

STAT504 Assignment1

Anuradha Ramachandran

2023-01-15

Problem 1

(a) Simulate 100 draws from $Y = 10 + 5X + \epsilon$ where $X \sim N(0, 1)$ and $\epsilon \sim N(0, 1)$

```
# Sample size n = 100

n = 100
# Generate 100 random X and e from No(0,1)
x = rnorm(n = n, mean = 0, sd = 1)
e = rnorm(n = n, mean = 0, sd = 1)
# y = 10+5x+e
y = 10 + 5 * x + e
summary(y)

##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
## -4.884   6.510   10.610   10.270   13.445   23.634
```

(b) Fit OLS model by regressing Y on X

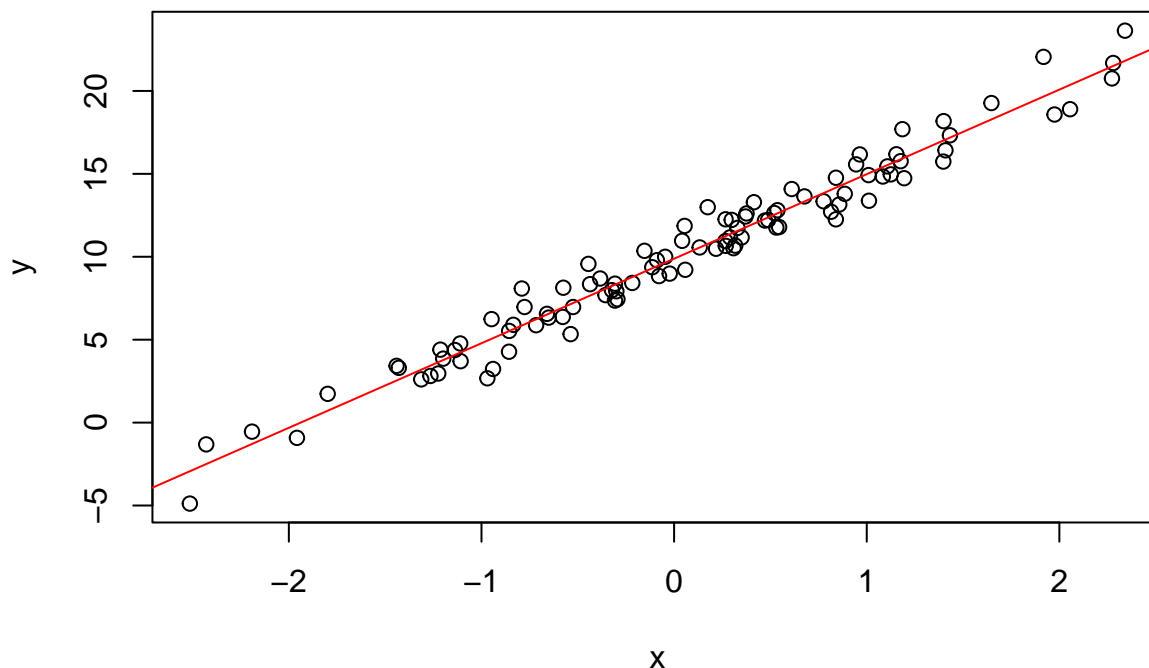
The coefficients from the regression is: Intercept = 10 Slope = 5

```
# OLS model for Y on X
ols = lm(y ~ x)
ols

##
## Call:
## lm(formula = y ~ x)
##
## Coefficients:
## (Intercept)          x
##      9.883       5.098
```

(c) Scatterplot of X & Y with the regression line

```
plot(y ~ x)
abline(ols, col = "red")
```



Problem 2

2(a)

Sample space $\Omega = \{1, 2, 3, 4, 5, 6\}$

Event space $S = \{\text{all subsets of } \Omega\} = 2^6$

Probability measure of event $A \in S = P(A) = \frac{\text{sizeof } A}{\text{sizeof } \Omega} = \frac{|A|}{6}$

$$P(\phi) = 0$$

$$P(\Omega) = 1$$

$$P(1) = P(2) = P(3) = P(4) = P(5) = P(6) = \frac{1}{6}$$

2(b)

Let D=Democrat, R=Republican, I=Independent.

