

# Week 5

2. Consider the integral:

$$\theta = \int_0^{10} x \ln(x) dx$$

- STA 380
- (a) Use the simple Monte Carlo estimator to estimate the expected value, where the importance function is the  $Uniform(0, 10)$  distribution. Call this  $\hat{\theta}$ .
- (b) Write an exact expression for  $\text{Var}[\hat{\theta}]$ . You do not need to evaluate this integral by hand. Then, compute this integral using `integrate()` in R.
- Disclaimer: you need to know how to integrate for the tests; we also just want you to learn the `integrate()` function as well. You will not be using the `integrate()` function in R for the term tests or exam.
- (c) Recall in lecture, we said the Monte Carlo estimate for the variance of the estimator for the integral is:

$$\widehat{\text{Var}}[\hat{\theta}] = \frac{\hat{\sigma}^2}{n} = \frac{\sum_{i=1}^n [g(x_i) - \bar{g}(x)]^2}{n^2}$$

Using this equation, compute  $\widehat{\text{Var}}[\hat{\theta}]$  in R. Use  $n = 10^4$ .

- (d) Compare the MC estimator against the theoretical variance of  $\hat{\theta}$ . Use `all.equal()`, which will report the mean relative difference.

(a)  $\theta = 10 \int_0^{10} \left( \frac{1}{10} \cdot x \cdot \ln(x) \right) dx$

$= 10 E_{U(0,10)} [x \cdot \ln(x)]$

samples from  $U \sim \text{Uniform}(0, 10)$ .

by SL2N, the sample MC estimator is

$$\hat{\theta} = 10 \times \frac{1}{n} \sum_{i=1}^n u_i \cdot \ln(u_i)$$

(b).

$$\begin{aligned} \text{Var}[\hat{\theta}] &= \text{Var} \left[ \frac{10}{n} \sum_{i=1}^n u_i \cdot \ln(u_i) \right] \\ &= \frac{100}{n^2} \text{Var} \left[ \sum_{i=1}^n u_i \cdot \ln(u_i) \right] \\ &= \frac{100}{n^2} \sum_{i=1}^n \text{Var}[u_i \cdot \ln(u_i)] \end{aligned}$$

$$= \frac{100}{n^2} \sum_{i=1}^n \left( E[(u_i \cdot \ln(u_i))^2] - (E[u_i \cdot \ln(u_i)])^2 \right)$$

$\text{Var}(X+Y)$

$$= \text{Var}(X) + \text{Var}(Y)$$

$$\text{Var}(Y) = E(Y^2) - (E(Y))^2$$

$$E(Y) = \int y \cdot f(y) dy$$

$$= \frac{100}{n^2} \sum_{i=1}^n \left( \int_0^{10} (u_i \cdot \ln(u_i))^2 \cdot \frac{1}{10} du_i - \left( \int_0^{10} u_i \cdot \ln(u_i) \cdot \frac{1}{10} du_i \right)^2 \right)$$

C

$$= \frac{100}{n^2} n \cdot C = \frac{100}{n} \left( \int_0^{10} (u_i \cdot \ln(u_i))^2 \cdot \frac{1}{10} du_i - \left( \int_0^{10} u_i \cdot \ln(u_i) \cdot \frac{1}{10} du_i \right)^2 \right)$$



(C)

$$\widehat{\text{Var}}[\hat{\theta}] = \frac{\hat{\sigma}^2}{n} = \frac{\sum_{i=1}^n \left[ g(x_i) - \bar{g}(x) \right]^2}{n^2}$$

$$n = 10^4$$

$$n \leftarrow 10^4$$

$$u \leftarrow \text{runif}(n, 0, 10)$$

Uniform (0, 1)  
 $u \leftarrow \text{runif}(n)$

$$\text{mc-est} \leftarrow \text{sum} \left( (10 * \log(u) * u - \text{mean}(10 * \log(u) * u))^2 \right) / n^2$$

$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^n 10 u_i \cdot \ln(u_i)$$

$f(x)$

# (b) Integrate ( )

$$\text{fun1} = \int_0^{10} (u_i \ln(u_i))^2 \cdot \frac{1}{10} du_i \quad \text{E}(Y^2)$$

$$\text{fun2} = \int_0^{10} (u_i \ln(u_i)) \cdot \frac{1}{10} du_i \quad \text{E}(Y)$$

```
fun1 <- function(x){log(x)^2*x^2/10}  
fun2 <- function(x){log(x)*x/10}
```

ex2 <- integrate(fun1, 0, 10)

ex\_sq <- integrate(fun2, 0, 10)

exact <- 100 \* (ex2\$value - (ex\_sq\$value)^2) / n

↗ .

$$\frac{100}{n} \left( \int_0^{10} (u_i \ln(u_i))^2 \cdot \frac{1}{10} du_i - \left( \int_0^{10} u_i \ln(u_i) \cdot \frac{1}{10} du_i \right)^2 \right)$$

# Test 1

Q3.

