Chapter 5 Some Discrete Probability Distributions

吳明龍 副教授 成大資訊系/醫資所 minglong.wu@csie.ncku.edu.tw 辦公室:資訊系館 12 樓

5.1 Introduction

 Often, the observations generated by different statistical experiments have the same general type of behavior.

 In fact, one needs only a handful of important probability distributions to describe many of the discrete random variables encountered in practice.

5.1 Introduction

- <u>Binomial distribution</u> (Section 5.3): test the effectiveness of a new drug.
- <u>Hypergeometric distribution</u> (Section 5.4): test the number of defective items from a batch of production.
- Negative binomial distribution (Geometric distribution) (Section 5.5): the number of samples required to produce a false alarm
- Poisson distribution (Section 5.6): the number of white cells from a fixed amount of an individual's blood sample.

5.2 Discrete Uniform Distribution (8th Ed. only)

• Discrete Uniform Distribution: If the random variable X assumes the values $x_1, x_2, ..., x_k$, with equal probabilities, then the discrete uniform distribution is given by

$$f(x;k) = \frac{1}{k}, \quad x = x_1, x_2, \dots, x_k.$$

 Example (5.1): When a light bulb is selected at random from a box that contains a 40-watt bulb, a 60-watt bulb, a 75-watt bulb, and a 100-watt bulb, each element of the sample space S = {40, 60, 75, 100} occurs with probability 1/4. Therefore, we have a uniform distribution, with

$$f(x;4) = \frac{1}{4}, \quad x = 40, 60, 75, 100$$

• Example (5.2): When a die is tossed, each of the sample space $S = \{1, 2, 3, 4, 5, 6\}$ occurs with probability 1/6. Therefore, we have a uniform distribution, with

$$f(x;6) = \frac{1}{6}, \quad x = 1,2,3,4,5,6.$$

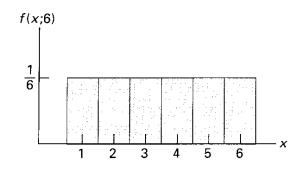


Figure 5.1 Histogram for the tossing of a die.

Discrete Uniform Distribution (8th Ed. only)

- Theorem 5.1: The mean and variance of the discrete uniform distribution f(x; k) are
 - $\mu = \frac{\sum_{i=1}^{k} x_i}{k}, \quad \sigma_X^2 = \frac{\sum_{i=1}^{k} (x_i \mu)^2}{k}.$
 - **Proof** $\mu = E(X) = \sum_{i=1}^{k} x_i f(x;k) = \sum_{i=1}^{k} \frac{x_i}{k} = \frac{\sum_{i=1}^{k} x_i}{k}$

$$\sigma^{2} = E[(X - \mu)^{2}] = \sum_{i=1}^{k} (x_{i} - \mu)^{2} f(x; k) = \sum_{i=1}^{k} \frac{(x_{i} - \mu)^{2}}{k} = \frac{\sum_{i=1}^{k} (x_{i} - \mu)^{2}}{k}.$$

Example 5.3: Referring to Example 5.2 (tossing a die), we find that

$$\mu = \frac{1+2+3+4+5+6}{6} = 3.5$$

$$\sigma^2 = \frac{(1-3.5)^2 + (2-3.5)^2 + \dots + (6-3.5)^2}{6} = \frac{35}{12}$$

- An experiment often consists of repeated trials, each with two possible outcomes that may be labeled success or failure.
- The most obvious application deals with the testing of items as they come off an assembly line, where each test/trial may indicate a defective or a nondefective item.
- The Bernoulli Process
 - The experiment consists of *n* repeated trials.
 - Each trial results in an outcome that may be classified as <u>a success</u> or a failure.
 - The probability of success, denoted by p, remains constant from trial to trial.
 - The repeated trials are independent.