Sample space $\Omega = \{R, I, D\}$, $|\Omega| = 1000$ since $|R| = 200$, $|D| = 400$, $|I| = 400$

Event space $S = \{\text{all subsets of } \Omega\} = \{\phi, R, I, D, \{R, I\}, \{R, D\}, \{I, D\}, \{R, I, D\}\}$

Probability measure of event $A \in S = P(A) = \text{Total people in event } A / 1000$

$$P(\phi) = 0$$

$$P(\Omega) = 1$$

$$P(D) = P(I) = \frac{400}{1000} = \frac{2}{5}$$

$$P(R) = \frac{200}{1000} = \frac{1}{5}$$

$$P(D, I) = \frac{800}{1000} = \frac{4}{5}$$

$$P(D, R) = \frac{600}{1000} = \frac{3}{5}$$

$$P(R, I) = \frac{600}{1000} = \frac{3}{5}$$

$$P(D, R, I) = \frac{1000}{1000} = 1$$

Problem 3

3.1

(a) $E[X] = \int_{-\infty}^{\infty} xf(x)dx$ if x is continuous

(b) $Var(X) = E[(X - EX)^2]$

(c) From definition of variance, $var(X) = E[(X - EX)^2]$
 $var(X) = E[X^2 - 2XE[X] + E[X]^2] = E[X^2] - 2E[X]^2 + E[X]^2$
 $var(X) = E[X^2] - E[X]^2$

(d) $SD(X) = \sqrt{var(X)}$

(e) Assume X is discrete a random variable. Let $Y = g(x)$

$$E[Y] = \sum_{y \in Y} yP(Y = y)$$

$$E[Y] = \sum_{y \in Y} yP(x = g^{-1}(y))$$

Since, $P(x = g^{-1}(y)) = f_X(x)$

$$E[Y] = \sum_{y \in Y} \sum_{x=g^{-1}(y)} yf_X(x)$$

$$E[Y] = \sum_x g(x)f_X(x)$$

(f) $E[a + bX] = \int_{-\infty}^{\infty} (a + bx)f(x)dx$

$$E[a + bX] = a \int_{-\infty}^{\infty} f(x)dx + b \int_{-\infty}^{\infty} xf(x)dx$$

Since $\int_{-\infty}^{\infty} f(x)dx = 1$

$$E[a + bX] = a + b \int_{-\infty}^{\infty} xf(x)dx$$

$$E[a + bX] = a + bE[X]$$

(g) By definition, $var(a + bX) = E[(a + bX - E[a + bX])^2]$

$$var(a + bX) = E[(a + bX - a - bE[X])^2]$$

$$var(a + bX) = E[b^2(X - E[X])^2] = b^2E[(X - EX)^2] = b^2var(X)$$

(h) $SD[a + bX] = \sqrt{var(a + bX)}$

From (g), $SD[a + bX] = \sqrt{b^2var(X)} = |b|SD(X)$

3.2

Markov's inequality: Let X be a random variable that takes only non negative values, then for any $a > 0$, $P(X \geq a) \leq \frac{E[X]}{a}$, provided $E[X]$ exists.

Proof: Let X be a continuous random variable. $E[X] = \int_{-\infty}^{\infty} xf_X(x)dx$

$$E[X] = \int_{-\infty}^0 xf_X(x)dx + \int_0^{\infty} xf_X(x)dx$$

$$E[X] \geq \int_a^{\infty} xf_X(x)dx, \text{ since } a \geq 0$$

$$E[X] \geq a \int_a^{\infty} f_X(x)dx = aP(X \geq a) \text{ Hence, } P(X \geq a) \leq \frac{E[X]}{a}$$

Chebychev's inequality: Let X be a random variable with finite variance then for any $\epsilon > 0$, $P(|X - EX| \geq \epsilon) \leq \frac{var(X)}{\epsilon^2}$

Proof: Consider $(X - EX)^2$ to a random variable and it is strictly positive.

