The Macroeconomy as a Random Forest

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Final Destination

Modeling *flexibly* macro relationships without assuming what flexible means first. Take something fundamental: a Phillips' curve.

$$u_t^{\mathrm{gap}} \to \pi_t$$

The statistical characterization of " \rightarrow " has forecasting, policy and theoretical (!) implications. Better get it right.

One way out is getting " \rightarrow " from off-the-shelf nonparametric Machine Learning (ML) techniques. But:

- Likely too flexible and wildly inefficient for the short time series we have.
- No obvious parameter(s) to look at interpretation is fuzzy.

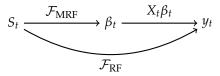
Another is assuming $\pi_t = \beta_t u_t^{\text{gap}} + \text{stuff}_t$. But:

- Rigid
- In-sample fit notoriously don't translate in out-of-sample gains.

Solution: *Generalized* **Time-Varying Parameters** via Random Forests.

(Machine) Learning β_t 's

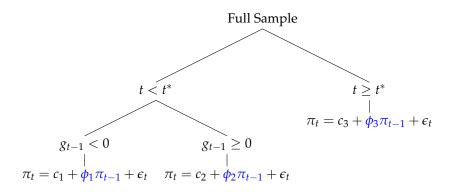
• I propose *Macroeconomic Random Forests* (MRF): fix the linear part X_t and let the coefficients β_t vary trough time according to a Random Forest.



- The core "mechanical" modification wrt plain RF is fitting an ensemble of trees which have a linear model in each leaf rather than a constant.
- MRF is nice "meeting halfway"
 - \Rightarrow Brings macro closer to ML by squashing many popular nonlinearities (structural change/breaks, thresholds, regime-switching, etc.) into an arbitrarily large S_t , handled easily by RF.
 - \Rightarrow The core output are β_t 's, Generalized Time-Varying Parameters (GTVPs).
 - \Leftarrow Brings ML closer to macro by adapting RF to the reality of economic time series. MRF \succ RF if the linear part is pervasive (like in a (V)AR).

Why Trees Make Sense (in Macro/Finance)

- Let π_t be inflation at time t.
- t^* is inflation targeting implementation date.
- Let g_t be some measure of output gap.



• The general model is

$$y_t = X_t \beta_t + \epsilon_t$$
$$\beta_t = \mathcal{F}(S_t)$$

where S_t are the state variables that determine time-variation.

- If we know the threshold variables ($S_t = [t, g_{t-1}]$) and values ($c = [t^*, 0]$): run OLS on subsamples.
- But we don't. So we need an algorithm to find out:

$$\min_{j \in \mathcal{J}^{-}, c \in \mathbb{R}} \left[\min_{\beta_{1}} \sum_{\{t \in l \mid S_{j,t} \leq c\}} (y_{t} - X_{t}\beta_{1})^{2} + \lambda \|\beta_{1}\|_{2} + \min_{\beta_{2}} \sum_{\{t \in l \mid S_{j,t} > c\}} (y_{t} - X_{t}\beta_{2})^{2} + \lambda \|\beta_{2}\|_{2} \right].$$

3 ingredients to go from a single tree to a forest

For each tree:

1. **Let the trees run deep**: even though that would surely imply overfitting for a single tree, let each tree run until leafs contain very few observations (usually < 5).

Diversifying the Portfolio (i.e., creating the ensemble)

- 2. **Bagging**: Create *B* nonparametric bootstrap samples of the data. That is, we are picking $[y_t \ X_t]$ pairs with replacement.
- 3. **De-correlated trees**: At each splitting point, we only consider a subset of all predictors ($\mathcal{J}^- \subset \mathcal{J}$) for the split.

(M)RF prediction is the simple average of all the *B* tree predictions.

Why does it not overfit? See *To Bag is to Prune*, a spin-off paper.