 A Bernoulli trial can result in a success with probability p and a failure with probability q = 1 - p. Then the probability distribution of the binomial random variable X, the number of successes in n independent trials, is

$$b(x; n, p) = \binom{n}{x} p^{x} q^{n-x}, \quad x = 0, 1, 2, \dots, n.$$

 Example 5.1: The probability that a certain kind of component will survive a given shock test is ¾. Find the probability that exactly 2 of the next 4 components tested survive.

$$b(2;4,\frac{3}{4}) = {4 \choose 2} \left(\frac{3}{4}\right)^2 \left(\frac{1}{4}\right)^2 = \frac{4!}{2!2!} \cdot \frac{3^2}{4^4} = \frac{27}{128}.$$

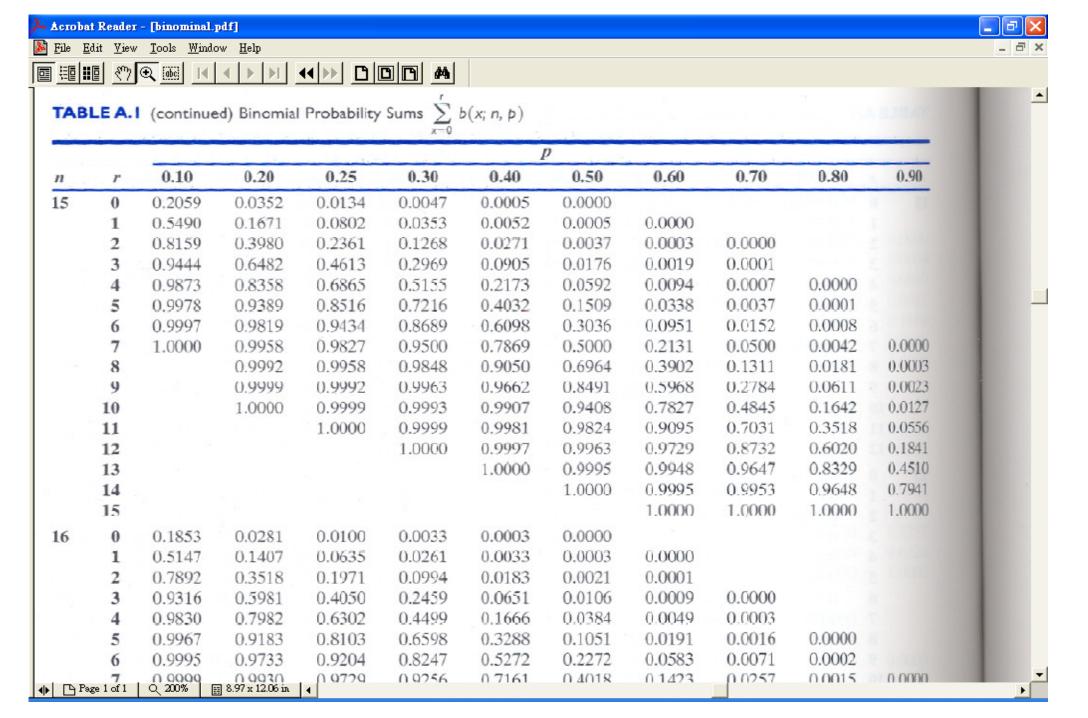
Binomial distribution corresponds to the binomial expansion of

$$(q+p)^n$$
, i.e., $(q+p)^n = \binom{n}{0}q^n + \binom{n}{1}pq^{n-1} + \binom{n}{2}p^2q^{n-2} + \dots + \binom{n}{n}p^n$
= $b(0;n,p) + b(1;n,p) + b(2;n,p) + \dots + b(n;n,p) = 1$

 Example 5.2: The probability that a patient recovers from a rare blood disease is 0.4. If 15 people are known to have contracted this disease, what is the probability that (a) at least 10 survive, (b) from 3 to 8 survive, and (c) exactly 5 survive?

