

Theorem:  $\{X_n\}$  is a sequence of uncorrelated random variables; such that  $E[X_n] = \mu$  and  $\text{Var}(X_n) = \sigma^2 < \infty$

$\uparrow$   
 not a r.v.  
 "Finite variance"

Sample average:  $S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i$

$S_n \xrightarrow{m.s.} \mu$

Proof:  $\lim_{n \rightarrow \infty} E[|S_n - \mu|^2] = 0$

$S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i$

$E[S_n] = \frac{1}{n} \sum_{i=0}^{n-1} E[X_i] = \frac{1}{n} \sum_{i=0}^{n-1} \mu = \mu$

$\rightarrow = E[(S_n - E[S_n])^2]$

$= E\left[\left(\frac{1}{n} \sum_{i=0}^{n-1} X_i - \frac{1}{n} \sum_{i=0}^{n-1} \mu\right)^2\right]$

$= \frac{1}{n^2} E\left[\left(\sum_{i=0}^{n-1} (X_i - \mu)\right)^2\right]$

$= \frac{1}{n^2} E\left[\sum_i (X_i - \mu)^2 + \sum_i \sum_{\substack{j \\ i \neq j}} (X_i - \mu)(X_j - \mu)\right]$

$= \frac{1}{n^2} \sum_i \underbrace{E[(X_i - \mu)^2]}_{\text{Var}(X_i) = \sigma^2} + \sum_i \sum_{\substack{j \\ i \neq j}} \underbrace{E[(X_i - \mu)(X_j - \mu)]}_0$

$= \frac{1}{n^2} \sum_i \sigma^2 = \frac{1}{n^2} \cdot n \cdot \sigma^2 = \frac{\sigma^2}{n}$

$$\lim_{n \rightarrow \infty} \mathbb{E}[(S_n - X)^2] = 0$$

Weak law of large numbers

Suppose  $\{X_n\}$  is a sequence of uncorrelated r.v.s such that  $\mathbb{E}[X_n] = X$  and  $\text{Var}(X_n) = \sigma^2 < \infty$

$$S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i \xrightarrow{P} X$$

Strong law of large numbers

Suppose  $\{X_n\}$  is an i.i.d sequence of random variables with finite mean and variance.

$$S_n = \frac{1}{n} \sum_{i=0}^{n-1} X_i \rightarrow X \text{ with probability 1}$$

Central Limit Theorem

Let  $\{X_n\}$  be a sequence of i.i.d random variables with finite mean ( $m$ ), and variance ( $\sigma^2$ )

$$R_n = \frac{1}{\sqrt{n}} \sum_{i=0}^{n-1} \left( \frac{X_i - m}{\sigma} \right) \quad \left( \text{e.g. } S_n \right)$$

"Normalization"

$$\text{then } R_n \xrightarrow{d} \mathcal{N}(0, 1)$$

$$\hat{X}_i = \frac{X_i - m}{\sigma}$$

$$\mathbb{E}[\hat{X}_i] = 0, \quad \text{Var}(\hat{X}_i) = 1 \quad [\text{Exercise 1}]$$

$$R_n = \frac{1}{\sqrt{n}} \sum_{i=0}^{n-1} \hat{X}_i$$

Need to show  $F_{R_n} \rightarrow \text{Standard Normal cdf}$

$M_{R_n} \rightarrow \text{MGF of standard normal}$

$$M_{R_n}(t) = \mathbb{E}[e^{t R_n}]$$

$$= \mathbb{E}\left[e^{\frac{t}{\sqrt{n}} \sum_{i=0}^{n-1} \hat{X}_i}\right]$$

$$= \mathbb{E}\left[e^{t \hat{X}_0 / \sqrt{n}} e^{t \hat{X}_1 / \sqrt{n}} \dots e^{t \hat{X}_{n-1} / \sqrt{n}}\right]$$

$$\stackrel{\text{independence}}{=} \mathbb{E}\left[e^{t \hat{X}_0 / \sqrt{n}}\right] \mathbb{E}\left[e^{t \hat{X}_1 / \sqrt{n}}\right] \dots \mathbb{E}\left[e^{t \hat{X}_{n-1} / \sqrt{n}}\right]$$

$$\boxed{\begin{aligned} M_{X_i}(t/\sqrt{n}) \\ = M(t/\sqrt{n}) \end{aligned}} = \prod_{i=0}^{n-1} \underbrace{\mathbb{E}\left[e^{t \hat{X}_i / \sqrt{n}}\right]}_{M_{X_i}(t/\sqrt{n})}$$

$$\stackrel{\text{identical}}{=} \left(M(t/\sqrt{n})\right)^n$$

$$\lim_{n \rightarrow \infty} \underbrace{\left(M(t/\sqrt{n})\right)^n}_{\rightarrow \text{MGF of } N(0,1)} \rightarrow \underbrace{e^{t^2/2}}$$

$$M(s) \underset{\uparrow}{\approx} M(0) + s \cdot M'(0) + \frac{s^2}{2} M''(0)$$