Useful Addition: Random Walk Regularization

- The above implements the prior $\beta_t \sim \mathcal{N}(0,.)$.
- However, $\beta_t \sim \mathcal{N}(\beta_{t-1}, .)$, i.e., time-smoothness, makes more sense.
- I implement it via WLS with rudimentary egalitarian Olympic podium weights $w(t;\zeta)$, where $\zeta < 1$ is a tuning parameter.
- The splitting rule becomes

$$\begin{split} \min_{j \in \mathcal{J}^{-}, \ c \in \mathbb{R}} \left[\min_{\beta_{1}} \sum_{t \in l_{1}^{RW}(j,c)} w(t;\zeta) \left(y_{t} - X_{t}\beta_{1} \right)^{2} + \lambda \|\beta_{1}\|_{2} \right. \\ + \min_{\beta_{2}} \sum_{t \in l_{2}^{RW}(j,c)} w(t;\zeta) \left(y_{t} - X_{t}\beta_{2} \right)^{2} + \lambda \|\beta_{2}\|_{2} \right]. \end{split}$$

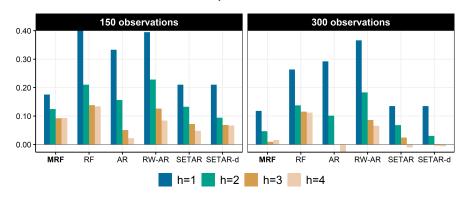
Inference

- Following (Taddy et al., 2015), interpret $\mathcal F$ as a the posterior mean of tree $\mathcal T$ which posterior distribution was obtained by (Rubin, 1981)'s Bayesian Bootstrap.
- Crucial advantage: no additional computations required, quantiles computed straight from the "bag" of trees.
- (Taddy et al., 2015)'s approach requires *iid* data: the bayesian model is multinomial with Dirichlet conjugate prior.
- I propose to rather use a Block Bayesian Bootstrap (BBB)

DGP 3: Persistent SETAR

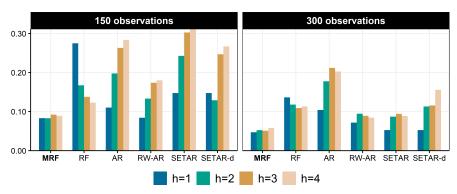
$$y_{t} = \phi_{0,t} + \phi_{1,t}y_{t-1} + \phi_{2,t}y_{t-2} + \varepsilon_{t}, \quad \varepsilon_{t} \sim N(0,0.5^{2})$$

$$\beta_{t} = [\phi_{0,t} \ \phi_{1,t} \ \phi_{2,t}] = \begin{cases} [2 \ 0.8 \ -0.2], & \text{if } y_{t-1} \geq 0\\ [0.25 \ 1.1 \ -0.4], & \text{otherwise} \end{cases}$$

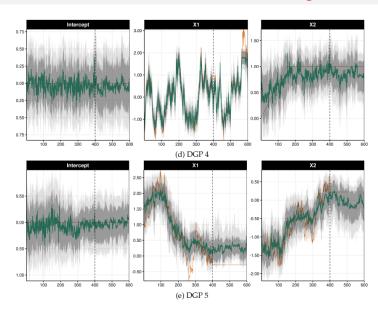


DGP 6: SETAR that morphs instantly in AR(2)

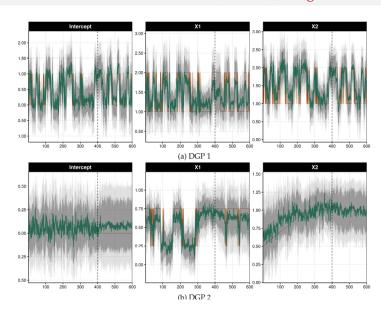
DGP 6 =
$$\begin{cases} SETAR, & \text{if } t < T/2 \\ Plain AR(2), & \text{otherwise} \end{cases}$$



A look at GTVPs under Different Contexts, when S_t is large



A look at GTVPs under Different Contexts, when S_t is large



Setup

- Data: FRED-QD, the SW data set update by (McCracken and Ng, 2016), 260 series
- POOS period starts on 2002Q1 and ends 2014Q1. Expanding window estimation from 1959Q3.
- Horizons: $h \in \{1, 2, 4, 6, 8\}$ quarters
- 6 variables of interest: GDP growth, Unemployment Rate (UNRATE) growth, Interest Rate (GS1), Inflation (Δlog (CPIAUCSL)), Housing Starts (HOUST) and some spread (T10YFFM).
- Evaluation metric is $RMSPE_{v,h,m} = \sqrt{\sum_{t \in OOS} (y_t^v \hat{y}_{t-h}^{v,h,m})^2}$

About the composition of S_t

- 1. 8 lags of y_t
- 2. *t* for structural breaks/exogenous time-variation
- 3. 2 lags of all variables in FRED
- 4. *F* to summarize the cross-section variation: 8 lags of 5 factors extracted from FRED by PCA

