

2) Rectangular distribution in

(uniform distribution)  
(continuous)

\* A very simple distribution for a continuous R.V is the uniform distribution.

\* It is particularly useful in theoretical statistics, bcz it is convenient to deal with mathematically

def →

~~that~~ continuous R.V 'x' is said to be rectangular distribution, if its pdf given by,

$$f(x) = \frac{1}{b-a}$$

$$a \leq x \leq b$$

= 0, elsewhere.



Remark  $\rightarrow$

\*  $a$  and  $b$  be (arb) are 2 parameter of uniform distrib. on  $a, b$ .

\* The distrib.  $\rightarrow$  rectangular distrib. Since the curve,  $y = f(x)$ , describes a  $\square$  over the  $x$ -axis and  $b/w$  the ordinates at  $x=a$  &  $x=b$ .

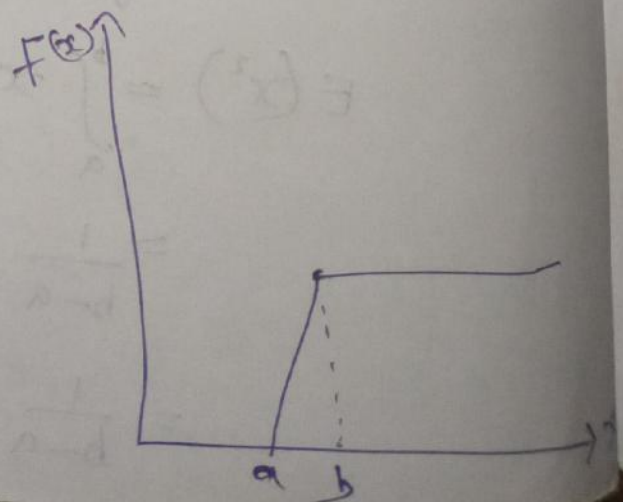
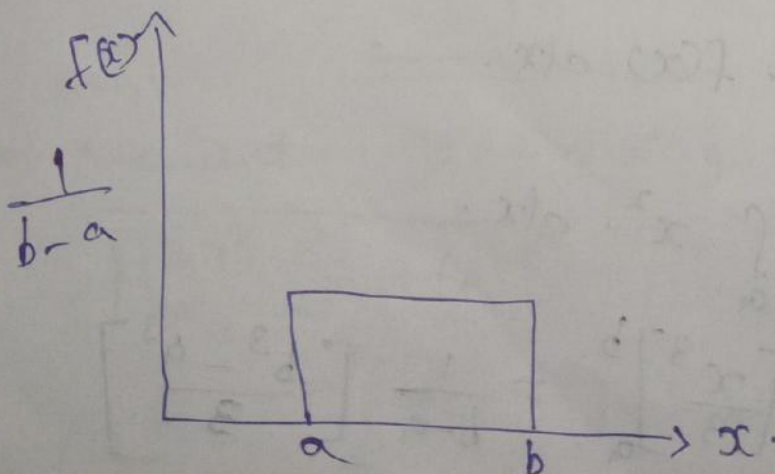
\* The distrib. ( )  $F(x) = \begin{cases} 0, & \text{if } -\infty < x < a \\ \frac{x-a}{b-a}, & a < x < b \\ 1, & b < x < \infty. \end{cases}$

Since  $F(x)$  is not continuous at  $x=a$  &  $x=b$ , it is not differentiable at this point. Thus,

$$\boxed{F'(x) = f(x) = \frac{1}{b-a} \neq 0}$$

exist everywhere except at the point  $x=a$  &  $x=b$ . & consequently we get pdf  $f(x)$

\* The graph of uniform pdf  $f(x)$  and the corresponding distrib. ( )  $F(x)$  are given below,





$$= \frac{b^3 - a^3}{3(b-a)}$$

$$b^3 - a^3 = (b-a)(b^2 + ab + a^2)$$

$$= \frac{(b-a)(b^2 + ab + a^2)}{3(b-a)}$$

$$E(x^2) = \frac{b^2 - ab + a^2}{3}$$

$$\therefore V(x) = \frac{b^2 - ab - a^2}{3} - \left[ \frac{b+a}{2} \right]^2$$

$$= \frac{b^2 - ab - a^2}{3} - \frac{(b+a)^2}{4}$$

$$= \frac{b^2 - ab - a^2}{3} - \frac{b^2 + 2ab + a^2}{4}$$

$$= \frac{4(b^2 - ab - a^2) - 3(b^2 + 2ab + a^2)}{12}$$

$$= \frac{4b^2 - 4ab - 4a^2 - 3b^2 - 6ab - 3a^2}{12}$$

$$= \frac{b^2 + 2ab - a^2}{12}$$

$$= \frac{(b-a)^2}{12}$$

$$(b-a)^2 = b^2 + 2ab + a^2$$

$\Rightarrow$  Moment generating function :-

$$M_x(t) = E(e^{tx})$$

$$= \int_a^b e^{tx} \cdot f(x) dx$$

\* for a  $\square$ -law / uniform variate 'x'  
pdf is given by,

$$f(x) = \begin{cases} \frac{1}{b-a} & , -a < x < a \\ 0 & , \text{elsewhere} \end{cases}$$