(a)
$$P(X \ge 10) = 1 - P(X < 10) = 1 - \sum_{x=0}^{9} b(x;15,0.4) = 1 - 0.9662 = 0.0338$$

(b) $P(3 \le X \le 8) = \sum_{x=3}^{8} b(x;15,0.4) = \sum_{x=0}^{8} b(x;15,0.4) - \sum_{x=0}^{2} b(x;15,0.4) = 0.9050 - 0.0271 = 0.8779$
(c) $P(X = 5) = b(5;15,0.4) = \sum_{x=0}^{5} b(x;15,0.4) - \sum_{x=0}^{4} b(x;15,0.4) = 0.4032 - 0.2173 = 0.1859$



- Example 5.3: A large chain retailer purchases a certain kind of electronic device from a manufacturer. The manufacturer indicates that the defective rate of the device is 3%
 - a) The inspector of the retailer randomly picks 20 items from a shipment. What is the probability that there will be at least one defective item among these 20?
 - b) Suppose that the retailer receives 10 shipments in a month and the inspector randomly tests 20 devices per shipment. What is the probability that there will be 3 shipments containing at least one defective device?

Solution

(a) Denote by X the number of defective devices among the 20, X follows b(x;20,0.03).

$$P(X \ge 1) = 1 - P(X = 0) = 1 - b(0;20,0.03) = 1 - 0.03^{\circ}0.97^{20-0} = 0.4562$$

(b) Denote by Y the number of shipments containing at least one defective item, Y follows b(y;10,0.4562) $P(Y=3) = \binom{10}{2} 0.4562^3 (1-0.4562)^{10-3} = 0.1602$

- Theorem 5.1: The mean and variance of the binomial distribution b(x; n, p) are $\mu = np$ and $\sigma^2 = npq$.
 - Proof

Let the outcome on the jth trial represented by a Bernoulli random variable I_j . The number of successes in a binomial experiment is denoted by

$$X = I_1 + I_2 + \dots + I_n$$

$$E(I_j) = 0 \cdot q + 1 \cdot p = p$$

$$E(X \pm Y) = E(X) \pm E(Y).$$

$$\mu_X = E(X) = E(I_1) + E(I_2) + \dots + E(I_n) = \underbrace{p + p + \dots + p}_{n \text{ terms}} = np.$$

$$\sigma_{I_j}^2 = E[(I_j - p)^2] = E(I_j^2) - p^2 = (0^2 \cdot q + 1^2 \cdot p) - p^2 = p(1 - p) = pq.$$

$$\sigma_X^2 = \sigma_{I_1}^2 + \sigma_{I_2}^2 + \dots + \sigma_{I_n}^2 = \underbrace{pq + pq + \dots + pq}_{n \text{ terms}} = npq.$$

$$\sigma_{a_1X_1+a_2X_2+\cdots+a_nX_n}^2$$

$$= a_1^2 \sigma_{X_1}^2 + a_2^2 \sigma_{X_2}^2 + \cdots + a_n^2 \sigma_{X_n}^2.$$

• Example 5.5: Find the mean and variance of the binomial random variable of Example 5.2 (n = 15, p = 0.4), and then use Chebyshev's theorem to interpret the interval $\mu \pm 2\sigma$.

(Example 5.2: The probability that a patient recovers from a rare blood disease is 0.4.)

Solution

Example 5.5 was a binomial experiment with n = 15 and p = 0.4

$$\mu = 15 \cdot 0.4 = 6$$

$$\sigma^2 = 15 \cdot 0.4 \cdot 0.6 = 3.6, \sigma = \sqrt{3.6} = 1.897$$

* Chebyshev's theorem:

$$P(\mu - k\sigma < X < \mu + k\sigma) \ge 1 - \frac{1}{k^2}$$

The interval $\mu \pm 2\sigma = 6 \pm 2.1.897 \Rightarrow 2.206$ to 9.794

has a probability of at least $\frac{3}{4}$.

- Example 5.6: It is conjectured that an impurity exists in 30% of all drinking wells in a city. Suppose 10 wells are randomly selected and 6 are found to contain the impurity. What does this imply about the conjecture? Use a probability statement.
 - Solution If the conjecture is correct, is it likely that we could have found 6 or more impure wells?

$$P(X \ge 6) = \sum_{x=0}^{10} b(x;10,0.3) - \sum_{x=0}^{5} b(x;10,0.3)$$
$$= 1 - 0.9527 = 0.0473$$

As a result, it is unlikely (4.7% chance) that 6 wells would be found impure if only 30% of all are impure.