Most importantly

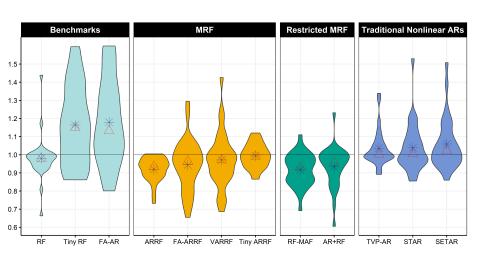
- 5. For each variable j, I generate two MAF $_{t,j}$ (Moving Average Factors) summarizing the information contained in its distributed lags.
 - Bypasses the need to penalize explicitly a nonexistent lag polynomial.
 - Done by PCA on 8 lags.
 - Further studied for many ML models (along with other transformations) in (Goulet Coulombe et al., 2020).

Forecasting Main Models

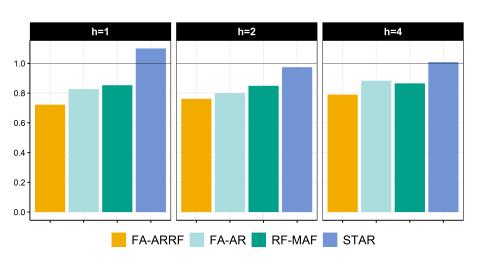
Table: Wild Horses

Acronym	Linear Part (X_t^m)	RF part
AR	$[1, y_{t-\{1:4\}}]$	Ø
FA-AR	$[1, y_{t-\{1:4\}}, F_{1,t-\{1:2\}}, F_{2,t-\{1:2\}}]$	Ø
RF	Ø	8 lags of all raw data
Tiny RF	Ø	$[y_{t-\{1:8\}}, t]$
RF-MAF	Ø	S_t
AR+RF	Filter y_t first with an AR(4)	S_t
ARRF	$[1, y_{t-\{1:2\}}]$	S_t
Tiny ARRF	$[1, y_{t-\{1:2\}}]$	$[y_{t-\{1:8\}}, t]$
FAARRF	$[1, y_{t-\{1:2\}}, F_{1,t-1}, F_{2,t-1}]$	S_t
VARRF	$[1, y_{t-\{1:2\}}, GDP_{t-1}, IR_{t-1}, INF_{t-1}]$	S_t

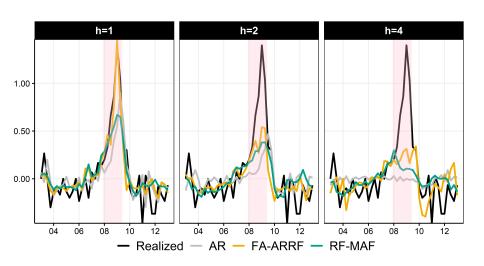
Visualizing the distribution of $RMSPE_{v,h,m}/RMSPE_{v,h,AR}$



 $RMSPE_{UR,h,m}/RMSPE_{UR,h,AR}$ in more detail



What do forecasts look like for UR? $\rightarrow R_{OOS}^2$ 80% for h = 1



GTVPs of the one-quarter ahead UR forecast

$$\Delta UR_{t+1} = \mu_t + \phi_t^1 y_t + \phi_t^2 y_{t-1} + \gamma_t^1 F_t^1 + \gamma_t^2 F_t^2 + e_{t+1}.$$

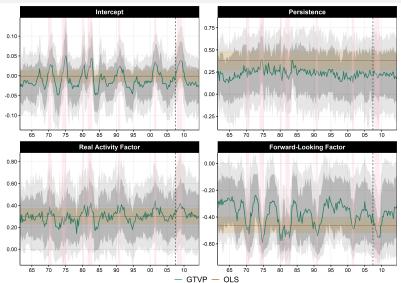


Figure: GTVPs of the one-quarter ahead UR forecast. The grey bands are the 68% and 90% credible region. The pale orange region is the OLS coefficient ± one standard error. The vertical dotted blue line is the end of the training sample. Pink shading corresponds to NBER recessions.

Dynamic β_t Learning

$$\Delta UR_{t+1} = \mu_t + \phi_t^1 y_t + \phi_t^2 y_{t-1} + \gamma_t^1 F_t^1 + \gamma_t^2 F_t^2 + e_{t+1}.$$

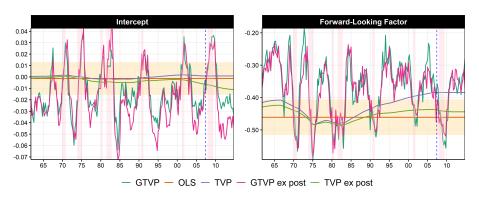
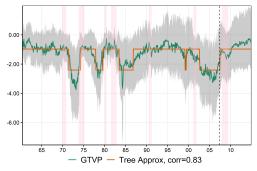


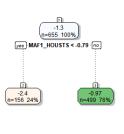
Figure: Comparing TVPs and GTVPs, ex-ante and ex-post.