By applying Markov's inequality for any $\epsilon > 0$

$$P((X - EX)^2 \geq \epsilon^2) \leq \frac{E[(X - EX)^2]}{\epsilon^2}$$

Since $(X - EX)^2 \geq \epsilon^2 \implies |X - EX| \geq \epsilon$

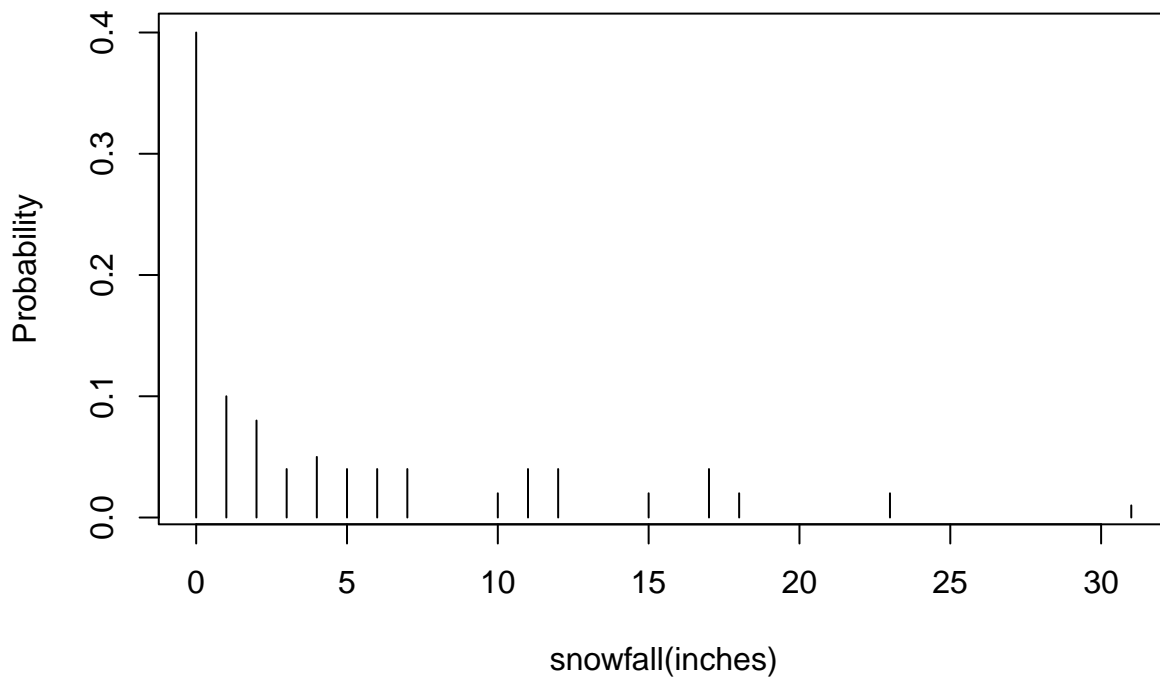
By (b), $P(|X - EX| \geq \epsilon) \leq \frac{\text{var}(X)}{\epsilon^2}$

According to Chebychev's inequality, the probability that the absolute deviation of a random variable from its mean will exceed a threshold ϵ times standard deviation is less than or equal to $\frac{1}{\epsilon^2}$.