This casts considerable doubt on the conjecture and suggests that the impurity problem is much more severe.

Multinomial Distribution: If a given trial can result in the k outcomes E_1 , E_2 ,..., E_k with probabilities p_1 , p_2 ,..., p_k , then the probability distribution of the random variables X_1 , X_2 ,..., X_k , representing the number of occurrences for E_1 , E_2 ,..., E_k in n independent trials is

$$f(x_1, x_2, \dots, x_k; p_1, p_2, \dots, p_k, n) = \binom{n}{x_1, x_2, \dots, x_k} p_1^{x_1} \cdot p_2^{x_2} \cdot \dots \cdot p_k^{x_k}$$
with $\sum_{i=1}^k x_i = n$ and $\sum_{i=1}^k p_i = 1$.

- Example 5.7: The complexity of arrivals and departures into an airport are such that computer simulation is often used to model the "ideal" conditions. For a certain airport containing three runways it is known that in the ideal setting the following are the probabilities that the individual runways are accessed by a randomly arriving commercial jet: Runway 1: p₁= 2/9, Runway 2: p₂= 1/6, Runway 3: p₃= 11/18. What is the probability that 6 randomly arriving airplanes are distributed in the following fashion? Runway 1: 2 airplanes, Runway 2: 1 airplanes, Runway 3: 3 airplanes.
 - Solution

Using multinomial distribution

$$f(2,1,3;\frac{2}{9},\frac{1}{6},\frac{11}{18},6) = {\binom{6}{2,1,3}} {\left(\frac{2}{9}\right)}^2 {\left(\frac{1}{6}\right)}^1 {\left(\frac{11}{18}\right)}^3 = \frac{6!}{2!1!3!} \cdot \frac{2^2}{9^2} \cdot \frac{1}{6} \cdot \frac{11^3}{18^3} = 0.1127$$

- Binomial distribution: the sampling with replacement
 Hypergeometric distribution: the sampling without replacement
- Hypergeometric experiment
 - 1. A random sample of size *n* is selected without replacement from *N* items.
 - 2. k of the N items may be classified as successes and N-k as failures.
- Hypergeometric random variable: the number X of successes of a hypergeometric experiment.
- Hypergeometric distribution: the probability distribution of the hypergeometric variable X, the number of successes in a

$$h(x; N; n, k) = \frac{\binom{k}{x} \binom{N-k}{n-x}}{\binom{N}{n}}$$

random sample of size \underline{n} selected from N items of which \underline{k} are labeled success and N-k labeled failure.

• Example 5.9: Lots of 40 components each are called unacceptable if they contain as many as 3 defective or more. The procedure for sampling the lot is to select 5 components at random and to reject the lot if a defective is found. What is the probability that exactly 1 defective is found in the sample if there are 3 defectives in the entire lot?

Solution

Using hypergeometric distribution with x = 1, N = 40, n = 5, and, k = 3

$$h(1;40,5,3) = \frac{\binom{3}{1}\binom{37}{4}}{\binom{40}{5}} = 0.3011.$$

So this plan is likely not desirable since it detects a bad lot (3 defectives) only about 30% of the time.

 Theorem 5.2: The mean and variance of the hypergeometric distribution h(x; N, n, k) are

$$\mu = \frac{nk}{N}$$
 and $\sigma^2 = \frac{N-n}{N-1} \cdot n \cdot \frac{k}{N} (1 - \frac{k}{N})$

(the proof is shown in Appendix A24)

- Exapmle 5.11: Find the mean and variance of the random variable of Example 5.9 (n = 5, N = 40, and k = 3) and then use Chebyshev's theorem to interpret the interval $\mu \pm 2\sigma$.
 - Solution

$$\mu = \frac{5.3}{40} = \frac{3}{8} = 0.375$$

$$\sigma^2 = (\frac{40-5}{39})(5)(\frac{3}{40})(1-\frac{3}{40}) = 0.3113 \Rightarrow \sigma = 0.558$$

$$\mu \pm 2\sigma = 0.375 \pm 2 \cdot 0.558$$

 \Rightarrow has a probability of at least 3/4 of falling between -0.741 and 1.491.

