\right) p_i, \ \ i \in \mathbb{N}.$$

2. Write the R code to simulate the algorithm.

In general:

$$X \Rightarrow F_X^{-1}(x) \Rightarrow F_X^{-1}(u)$$

However, this only works if we can define the inverse. We can think of many different functions where the inverse would be hard to find: the Gaussian distribution, beta, etc...

# The Acceptance-Rejection Method

Suppose we want to generate the random variable X with target density f using the acceptance-rejection method.

1. Find another random variable, Y with trial/candidate/envelope density g where there exists  $c \in \mathbb{R}$  such that:

$$\frac{f(t)}{g(t)} < c.$$

- 2. For each random variate required:
  - 2.1 Generate y from the distribution with density g.
  - 2.2 Generate u from the Uniform(0,1) distribution.
  - 2.3 If

$$u<\frac{f(y)}{cg(y)}$$

accept y and deliver x = y. Otherwise, reject y and generate a random variate again.

## Probability of Accepting

Given the acceptance-rejection algorithm, evaluate the probability of acceptance for any iteration (see part 2.3 in the previous slide).

Proof.

## Probability of Accepting

Given the acceptance-rejection algorithm, prove that the accepted sample has the same distribution as X.

Proof.

## The Choice of c

Given the acceptance-rejection algorithm, show that that  $c \geq 1$ .

- Ideally, you want g to be easy to sample from, consistent with the support of f.
- Theoretically, as long as you know that f indeed has longer tails than g, you can choose c to be ridiculously large and this will still yield a valid algorithm.
- Since the acceptance rate is equal to 1/c, you will have to, on average, generate  $c \times n$  draws from the trial distribution and from the uniform distribution just to get n draws from the target distribution.
- So choosing c to be too large will yield an inefficient algorithm.
- Ideally: choose c close to 1; so f and g are similar.

#### The Distribution of N

Let N represent the number of iterations that the acceptance-rejection algorithm needs to successfully generate one value of X. What is the distribution of N?

## Example

Use the acceptance-rejection method to simulate a random sample from the distribution with density:

$$f_X(x) = 3x^2, \quad 0 < x < 1.$$

Let the trial density be Uniform(0,1). Write the R code to simulate the algorithm. Compare the number of iterations to the value of  $\frac{1}{c}$ .

## Example

Use the acceptance-rejection method to simulate a random sample from the beta distribution with density:

$$f_X(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha - 1} (1 - x)^{\beta - 1}, \quad 0 < x < 1.$$

Assume  $\alpha=2,\beta=4$ . Let the trial density be Uniform(0,1). Write the R code to simulate the algorithm. Compare the number of iterations to the value of  $\frac{1}{c}$ .

## Example

Use the acceptance-rejection method to simulate a random sample from the distribution with density:

$$f_X(x) = egin{cases} rac{2}{\pi}\sqrt{1-x^2} & |x| < 1, \ 0 & ext{otherwise.} \end{cases}$$

Let the trial density be Uniform(-1,1). Write the R code to simulate the algorithm. Compare the number of iterations to the value of  $\frac{1}{6}$ .

You've done transformations or convolution methods before (or at least I hope!)

- If  $X_i \sim Exponential(a)$  with  $i \in \{1, 2, ..., n\}$  then  $\sum_i X_i \sim Gamma(n, a)$ .
- If  $U \sim Gamma(r, \lambda)$  and  $V \sim Gamma(s, \lambda)$  and  $U \perp V$  then

$$X := \frac{U}{U+V} \sim \textit{Beta}(r,s).$$

- If  $Z \sim N(0,1)$  then  $V := Z^2 \sim \chi^2(1)$ .
- If  $U \sim \chi^2(m)$  and  $V \sim \chi^2(n)$  and  $U \perp V$  then

$$\tilde{F}:=\frac{U/m}{V/n}\sim F(m,n).$$

• If  $Z \sim N(0,1)$  and  $V \sim \chi^2(n)$  and  $Z \perp V$  then:

$$\tilde{T}:=\frac{Z}{\sqrt{V/n}}\sim t(n).$$

## Example

Use rexp() to generate the  $Gamma(\alpha=10,\beta=1/2)$  distribution.