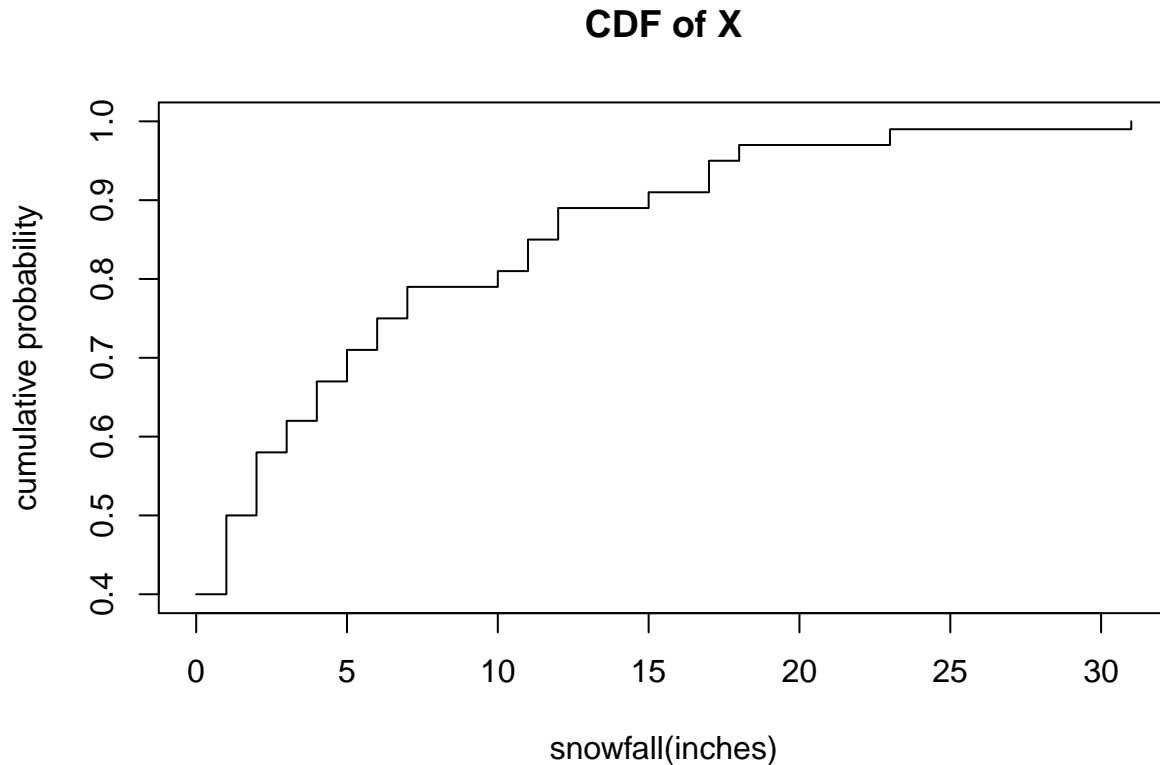
3.3(a) PMF and CDF of X

```
# Load CSV
snow = read.csv("/Users/anuram/Library/Mobile Documents/com~apple~CloudDocs/MS Stats/Winter 2023/STAT50
# PMF of X
pmf = plot(snow$prob ~ snow$snowfall, type = "h", xlab = "snowfall(inches)",
  ylab = "Probability", main = "PMF of X")
```

PMF of X



```
# cdf of X
snow$CDF = cumsum(snow$prob)
plot(snow$CDF ~ snow$snowfall, type = "s", xlab = "snowfall(inches)",
  ylab = "cumulative probability", main = "CDF of X")
```



3.3(b) Mean, median, mode, variance of X and 95% percentile of X

$$E[X] = \sum_x xP(X = x) = 4.53$$

Median(X) is m for which $P(X \leq m) = 0.5$.

