# Bad Control

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What goes wrong with bad controls? Everything!

### A Very Simple Example

Let's make some data in just the same way that we typically make data.

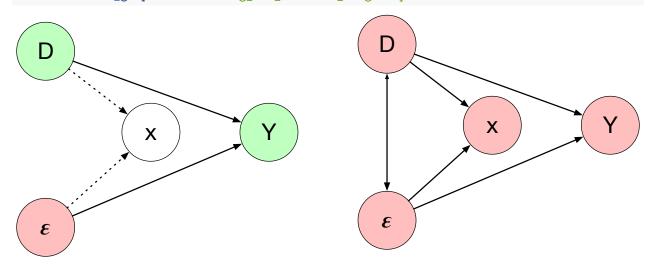
First, let's make and id variable, an epsilon variable that represents everything in the world that we haven't measured, a D variable that represents the assignment to treatment, and a tau variable that represents each individuals response to treatment if he or she or it is exposed to the treatment.

```
d <- data.table(
  id = 1:1000,
  epsilon = rnorm(1000),
  D = sample(0:1, 1000, replace = TRUE),
  tau = rnorm(n = 1000, mean = 2, sd = 1)
)</pre>
```

Now, lets make the *slightly* more complicated quantities, x, the bad control, and the potential outcomes. We're going to take some care to create the x-variable. As we show in the figure below, the x-variable is **caused** by both the treatment and other things in the system.

For example, consider the case of the white collar and blue collar jobs. If you're born the heir of a real-estate empire in NY (in this case represented by  $\epsilon$ ), you might be likely, no matter your college attendance or draft status, to end up in a white collar job (in this case represented in X). But, it may also be the case in the data that having been assigned to attend college (in this case represented in D) increases the probability that you get a white collar job.

knitr::include\_graphics("./coding\_bad\_controls\_diagram.pdf")



Suppose that we measure four variables,  $D, x, Y, \epsilon$  that all have a positive relationship between one another, shown in the figures above. The one exception, and this is important to make this example easy-ish to understand, is that x actually has no relationship on Y.

The way we've built this data, if someone is in treatment (D=1), then they're more likely to have a larger value for X and also a larger value for Y. Similarly, if someone has a high  $\epsilon$  value, then they're more likely to have a higher value for X and also a higher value for Y.

In the figure on the left, where we've included dotted-line relationships between D, X, and  $\epsilon$ , we're acknowledging that these relationships exist, but we're not going to condition on them. As a result, there is no covariance between D and  $\epsilon$ , which we represent with no line. In the figure on the right, consider that we've conditioned on the X variable. As a result, there is a solid line relationship between  $\epsilon$  and X; but so too is there a solid line relationship between  $\epsilon$  and D after we condition on X.

Perhaps think of it this way – if we don't condition on X then any value of  $\epsilon$  is possible, and so too is any value of D. But, after we set the X to a value by conditioning on it, some arrangements are more likely to have been the case than others.

- If we don't condition on X, and if  $\epsilon$  is low, it could still be the case that the unit was in treatment or was in control. We don't have any conditional information.
- But, if we condition so that the units we're considering have a X = 1, and if  $\epsilon$  is low, then we can know that the person was more likely to be in the treatment group than the control group.

Let's make these, carefully. Create the variable x, the realization of the variable X to be a binary indicator that takes the values of either 1 or 0. And, suppose that the probability of being a 1 increases if you're in treatment and also if you've got a high value of epsilon. (I'm going to scale these probabilities using a pnorm call, just so that the values for the prob argument to rbinom conforms to be a valid probability statement.

```
d[ , x := rbinom(n = 1000, size = 1, prob = pnorm(2*D + epsilon))]
```

### Questions for Understanding

- 1. If we examine the relationship between x and D with a regression, should the relationship be positive, negative, or we don't know?
- 2. If we examine the relationship between x and  $\epsilon$  with a regression, should the relationship be positive, negative, or we don't know?
- 3. **Challenge**: If we examine the regression of x ~ D + epsilon what will be the nature of each variables' relationship with x? Will they be the same, different, and will they be positive, negative, or we don't know?

```
mod_d <- d[ , lm(x ~ D)]
mod_x <- d[ , lm(x ~ epsilon)]
mod_both <- d[ , lm(x ~ D + epsilon)]

## I've commented this out so that you can uncomment it after you've answered the
## questions for understanding.
##
stargazer(mod_d, mod_x, mod_both, type = 'text')</pre>
```

\_\_\_\_\_

```
Dependent variable:
```

```
(1) (2) (3)

(0.026) (0.022)
epsilon 0.212*** 0.195***
(0.013) (0.011)
Constant 0.464*** 0.691*** 0.484***
(0.018) (0.013) (0.016)
```

#### Make Potential Outcomes

To this point, we haven't actually built any outcomes, we've been principally concerned with building the *upstream* data for our causal system. Let's make those potential outcomes now.