~~mean :-~~

$$* \text{ mean :- } E(x)$$

$$= \int_a^b x \cdot f(x) dx$$

$$= \int_a^b x \cdot \frac{1}{b-a} dx$$

$$= \frac{1}{b-a} \int_a^b x \cdot dx = \frac{1}{b-a} \left[ \frac{x^2}{2} \right]_a^b$$

$$= \frac{1}{b-a} \times \frac{b^2 - a^2}{2}$$

$$\left[ \frac{b^2 - a^2}{2} = (b+a)(b-a) \right]$$

$$= \frac{1}{(b-a)} \times \frac{(b+a)(b-a)}{2}$$

$$\text{mean} = \frac{b+a}{2}$$

\* Variance :-

$$V(x) = E(x^2) - (E(x))^2$$

$$E(x^2) = \int_a^b x^2 \cdot f(x) dx$$

$$= \frac{1}{b-a} \int_a^b x^2 \cdot dx$$

$$= \frac{1}{b-a} \left[ \frac{x^3}{3} \right]_a^b = \frac{1}{b-a} \left[ \frac{b^3 - a^3}{3} \right]$$



$$\sqrt{n+1} = n! = n(n-1)! = n! \quad (a)$$

$$\text{putting } m=1 \quad \xi_1 \quad p=\frac{1}{2}$$

$$f(x) = \frac{1}{\Gamma(p)} e^{-mx} x^{p-1}$$

$$\int_0^\infty e^{-mx} x^{p-1} dx = \frac{\Gamma(p)}{m^p}$$

$$\Rightarrow \int_0^\infty e^{-x} x^{p-1} dx = \Gamma(p) = \sqrt{\pi}$$

$\Rightarrow$  moment

$$(1) \text{ mean } = E(x) = \int_0^\infty x \cdot f(x) dx$$

$$= \int_0^\infty x \frac{1}{\Gamma(p)} e^{-mx} x^{p-1} dx$$

$$= \frac{1}{\Gamma(p)} \int_0^\infty x^p e^{-mx} dx$$

$$= \frac{1}{\Gamma(p)} \int_0^\infty e^{-mx} x^{(p+1)-1} dx$$

$$= \frac{1}{\Gamma(p)} \cdot \frac{\Gamma(p+1)}{m^{p+1}}$$

$$= \frac{1}{\Gamma(p)} \cdot \frac{p!}{m^{p+1}}$$

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$$= \int_0^\infty e^{-bx} \cdot \frac{1}{b-a} dx$$

$$= \frac{1}{b-a} \left[ \frac{e^{-bx}}{-b} \right]_0^\infty = \frac{1}{b-a} \left[ \frac{e^{-bx}}{-b} - \frac{e^{-ax}}{-a} \right]$$

$$= \frac{1}{b-a} \left[ \frac{e^{-bx}}{-b} - \frac{e^{-ax}}{-a} \right]$$

$\Rightarrow$  Gamma distribution

is said to be a gamma distribution if its pdf is given by

$$f(x) = \frac{1}{\Gamma(p)} e^{-mx} x^{p-1}$$

where  $m > 0$ ,  $p > 0$  are parameters & conditions.

Note

Being a pdf we know that  $\int_0^\infty f(x) dx = 1$ .

if  $p = n$ , a true integer

if  $p = n$ , then  $\Gamma(n) = (n-1)!$

using integration by parts we can get

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using integration by parts we can get

if  $p = n$ , a true integer

$$\Rightarrow \int_0^\infty e^{-mx} x^{p-1} dx = \frac{\Gamma(p)}{m^p}$$

$$= \frac{1}{\Gamma(p)} \int_0^\infty e^{-mx} x^{p-1} dx = 1$$

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$$= \int_0^{\infty} e^{tx} \cdot \frac{n!}{p} \cdot e^{-mx} \cdot x^{p-1} dx$$

$$= \frac{n!}{p} \int_0^{\infty} e^{tx} \cdot e^{-mx} \cdot x^{p-1} dx$$

$$= \frac{n!}{p} \int_0^{\infty} e^{-(m-t)x} \cdot x^{p-1} dx$$

$$= \frac{n!}{p} \cdot \frac{1}{(m-t)^p}$$

$$= \frac{n!}{(m-t)^p} = \left( \frac{n!}{m-t} \right)^p$$

$$= \left( \frac{n!}{m(1-\frac{t}{m})} \right)^p$$

$$= \left( \frac{1}{1-\frac{t}{m}} \right)^p = \left( \frac{1-\frac{t}{m}}{1} \right)^p$$