Hence Median(X) = 1

Mode(X) is the X that has highest probability. Mode of this distribution is $X = 0$ inches

$$Var(X) = E[X^2] - EX^2 = 40.79$$

From the CDF, it can be seen that the 95% percentile of X is 17 inches.

```
# mean of X
mean = sum(snow$snowfall * snow$prob)
mean

## [1] 4.53

median = subset(snow$snowfall, snow$CDF == 0.5)
median

## [1] 1

# mode of X
mode = subset(snow$snowfall, snow$prob == max(snow$prob))
mode

## [1] 0

# variance of X
exp_sq = sum((snow$snowfall^2) * snow$prob)
variance = exp_sq - mean^2
variance
```

```
## [1] 40.7891
```

```
# 95% percentile of X
percentile = subset(snow, snow$CDF >= 0.95)
percentile
```

```
##      X snowfall prob CDF
## 13 13      17 0.04 0.95
## 14 14      18 0.02 0.97
## 15 15      23 0.02 0.99
## 16 16      31 0.01 1.00
```

3.3(c) Odds of snowing

Odd of snowing = $\frac{P(\text{snowfall} > 0)}{P(\text{snowfall} = 0)} = 1.5$

```
# odds of snowing
nosnow = subset(snow$prob, snow$snowfall == 0)
odds = (1 - nosnow)/nosnow
odds
```

```
## [1] 1.5
```

3.3(d)

The best predictors could be any of the summary statistics like mean, median, mode calculated above depending on the definition of the loss function. Eg: If MSE is the loss function, then $\text{mean}(X)$ would be the best predictor that minimizes MSE and gives the best prediction of snowfall.

3.3(e) MSE when $E(X)$ is used

$MSE = E[(X - EX)^2] = \text{var}(X) = 40.7891$

3.3(f) 95% prediction interval

95% Prediction Interval = $[E[X] - 1.96 * SD[X], E[X] + 1.96 * SD[X]]$

Lower_limit = -7.987. Since X represents snowfall in inches, the lowest value it can take is 0. Hence the lower limit in the prediction interval is 0. Upper_limit = 17 inches

95% prediction interval of snowfall = [0,17]

```
Lower_limit = mean - 1.96 * sqrt(variance)
Lower_limit
```

```
## [1] -7.987804
```

```
Upper_limit = mean + 1.96 * sqrt(variance)
Upper_limit
```

```
## [1] 17.0478
```

Problem 4

4(a)

$$\begin{aligned} E[(X - c)^2] &= E[X^2 - 2cX + c^2] = E[X^2] - 2cE[X] + c^2 \\ E[(X - c)^2] &= E[X^2] - E[X]^2 + E[X]^2 - 2cE[X] + c^2 \\ E[(X - c)^2] &= E[X^2] - E[X]^2 + (E[X] - c)^2 \\ E[(X - c)^2] &= \text{var}(X) + (E[X] - c)^2 \end{aligned}$$

4(b)

Let $Y = \operatorname{argmin}_{c \in R} E[(X - c)^2]$

From 4(a), $Y = \operatorname{argmin}_{c \in R} \operatorname{var}(X) + (E[X] - c)^2$

Differentiating Y with respect to c and equating to 0, $c = E[X]$

Hence $E[X]$ is the best predictor of X when MSE is the loss function.

4(c)

X is continuous random variable. Let $\phi = E[|X - c|] = \int_{-\infty}^c (c - x)f(x)dx + \int_c^{\infty} (x - c)f(x)dx$

Differentiating ϕ with respect to c and equating to 0 by using Leibniz's rule, $\frac{d\phi}{dc} = \int_{-\infty}^c \frac{\partial}{\partial c}(c - x)f(x)dx +$

$$\int_c^{\infty} \frac{\partial}{\partial c}(x - c)f(x)dx$$

$$\frac{d\phi}{dc} = \int_{-\infty}^c f(x)dx - \int_c^{\infty} f(x)dx = 0$$

$$\implies P(X \leq c) = P(X > c)$$

But $P(X \leq c) + P(X > c) = 1$

$$\implies P(X \leq c) = P(X > c) = 1/2$$

Hence Median $[X]$ is the best predictor when mean absolute error is the loss function.

