

Problem Set 6 Solutions

May 7, 2018

1. Associate each node $i \in V$ with a random variable x_i , such that $x_i = 1$ if $i \in A$, otherwise $x_i = 0$.

Define singleton potential $\phi(x_i) = e^{\theta_i x_i}$, edge potential $\psi(x_i, x_j) = \mathbb{1}\{x_i + x_j > 0\}$, then the MRF over the graph is:

$$P(x) = \frac{1}{Z} \prod_{i \in V} \phi(x_i) \prod_{(i,j) \in E} \psi(x_i, x_j) = \frac{1}{Z} \prod_{i \in V} e^{\theta_i x_i} \prod_{(i,j) \in E} \mathbb{1}\{x_i + x_j > 0\}$$

It's easy to see that for any valid vertex cover A , we have $P(A) \propto w(A)$.

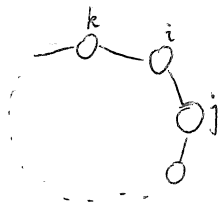
$$2. \arg \min_{A \subseteq V} P(A) = \arg \min_{x: P(x) > 0} P(x) = \arg \min_{x: P(x) > 0} \prod_{i \in V} e^{\theta_i x_i} = \arg \min_{x: P(x) > 0} \sum_{i \in V} \theta_i x_i = \arg \max_{x: P(x) > 0} \sum_{i \in V} (-\theta_i) x_i$$

So we can ~~redefine~~ negate the weights ($\theta_i := -\theta_i$) and simply run max-product. Alternatively, we can directly apply min-product on the original model, where the MIN operation is used (instead of MAX or SUM).

3. The sum-product algorithm attempts to find a fixed point m^* of the update equations:

$$m_{ij}(x_j) \propto \sum_{x_i} \phi_i(x_i) \psi_{ij}(x_i, x_j) \prod_{k \in N_b(i) \setminus j} m_{ki}(x_i)$$

Consider a K -cycle graph and the message from i to j , with k being the other neighbor of i :



$$m_{ij}(x_j) \propto \sum_{x_i} \phi_i(x_i) \psi_{ij}(x_i, x_j) m_{ki}(x_i)$$

Plugging in $\phi_i(x_i) = e^{x_i}$, $\psi_{ij}(x_i, x_j) = \mathbb{1}\{x_i + x_j > 0\}$ gives

$$m_{ij}(1) \propto m_{ki}(0) + e m_{ki}(1), \quad m_{ij}(0) \propto e m_{ki}(1)$$

$$\text{So } m_{ij}(1) = \frac{m_{ki}(0) + e m_{ki}(1)}{2e m_{ki}(1) + m_{ki}(0)}, \quad m_{ij}(0) = \frac{e m_{ki}(1)}{2e m_{ki}(1) + m_{ki}(0)}$$

At convergence, a fixed point m^* should satisfy $m_{ij}^* = m_{ki}^*$ due to symmetry in the graph,

thus

$$\begin{cases} m_{ki}^*(1) = \frac{m_{ki}^*(0) + e m_{ki}^*(1)}{2e m_{ki}^*(1) + m_{ki}^*(0)} \\ m_{ki}^*(0) = \frac{e m_{ki}^*(1)}{2e m_{ki}^*(1) + m_{ki}^*(0)} \\ m_{ki}^*(1) \geq 0 \\ m_{ki}^*(0) \geq 0 \end{cases}$$

Solving the system of equations gives the converged messages (identical for all neighboring nodes $k, i \in V$):

$$m_{ki}^*(0) = \frac{3e - \sqrt{e(4+e)}}{4e-2} \approx 0.437$$

$$m_{ki}^*(1) = \frac{e - 2 + \sqrt{e(4+e)}}{4e-2} \approx 0.563$$

The answer is independent of K , the size of the cycle.