$$= \left( 1 - \frac{t}{m} \right)^{-p}$$

Exponential distribution

It is continuous r.v.  $x$  has a pdf

$$f(x) = \theta e^{-\theta x} \quad x > 0, \theta > 0$$

\* Mean :-

$$E(x) = \frac{1}{\theta}$$

$$E(x) = \int_0^{\infty} x \cdot f(x) dx = \int_0^{\infty} x \cdot \theta e^{-\theta x} \cdot dx$$

$$= \int_0^{\infty} x^{2-1} \cdot \theta e^{-\theta x} dx = \theta \int_0^{\infty} \underbrace{x^{2-1}}_{x^1} \cdot e^{-\theta x} \cdot dx$$

(2) Variance  $\rightarrow E(x^2) - (E(x))^2$

$$V(x) = E(x^2) - (E(x))^2$$

$$E(x^2) = \int_0^{\infty} x^2 \cdot \frac{n!}{p} \cdot e^{-mx} \cdot x^{p-1} dx$$

$$= \frac{n!}{p} \int_0^{\infty} x^2 \cdot e^{-mx} \cdot x^{p-1} dx$$

$$= \frac{n!}{p} \int_0^{\infty} e^{-mx} \cdot x^{(p+2)-1} dx$$

$$\Rightarrow \frac{n!}{p} \times \frac{1}{m^{p+2}} = \frac{n!}{p} \cdot \frac{(p+1)p!}{m^{p+2}}$$

$$\Rightarrow \frac{n!}{p} \cdot \frac{(p+1)p!}{m^{p+2}} = \frac{(p+1)p!}{m^{p+2}}$$

$$\Rightarrow \frac{(p+1)p!}{m^{p+2}} = \frac{(p+1)!}{m^{p+2}}$$

$$E(x^2) = \frac{(p+1)!}{m^{p+2}}$$

$$V(x) = \frac{(p+1)!}{m^{p+2}} - \left( \frac{p!}{m^p} \right)^2$$

$$= \frac{(p+1)!}{m^{p+2}} - \frac{p^2}{m^2} = \frac{p^2 + p - p^2}{m^2}$$

$$V(x) = \frac{p}{m^2}$$

(3) Moment generating (\*) :-

$$M(x) = E(e^{tx})$$

$$= \int_0^{\infty} e^{tx} \cdot f(x) dx$$



\* Mgf :-

$$M_x(t) = \left[ 1 - \frac{t}{\theta} \right]^{-1}$$

$$M_x(t) = E(e^{tx}) = \int_0^\infty e^{tx} \cdot f(x) dx$$

$$= \int_0^\infty e^{tx} \cdot \theta e^{-\theta x} dx$$

$$= \theta \int_0^\infty e^{tx} \cdot e^{-\theta x} dx = \theta \int_0^\infty e^{(t-\theta)x} dx$$

$$= \theta \int_0^\infty e^{-(\theta-t)x} dx$$

$$= \theta \cdot \left[ \frac{e^{-(\theta-t)x}}{-(\theta-t)} \right]_0^\infty \quad (\text{if } \theta > t)$$

$$\bullet \text{ If } t = \theta$$

$$= \theta \cdot \left[ \frac{e^{-(\theta-t)x}}{-(\theta-t)} \right]_0^\infty = \theta \cdot \left[ \frac{e^{-(\theta-t)x}}{-(\theta-t)} \right]_0^\infty$$

$$= \theta \cdot \left[ \frac{e^{-(\theta-t)x}}{-(\theta-t)} \right]_0^\infty = \theta \cdot \left[ \frac{e^{-(\theta-t)x}}{-(\theta-t)} \right]_0^\infty$$

$$= \theta \cdot \frac{1}{\theta-t} = \frac{\theta}{\theta-t} \quad (\text{if } t < \theta)$$

$$= \left[ \frac{\theta}{\theta-t} \right]^{-1} = \left[ \frac{\theta}{\theta-t} \right]^{-1}$$

$$M_x(t) = \left[ 1 - \frac{t}{\theta} \right]^{-1}$$

\* For distribution :-

Exponential distribution

•

$$\int_0^\infty e^{-nx} \cdot x^{p-1} dx = \frac{\Gamma(p)}{n^p}$$

$$p = 2 \quad \therefore \Rightarrow \theta \cdot \frac{\sqrt{2}}{\theta^2} = \frac{\sqrt{2}}{\theta}$$