Taylor series expansion around  $s=0$

$M$ : MGF of  $\hat{X}_i$

$$M(0) = \mathbb{E}[e^{0 \cdot \hat{x}_i}] = \mathbb{E}[1] = 1$$

$$M'(0) = \left. \frac{d}{dt} \mathbb{E}[e^{t \hat{x}_i}] \right|_{t=0} = \mathbb{E}[\hat{x}_i] = 0$$

$$M''(0) = \left. \frac{d^2}{dt^2} \mathbb{E}[e^{t \hat{x}_i}] \right|_{t=0} = \mathbb{E}[\hat{x}_i^2] = 1$$

$$M(s) \approx 1 + s^2/2$$

$$\lim_{n \rightarrow \infty} \left( M(t/\sqrt{n}) \right)^n = \lim_{n \rightarrow \infty} \left( 1 + \frac{t^2}{2n} \right)^n$$

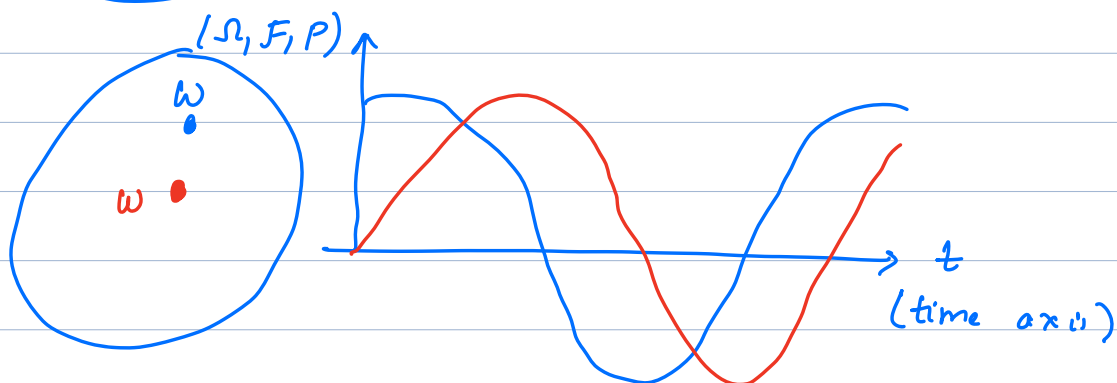
$$= e^{t^2/2} //$$

"MGF of standard normal"

### Random Process



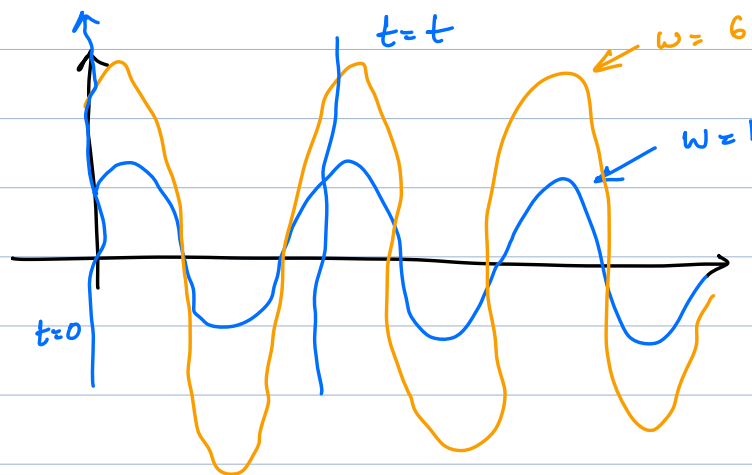
random vectors  
 $= (x_0, x_1, \dots, x_{n-1})$



Example: Random Experiment: Throw a dice.

$$\omega \in \{1, 2, \dots, 6\}$$

$$X(\omega, t) = \omega \cdot \cos(t)$$



"Sample paths"

"Realization"

$X_0$  is a random variable

$X_t$  is a random variable.

$$= \begin{cases} \cos t & \text{w.p. } \frac{1}{6} \\ 2 \cdot \cos t & \text{w.p. } \frac{1}{6} \\ \vdots \\ 6 \cos t & \text{w.p. } \frac{1}{6} \end{cases}$$

$X_{t_0} = X(\omega, t=t_0)$  is a random variable.

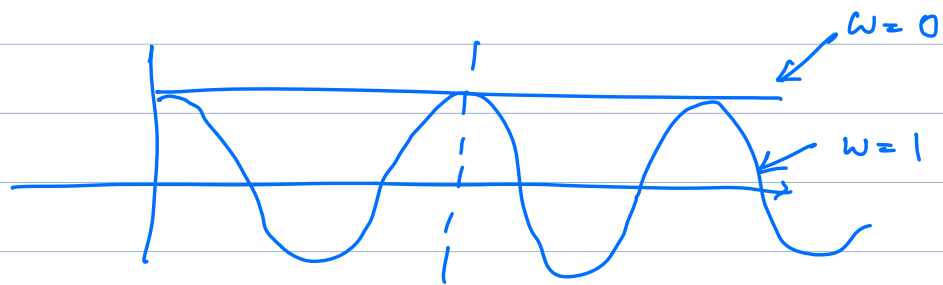
Example:  $\Omega = \{0, 1\}$

$$P(0) = \alpha$$

$$P(1) = 1 - \alpha$$

$$X(t, \omega) = \cos(\omega t)$$

What are the sample paths?



$$X_t = \begin{cases} 1 & \text{w.p. } \alpha \\ \cos t & \text{w.p. } 1-\alpha \end{cases} \quad X_t$$