PROOF To find the mean of the hypergeometric distribution, we write

$$E(X) = \sum_{x=0}^{n} x \frac{\binom{k}{x} \binom{N-k}{n-x}}{\binom{N}{n}} = k \sum_{x=1}^{n} \frac{(k-1)!}{(x-1)!(k-x)!} \cdot \frac{\binom{N-k}{n-x}}{\binom{N}{n}}$$
$$= k \sum_{x=1}^{n} \frac{\binom{k-1}{x-1} \binom{N-k}{n-x}}{\binom{N}{n}}.$$

Letting y = x - 1, we find that this becomes

$$E(X) = k \sum_{y=0}^{n-1} \binom{k-1}{y} \binom{N-k}{n-1-y}.$$

Writing

$$\begin{pmatrix} N-k \\ n-1-y \end{pmatrix} = \begin{pmatrix} (N-1)-(k-1) \\ n-1-y \end{pmatrix}$$

and

$$\binom{N}{n} = \frac{N!}{n!(N-n)!} = \frac{N}{n} \binom{N-1}{n-1},$$

we obtain

$$E(X) = \frac{nk}{N} \sum_{y=0}^{n-1} \frac{\binom{k-1}{y} \binom{(N-1)-(k-1)}{n-1-y}}{\binom{N-1}{n-1}} = \frac{nk}{N},$$

since the summation represents the total of all probabilities in a hypergeometric experiment when n-1 items are selected at random from N-1, of which k-1 are labeled success.

- Relationship to the Binomial Distribution
 - If n is small compared to N, the nature of the N items changes very little in each draw. (when $\frac{n}{N} \le 0.05$)
 - little in each draw. (when $\frac{n}{N} \le 0.05$) $\mu = np = \frac{nk}{N}, \ \sigma^2 = npq = n \cdot \frac{k}{N} (1 \frac{k}{N})$ $(\frac{N-n}{N-1} \text{ is negligible when } n \text{ is small relative to } N).$
 - The binomial distribution may be viewed as a large population edition of the hypergeometric distributions.
- Example 5.12: A manufacture of automobile tires reports that among <u>a shipment of 5000</u> sent to a local distributor, 1000 are slightly blemished. If one purchases <u>10 of these tires</u> at random from the distributor, what is the probability that exactly 3 are blemished?
 - Solution

$$h(3;5000,10,1000) \approx \underline{b(3;10,0.2)} = \sum_{x=0}^{3} b(x;10,0.2) - \sum_{x=0}^{2} b(x;10,0.2)$$
$$= 0.8791 - 0.6778 = 0.2013$$

20

Multivariate Hypergeometric Distribution:

If N items can be partitioned into the k cells $A_1, A_2, ..., A_k$ with a_1 , a_2, \ldots, a_k elements, respectively, then the probability distribution of the random variable $X_1, X_2, ..., X_k$, representing the number of elements selected from $A_1, A_2, ..., A_k$ in a random sample of size n, is

$$f(x_1, x_2, \dots, x_k; a_1, a_2, \dots, a_k, N, n) = \frac{\binom{a_1}{x_1}\binom{a_2}{x_2} \dots \binom{a_k}{x_k}}{\binom{N}{n}} \quad \text{with } \sum_{i=1}^k x_i = n \text{ and } \sum_{i=1}^k a_i = N$$

Example 5.13: A group of 10 individuals are used for a biological case study. The group contains 3 people with blood type O, 4 with blood type A, and 3 with blood type B. What is the probability that a random sample of 5 will contain 1 person with blood type O, 2 with blood type A, and 2 with blood type B?

- Solution $f(1,2,2;3,4,3,10,5) = \frac{\binom{3}{1}\binom{4}{2}\binom{3}{2}}{\binom{10}{5}} = \frac{3}{14}$

$$f(1,2,2;3,4,3,10,5) = \frac{\binom{3}{1}\binom{4}{2}\binom{3}{2}}{\binom{10}{5}} = \frac{3}{14}$$