4(d)

Mode(X) is the value $x \in X$ for which the marginal distribution of X (PDF if X is continuous or PMF if X is discrete) is maximum. $\operatorname{Mode}(X) = \operatorname{argmax}_{x \in X} P(X = x)$

\implies the value of c that maximizes $P(X = c)$ is Mode(X). Hence $\operatorname{Mode}[X] = \operatorname{argmax}_{c \in R} P(X = c)$

Problem 5**5.1**

(a) From definition of expectation, $E[a + bX + cY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (a + bX + cY)f(x, y)dxdy$

$$E[a + bX + cY] = a \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y)dxdy + b \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xf(x, y)dxdy + c \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} yf(x, y)dxdy$$

since, $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y)dxdy = 1$

$$\int_{-\infty}^{\infty} f(x, y)dy = f(x)$$

$$\int_{-\infty}^{\infty} f(x, y)dx = f(y)$$

$$\implies E[a + bX + cY] = a + b \int_{-\infty}^{\infty} xf(x)dx + c \int_{-\infty}^{\infty} yf(y)dy$$

$$\implies E[a + bX + cY] = a + bE[X] + cE[Y]$$

$$(b) \quad E[Y|X = x] = \int_{-\infty}^{\infty} yf(y|x)dy$$

$E[Y|X = x]$ is the expected values of Y given that a certain set of $X = x$ is known to occur.

$$(c) \quad \operatorname{var}[Y|X = x] = E[(Y - E[Y|X])^2|X]$$

$\operatorname{var}[Y|X = x]$ is the variance of Y given that a certain set of $X = x$ is known to occur.

$$(d) \quad \operatorname{cov}(X, Y) = E[(X - EX)(Y - EY)]$$

(e) From (d), $\operatorname{cov}(X, Y) = E[XY - XEY - YEX + EXEY]$

$$\operatorname{cov}(X, Y) = E[XY] - EXEY - EYEX + EXEY$$

$$\operatorname{cov}(X, Y) = E[XY] - EXEY$$

From the above results, $\operatorname{cov}(X, X) = E[X^2] - E[X]^2 = \operatorname{var}(X)$

$$\begin{aligned} \text{(f)} \quad \text{cov}(bX, cY) &= E[bXcY] - E[bX]E[cY] \\ \text{cov}(bX, cY) &= bcE[XY] - bcE[X]E[Y] = bc(E[XY] - E[X]E[Y]) = bc\text{Cov}(X, Y) \end{aligned}$$

$$\begin{aligned} \text{(g)} \quad \text{var}(a + bX + cY) &= \text{var}(bX + cY) = E[(bX + cY)^2] - E[(bX + cY)]^2 \\ \text{By expanding and grouping terms we get,} \\ \text{var}(a + bX + cY) &= b^2[E[X^2] - E[X]^2] + c^2[E[Y^2] - E[Y]^2] + 2bc(E[XY] - E[X]E[Y]) \\ \text{var}(a + bX + cY) &= b^2\text{var}(X) + c^2\text{var}(Y) + 2bc\text{Cov}(X, Y) \end{aligned}$$

$$\begin{aligned} \text{(h)} \quad \text{cov}(Y + X, Z) &= E[(Y + X)Z] - E[Y + X]E[Z] \\ \text{cov}(Y + X, Z) &= E[YZ + XZ] - E[Y]E[Z] - E[X]E[Z] \\ \text{cov}(Y + X, Z) &= E[YZ] - E[Y]E[Z] + E[XZ] - E[X]E[Z] \\ \text{cov}(Y + X, Z) &= \text{cov}(Y, Z) + \text{cov}(X, Z) \end{aligned}$$

$$\text{(i)} \quad \text{cor}(X, Y) = \frac{\text{cov}(X, Y)}{SD(X)SD(Y)}$$

$$\begin{aligned} \text{(j)} \quad \text{cor}(a + bX, c + dY) &= \frac{\text{cov}(a + bX, c + dY)}{SD(a + bX)SD(c + dY)} \\ \text{cov}(a + bX, c + dY) &= E[b(X - EX)d(Y - EY)] = bdcov(X, Y) \\ SD(a + bX) &= |b|SD(X) \\ SD(c + dY) &= |d|SD(Y) \\ \text{cor}(a + bX, c + dY) &= \frac{bdcov(X, Y)}{|bd|SD(X)SD(Y)} \end{aligned}$$

