

## Lecture 30: Chi-Square, Student-t, Multivariate Normal

### Chi-Square distribution (Univariate distribution)

$\chi^2(n)$  (Chi-square) ( $n$  is degree of freedom)

Let  $V = Z_1^2 + Z_2^2 + \dots + Z_n^2$ ,  $Z_i$  i.i.d.  $N(0,1)$

Then (by defn.)  $V \sim \chi^2(n)$ .

How many  
normals  
squared did  
we add up?

Fact:  $\chi^2(1)$  is Gamma( $\frac{1}{2}, \frac{1}{2}$ ).

So,  $\chi^2(n)$  is Gamma( $\frac{n}{2}, \frac{1}{2}$ ).

### Student-t distribution (Univariate distribution) (Gosset, 1908)

Let  $T = \frac{Z}{\sqrt{V/n}}$ , where  $Z \sim N(0,1)$   
 $V \sim \chi^2(n)$  indep.  $Z, V$ .

Then  $T \sim t_n$  ( $n$  is degree of freedom)

properties:

- ① Symmetric distribution,  
i.e.,  $-T \sim t_n$ .
- ②  $n=1 \Rightarrow$  Cauchy, mean doesn't exist.
- ③  $n \geq 2 \Rightarrow E(T) = E(Z) E\left(\frac{1}{\sqrt{V/n}}\right)$

$$= 0$$

$$E(Z^2) = 1, E(Z^4) = 3, E(Z^6) = 5, \dots$$

Use MGFs.

Another way:

$$E(Z^{2n}) = E((Z^2)^n)$$

$$\neq \chi_1^2 \sim \text{Gamma}(1/2, 1/2)$$

After finding  $n$ th moment of Gamma  $(1/2, 1/2)$

↓ then,

(Use LOTUS)

- ④ Heavier-tailed, i.e., extreme values are relatively more likely than they would be for the normal. That's especially true if  $n$  is small, like in the Cauchy has very heavy tails. Normal is decaying much much faster than a Cauchy.

- ⑤ For  $n$  large (like 30, 40, 50 or larger),  $t_n$  looks very much like  $N(0, 1)$ .  
distribution of  $t_n$  (either CDF or PDF) goes to  $N(0, 1)$  as  $n \rightarrow \infty$ .

Let  $T_n = \frac{Z}{\sqrt{V_n/n}}$ , with  $Z_1, Z_2, \dots$  i.i.d.  $N(0, 1)$

$$V_n = Z_1^2 + \dots + Z_n^2$$

$Z \sim N(0, 1)$  indep. of  $Z_j$ 's.

$\xrightarrow{n \rightarrow \infty} \frac{V_n}{n} \rightarrow 1$  with prob. 1 by Law of Large numbers since  $E Z_1^2 = 1$  (variance of  $Z_1$ )

$$\sqrt{\frac{V_n}{n}} \rightarrow 1$$

with prob. 1

(LLNs says that the average of the same converges to the true theoretical average,

So,  $T_n \rightarrow Z$  with prob. 1. which is just the expected value of one of  $\{Z_1^2, Z_2^2, \dots, Z_n^2\}$

Hence,  $t_n$  Converges to  $N(0, 1)$  distribution.  
distribution



## Multivariate Normal (MVN)

Defn. Random vector  $(X_1, X_2, \dots, X_k) = \vec{X}$  is

Multivariate Normal if every linear combination  $t_1 X_1 + t_2 X_2 + \dots + t_k X_k$  is Normal.

Example Let  $Z, W$  be i.i.d.  $N(0,1)$ . Then  $(\underline{Z+2W}, 3Z+5W)$  is MVN, since

$$\begin{aligned} & s(Z+2W) + t(3Z+5W), \quad s, t \text{ constants} \\ &= \underbrace{(s+3t)Z + (2s+5t)W}_{\text{is Normal.}} \end{aligned}$$

Sum of two indep. normals is normal.

Example  
(Not a Multivariate Normal)

Let  $Z \sim N(0,1)$ , let  $S$  be random sign, indep. of  $Z$ .

Then  $Z, SZ$  are marginally,  $N(0,1)$

(look at  $Z$  on its own, look at  $SZ$  on its own)

But,  $(Z, SZ)$  pair is not MVN.

Look at  $Z + SZ$  (linear combination)

Half the time,  $S$  is  $-1$ , we get 0. The other half the time, we get some continuous thing. So,  $(Z + SZ)$  is actually a mixture of discrete and continuous. We will never find a normal distribution that equals 0 with prob. equal to  $1/2$ . That's not a property of the normal distribution.

MGF of  $\vec{X}$  (MVN) is  $E(e^{t' \vec{X}})$   $\rightarrow$  dot vector product of vector of  $t$  and  $\vec{X}$

Let  $E_{X_j} = \mu_j$

$$= E(e^{t_1 X_1 + \dots + t_k X_k})$$

$$= e^{((t_1 \mu_1 + \dots + t_k \mu_k) + \frac{1}{2} \text{Var}(t_1 X_1 + \dots + t_k X_k))} \quad X \sim N(\mu, \sigma^2)$$

MGF(X) =

Theorem: Within MVM,  
uncorrelated implies indep.

(Note:- In general,  
indep implies uncorrelation,  
not other way around)

$$E(e^{tX})$$

$$= e^{\mu t + \frac{1}{2} \sigma^2 t^2}$$

(MGF of any univariate normal)

$\vec{X} = \begin{pmatrix} \vec{X}_1 \\ \vec{X}_2 \end{pmatrix}$  MVN, if every component of  $\vec{X}_1$

is uncorrelated with every component of  $\vec{X}_2$ ,  
then  $\vec{X}_1$  is independent of  $\vec{X}_2$ .

Example

Let  $X, Y$  be i.i.d.  $N(0, 1)$ . Then,  $(X+Y, X-Y)$   $\nearrow$  2D (bivariate normal)  
is MVN.

They are uncorrelated:

$$\text{Cov}(X+Y, X-Y) = \text{Var}(X) + \text{Cov}(X, Y) - \text{Cov}(X, Y) - \text{Var}(Y)$$

$$= 0$$

So,  $X+Y, X-Y$  are independent.