Once we have m^* , we can also compute the converged beliefs easily:

$$b_i^*(x_i) = \gamma_i \phi_i(x_i) \prod_{k \in N_b(i)} m_{ki}^*(x_i)$$

$$b_{ij}^*(x_i, x_j) = \gamma_{ij} \psi_{ij}(x_i, x_j) \phi_i(x_i) \phi_j(x_j) \prod_{k \in N_b(i) \setminus j} m_{ki}^*(x_i) \prod_{k' \in N_b(j) \setminus i} m_{k'j}^*(x_j)$$

4. In Gibbs sampling, we sample x_i from $P(x_i | x_{\setminus i})$ (where $x_{\setminus i}$ is the configuration of all the other variables sampled in previous iterations) with the following probability: (see Koller PBM text section 12.3.3 for more details)

$$P(x_i | x_{\setminus i}) \underset{\substack{\text{by local Markov} \\ \text{property}}}{=} P(x_i | x_{nb(i)}) = \frac{\phi_i(x_i) \prod_{j \in nb(i)} \psi_{ij}(x_i, x_j)}{\sum_{x_i'} \phi_i(x_i') \prod_{j \in nb(i)} \psi_{ij}(x_i', x_j)}$$

where $\phi_i(x_i) = e^{\theta_i x_i}$, $\psi_{ij}(x_i, x_j) = \mathbb{1}\{x_i + x_j > 0\}$. To sample from $P(x)$, we simply iteratively sample each $x_i, i \in V$ from $P(x_i | x_{\setminus i})$ while holding the other variables fixed.

5. $\ell(x^{(1)}, \dots, x^{(m)} | \theta) = \frac{1}{m} \sum_m \log P_\theta(x^{(m)}) = \frac{1}{m} \sum_m \left(\sum_{i \in V} \theta_i x_i^{(m)} + \sum_{(i,j) \in E} \log \mathbb{1}\{x_i^{(m)} + x_j^{(m)} > 0\} \right) - \log Z$

$$\frac{\partial \ell}{\partial \theta_i} = \frac{1}{m} \sum_m x_i^{(m)} - \frac{\mathbb{E}_{P_\theta(x_i=1)} [x_i]}{P_\theta(x_i=1)} = \frac{\mathbb{E}_P[x_i] - \mathbb{E}_\theta[x_i]}{P_\theta(x_i=1)} \quad (\text{see Koller text section 20.3.1})$$

6. $\ell_{PL}(x^{(1)}, \dots, x^{(m)} | \theta) = \frac{1}{m} \sum_m \sum_i \log P(x_i^{(m)} | x_{\setminus i}^{(m)}, \theta)$

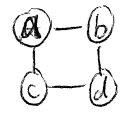
We have derived each conditional probability $P(x_i^{(m)} | x_{\setminus i}^{(m)})$ in problem 4. The general form of the derivative is given in section 20.6.1 of Koller text, equation (20.23).

$$\frac{\partial \ell_{PL}}{\partial \theta_i} = \sum_{j: X_j \in \text{scope}[f_i]} \left(\frac{1}{m} \sum_m f_i(x^{(m)}) - \mathbb{E}_{x_j' \sim P_\theta(x_j | x_{\setminus j}^{(m)})} [f_i(x_j', x_{\setminus j}^{(m)})] \right)$$

In our specific problem, $f_i(x_i) = x_i$, $\text{scope}[f_i] = \{x_i\}$, (i.e., the factor f_i associated with θ_i involves x_i only), so

$$\frac{\partial \ell_{PL}}{\partial \theta_i} = \frac{1}{m} \sum_m f_i(x_i^{(m)}) - \mathbb{E}_{x_i' \sim P_\theta(x_i | x_{\setminus i}^{(m)})} [f_i(x_i')] = \frac{1}{m} \sum_m x_i^{(m)} - P_\theta(x_i=1 | x_{\setminus i}^{(m)})$$

7. Clearly, any graph structure (particularly, the edge set) inconsistent with the data yields a log likelihood of $-\infty$, so we only need to consider graphs consistent with data. The observations $\{a, d\}, \{b, c\}$ imply $\{b, c\} \notin E$ and $\{a, d\} \notin E$, respectively. To achieve maximum likelihood, we therefore pick the most complex model under these constraints, i.e., $E = \{(a, b), (b, d), (a, c), (c, d)\}$ (see Koller text section 20.7.3.1 for justification).