An Interesting Observation for (monthly) *Inflation*

$$\pi_{t+1} = \mu_t + \phi_t^1 y_t + \phi_t^2 y_{t-1} + \gamma_t^1 F_t^1 + \gamma_t^2 F_t^2 + e_{t+1}.$$



(a) Surrogate Model Replication

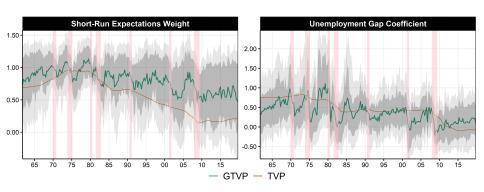


(b) Corresponding Tree

A more traditional Phillips' Curve

À la (Blanchard et al., 2015) and many others

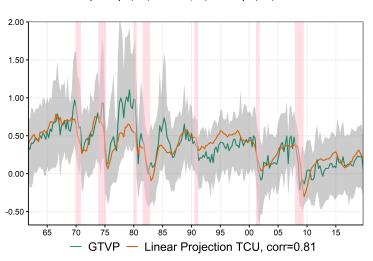
$$\pi_t = \mu_t + \beta_{1,t} \hat{\pi}_t^{SR} + \beta_{2,t} u_t^{GAP} + \beta_{3,t} \pi_t^{IMP} + \varepsilon_t$$



A more traditional Phillips' Curve

 $\beta_{2,t}$ looks like Total Capacity Utilization (TCU)

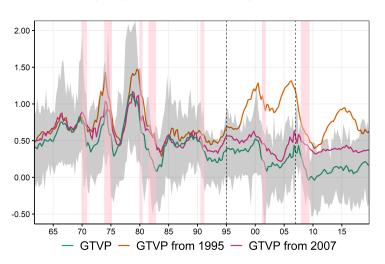
$$\pi_t = \mu_t + \beta_{1,t} \hat{\pi}_t^{SR} + \beta_{2,t} u_t^{GAP} + \beta_{3,t} \pi_t^{IMP} + \varepsilon_t$$



Dynamic Phillips' Curve Learning

Comparing "out-of-sample" predictions of GTVPs at different points in time

$$\pi_t = \mu_t + \beta_{1,t} \hat{\pi}_t^{SR} + \beta_{2,t} u_t^{GAP} + \beta_{3,t} \pi_t^{IMP} + \varepsilon_t$$



Conclusion

I proposed a new time series model that

- 1. works;
- 2. is interpretable;
- 3. is highly versatile;
- 4. off-the-shelf (R package is available);

Extensions/applications:

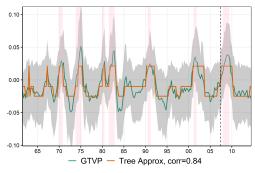
- VARs
- Conditional CAPM
- HAR volatility
- DSGEs?
- Anything goes

Try it with your favorite X_t today!

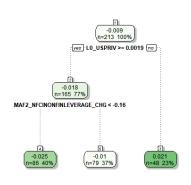
Misc

- GTVPs ➤ Random Walk TVPs since it implies an adaptive kernel rather than a fixed one
 - The intercept itself is a RF rather than a RW (e.i., a bad X_t choice can be rescued)
 - Less reliant (or not all) on $t \rightarrow$ less boundary problems (or none) when forecasting.
- 2. Compress lag polynomials $S_{t,j}^{1:P}$ ex-ante with Moving Average Factors
 - Get MAFs by running PCA on the panel $[S_{t-1,j} \dots S_{t-P,j}]$ of P lags of variable j.
 - Boost splits' meaningfulness (not wasting splits on 12 individual lags)
 - Reduce computing time

Cutting Down the Forest, One Tree at a Time $(\mu_t^{UR,h=1})$



(a) Surrogate Model Replication



(b) Corresponding Tree

Cutting Down the Forest, One Tree at a Time $(\gamma_{t,F_1}^{\mathit{INF},h=12}, \mathsf{monthly})$

