

Theorem:  $\{X_n\}$  is a sequence of uncorrelated random variables; such that  $\mathbb{E}[X_n] = \bar{x}$   
 and  $\text{Var}(X_n) = \sigma^2 < \infty$

"Finite variance"

$$\text{Sample average: } S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i$$

$$S_n \xrightarrow{n \rightarrow \infty} \bar{x}$$

$$\text{Proof: } \lim_{n \rightarrow \infty} \mathbb{E}[|S_n - \bar{x}|^2] = 0$$

$$S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i$$

$$\mathbb{E}[S_n] = \frac{1}{n} \sum_{i=0}^{n-1} \mathbb{E}[X_i] = \frac{1}{n} \sum_{i=0}^{n-1} \bar{x} = \bar{x}$$

$$\rightarrow = \mathbb{E}[(S_n - \mathbb{E}[S_n])^2]$$

$$= \mathbb{E}\left[\left(\frac{1}{n} \sum_{i=0}^{n-1} X_i - \frac{1}{n} \sum_{i=0}^{n-1} \bar{x}\right)^2\right]$$

$$= \frac{1}{n^2} \mathbb{E}\left[\left(\sum_{i=0}^{n-1} (X_i - \bar{x})\right)^2\right]$$

$$= \frac{1}{n^2} \mathbb{E}\left[\sum_i (X_i - \bar{x})^2 + \sum_i \sum_{j \neq i} (X_i - \bar{x})(X_j - \bar{x})\right]$$

$$= \frac{1}{n^2} \sum_i \underbrace{\mathbb{E}[(X_i - \bar{x})^2]}_{\text{Var}(X_i)} + \sum_i \sum_{j \neq i} \underbrace{\mathbb{E}[(X_i - \bar{x})(X_j - \bar{x})]}_0$$

$$= \frac{1}{n^2} \sum_i \sigma^2 = \frac{1}{n^2} \cdot n \cdot \sigma^2 = \frac{\sigma^2}{n}$$

$$\lim_{n \rightarrow \infty} \mathbb{E}[(S_n - X)^2] = 0$$

weak law of large numbers

Suppose  $\{X_n\}$  is a sequence of uncorrelated r.v.s such that  $\mathbb{E}[X_n] = X$  and  $\text{Var}(X_n) = \sigma^2 < \infty$

$$S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i \xrightarrow{P} X$$

Strong law of large numbers

Suppose  $\{X_n\}$  is an i.i.d sequence of random variables with finite mean and variance.

$$S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i \rightarrow X \text{ with probability 1}$$

Central Limit Theorem

let  $\{X_n\}$  be a sequence of i.i.d random variables with finite mean ( $m$ ), and variance ( $\sigma^2$ )

$$R_n = \frac{1}{\sqrt{n}} \sum_{i=0}^{n-1} \left( \frac{X_i - m}{\sigma} \right) \left( \in \mathcal{N}(0, 1) \right)$$

"Normalization"

$$\text{then } R_n \xrightarrow{d} \mathcal{N}(0, 1)$$

$$\hat{X}_i = \frac{X_i - m}{\sigma}$$

$$\mathbb{E}[\hat{X}_i] = 0, \quad \text{Var}(\hat{X}_i) = 1 \quad [\text{Exercise}]$$

$$R_n = \frac{1}{\sqrt{n}} \sum_{i=0}^{n-1} \hat{X}_i$$

Need to show  $F_{R_n} \rightarrow \text{Standard Normal Cdf}$

$M_{R_n} \rightarrow \text{MGF of standard normal}$

$$M_{R_n}(t) = \mathbb{E}[e^{t R_n}]$$

$$= \mathbb{E}\left[e^{\frac{t}{\sqrt{n}} \sum_{i=0}^{n-1} \hat{X}_i}\right]$$

$$= \mathbb{E}\left[e^{t \hat{X}_0 / \sqrt{n}} e^{t \hat{X}_1 / \sqrt{n}} \dots e^{t \hat{X}_{n-1} / \sqrt{n}}\right]$$

$$= \stackrel{\text{independence}}{\mathbb{E}}[e^{t \hat{X}_0 / \sqrt{n}}] \mathbb{E}[e^{t \hat{X}_1 / \sqrt{n}}] \dots \mathbb{E}[e^{t \hat{X}_{n-1} / \sqrt{n}}]$$

$$\left. \begin{aligned} M_{X_i}(t/\sqrt{n}) \\ = M(t/\sqrt{n}) \end{aligned} \right\} = \prod_{i=0}^{n-1} \underbrace{\mathbb{E}[e^{t \hat{X}_i / \sqrt{n}}]}_{M_{X_i}(t/\sqrt{n})}$$

$$= \stackrel{\text{identical}}{\underbrace{(M(t/\sqrt{n}))^n}}$$

$$\lim_{n \rightarrow \infty} \underbrace{(M(t/\sqrt{n}))^n} \rightarrow \underbrace{\text{MGF of } N(0, 1)}_{e^{-t^2/2}}$$

$$M(s) \approx M(0) + s \cdot M'(0) + \frac{s^2}{2} M''(0)$$

Taylor series expansion around  $s=0$

$M$ : MGF of  $\hat{X}_i$

$$M(0) = \mathbb{E}[e^{0 \cdot \hat{X}_i}] = \mathbb{E}[1] = 1$$

$$M'(0) = \frac{d}{dt} \mathbb{E}[e^{t \cdot \hat{X}_i}] \Big|_{t=0} = \mathbb{E}[\hat{X}_i] = 0$$

