## MTH210: Lab 10

## MM Algorithm

1. Run the code in classDemonstration.R which contains the MM-algorithm implementation for the location Cauchy example. Change the seed at the top, and rerun the code for different starting values.

No solution required for this.

- 2. **Ridge Regression:** Show that the ridge regression estimator in biased for  $\beta$  but has lower variance than the MLE estimator. (No need to use R for this).
- 3. Bridge Regression: In this problem you will write a function to calculate the Bridge Regression estimate using MM algorithm for any given data set.

Recall that the Bridge objective function to minimize is

$$Q_B(\beta) = \frac{(y - X\beta)^T (y - X\beta)}{2} + \frac{\lambda}{2\alpha} \sum_{i=1}^p |\beta_j|^\alpha$$

We found the minorizing function to be

$$\tilde{f}(\beta|\beta_{(k)}) = \text{constants} + \frac{(y - X\beta)^T (y - X\beta)}{2} + \frac{\lambda}{2\alpha} \sum_{i=1}^p m_{j,(k)} \beta_j^2$$

where  $m_{j,(k)} = \alpha |\beta_{j,(k)}|^{\alpha-2}$ . Here  $\beta_j$  and  $\beta_{j,(k)}$  denoted the jth component of  $\beta$  and  $\beta_{(k)}$  respectively. The maximizer of this minorizing function gives us the next iterate of the MM algorithm, which is

$$\beta_{(k+1)} = \left(X^TX + \frac{\lambda}{\alpha}M_{(k)}\right)^{-1}X^Ty\,,$$

where  $M_{(k)}$  is the diagonal matrix of  $m_{j,(k)}$ s.

Now, you will write a function that takes data y, X as input, along with values of  $\lambda$  and  $\alpha$ , and returns the Bridge regression estimate. The function should look like:

```
# y = response
# X = covariate matrix
# lambda = penlaty term
# alpha = bridge term
# max.ter = maximum iterations for the MM algorithm. if max iteration has been reached, the fund
# tol = tolerance level for when to stop MM
bridgeReg <- function(y, X, alpha, lambda, max.iter, tol)</pre>
{
  . . .
  while(... < tol)
    if(iter > max.iter)
      print("Maximum iterations reached")
      stop
    }
    # MM steps
  return(...)
}
```

Below is a version of the function. In the function below I have set the starting value to be the ridge solution as that is available tractably, can be easily calculated, and it expected to be close to the bridge solution.

```
# y = response
# X = covariate matrix
# lambda = penlaty term
# alpha = bridge term
# max.ter = maximum iterations for the MM algorithm. if max iteration has been reached, the fund
# tol = tolerance level for when to stop MM
bridgeReg <- function(y, X, alpha, lambda, max.iter, tol)
{
    # a value larger than tol
    distance <- tol + 1
    iter <- 0
    p <- dim(X)[2]

# starting from the ridge solution
    current <- matrix(1, ncol = 1, nrow = p)
    #(t(X) %*% X + lambda*diag(p)) %*% t(X) %*%y</pre>
```

```
while(distance > tol)
{
   iter <- iter + 1
   if(iter > max.iter)
   {
      print("Maximum iterations reached")
      stop
   }

# MM steps
previous <- current
mjs <- alpha* abs(current)^(alpha - 2)

## using qr.solve since that is more stable than solve
current <- qr.solve(t(X)%*%X + lambda/alpha * diag(mjs) ) %*% t(X) %*% y
   distance <- norm(previous - current, "2")
}
#returning the last iterate of the
return(current)
}</pre>
```

## 4. Simulate the following data:

```
set.seed(1)
n <- 100
p <- 5
beta.star <- c(3, 1, .01, -2, -.003)

# X matrix with intercept
X <- cbind(1, matrix(rnorm(n*(p-1)), nrow = n, ncol = (p-1)))

# generating response
y <- X %*% beta.star + rnorm(n)</pre>
```

Call your bridgeReg() function from the previous question on this dataset for the following values of  $\alpha$  and  $\lambda$ , and the print the estimate. See if the estimate is close to the true value of  $\beta$  from which the data was simulated.

