

MTH211A: Theory of Statistics

Quiz 4

Name: _____

Time: 25 minutes

Roll number: _____

Total marks: ~~4+5+1+4~~ 4+6 = 10

Q.1 Based on a random sample X of size one from $\text{Cauchy}(\theta, 1)$ distribution, find the shortest length 5% confidence interval for θ . [4]

Let $X \sim \text{Cauchy}(\theta, 1)$ distribution. Then $f_X(x) = \frac{1}{\pi(1+(x-\theta)^2)}$; $x \in \mathbb{R}$.

Define $T(X, \theta) = X - \theta$ as the pivot.

Clearly, $T(X, \theta) \sim \text{Cauchy}(0, 1)$ distribution, i.e., $f_T(t) = \frac{1}{\pi(1+t^2)}$; $t \in \mathbb{R}$.

Let c_1, c_2 be such that $\int_{c_1}^{c_2} \frac{1}{\pi(1+t^2)} dt = 0.05$. (*)

Then $P(c_1 < T(X, \theta) < c_2) = 0.05$

$\Rightarrow P(c_1 < X - \theta < c_2) = P(X - c_2 < \theta < X - c_1) = 0.05$.

Thus, $(\underbrace{X - c_2}_{L(X)}, \underbrace{X - c_1}_{U(X)})$ is a 5% confidence interval.

Length of the interval, $U(X) - L(X) = c_2 - c_1$ is lowest if

$$(i) \quad \frac{1}{\pi(1+c_1^2)} = \frac{1}{\pi(1+c_2^2)} \Leftrightarrow c_1^2 = c_2^2$$

and (ii) $c_1 < \text{mode of } f_T(t) < c_2$; and mode $f_T(t)$ is 0.

Thus, $c_1 = -c_2$.

Further, from (*) we get: $\int_{-c_2}^{c_2} \frac{dt}{\pi(1+t^2)} = \frac{\tan^{-1} t}{\pi} \Big|_{-c_2}^{c_2} = 0.05$

$$\Rightarrow 2 \tan^{-1} c_2 = \pi \times 0.05 \Rightarrow \tan^{-1} c_2 = \frac{\pi}{2} \times 0.05$$

$$\Rightarrow c_2 = \tan\left(\frac{\pi}{2} \times 0.05\right)$$

Thus, $(X - \tan(\frac{\pi}{2} \times 0.05), X + \tan(\frac{\pi}{2} \times 0.05))$ is the shortest length C.I.

Q.2 The lifetime of an electric bulb, denoted as X , is assumed to follow an exponential distribution with parameter λ . However, in practice, the lifetime is rounded to the nearest number of hours. Given a data point indicating that a bulb's lifetime is y_0 hours (where y_0 is a positive integer greater than 0), find the maximum likelihood estimate for λ . 6

Let Y denote the random variable indicating integer valued lifetime (in nearest hours).

PMF of Y :

$$P(Y=0) = \int_0^{0.5} \lambda e^{-\lambda x} dx = \left[-e^{-\lambda x} \right]_0^{0.5} = 1 - e^{-\lambda/2}$$

$$P(Y=y) = \int_{y-1/2}^{y+1/2} \lambda e^{-\lambda x} dx = \left[-e^{-\lambda x} \right]_{y-1/2}^{y+1/2} = e^{-\lambda(y-1/2)} - e^{-\lambda(y+1/2)}$$

$$= e^{-\lambda y} \left[e^{\lambda/2} - e^{-\lambda/2} \right]$$

$$= e^{-\lambda y} \left(\frac{e^{\lambda} - 1}{e^{\lambda/2}} \right)$$

for $y = 1, 2, \dots$

Log Likelihood of λ given $y_0 > 0$:

$$l(\lambda) = -\lambda y_0 - \frac{\lambda}{2} + \log(e^{\lambda} - 1)$$

$$\frac{\partial l(\lambda)}{\partial \lambda} = -y_0 - \frac{1}{2} + \frac{e^{\lambda}}{(e^{\lambda} - 1)} \Rightarrow$$

First order condition (FOC) gives

$$\frac{e^{\lambda}}{e^{\lambda} - 1} = \underbrace{(y_0 + 1/2)}_{c_0, \text{ say}}$$

$$\Rightarrow e^{\lambda} = c_0 e^{\lambda} - c_0$$

$$\Rightarrow (c_0 - 1) e^{\lambda} = c_0$$

$$\Rightarrow e^{\lambda} = \frac{c_0}{c_0 - 1}$$

$$\Rightarrow \lambda = \log c_0 - \log(c_0 - 1)$$

$\therefore \hat{\lambda} = [\log c_0 - \log(c_0 - 1)]$ is a critical point.

Second order condition:

$$\frac{\partial^2 l(\lambda)}{\partial \lambda^2} = \frac{(e^{\lambda} - 1)e^{\lambda} - e^{\lambda}e^{\lambda}}{(e^{\lambda} - 1)^2}$$

$$= -\frac{e^{\lambda}}{(e^{\lambda} - 1)^2} < 0 \text{ for any } \lambda > 0$$

$$\therefore \hat{\lambda}_{MLE} = \log \frac{c_0}{c_0 - 1}$$

$$= \log \left(\frac{y_0 + 1/2}{y_0 - 1/2} \right)$$