For everybody, their potential outcomes to control are a combination of some individual idiosyncratic value (represented by the random draw in this code), plus their x value, plus all the other features in the world that we haven't measured, their  $\epsilon$ .

```
d[ , y0 := runif(1000, min = 0, max = 10) + epsilon]
```

#### Question for understanding

Do we **need** in a strict sense to define the potential outcomes to treatment on this data? Or, would that information simply be a linear combination of some data that we already have in hand?

Finally, let's produce a measured outcome value for Y.

```
d[, Y := as.numeric(NA)]
d[D==0, Y := y0]
d[D==1, Y := y0 + tau]
```

### Finally, Estimate Causal Effects

In the world that we've occupied to this point, whether we do or do not condition on a variable, an experiment that we've executed that has successfully randomized the treatment should produce a reliable causal estimate. That is, in this setup, if we estimate a model, it should provide us with an unbiased estiamte of  $\tau$ , or 2.

```
unconditional_model <- d[ , lm(Y ~ D)]
```

But, what is going to happen if we use a *cough cough bad...* control variable to "clean up our estimate". After all, if controls only *increase the accuracy of our predictions*, then this will improve our model right?

```
bad_control_model <- d[ , lm(Y ~ D + x)]</pre>
```

What are other ways that we could see this relationship? Well, earlier we talked about "setting x to be equal to 1". Let's try that on this data. Among the people who have an x value that is one, what is the causal relationship between D and Y?

Table 1:

	Dependent variable: Y			
	(1)	(2)	(3)	
D	2.073***	1.515***	1.630***	
	(0.203)	(0.230)	(0.258)	
X		1.220***		
		(0.248)		
Constant	4.878***	4.311***	5.457***	
	(0.141)	(0.181)	(0.209)	
Data Subset	All	All	x == 1	
Observations	1,000	1,000	686	
$\mathbb{R}^2$	0.095	0.116	0.055	
Adjusted $R^2$	0.094	0.115	0.054	
Residual Std. Error	3.204 (df = 998)	3.167 (df = 997)	3.224 (df = 684)	
F Statistic	$104.585^{***} (df = 1; 998)$	$65.616^{***} (df = 2; 997)$	$39.791^{***} (df = 1; 684)$	
Residual Std. Error F Statistic	` ,	$65.616^{***}$ (df = 2; 997)		

*Note:* \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

```
d <- data.table(id = 1:nRows)
setkeyv(d, cols = "id")</pre>
```

Create Potential Outcomes to control

```
d[ , y0 := runif(min=-10, max=10, n = .N)]
```

Create a Treatment Effect with a mean 4. This is the value that we're looking to recover with our estimator.

```
d[, tau := rnorm(n = .N, mean = 4)]
```

Randomly Assign Treatment

```
d[ , D := sample(0:1, size = .N, replace = TRUE)]
```

Create a bad control that is correlated with treatment

```
d[, bc := 2 + 3 * D + rnorm(n = .N, mean = 0)]
```

And, finally, reate the observed outcomes.

```
d[D == 0, Y := y0 + bc]
d[D == 1, Y := y0 + bc + tau]
```

#### What is the causal model we hold?

When we are thinking about the causal model here, we're saying, "I think that the conditional expectation of Y depends on the treatment status". But, maybe I think I want to also control for the variable bc.

In fact, we can estimate a reliable causal effect for either D on Y, or D on bc, but not the two togheter.