$$M''(0) = \frac{d^2}{dt^2} \mathbb{E}[e^{t \cdot \hat{X}_i}] \Big|_{t=0} = \mathbb{E}[\hat{X}_i^2] = 1$$

$$M(s) \approx 1 + s^2/2$$

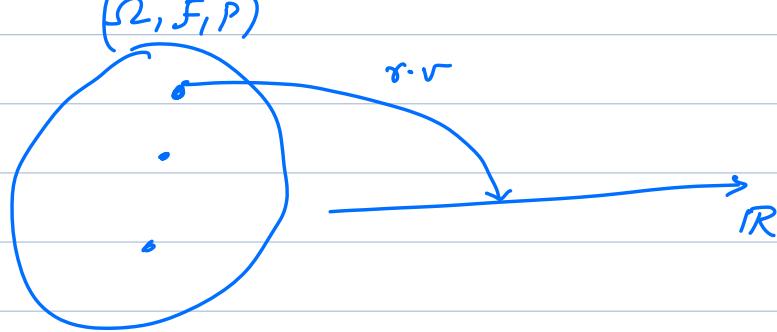
$$\lim_{n \rightarrow \infty} (M(t/\sqrt{n}))^n = \lim_{n \rightarrow \infty} \left(1 + \frac{t^2}{2n}\right)^n$$

$$= e^{t^2/2} //$$

"MGF of standard normal"

### Random Process

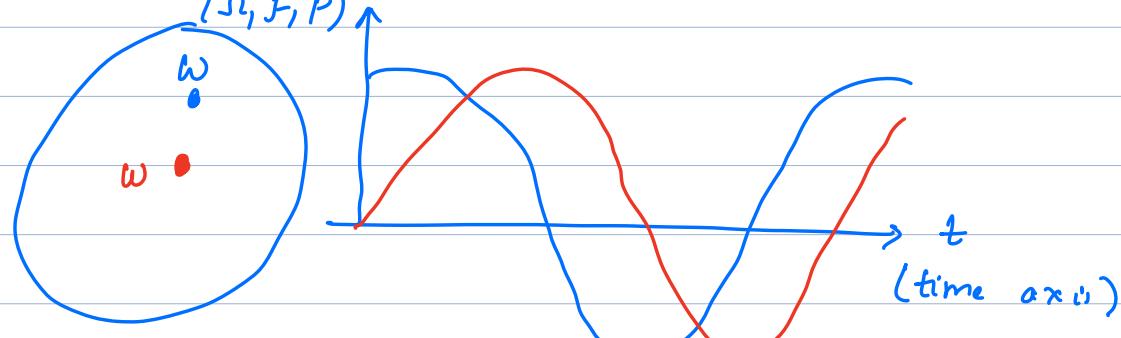
$(\Omega, \mathcal{F}, P)$



random vectors

$$= (X_0, X_1, \dots, X_{n-1})$$

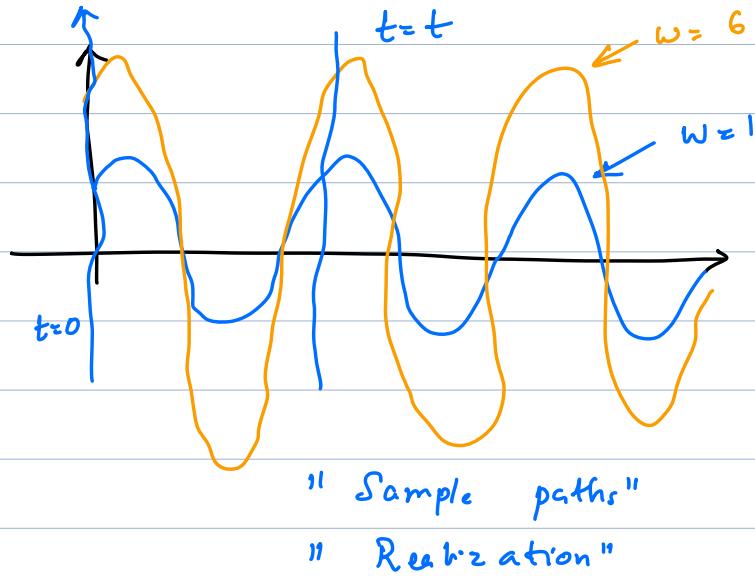
$(\Omega, \mathcal{F}, P)$



Example: Random Experiment: Throw a dice.

$$\omega = \{1, 2, \dots, 6\}$$

$$X(\omega, t) = \omega \cdot \cos(t)$$



$X_0$  is a random variable

$X_t$  is a random variable.

$$= \begin{cases} \cos t & \text{w.p. } \frac{1}{6} \\ 2 \cdot \cos t & \text{w.p. } \frac{1}{6} \\ \vdots \\ 6 \cdot \cos t & \text{w.p. } \frac{1}{6} \end{cases}$$

$X_{t_0} = X(\omega, t=t_0)$  is a random variable.

Example:  $\Omega = \{0, 1\}$

$$P(0) = \alpha$$

$$P(1) = 1 - \alpha$$

$$X(t, \omega) = \cos(\omega t)$$

What are the sample paths?



$$X_t = \begin{cases} 1 & \text{w.p } \alpha \\ \cos t & \text{w.p } 1 - \alpha \end{cases}$$