## Indian Institute of Technology Bombay

## SI 422: Regression Analysis

## Formula sheet for Problem Set 1

## January 22, 2020

Data: We observe a paired random sample  $\{(x_i, y_i) : i = 1, 2, ..., n\}$  of size n on the response variable y and the predictor variable x.

Simple linear regression model and assumptions:  $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i \ \forall i = 1, 2, ..., n$  where  $\varepsilon_i \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, \sigma^2)$  is random error corresponding to the *i*-th observation and  $\beta_0, \beta_1$  are the unknown intercept and slope parameter of the model. Additionally, we assume that the predictor variable x is non-stochastic.

- 1.  $\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i, \ \bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i, \ S_{xy} = \sum_{i=1}^{n} (y_i \bar{y})(x_i \bar{x}_i) = \sum_{i=1}^{n} y_i x_i n\bar{y}\bar{x},$  $S_{xx} = \sum_{i=1}^{n} (x_i - \bar{x}_i)^2 = \sum_{i=1}^{n} x_i^2 - n\bar{x}^2$
- 2.  $t_{\nu}$  is a t-distributed random variable or t-distribution with degrees of freedom (DF)  $\nu$ .
- 3.  $t'_{\nu,\delta}$  is a non-central t-distribution with DF  $\nu$  and non-centrality parameter  $\delta$ .
- 4.  $(t_{\nu})_{\alpha}$  is the upper  $\alpha$ -point or  $100(1-\alpha)$ -th percentile of the  $t_{\nu}$  distribution.
- 5.  $F_{\nu,\eta}$  is a F-distributed random variable or F-distribution with DF  $\nu, \eta$ .
- 6.  $F'_{\nu,\eta,\lambda}$  is a non-central F-distribution with DF  $\nu,\eta$  and non-centrality parameter  $\lambda$ .
- 7.  $(F_{\nu,\eta})_{\alpha}$  is the upper  $\alpha$ -point or  $100(1-\alpha)$ -th percentile of the  $F_{\nu,\eta}$  distribution.
- 8. Let X be a random variable and  $f(\cdot)$  be a measurable function. Then  $f(X)_{\text{observed}}$  is the observed value of f(X).
- 9. LSE of the slope parameter  $\beta_1$  and the intercept parameter  $\beta_0$  are respectively  $\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}}$  and  $\hat{\beta}_0 = \bar{y} \hat{\beta}_1 \bar{x}$ .
- 10. Fitted simple linear regression model is  $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$ .
- 11. Fitted values are  $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$  for  $i = 1, 2, \dots, n$ .
- 12. Residuals due to fit are  $e_i = y_i \hat{y}_i$  for i = 1, 2, ..., n.
- 13. Total sum of squares  $TSS = \sum_{i=1}^{n} (y_i \bar{y})^2$  measures total variation of y in the observed data.

- 14. Sum of squares due to regression fit  $SS_R = \sum_{i=1}^n (\hat{y}_i \bar{y})^2$  measures the variation of y explained by the underlying regression model.
- 15. Error sum of squares or residual sum of squares  $SSE = \sum_{i=1}^{n} e_i^2 = \sum_{i=1}^{n} (\hat{y}_i y_i)^2$  measures the variation of y which cannot be explained by the underlying regression model.
- 16. Fundamental ANOVA identity for regression model is  $TSS = SS_R + SSE$ .
- 17. Coefficient of determination  $R^2 = \frac{SS_R}{TSS} = 1 \frac{SSE}{TSS}$  is the proportion of variation of y explained by the underlying regression model. Note that  $0 \le R^2 \le 1$ .
- 18. ANOVA table

| Source of variation | SS     | DF  | MS              | F stat                   | p-value                                  |
|---------------------|--------|-----|-----------------|--------------------------|--|
| Regression          | $SS_R$ | 1   | $MS_R = SS_R/1$ | $F_0 = \frac{MS_R}{MSE}$ | $P(F_{1,n-2} > (F_0)_{\text{observed}})$ |
| Error               | SSE    | n-2 | MSE = SSE/(n-2) |                          |  |
| Total               | TSS    | n-1 |                 |                          |  |

19. To test  $H_0: \beta_0 = \beta_{00}$  against  $H_1: \beta_0 \neq \beta_{00}$ , we use the test statistic

$$T_0 = \frac{\hat{\beta}_0 - \beta_{00}}{\sqrt{MSE\left(\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}\right)}}$$

which follows  $t_{n-2}$  and  $t'_{n-2,\delta_0}$  with  $\delta_0 = \frac{\beta_0 - \beta_{00}}{\sigma \sqrt{\frac{1}{n} + \frac{\vec{x}^2}{S_{xx}}}}$  under  $H_0$  and  $H_1$  respectively. We reject  $H_0$  at  $100\alpha\%$  level of significance if

$$|T_0|_{\text{observed}} > (t_{n-2})_{\alpha/2}$$
 or equivalently  $P(|t_{n-2}| > |T_0|_{\text{observed}}) < \alpha$ .