First, the caual effect of D on Y in Table 2.

```
m1 <- lm(Y ~ D, data = d)
stargazer(m1, type = 'latex', label = 'goodModel', header = F)</pre>
```

Table 2:

Table 2.		
	Dependent variable:	
	Y	
D	6.700***	
	(0.385)	
Constant	2.200***	
	(0.264)	
Observations	1,000	
$\mathbb{R}^2$	0.233	
Adjusted R <sup>2</sup>	0.232	
Residual Std. Error	6.081 (df = 998)	
F Statistic	302.328*** (df = 1; 998)	
Note:	*p<0.1; **p<0.05; ***p<0.01	

Then, the causal effect of D on bc in Table 3.

```
m2 <- lm(bc ~ D, data = d)
stargazer(m2, type = 'latex', label = 'goodModelOnControl', header = F)</pre>
```

Table 3:

	Dependent variable:
	bc
D	3.036***
	(0.065)
Constant	2.049***
	(0.044)
Observations	1,000
$\mathbb{R}^2$	0.689
Adjusted R <sup>2</sup>	0.688
Residual Std. Error	1.019 (df = 998)
F Statistic	2,209.019*** (df = 1; 998)
Note:	*p<0.1; **p<0.05; ***p<0.01

So does this. But what happens if we attempt to control for the bc indicator at the same time that we estimate the effect of D on Y? This isn't going to work-out, as we can see in Table 4

```
m3 <- lm(Y ~ D + bc, data = d)
stargazer(m3, type = 'latex', label = 'badModel', header = F)</pre>
```

Now we're left with an under-estiamte of the causal effect of D on Y, and we've got some estimate of the effect of bc on Y. But, when we built the data, there wasn't such an effect! In fact, what is in here is the relationship between y0 and Y, but through a mangled causal pipeline.

Table 4:

	Dependent variable:	
	Y	
D	3.368***	
	(0.679)	
bc	1.098***	
	(0.186)	
Constant	-0.050	
	(0.461)	
Observations	1,000	
$\mathbb{R}^2$	0.258	
Adjusted R <sup>2</sup>	0.257	
Residual Std. Error	5.980 (df = 997)	
F Statistic	$173.774^{***} (df = 2; 997)$	
Note:	*p<0.1; **p<0.05; ***p<0.01	

The thing of it is, depending on the relationship between the bad control and the treatment, and the bad control and the treatment, there's *just no structured way to know* which direction the bias is going to be when you include a bad control.

stargazer(m1, m2, m3, type = 'latex', header = F)

Table 5:

	Dependent variable:		
	Y	bc	Y
	(1)	(2)	(3)
D	6.700***	3.036***	3.368***
	(0.385)	(0.065)	(0.679)
bc			1.098***
			(0.186)
Constant	2.200***	2.049***	-0.050
	(0.264)	(0.044)	(0.461)
Observations	1,000	1,000	1,000
$\mathbb{R}^2$	0.233	0.689	0.258
Adjusted R <sup>2</sup>	0.232	0.688	0.257
Residual Std. Error	6.081 (df = 998)	1.019 (df = 998)	5.980 (df = 997)
F Statistic	$302.328^{***} (df = 1; 998)$	$2,209.019^{***} (df = 1; 998)$	$173.774^{***} (df = 2; 997)$
Note:		*p	<0.1; **p<0.05; ***p<0.01

I can't get this written before class Fall 2017. Let me know if you figure it out.

### An example with Education

What if we build an example that is a little more complicated, but that fits with a story that we believe about the world.

Suppose that people have potential outcomes to earnings, and they are deciding whether they want to pursue a MIDS degree or not. We've been working on this one for a while.

The Dean, in her wisdom, agrees to let current students and 241 faculty run an expriment where they randomly admit people to the program. Hurrah! We can evaluate this! And we can make a causal claim.

Suppose we have the following as a setup:

- Some students live in Bay Area. Because wages are stupid here (so are mortages), these students are paid on average \$25,000 more than people who do not live in the Bay Area.
- The *true* state of the world is that graduating with a MIDS degree leads individuals to earn \$50,000 more, on average, if they graduate with a MIDS degree than if they didn't. This isn't a selection effect, but instead is actually the causal effect.
- There are some students who are *really* motivated to get that job at Pintrest. And, they think that actually have a shot at it once they're in MIDS! So, of the folks who are admitted, suppose that 25% of them move to the Bay Area. None of the people who are not admitted to MIDS move to the Bay Area.

We could make data that is consistent with this story, right? Suppose that we have 500 students, at baseline the mean earnings for people who live in SF is \$75,000 and the mean earnings for people who do not live in SF is \$50,000. Furthermore, at baseline, 25% of the students live in SF.

