

## AI1103 Assignment 2

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Q59. Suppose that  $\begin{pmatrix} X \\ Y \end{pmatrix}$  has a bivariate density  $f = \frac{1}{2}f_1 + \frac{1}{2}f_2$ , where  $f_1$  and  $f_2$  are respectively, the densities of bivariate normal distribution  $N(\mu_1, \Sigma)$ , and  $N(\mu_2, \Sigma)$ , with  $\mu_1 = \begin{pmatrix} 1 \\ 1 \end{pmatrix}$ ,  $\mu_2 = \begin{pmatrix} -1 \\ -1 \end{pmatrix}$  and  $\Sigma = \mathbb{I}_2$ , the  $2 \times 2$  identity matrix. Then which of the following is correct ?

- a)  $X$  and  $Y$  are positively correlated
- b)  $X$  and  $Y$  are negatively correlated
- c)  $X$  and  $Y$  are uncorrelated but they are not independent
- d)  $X$  and  $Y$  are independent

Answer:

To check the correlation between  $X$  and  $Y$  we can calculate the covariance of  $X$  and  $Y$ . If the value of covariance is  $+ve$  then  $X$  and  $Y$  are positively correlated, if the value is  $-ve$  then  $X$  and  $Y$  are negatively correlated and if the value is 0 then  $X$  and  $Y$  are independent.

For a bivariate, covariance is defined as

$$cov(X, Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - \mu_x)(y - \mu_y)f(x, y)dxdy \quad (1)$$

The bivariate normal distribution is defined as

$$f(\underline{X}) = \left(2\pi\sqrt{|\Sigma|}\right)^{-1} e^{-\frac{1}{2}[(\underline{X}-\underline{\mu})^T \Sigma^{-1}(\underline{X}-\underline{\mu})]} \quad (2)$$

where

$$\underline{X} = \begin{pmatrix} X \\ Y \end{pmatrix}, \Sigma = \begin{pmatrix} Var[X] & Cov(X, Y) \\ Cov(X, Y) & Var[Y] \end{pmatrix} = \begin{pmatrix} \sigma_X^2 & \rho\sigma_X\sigma_Y \\ \rho\sigma_X\sigma_Y & \sigma_Y^2 \end{pmatrix}$$

$\rho$  is the correlation coefficient.  
and standard normal distribution is given by

$$g(\underline{X}) = \frac{1}{2\pi} e^{-\frac{1}{2} [\underline{X}^T \underline{X}]} \quad (3)$$

where,

$$\underline{\Sigma} = \mathbb{I}_2, \underline{\mu} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

For the given problem mean can be obtained as such

$$\mathbb{E}_f \left[ \begin{pmatrix} X \\ Y \end{pmatrix} \right] = \frac{1}{2} \mathbb{E}_{f_1} \left[ \begin{pmatrix} X \\ Y \end{pmatrix} \right] + \frac{1}{2} \mathbb{E}_{f_2} \left[ \begin{pmatrix} X \\ Y \end{pmatrix} \right] = \frac{1}{2} \underline{\mu}_1 + \frac{1}{2} \underline{\mu}_2 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

and

$$\begin{aligned} cov(X, Y) &= \mathbb{E}_f [(X - \mu_x)(Y - \mu_y)] \\ &= \frac{1}{2} \mathbb{E}_{f_1} [XY] + \frac{1}{2} \mathbb{E}_{f_2} [XY] \quad (\mu_x = 0, \mu_y = 0) \end{aligned} \quad (4)$$

Now,

$$\begin{aligned} \mathbb{E}_{f_1} [X, Y] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_1 dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \frac{1}{2\pi} e^{-\frac{1}{2} [(\underline{X} - \underline{\mu}_1)^T (\underline{X} - \underline{\mu}_1)]} dx dy \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy e^{-\frac{1}{2} (\|\underline{X}\|^2 + \|\underline{\mu}_1\|^2 - 2\underline{X}^T \underline{\mu}_1)} dx dy \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy e^{-\frac{1}{2} (x^2 + y^2 + 2 - 2x - 2y)} dx dy \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy e^{-\frac{1}{2} [(x-1)^2 + (y-1)^2]} dx dy \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x' + 1)(y' + 1) e^{-\frac{1}{2} [(x')^2 + (y')^2]} dx dy \quad (x' = x - 1, y' = y - 1) \\ &= \mathbb{E}_g [X' + 1, Y' + 1] \quad (\text{comparing with (3)}) \end{aligned}$$

Similarly,

$$\mathbb{E}_{f_2} [X, Y] = \mathbb{E}_g [X'' + 1, Y'' + 1] \quad (X'' = X + 1, Y'' = Y + 1)$$

$$\begin{aligned}
\mathbb{E}_{f_1} [XY] &= \mathbb{E}_g [(X' + 1)(Y' + 1)] \\
&= \mathbb{E}_g [X'Y'] + \mathbb{E}_g [X'] + \mathbb{E}_g [Y'] + \mathbb{E}_g [1] \\
&= 0 + 0 + 0 + 1 \\
&= 1
\end{aligned} \tag{5}$$

$$\begin{aligned}
\mathbb{E}_{f_2} [XY] &= \mathbb{E}_g [(X'' - 1)(Y'' - 1)] \\
&= \mathbb{E}_g [X''Y''] - \mathbb{E}_g [X''] - \mathbb{E}_g [Y''] + \mathbb{E}_g [1] \\
&= 0 + 0 + 0 + 1 \\
&= 1
\end{aligned} \tag{6}$$

Substituting (5) and (6) back in (4)

$$\begin{aligned}
cov(X, Y) &= \frac{1}{2} \mathbb{E}_{f_1} [XY] + \frac{1}{2} \mathbb{E}_{f_2} [XY] \\
&= \frac{1}{2} + \frac{1}{2} \\
&= 1
\end{aligned} \tag{$(X, Y \text{ are positively correlated})$}$$