Below is the completed code

```
alpha.vec \leftarrow seq(1, 2, length = 5)
  lambda <- c(.01, .1, 1, 10, 100)
  for(i in 1:length(alpha.vec))
    for(j in 1:length(lambda))
      bridge.est <- bridgeReg(y, X, alpha = alpha.vec[i], lambda = lambda[j], max.iter = 100, tol</pre>
      print(bridge.est)
    }
  }
            [,1]
[1,] 2.94295600
[2,] 1.19323492
[3,] 0.17953850
[4,] -1.96632131
[5,] 0.03960107
            [,1]
[1,] 2.94233976
[2,] 1.19239256
[3,] 0.17862704
[4,] -1.96694066
[5,] 0.03887208
            [,1]
[1,] 2.93626765
[2,] 1.18409239
[3,] 0.16964596
[4,] -1.97304337
[5,] 0.03168906
            [,1]
[1,] 2.88354026
[2,] 1.11201757
[3,] 0.09165849
[4,] -2.02603647
[5,] -0.03068493
           [,1]
[1,] 2.6836096
[2,] 0.8387257
[3,] -0.2040529
[4,] -2.2269747
[5,] -0.2671934
            [,1]
```

- [1,] 2.94293477
- [2,] 1.19320589
- [3,] 0.17950709
- [4,] -1.96634265
- [5,] 0.03957594
  - [,1]
- [1,] 2.9421281
- [2,] 1.1921032
- [3,] 0.1783140
- [4,] -1.9671534
- [5,] 0.0386217
  - [,1]
- [1,] 2.93422146
- [2,] 1.18129539
- [3,] 0.16661951
- [4,] -1.97509987
- [5,] 0.02926852
  - [,1]
- [1,] 2.86865134
- [2,] 1.09166542
- [3,] 0.06963675
- [4,] -2.04100042
- [5,] -0.04829781
  - [,1]
- [1,] 2.6619496
- [2,] 0.8091180
- [3,] -0.2360894
- [4,] -2.2487438
- [5,] -0.2928161
  - [,1]
- [1,] 2.94290695
- [2,] 1.19316787
- [3,] 0.17946595
- [4,] -1.96637061
- [5,] 0.03954304
  - [,1]
- [1,] 2.94185125
- [2,] 1.19172480
- [3,] 0.17790449
- [4,] -1.96743163
- [5,] 0.03829419
  - [,1]
- [1,] 2.93157660

- [2,] 1.17768004
- [3,] 0.16270758
- [4,] -1.97775806
- [5,] 0.02613978 [,1]
- [1,] 2.85120433
- [2,] 1.06781652
- [3,] 0.04383141
- [4,] -2.05853535
- [5,] -0.06893679
  - [,1]
- [1,] 2.6427868
- [2,] 0.7829236
- [3,] -0.2644326
- [4,] -2.2680032
- [5,] -0.3154848
  - [,1]
- [1,] 2.94287053
- [2,] 1.19311809
- [3,] 0.17941208
- [4,] -1.96640721
- [5,] 0.03949996
  - [,1]
- [1,] 2.94148929
- [2,] 1.19123002
- [3,] 0.17736913
- [4,] -1.96779542
- [5,] 0.03786601 [,1]
  - \_,\_\_
- [1,] 2.92817312
- [2,] 1.17302770
- [3,] 0.15767358
- [4,] -1.98117869
- [5,] 0.02211362
  - [,1]
- [1,] 2.83131566
- [2,] 1.04063003
- [3,] 0.01441468
- [4,] -2.07852425
- [5,] -0.09246415
  - [,1]
- [1,] 2.6261654
- [2,] 0.7602033

```
[3,] -0.2890168
[4,] -2.2847083
[5,] -0.3351471
            [,1]
[1,]
     2.94282284
[2,] 1.19305289
[3,] 0.17934154
[4,] -1.96645515
[5,]
     0.03944354
            [,1]
[1,]
     2.94101636
[2,]
     1.19058356
[3,] 0.17666963
[4,] -1.96827073
[5,] 0.03730656
            [,1]
[1,]
     2.92381895
[2,] 1.16707585
[3,] 0.15123347
[4,] -1.98555480
[5,]
     0.01696285
            [,1]
[1,] 2.80932708
[2,] 1.01057312
[3,] -0.01810795
[4,] -2.10062363
[5,] -0.11847558
           [,1]
[1,] 2.6119778
[2,] 0.7408099
[3,] -0.3100012
[4,] -2.2989674
[5,] -0.3519304
```