$$\sqrt{n} = (n-1) \sqrt{n-1}$$

$$\therefore \Rightarrow \sqrt{2} = (2-1) \sqrt{2} = \sqrt{2}$$

$$\therefore \Rightarrow \frac{1}{\theta}$$

$$\text{mean} = \frac{1}{\theta}$$

\* Variance :-

$$V(x) = \frac{1}{\theta^2}$$

$$V(x) = E(x^2) - (E(x))^2$$

$$E(x^2) = \int_0^\infty x^2 \cdot f(x) dx = \int_0^\infty x^2 \cdot \theta e^{-\theta x} dx$$

$$= \theta \int_0^\infty x^{3-1} \cdot e^{-\theta x} dx = \theta \int_0^\infty x^{3-1} \cdot e^{-\theta x} dx = \frac{\Gamma(3)}{\theta^3}$$

$$= \theta \cdot \frac{\sqrt{3}}{\theta^3} = \frac{\sqrt{3}}{\theta^2}$$

$$\sqrt{3} = (3-1) \sqrt{3-1} = 2\sqrt{2}$$

$$E(x^2) = \frac{\sqrt{3}}{\theta^2}$$

$$V(x) = E(x^2) - (E(x))^2$$

$$= \frac{\sqrt{3}}{\theta^2} - \frac{1}{\theta^2} = \frac{2-1}{\theta^2} = \frac{1}{\theta^2}$$



1) Mean :-

$$E(x) = \int_0^1 x \cdot f(x) dx = \int_0^1 x \cdot \frac{1}{\beta(m,n)} x^{m-1} (1-x)^{n-1} dx$$

$$= \frac{1}{\beta(m,n)} \int_0^1 x^m (1-x)^{n-1} dx$$

$$= \frac{1}{\beta(m,n)} \int_0^1 x^{(m+1)-1} (1-x)^{n-1} dx$$

$$= \frac{\beta(m+1, n)}{\beta(m, n)}$$

$$\frac{\beta(m+1, n)}{\beta(m, n)} = \frac{\frac{\Gamma(m+1)\Gamma(n)}{\Gamma(m+n+1)}}{\frac{\Gamma(m)\Gamma(n)}{\Gamma(m+n)}} = \frac{m}{m+n}$$

$$= \frac{\frac{\Gamma(m+1)\Gamma(n)}{\Gamma(m+n+1)}}{\frac{\Gamma(m)\Gamma(n)}{\Gamma(m+n)}} = \frac{m}{m+n}$$

$$\frac{\Gamma(m+1)\Gamma(n)}{\Gamma(m+n+1)} = \frac{m\Gamma(m)\Gamma(n)}{(m+n)\Gamma(m+n)}$$

$$= \frac{m}{m+n} \frac{\Gamma(m)\Gamma(n)}{\Gamma(m+n)}$$

$$E(x) = \frac{m}{m+n}$$

2) If  $x$  is a r.v with continuous distribution  $F$  the  $F(x)$  has a uniform distribution on  $[0,1]$ ?

Ans) Let  $f(x)$  be the pdf of  $x$ .

Let  $y = F(x)$ .

$$\therefore, f(x) = \frac{dF(x)}{dx} = \frac{dy}{dx}$$

Let  $g(y)$  be the pdf of  $y$ .

$$g(y) = F(x) \cdot \left| \frac{dx}{dy} \right|$$

$$= F(x) \cdot \frac{1}{\frac{dy}{dx}} = \frac{F(x)}{f(x)} = 1$$

$$0 < y < 1$$

$\therefore y \rightarrow$  a uniform distribution in  $[0,1]$ .

$\Rightarrow$  Beta distribution is [1st kind]

If a r.v 'x' has pdf given by

$$f(x) = \frac{1}{\beta(m,n)} x^{m-1} (1-x)^{n-1} \quad 0 < x < 1$$

$m > 0$  &  $n > 0$  then  $x$  is said to have beta distribution.

$$\beta(m,n) = \int_0^1 x^{m-1} (1-x)^{n-1} dx \rightarrow \beta(m,n)$$

$$\beta(m,n) \text{ or } \beta_1(m,n)$$

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$$= \frac{(n+1)n}{n+2+n} = \frac{(n+1)n}{n+n+1}$$

$$E(x^2) = \frac{(n+1)n}{(n+1)(n+n)} = \frac{(n+1)n}{(n+n+1)(n+n)}$$

$$V(x) = E(x^2) - (E(x))^2$$

$$= \frac{(n+1)n}{(n+n+1)(n+n)} - \left(\frac{n}{n+n}\right)^2$$

$$= \frac{n}{n+n} \left[ \frac{(n+1)}{n+n+1} - \frac{n}{n+n} \right]$$

$$= \frac{n}{n+n} \left[ \frac{(n+1)(n+n) - n(n+n+1)}{(n+n+1)(n+n)} \right]$$

$$= \frac{n}{n+n} \left[ \frac{n^2 + n^2 + n + n - n^2 - n^2 - n}{(n+n+1)(n+n)} \right]$$