The set of valid configurations under this model are $\{a, d\}, \{b, c\}, \{a, b, c\}, \{b, c, d\}, \{a, b, d\}, \{a, c, d\}, \{a, b, c, d\}$.

So $Z = \sum_x \prod_{i \in V} \phi(x_i) \prod_{(i,j) \in E} \psi_{ij}(x_i, x_j) = e^{\theta_a + \theta_d} + e^{\theta_b + \theta_c} + e^{\theta_a + \theta_b + \theta_c} + e^{\theta_b + \theta_c + \theta_d} + e^{\theta_a + \theta_c + \theta_d} + e^{\theta_a + \theta_b + \theta_d} + e^{\theta_a + \theta_b + \theta_c + \theta_d}$

Given $M=3$ samples $\{a, d\}, \{a, d\}, \{b, c\}$, with ℓ_2 regularizer $-\frac{\lambda}{2} \|\theta\|_2^2, \lambda=100$

$$\frac{\partial \ell}{\partial \theta_a} = \frac{1}{m} \sum_m x_a^{(m)} - \frac{\partial}{\partial \theta_a} \log Z - \lambda \theta_a = \frac{2}{3} - \frac{1}{Z} (e^{\theta_a + \theta_d} + e^{\theta_a + \theta_b + \theta_c} + e^{\theta_a + \theta_c + \theta_d} + e^{\theta_a + \theta_b + \theta_d}) - 100 \theta_a$$

Similarly the derivatives w.r.t. θ_b, θ_c , and θ_d can be obtained. Running gradient ascent gives $\theta^* = [-4.8 \times 10^{-4}, -3.8 \times 10^{-3}, -3.8 \times 10^{-3}, -4.8 \times 10^{-4}]$ with optimal log likelihood $\ell^* = -1.95$.

② Given a single sample $\{a\}$, the log-likelihood is $\ell = \log P(x_a=1, x_b=x_c=x_d=0)$

i) if the edge set is empty, then the MRF distribution factorizes over each variable, with $P(x_a) = \frac{e^{\theta_a x_a}}{(e^{\theta_a} + 1)}$

$$\frac{\partial \ell}{\partial \theta_a} = 1 - \mathbb{E}_\theta[x_a] = 1 - P_\theta(x_a=1) = 1 - \frac{e^{\theta_a}}{(e^{\theta_a} + 1)}$$

Gradient ascent yields $\theta_a \rightarrow \infty$

ii) if the edge set is not empty, with the sample $\{a\}$, i.e., $E = \{(a, b), (a, c), (a, d)\}$

Like before, to maximize the likelihood, we choose the largest edge set compatible

7. (2) ^{contd} The valid configurations are $\{a\}$, $\{a, b\}$, $\{a, c\}$, $\{a, d\}$, $\{a, b, c\}$, $\{a, b, d\}$, $\{a, c, d\}$, $\{a, b, c, d\}$, $\{b, c, d\}$

(a) — (b)
| |
(c) — (d)

$$\text{So } P_\theta(X_a=1) = \frac{\partial}{\partial \theta_a} \log Z = \frac{\frac{\partial}{\partial \theta_a} Z}{Z} = \frac{e^{\theta_a} + e^{\theta_a + \theta_b} + e^{\theta_a + \theta_c} + e^{\theta_a + \theta_d} + e^{\theta_a + \theta_b + \theta_c} + e^{\theta_a + \theta_c + \theta_d} + e^{\theta_a + \theta_b + \theta_c + \theta_d}}{e^{\theta_a} + e^{\theta_a + \theta_b} + e^{\theta_a + \theta_c} + e^{\theta_a + \theta_d} + e^{\theta_a + \theta_b + \theta_c} + e^{\theta_a + \theta_c + \theta_d} + e^{\theta_a + \theta_b + \theta_c + \theta_d} + e^{\theta_b + \theta_c + \theta_d}}$$

To find the MLE θ_a^* , the moment matching condition

$$\frac{\partial \ell}{\partial \theta_a} = 1 - E_\theta[X_a] = 1 - P_\theta(X_a=1) = 0 \quad \text{would require } \theta_a^* \rightarrow \infty$$