When the above code runs, the estimate of  $\beta$  is printed for each value of  $\alpha$  and  $\lambda$ . You can see that for larger values of  $\lambda$ , the larger estimates are pushed closer to 0.

5. In the previous question, suppose we fix  $\lambda=10$  and for different values of  $\alpha$ , we obtain a different estimation of bridge regression coefficient:  $\hat{\beta}_{\alpha,\lambda}$ . The distance of this estimate from the true  $\beta$  can be measured with  $\|\hat{\beta}_{\alpha,\lambda}-\beta\|$ .

```
for(i in 1:length(alpha.vec))
{
    bridge.est <- bridgeReg(.....)</pre>
```

```
norm(bridge.est - beta.star, "2")
}
```

This is a random quantity (since  $\hat{\beta}_{\alpha,\lambda}$  is random). Suppose I want to find the average of expected distance of  $\hat{\beta}_{\alpha,\lambda}$  from  $\beta$ .

$$E\left[\|\hat{\beta}_{\alpha,\lambda}-\beta\|\right]$$

The expectation is with respect to the distribution of  $\hat{\beta}_{\alpha,\lambda}$ , which we don't know. But since this is a simulated data setup, we can simulate multiple such datasets (keeping the same value of  $\beta$ ), and obtain different realizations of  $\hat{\beta}_{\alpha,\lambda}$ . Taking an average of all the resulting distances will give us an estimate of the above expectation. Implement this exercise and see which value of  $\alpha$  gives the lowest expected distance. Below is code that might help you.

```
alpha.vec \leftarrow seq(1, 2, length = 5)
reps <- 100
dist <- matrix(0, nrow = reps, ncol = length(alpha.vec))</pre>
for(r in 1:reps)
{
  # generate X and y again here
  # X matrix with intercept
    X \leftarrow cbind(1, matrix(rnorm(n*(p-1)), nrow = n, ncol = (p-1)))
    # generating response
    y <- X %*% beta.star + rnorm(n)
  for(i in 1:length(alpha.vec))
  {
    bridge.est <- bridgeReg(y, X, alpha = alpha.vec[i], lambda = 10, max.iter = 100, tol = 1e-5)
    dist[r, i] <- norm(bridge.est - beta.star, "2")</pre>
  }
}
```

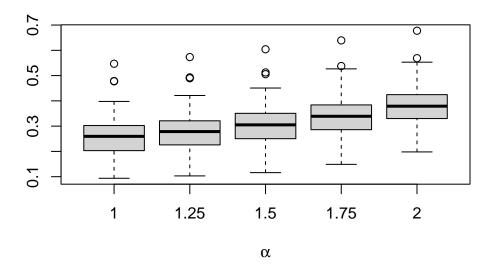
Now **dist** has the distance from the true  $\beta$  for every value of  $\alpha$  in the each column, and replications are in rows. Thus to estimate the expectation, we can take column means:

```
# naming the columns of dist
colnames(dist) <- alpha.vec
# expectation
colMeans(dist)

1  1.25  1.5  1.75  2
0.2590982 0.2781903 0.3047150 0.3390168 0.3803636</pre>
```

We can also plot the column-wise boxplot:

boxplot(dist, xlab = expression(alpha))



From the looks of it,  $\alpha=1$  (Lasso) has the least expected error.