First, create baseline earnings for people.

```
d <- data.table(id = 1:500)
d[, y0 := rnorm(.N, mean = 50000, 10000)]
d[, lives_in_sf_before := sample(c(0,1), size = .N, replace = TRUE, prob = c(0.75, 0.25))]
d[, y0 := y0 + lives_in_sf_before * rnorm(.N, 25000, 10000)]</pre>
```

Then, randomly assign people to receive the MIDS admission.

```
d[ , MIDS_admission := sample(c(0,1), size = .N, replace = TRUE)]
```

For the people who are admitted to MIDS and don't live in Bay Area, suppose the hardest working among them move to the Bay Area (and in doing so, they earn \$25,000 more just because wages are crazy). They the folks who just needed to catch a break!

Did that seem to work? Yes.

```
d[ , table(lives_in_sf_before, lives_in_sf_after)]

## lives_in_sf_after

## lives_in_sf_before 0 1

## 0 325 56

## 1 0 119

Make the after MIDS earnings data.
```

```
d[MIDS_admission == 0, Y := y0]
d[MIDS_admission == 1 & lives_in_sf_before == 1, Y := y0 + rnorm(.N, 50000, 10000)]
```

```
d[MIDS_admission == 1 & lives_in_sf_before == 0 & lives_in_sf_after == 1,
 Y := y0 + rnorm(.N, 50000, 10000) + rnorm(.N, 25000, 10000)]
d[MIDS_admission == 1 & lives_in_sf_before == 0 & lives_in_sf_after == 0,
 Y := y0 + rnorm(.N, 50000, 10000)]
mod1 <- d[ , lm(Y ~ MIDS_admission)]</pre>
mod2 <- d[ , lm(Y ~ MIDS_admission + lives_in_sf_before)]</pre>
mod3 <- d[ , lm(Y ~ MIDS_admission + lives_in_sf_after)]</pre>
stargazer(mod1, mod2, mod3, type = 'text')
##
                                            Dependent variable:
##
##
                                                   (2)
##
                            (1)
                                                                         (3)
                     56,209.760***
                                             55,515.650***
                                                                   50,049.570***
## MIDS_admission
##
                        (1,533.294)
                                               (1,343.733)
                                                                     (1,243.839)
##
                                               19,483.750***
## lives_in_sf_before
                                                (1,577.471)
##
##
## lives_in_sf_after
                                                                     23,399.010***
##
                                                                      (1,303.731)
##
                       55,760.640***
                                            51,465.020***
## Constant
                                                                    50,601.810***
                        (1,075.494)
##
                                              (1,003.876)
                                                                      (886.482)
##
## -----
## Observations
                            500
                                                  500
                                                                         500
## R2
                           0.730
                                                 0.793
                                                                        0.836
                                                 0.792
                           0.729
## Adjusted R2
                                                                        0.835
## Residual Std. Error 17,140.560 (df = 498) 15,008.320 (df = 497) 13,364.900 (df = 497)
## F Statistic 1,343.917*** (df = 1; 498) 952.729*** (df = 2; 497) 1,266.312*** (df = 2; 497)
*p<0.1; **p<0.05; ***p<0.01
## Note:
## The people who newly have moved to SF now also get the SF pay-bump.
## People who newly moved to SF now get the SF bump, and the MIDS bump
d[lives_in_sf_before == 0 & lives_in_sf_after == 1,
 y0 := y0 + rnorm(.N, 25000, 10000) + rnorm(.N, 50000, 10000)]
## For the people who didn't move, they get the treatment effect
d[lives_in_sf_before == 0 & lives_in_sf_after == 0,
 y1 := y0 + rnorm(.N, 50000, 10000)]
## And for the people who already lived in SF, they get the treatmetn effect
d[lives_in_sf_before == 1,
 y1 := y0 + rnorm(.N, 50000, 10000)]
## For everyone, if they were to be admitted to MIDS, they would realize
## a $50,000 increase
```

```
## d[ , y1 := y0 + rnorm(.N, 50000, 10000)]
## The people who newly have moved to SF now also get the SF pay-bump.
## People who newly moved to SF now get the SF bump
d[lives_in_sf_before == 0 & lives_in_sf_after == 1,
  y0 := y0 + rnorm(.N, 25000, 10000)]
d[, y1 := y0 + rnorm(.N, 50000, 10000)]
## For the people who didn't move, they get the treatment effect
##d[lives_in_sf_before == 0 & lives_in_sf_after == 0,
## y1 := y0 + rnorm(.N, 50000, 10000)]
## And for the people who already lived in SF, they get the treatmetn effect
##d[lives_in_sf_before == 1,
## y1 := y0 + rnorm(.N, 50000, 10000)]
## For everyone, if they were to be admitted to MIDS, they would realize
## a $50,000 increase
## d[ , y1 := y0 + rnorm(.N, 50000, 10000)]
Then, the last thing is just to reveal these potential outcomes.
d[ , Y := ifelse(MIDS_admission == 1, y1, y0)]
mod1 <- d[ , lm(Y ~ MIDS_admission)]</pre>
mod2 <- d[ , lm(Y ~ MIDS_admission + lives_in_sf_before)]</pre>
mod3 <- d[ , lm(Y ~ MIDS_admission + lives_in_sf_after)]</pre>
stargazer(mod1, mod2, mod3, type = 'text')
Dependent variable:
(1)(2)(3)
                                                               — MIDS admission 74,128.790***
73,894.710*** 62,540.350***
(2,899.300) (2,893.886) (2,359.770)
lives in sf before 6.570.597*
(3,397.269)
lives in sf after 44,017.750***
(2,473.394)
Constant 55,760.640*** 54,312.010*** 46,055.940***
(2,033.647) (2,161.965) (1,681.803)
Observations 500 500 500
R2 0.568 0.571 0.736
Adjusted R2 0.567 0.569 0.735
Residual Std. Error 32,411.010 (df = 498) 32,322.190 (df = 497) 25,355.420 (df = 497)
F Statistic 653.714*** (df = 1; 498) 330.526*** (df = 2; 497) 692.431*** (df = 2; 497) =======
Note: p < 0.1; p < 0.05; p < 0.01
```