20. To test  $H_0: \beta_1 = \beta_{10}$  against  $H_1: \beta_1 \neq \beta_{10}$ , we use the test statistic

$$T_1 = \frac{\hat{\beta}_1 - \beta_{10}}{\sqrt{\frac{MSE}{S_{xx}}}}$$

which follows  $t_{n-2}$  and  $t'_{n-2,\delta_1}$  with  $\delta_1 = \frac{\beta_1 - \beta_{10}}{\sigma \sqrt{\frac{1}{S_{xx}}}}$  under  $H_0$  and  $H_1$  respectively. We reject  $H_0$  at  $100\alpha\%$  level of significance if

$$|T_1|_{\text{observed}} > (t_{n-2})_{\alpha/2}$$
 or equivalently  $P(|t_{n-2}| > |T_1|_{\text{observed}}) < \alpha$ .

21. A  $100(1-\alpha)\%$  confidence interval of  $\beta_0$  is

$$\left(\hat{\beta}_0 - (t_{n-2})_{\alpha/2} \sqrt{MSE\left(\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}\right)}, \hat{\beta}_0 + (t_{n-2})_{\alpha/2} \sqrt{MSE\left(\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}\right)}\right).$$

22. A  $100(1-\alpha)\%$  confidence interval of  $\beta_1$  is

$$\left(\hat{\beta}_1 - (t_{n-2})_{\alpha/2} \sqrt{\frac{MSE}{S_{xx}}}, \hat{\beta}_1 + (t_{n-2})_{\alpha/2} \sqrt{\frac{MSE}{S_{xx}}}\right).$$

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23. Mean of Y when  $x = x_0$  is  $E(Y|x_0) = \beta_0 + \beta_1 x_0$  and its unbiased estimator is  $\widehat{E(Y|x_0)} = \hat{\beta}_0 + \hat{\beta}_1 x_0$ . Moreover,

$$\widehat{E(Y|x_0)} \sim \mathcal{N}\left(E(Y|x_0), \sigma^2\left(\frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}\right)\right).$$

Thus a  $100(1-\alpha)\%$  confidence interval of  $E(Y|x_0)$  is

$$\left(\widehat{E(Y|x_0)} - (t_{n-2})_{\alpha/2} \sqrt{MSE\left(\frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}\right)}, \widehat{E(Y|x_0)} + (t_{n-2})_{\alpha/2} \sqrt{MSE\left(\frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}\right)}\right).$$

24. Let  $y_0$  be the observed value of Y when  $x = x_0$ . The fitted value of Y at  $x = x_0$  is  $\hat{y}_0 = \hat{\beta}_0 + \hat{\beta}_1 x_0$ . Moreover, we have

$$y_0 - \hat{y}_0 \sim \mathcal{N}\left(0, \sigma^2\left(1 + \frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}\right)\right).$$

Thus a  $100(1-\alpha)\%$  prediction interval of  $y_0$  is

$$\left(\hat{y}_0 - (t_{n-2})_{\alpha/2} \sqrt{MSE\left(1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}\right)}, \hat{y}_0 + (t_{n-2})_{\alpha/2} \sqrt{MSE\left(1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{S_{xx}}\right)}\right).$$

25. Consider an equidistant partition of  $(x_{\min}, x_{\max})$  of length k (say). Suppose it is  $x_{new} = ((x_{new})_1, (x_{new})_2, \dots, (x_{new})_k)$ . Then a  $100(1-\alpha)\%$  confidence band is

$$\left(E(\widehat{Y|(x_{new})_j}) - (t_{n-2})_{\alpha/2} \sqrt{MSE\left(\frac{1}{n} + \frac{((x_{new})_j - \bar{x})^2}{S_{xx}}\right)}, E(\widehat{Y|(x_{new})_j}) + (t_{n-2})_{\alpha/2} \sqrt{MSE\left(\frac{1}{n} + \frac{((x_{new})_j) - \bar{x})^2}{S_{xx}}\right)}\right), \quad j = 1, 2, \dots, k.$$

26. A  $100(1-\alpha)\%$  prediction band is

$$\left(E(\widehat{Y|(x_{new})_j}) - (t_{n-2})_{\alpha/2} \sqrt{MSE\left(1 + \frac{1}{n} + \frac{((x_{new})_j - \bar{x})^2}{S_{xx}}\right)}, E(\widehat{Y|(x_{new})_j}) + (t_{n-2})_{\alpha/2} \sqrt{MSE\left(1 + \frac{1}{n} + \frac{((x_{new})_j) - \bar{x})^2}{S_{xx}}\right)}\right), \quad j = 1, 2, \dots, k.$$