$$= \frac{n}{n+n} \left[ \frac{n}{(n+n+1)(n+n)} \right]$$

$$V(x) = \frac{n}{(n+n+1)(n+n)^2}$$

2) Variance :-

$$V(x) = E(x^2) - (E(x))^2$$

$$E(x^2) = \int_0^1 x^2 \cdot \frac{1}{\beta(m,n)} x^{m-1} (1-x)^{n-1} dx$$

$$= \frac{1}{\beta(m,n)} \int_0^1 x^2 x^{m-1} (1-x)^{n-1} dx$$

$$= \frac{1}{\beta(m,n)} \int_0^1 x^{m+2-1} (1-x)^{n-1} dx$$

$$\beta(m,n) = \int_0^1 x^{m-1} (1-x)^{n-1} dx$$

$$\therefore \beta(m+2,n) = \int_0^1 x^{m+2-1} (1-x)^{n-1} dx$$

$$\Rightarrow \frac{1}{\beta(m,n)} \cdot \beta(m+2,n)$$

$$\Rightarrow \frac{\beta(m+2,n)}{\beta(m,n)} = \frac{\int_0^1 x^{m+2-1} (1-x)^{n-1} dx}{\int_0^1 x^{m-1} (1-x)^{n-1} dx}$$

$$\frac{\int_0^1 x^{m+2-1} (1-x)^{n-1} dx}{\int_0^1 x^{m-1} (1-x)^{n-1} dx}$$

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=> Beta Distribution of 2nd kind :-

Let 'x' be a continuous r.v assuming values from 0 to 1, if the pdf of x is given by,

$$f(x) = \frac{1}{B(m, n)} \frac{x^{m-1}}{(1+x)^{m+n}}, \quad m, n > 0, \quad 0 < x < 1$$

then x is said to follow a beta distribution of 2nd kind & is denoted by  $\beta_2(m, n)$ ,

Note →

\*  $\beta$  distribution of 2nd kind can be transformed to  $\beta$  distribution of 1st kind by the transformation,

$$1+x = \frac{1}{y} \quad (\text{i.e.})$$

if  $x \sim \beta_2(m, n)$  then y is defined above follow  $\beta_1(m, n)$

\* As above we can show that,

$$\begin{aligned} \text{mean} = E(x) &= \frac{m}{n-1} \\ \text{variance} &= \frac{m(m+n-1)}{(n-1)^2(n-2)} \end{aligned}$$



Q) Rectangular distribution

→ Moments of normal distribution =

1) Mean :-

$$E(X) = \mu$$

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_{-\infty}^{\infty} x \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \int_{-\infty}^{\infty} [x - \mu + \mu] \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \int_{-\infty}^{\infty} (x - \mu) \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx + \mu \int_{-\infty}^{\infty} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$



$$\begin{aligned}
 &= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} (x-\mu) e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx + \mu \int_{-\infty}^{\infty} f(x) dx \\
 &= \frac{1}{\sigma \sqrt{2\pi}} \underbrace{\int_{-\infty}^{\infty} (x-\mu) e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx}_{=0} + \mu \cdot 1 \\
 &= \frac{1}{\sigma \sqrt{2\pi}} \times 0 + \mu
 \end{aligned}$$

$E(x) = \mu$

$\int_{-\infty}^{\infty} (x-\mu) e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \rightarrow$  odd f(x)  
 $\Rightarrow 0$   
 $f(x) = -f(x)$  by the prop. of odd f(x) = 0

3) Variance :-

$V(x) = \sigma^2$

$$\begin{aligned}
 V(x) &= E(x - E(x))^2 \\
 V(x) &= E(x - \mu)^2
 \end{aligned}$$

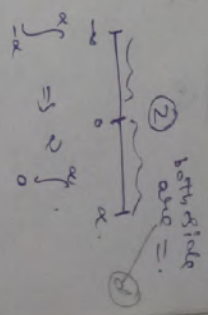
$$\begin{aligned}
 &= \int_{-\infty}^{\infty} (x-\mu) \cdot f(x) dx \\
 &= \int_{-\infty}^{\infty} (x-\mu) \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx
 \end{aligned}$$

by rule method

$z = \frac{x-\mu}{\sigma} \Rightarrow x-\mu = z\sigma$   
 $dx = \sigma dz$   
 $\therefore x-\mu = z\sigma$

$$\begin{aligned}
 d(x-\mu) &= d(z\sigma) \\
 dx &= \sigma dz
 \end{aligned}$$

$$\begin{aligned}
 &= \int_{-\infty}^{\infty} (z\sigma)^2 \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{z^2 \sigma^2}{2\sigma^2}} \sigma dz \\
 &= \int_{-\infty}^{\infty} (z\sigma)^2 \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{z^2}{2}} \sigma dz \\
 &= \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^2 e^{-\frac{z^2}{2}} dz
 \end{aligned}$$