## Example

Use rexp() to generate the  $Beta(\alpha=2,\beta=3)$  distribution.

# **Mixture Methods**

## Mixture Methods

#### Finite Mixture Model

A finite mixture model is a statistical model that represents a probability distribution as a mixture of several component distributions. Mathematically, given k component distributions  $f_1(x), ..., f_k(x)$ , each with associated mixing probabilities (also known mixing weights)  $\pi_1, ..., \pi_k$ , a finite mixture model f(x) is defined as:

$$f(x) := \sum_{i=1}^k \pi_i f_i(x),$$

where  $0 \le \pi_i \le 1$  and  $\sum_{i=1}^k \pi_i = 1$ .

## Mixture Model Properties

We want to show that f(x) satisfies the properties of a density function. That is, show that f(x) is non-negative and  $\int_{-\infty}^{\infty} f(x) dx = 1$ .

## Mixture Model CDF

Compute F(x).

## Mixture Model Expected Value

Let  $\mu_i := \mathbb{E}[X|i] = \int_{-\infty}^{\infty} x f_i(x) dx$ ; this represents the mean for the *i*-th component distribution. Compute  $\mu := \mathbb{E}[X]$ . What does it represent?

#### Mixture Model Variance

Let  $\sigma_i^2 := \mathbb{V}[X|i]$  represent the standard deviation for the *i*-th component distribution. Show that:

$$\sigma^2 := \mathbb{V}[X] = \sum_{i=1}^k \pi_i \sigma_i^2 + \sum_{i=1}^k \pi_i (\mu_i - \mu)^2$$

What does it represent?

Convolutions and mixtures look similar but the represented distributions differ!

## Example

Suppose  $X_1 \sim N(0,1)$ ,  $X_2 \sim N(3,1)$ , and  $X_1 \perp X_2$ .

Convolution representation:

$$S := 0.4X_1 + 0.6X_2$$

Mixture representation:

$$F_X(x) := 0.4F_{X_1}(x_1) + 0.6F_{X_2}(x_2)$$

Note: for the convolution, the coefficients in front of the random variables do not necessarily have to add to one; we could have done  $S := aX_1 + bX_2$  for any  $a, b \in \mathbb{R}$ . However, for finite mixture models, these coefficients must add to 1.

Simulating a variable from a finite k-mixture distribution is typically carried out by the composition method: Consider  $F(x) = \sum_{i=1}^{k} \pi_i F_{X_i}(x)$ . Then,

## Composition Method

• Generate an integer  $I \in \{1, ..., k\}$  such that:

$$\mathbb{P}(I=i)=\pi_i, \quad i\in\{1,2,\ldots,k\}.$$

• Deliver X with cdf  $F_{X_l}(x)$ .

## Example

Suppose  $X_1 \sim N(0,1)$ ,  $X_2 \sim N(3,1)$ , and  $X_1 \perp X_2$ . Simulating the following using R:

$$F_X(x) := 0.4F_{X_1}(x_1) + 0.6F_{X_2}(x_2)$$

Then, compare the above to the following convolution counterpart:

$$S := 0.4X_1 + 0.6X_2$$

The first two methods we discussed, inverse-transform and acceptance-rejection method, depend on the uniform distribution. Similarly, we can do the same for generating finite mixture models. Consider  $F(x) = \sum_{i=1}^{k} \pi_i F_{X_i}(x)$ . Then:

## Modified Composition Method

- Generate u from the Uniform(0,1) distribution.
- If:

$$\sum_{i=1}^{l-1} \pi_i \le u < \sum_{i=1}^{l} \pi_i$$

then generate a random x from  $F_{X_l}(x)$  where  $l=1,\ldots,k$  with the convention  $\sum_{i=1}^{0} \pi_i = 0$ .

## Example

Using the alternative method... Suppose  $X_1 \sim N(0,1)$ ,  $X_2 \sim N(3,1)$ ,  $X_3 \sim N(5,1)$ , and assume they're all independent of each other. Code the following algorithm using R:

$$F_X(x) := 0.4F_{X_1}(x_1) + 0.3F_{X_2}(x_2) + 0.3F_{X_3}(x_3)$$