$z_1, z_2 \dots z_n$  i.i.d.  $z_i \sim N(0, 1)$  for all  $i = 1, 2 \dots n$ .

chi-squared distribution

$$\sum_{i=1}^n z_i^2 \sim \chi_n^2$$

Use `rnorm()` ...  $n = 10^4$  samples from the  $\chi_{(30)}^2$ .

R codes

$$n \leftarrow 10^4$$

# vector. values  $\leftarrow$  numeric (n)

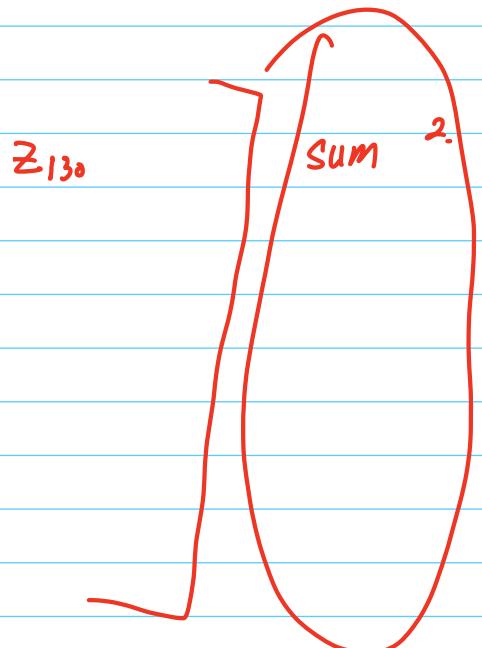
# loop for (i in 1:n) {

values[i]  $\leftarrow$  sum ( rnorm (30, mean=0, sd=1) ^ 2 )

}

SumRow

1  $z_1 z_2 \dots \dots$   
2  
3  
:  
 $10^4$



Q4.  $f_i(x)$  density  $i = 1, 2 \dots k$ .

$f_1(x), \dots, f_k(x)$ .

$\pi_1, \dots, \pi_k$  weights.

$$f(x) := \sum_{i=1}^k \pi_i f_i(x)$$

where  $0 \leq \pi_i \leq 1$  &  $\sum_{i=1}^k \pi_i = 1$

(a)  $f(x)$  non-negative

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

(1) Non-negative.

Since  $f_i(x)$  is valid density  $\Rightarrow 0 \leq f_i(x)$

And  $0 \leq \pi_i \leq 1$

$$\Rightarrow 0 \leq \sum_{i=1}^k \pi_i f_i(x)$$

$$\begin{aligned} \int(A+B) &= \int A + \int B \end{aligned}$$

$$\textcircled{2} \quad \int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^{\infty} \sum_{i=1}^k \pi_i f_i(x) dx$$

$$= \sum_{i=1}^n \int_{-\infty}^{\infty} \pi_i \cdot f_i(x) dx.$$

$$= \sum_{i=1}^n \pi_i \underbrace{\int_{-\infty}^{\infty} f_i(x) dx}_{1} = \sum_{i=1}^n \pi_i = 1$$

(b) suppose  $X_i$  has  $f_i$ .

$$\mu_i = E[X_i] = \int_{-\infty}^{\infty} x \cdot f_i(x) dx.$$

$$\sigma_i^2 := V[X_i]$$

$$WTS \Rightarrow \sigma^2: V[X] = \sum_{i=1}^k \pi_i \sigma_i^2 + \sum_{i=1}^k \pi_i (\mu_i - \mu)^2$$

$$V[X] = E[(X - \mu)^2]$$

$$= \int_{-\infty}^{\infty} (x - \mu)^2 \cdot \sum_{i=1}^k \pi_i \cdot f_i(x) dx$$

$$= \int_{-\infty}^{\infty} (x^2 - 2x\mu + \mu^2) \cdot \sum_{i=1}^k \pi_i \cdot f_i(x) dx$$

$$= \sum_{i=1}^k \pi_i \int_{-\infty}^{\infty} (x^2 - 2x\mu + \mu^2) f_i(x) dx$$

$$= \sum_{i=1}^k \pi_i \left[ \int_{-\infty}^{\infty} x^2 \cdot f_i(x) dx - \int_{-\infty}^{\infty} 2x\mu \cdot f_i(x) dx + \int_{-\infty}^{\infty} \mu^2 \cdot f_i(x) dx \right]$$

$$= \sum_{i=1}^k \pi_i \left[ E[X^2|i] - 2\mu\mu_i + \mu^2 \right]$$

$$= \sum_{i=1}^k \pi_i \left[ V[X|i] + (E[X|i])^2 - 2\mu\mu_i + \mu^2 \right]$$

$$= \sum_{i=1}^k \pi_i \left[ \sigma_i^2 + \mu_i^2 - 2\mu\mu_i + \mu^2 \right]$$

$$= \sum_{i=1}^k \pi_i (\sigma_i^2 + (\mu_i - \mu)^2)$$

$$= \sum_{i=1}^k \pi_i \sigma_i^2 + \sum_{i=1}^k \pi_i (\mu_i - \mu)$$