Rule method

let  $u = \frac{z^2}{2}$

$z^2 = 2u \Rightarrow z = \sqrt{2u}$   
 $d(z^2) = d(2u)$

$2z dz = 2 du$   
 $z dz = du \Rightarrow \frac{du}{z}$

$$= \frac{\sigma^2}{\sqrt{2\pi}} \int_0^{\infty} z^2 e^{-z^2/2} dz$$

$$\Rightarrow \frac{\sigma^2}{\sqrt{2\pi}} \int_0^{\infty} 2u e^{-u} \frac{du}{\sqrt{2u}}$$

$$= \frac{\sigma^2}{\sqrt{2\pi}} \int_0^{\infty} \sqrt{2u} e^{-u} du$$

$\sqrt{2} \times \sqrt{2} = 2$



$$= \frac{2\sigma^2}{\sqrt{\pi}} \int_0^\infty u e^{-u} \frac{du}{u^{3/2}}$$

$$\frac{u}{u^{3/2}} \Rightarrow u^{1-3/2} \Rightarrow u^{-1/2}$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \int_0^\infty u^{1/2} e^{-u} du$$

$$u^{1/2} \Rightarrow u^{3/2-1} \Rightarrow u^{3/2}$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \int_0^\infty u^{3/2-1} e^{-u} du$$

$$\Rightarrow \frac{2\sigma^2}{\sqrt{\pi}} \int_0^\infty e^{-u} u^{3/2-1} du$$

$$\int_0^\infty e^{-u} u^{p-1} du = \frac{\Gamma(p)}{p!} \quad n=1, p=3/2$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \cdot \frac{\Gamma(3/2)}{1^{3/2}}$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \cdot \frac{1}{2} \cdot \frac{\Gamma(1/2)}{1}$$

$$= \frac{2\sigma^2}{\sqrt{\pi}} \cdot \frac{1}{2}$$

$$\frac{1}{2} \Gamma(1/2) = \frac{1}{2} \sqrt{\pi}$$

$$\boxed{V(x) = \sigma^2}$$

$$SD(x) = \sqrt{V(x)}$$

$$= \sqrt{\sigma^2}$$

$$\boxed{SD(x) = \sigma}$$

$\Rightarrow$  odd order moments about mean = 0

$$\boxed{\mu_{x+1} = 0}$$

for  $x = 1, 2, 3, \dots$

$$\mu_{x+1} = E(x-\mu)^{x+1} \cdot f(x) dx$$

$$= \int_{-\infty}^{\infty} (x-\mu)^{x+1} \cdot f(x) dx$$

$$= \int_{-\infty}^{\infty} (x-\mu)^{x+1} \cdot \frac{1}{\sigma \sqrt{\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

follow method

$$\text{put } z = \frac{x-\mu}{\sigma}$$

$$dz = \frac{dx}{\sigma} \Rightarrow dx = \sigma dz$$

$$\sigma dz = dx$$

$$\Rightarrow \frac{1}{\sigma \sqrt{\pi}} \int_{-\infty}^{\infty} (\sigma z)^{x+1} e^{-1/2 z^2} \sigma dz$$

$$= \frac{1}{\sigma \sqrt{\pi}} \int_{-\infty}^{\infty} \sigma^{x+1} \cdot z^{x+1} \cdot e^{-1/2 z^2} \sigma dz$$

$$= \frac{\sigma^{x+1}}{\sqrt{\pi}} \int_{-\infty}^{\infty} z^{x+1} \cdot e^{-1/2 z^2} dz$$

$$= \frac{\sigma^{x+1}}{\sqrt{\pi}} \times 0$$

$$\boxed{\mu_{x+1} = 0}$$

$f(x) = -f(x)$  by the property of odd function  $\int_{-a}^a f(x) dx = 0$



1) Even order central moment of

W.K.M.  $x \rightarrow N(\mu, \sigma)$

$$\mu_{2r} = 1 \cdot 3 \cdot 5 \dots (2r-1) \cdot \sigma^{2r}$$

$$\mu_{21} = E(x-\mu)^{2r}$$

$$= \int_{-\infty}^{\infty} (x-\mu)^{2r} \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$\mu_{21} = 2r-1 \cdot \sigma^2$$

$$\mu_{22} = 2r-1 \cdot 3 \cdot \sigma^4$$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} (x-\mu)^{2r} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} (x-\mu)^{2r} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$\text{put } z = \frac{x-\mu}{\sigma}$$

$$x-\mu = \sigma z$$

$$dx = \sigma dz$$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} (\sigma z)^{2r} \cdot e^{-\frac{1}{2} z^2} \sigma dz$$