# People who didn't move to SF as a result of MIDS\_admission just get the MIDS bump

d[lives\_in\_sf\_before == 0 & lives\_in\_sf\_after == 0, y1 := y0 + rnorm(.N, 50000, 10000)] ## People who already live in SF get the MIDS bump. d[lives\_in\_sf\_before == 1 & MIDS\_admission == 1, y1 := y0 + rnorm(.N, 50000, 10000)]

## People who didn't get admitted to MIDS just show their potential outcomes to control

```
d[MIDS \ admission == 0, Y := y0] "'
Phew! That's a lot of data to have made!
mod <- d[ , lm(Y ~ MIDS_admission)]</pre>
summary(mod)
##
## Call:
## lm(formula = Y ~ MIDS_admission)
##
## Residuals:
##
              1Q Median
                            3Q
     Min
                                  Max
## -66437 -18866 -5173
                          9837 129470
##
## Coefficients:
                  Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                     55761
                                 2034
                                        27.42
                                                 <2e-16 ***
## MIDS admission
                    74129
                                 2899
                                        25.57
                                                 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 32410 on 498 degrees of freedom
## Multiple R-squared: 0.5676, Adjusted R-squared: 0.5667
## F-statistic: 653.7 on 1 and 498 DF, p-value: < 2.2e-16
mod2 <- d[ , lm(Y ~ MIDS_admission + lives_in_sf_before)]</pre>
summary(mod2)
##
## lm(formula = Y ~ MIDS_admission + lives_in_sf_before)
##
## Residuals:
##
     Min
             1Q Median
                            3Q
                                  Max
## -64755 -18316 -4385
                         8105 131153
##
## Coefficients:
                      Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                         54312
                                     2162 25.122
                                                    <2e-16 ***
## MIDS admission
                         73895
                                     2894 25.535
                                                    <2e-16 ***
## lives_in_sf_before
                          6571
                                     3397
                                            1.934
                                                    0.0537 .
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
```

```
## Residual standard error: 32320 on 497 degrees of freedom
## Multiple R-squared: 0.5708, Adjusted R-squared: 0.5691
## F-statistic: 330.5 on 2 and 497 DF, p-value: < 2.2e-16
mod3 <- d[ , lm(Y ~ MIDS_admission + lives_in_sf_after)]</pre>
summary(mod3)
##
## Call:
## lm(formula = Y ~ MIDS_admission + lives_in_sf_after)
## Residuals:
     Min
              1Q Median
                            3Q
                                  Max
## -76072 -13988 -1246 10319 106746
##
## Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                        46056
                                    1682
                                           27.39
                                                   <2e-16 ***
## MIDS_admission
                        62540
                                    2360
                                           26.50
                                                   <2e-16 ***
## lives_in_sf_after
                        44018
                                    2473
                                           17.80
                                                   <2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 25360 on 497 degrees of freedom
## Multiple R-squared: 0.7359, Adjusted R-squared: 0.7348
## F-statistic: 692.4 on 2 and 497 DF, p-value: < 2.2e-16
```