$$= \frac{\sigma^{2r}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^{2r} \cdot e^{-\frac{1}{2} z^2} dz$$

$$= \frac{\sigma^{2r}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^{2r} \cdot e^{-\frac{1}{2} z^2} dz$$

$$\int_{-\infty}^{\infty} z^{2r} \cdot e^{-\frac{1}{2} z^2} dz = 2 \int_0^{\infty} z^{2r} \cdot e^{-\frac{1}{2} z^2} dz$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \int_0^{\infty} z^{2r} \cdot e^{-\frac{1}{2} z^2} dz$$

put

$$u = \frac{z^2}{2}$$

$$z^2 = 2u \Rightarrow z = \sqrt{2u}$$

$$2z dz = du$$

$$dz = \frac{du}{2}$$

$$z = \sqrt{2u} \Rightarrow \frac{dz}{du} = \frac{1}{\sqrt{2u}}$$

$$\frac{dz}{du} = \frac{1}{\sqrt{2u}}$$

$$u^{r-1/2} \Rightarrow u^{r-1/2}$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \int_0^{\infty} u^{r-1/2} \cdot e^{-u} du$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \int_0^{\infty} u^{r-1/2} \cdot e^{-u} du$$

$$p = r + 1/2$$

$$m = 1$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \cdot \frac{\Gamma(r+1/2)}{1^{r+1/2}}$$

$$\Gamma(r+1/2) = (r-1/2) \Gamma(r-1/2)$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \cdot (r-1/2) \Gamma(r-1/2) \cdot \sqrt{r+1/2-1}$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \cdot (r-1/2) \Gamma(r-1/2) \cdot \sqrt{r-1/2}$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \cdot (r-1/2) \Gamma(r-1/2) \cdot \sqrt{r-1/2}$$

$$\frac{1}{2} \Gamma(r-1/2)$$

$$= \frac{2 \sigma^{2r}}{\sqrt{2\pi}} \cdot (r-1/2) \Gamma(r-1/2) \cdot \frac{1}{2} \Gamma(r-1/2)$$



$$= \frac{2^x \sigma^{2x} (2x-1) (2x-3) \dots 1}{2^x} \\ = 1 \cdot 3 \cdot 5 \dots (2x-3) (2x-1) \sigma^{2x}$$

Recurrence relation for even order centered moments :-

$$\begin{aligned} \mu_{2x} &= 1 \cdot 3 \cdot 5 \dots (2x-1) \sigma^{2x} \\ \mu_{2x+2} &= 1 \cdot 3 \cdot 5 \dots (2x-1) (2x+1) \sigma^{2x+2} \end{aligned}$$

$$\begin{aligned} \frac{\mu_{2x+2}}{\mu_{2x}} &= \frac{1 \cdot 3 \cdot 5 \dots (2x-1) (2x+1) \sigma^{2x+2}}{1 \cdot 3 \cdot 5 \dots (2x-1) \sigma^{2x}} \\ &= \frac{2x+1 \cdot \cancel{\sigma^{2x}} \cdot \sigma^2}{\cancel{\sigma^{2x}}} \end{aligned}$$

$$\mu_{2x+2} = (2x+1) \cdot \sigma^2$$

$$\boxed{\mu_{2x+2} = (2x+1) \sigma^2 \cdot \mu_{2x}}$$

\$\Rightarrow\$ Moment generating ( ) :- (normal mgf) / (non-normal mgf)

$$\boxed{M_x(t) = e^{t\mu + \frac{1}{2} t^2 \sigma^2}}$$

$$M_x(t) = E(e^{tx}) = \int_{-\infty}^{\infty} e^{tx} \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} e^{tx} \cdot e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2} dx$$

$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t(\mu + \sigma z)} \cdot e^{-\frac{1}{2} z^2} \sigma dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t\sigma z + t\mu} \cdot e^{-\frac{1}{2} z^2} dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t\sigma z} \cdot e^{t\mu} \cdot e^{-\frac{1}{2} z^2} dz$$

$$= \frac{e^{t\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t\sigma z} \cdot e^{-\frac{1}{2} z^2} dz$$

$$= \frac{e^{t\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t\sigma z - \frac{1}{2} z^2} dz$$

$$= \frac{e^{t\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} (z - t\sigma)^2 + \frac{1}{2} t^2 \sigma^2} dz$$

$$= \frac{e^{t\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} (z - t\sigma)^2} \cdot e^{\frac{1}{2} t^2 \sigma^2} dz$$

$$= \frac{1}{\sqrt{2\pi}} \left[ -2\sigma t z + z^2 + t^2 \sigma^2 \right] + \frac{1}{2} t^2 \sigma^2$$

$$= \frac{e^{t\mu + \frac{1}{2} t^2 \sigma^2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} (z - t\sigma)^2} dz$$

$$= \frac{e^{t\mu + \frac{1}{2} t^2 \sigma^2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2} (z - t\sigma)^2} dz$$



$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}u^2} du$$

$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{2\pi}} \int_0^{\infty} e^{-u^2/2} du$$

Let  $v = \frac{u^2}{2}$

$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{2\pi}} \int_0^{\infty} e^{-v} \frac{dv}{\sqrt{v}}$$

$u = \sqrt{2v}$   
 $2v = u^2$   
 $2dv = u du$   
 $dv = u du$   
 $du = \frac{dv}{u}$   
 $du = \frac{dv}{\sqrt{2v}}$

$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{\pi}} \int_0^{\infty} e^{-v} \frac{dv}{\sqrt{v}}$$

$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{\pi}} \int_0^{\infty} v^{-1/2} e^{-v} dv$$

$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{\pi}} \int_0^{\infty} v^{-1/2} e^{-v} dv$$

$$\int_0^{\infty} x^{p-1} e^{-x} dx = \frac{\Gamma(p)}{\Gamma(p)}$$

$p = \frac{1}{2}, m = 1$

$$\boxed{\sqrt{1/2} = \sqrt{\pi}}$$

$$= \frac{e^{tH + \frac{1}{2}t^2\sigma^2}}{\sqrt{\pi}}$$

$$M_n(t) = e^{tH + \frac{1}{2}t^2\sigma^2}$$

Central mgf :-

$$M_x(t) = E(e^{t(x-\mu)})$$

$$= E(e^{tx - t\mu})$$

$$= E(e^{tx} \cdot e^{-t\mu})$$

$$= e^{-t\mu} \cdot E(e^{tx})$$

$$= e^{-t\mu} \cdot e^{tH + \frac{1}{2}t^2\sigma^2}$$

$$= e^{tH + \frac{1}{2}t^2\sigma^2 - t\mu}$$

$$M_n(t) = e^{\frac{1}{2}t^2\sigma^2}$$

Mean deviation of normal distribution :-

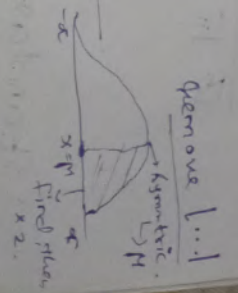
$$MD = E|x - \mu|$$

$$= E|x - \mu|$$

$$= \int_{-\infty}^{\infty} |x - \mu| \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

Curve is symmetric

$$= 2 \int_0^{\infty} (x - \mu) \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$



Put  $z = \frac{x-\mu}{\sigma}$

$\sigma z = x - \mu$

when  $x = \mu$  so  $z = 0$

$x = \mu, z = 0$

(variate)



$$\begin{aligned}
 &= \frac{2}{\sigma \sqrt{2\pi}} \int_0^\infty z e^{-\frac{1}{2} z^2} \sigma dz \\
 &= \frac{2\sigma}{\sigma \sqrt{2\pi}} \int_0^\infty z e^{-\frac{1}{2} z^2} dz \quad \text{put } u = \frac{z^2}{2} \\
 &\quad \text{when } z=0, u=0 \\
 &\quad \text{when } z=\infty, u=\infty \\
 &\quad \frac{du}{dz} = z \Rightarrow du = z dz \\
 &= \frac{2\sigma}{\sigma \sqrt{2\pi}} \int_0^\infty e^{-u} du \\
 &= \frac{2\sigma}{\sigma \sqrt{2\pi}} \left[ -e^{-u} \right]_0^\infty = \frac{2\sigma}{\sigma \sqrt{2\pi}} \left[ 0 - (-1) \right] \\
 &= \frac{2\sigma}{\sigma \sqrt{2\pi}} = \frac{2}{\sqrt{2\pi}} = \frac{\sqrt{2}}{\sqrt{\pi}} = \frac{1}{\sqrt{\pi/2}}
 \end{aligned}$$

$$\begin{aligned}
 &= \frac{2\sigma}{\sigma \sqrt{2\pi}} \left[ 0 + 1 \right] = \frac{2\sigma}{\sigma \sqrt{2\pi}} = \frac{2}{\sqrt{2\pi}} \\
 &= \frac{1}{\sqrt{\pi/2}} = 0.79788 \sigma \approx \frac{4}{5} \sigma
 \end{aligned}$$

⇒ Standard Normal distribution

$x \rightarrow N(\mu, \sigma)$  (For doing probm, we have to do in s.n.d)   
 $(x-\mu) \sim N(0,1)$    
 $z \rightarrow N(0,1)$    
 $f(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$    
 $z = \frac{x-\mu}{\sigma}$    
 $\sigma z = x - \mu$    
 $\sigma dz = dx$

$\therefore f(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} z^2}$    
 $f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} z^2}$    
 $\sigma = \frac{dx}{dz}$