5.2(a)

Joint distribution is the probability of two events occurring together. $P(X, Y) = P(X \cap Y)$

5.2(b)

```
# load the data set
income = read.csv("/Users/anuram/Library/Mobile Documents/com~apple~CloudDocs/MS Stats/Winter 2023/STAT1
# setting the column name and row name of the dataframe
# names(income)[1] = 'y'
colnames(income) = c("y", "0.5", "1.5", "2.5", "3.5", "4.5",
                    "5.5", "6.7", "8.8", "12.5", "17.5")
rownames(income) = income$y
income = income[, -1]

# PMF of Y
PMF_Y = rowSums(income)

# marginal distribution of Y table
dist_Y = data.frame(as.list(rowSums(income)))
colnames(dist_Y) = rownames(income)
rownames(dist_Y) = "P(Y=y)"
dist_Y

##           0.5   0.4   0.25  0.15  0.05      0 -0.05 -0.18 -0.25
## P(Y=y) 0.07 0.065 0.098 0.208 0.302 0.029 0.095  0.07 0.063
```



```

# PMF of X
PMF_X = numeric(ncol(income))
for (i in (1:ncol(income))) {
  PMF_X[i] = sum(income[i])
}

# marginal distribution of X table
dist_X = cbind(X = as.numeric(colnames(income)), `P(X=x)` = as.numeric(PMF_X))
dist_X

```

```

##           X P(X=x)
## [1,] 0.5 0.041
## [2,] 1.5 0.093
## [3,] 2.5 0.093
## [4,] 3.5 0.082
## [5,] 4.5 0.113
## [6,] 5.5 0.103
## [7,] 6.7 0.155
## [8,] 8.8 0.155
## [9,] 12.5 0.113
## [10,] 17.5 0.052

```

5.2(c) Conditional distribution of Y given X for all values of X=x

$$P(Y|X=x) = \frac{P(X=x, Y=y)}{P(X=x)} \forall x \in X$$

```

# Conditional distribution of Y given X=x for all value of
# X=x
cond_dist = data.frame(matrix(ncol = ncol(income), nrow = nrow(income)))
for (i in (1:ncol(income))) {
  cond_dist[i] = income[i]/PMF_X[i]
}
rownames(cond_dist) = rownames(income)
colnames(cond_dist) = colnames(income)
# conditional distribution table for all Y=y given X=x
cond_dist

```

```

##           0.5           1.5           2.5           3.5           4.5           5.5
## 0.5 0.02439024 0.11827957 0.07526882 0.07317073 0.044247788 0.04854369
## 0.4 0.02439024 0.02150538 0.06451613 0.08536585 0.088495575 0.06796117
## 0.25 0.04878049 0.06451613 0.04301075 0.08536585 0.088495575 0.10679612
## 0.15 0.04878049 0.09677419 0.09677419 0.14634146 0.141592920 0.19417476
## 0.05 0.24390244 0.24731183 0.35483871 0.37804878 0.362831858 0.28155340
## 0 0.31707317 0.13978495 0.00000000 0.02439024 0.008849558 0.00000000
## -0.05 0.02439024 0.12903226 0.11827957 0.06097561 0.106194690 0.15533981
## -0.18 0.04878049 0.08602151 0.13978495 0.07317073 0.079646018 0.07766990
## -0.25 0.21951220 0.09677419 0.10752688 0.07317073 0.079646018 0.06796117
##           6.7           8.8          12.5          17.5
## 0.5 0.05161290 0.05806452 0.12389381 0.07692308
## 0.4 0.05161290 0.05806452 0.07079646 0.13461538
## 0.25 0.12903226 0.12258065 0.11504425 0.11538462
## 0.15 0.27096774 0.34838710 0.21238938 0.38461538
## 0.05 0.30322581 0.25161290 0.37168142 0.13461538
## 0 0.00000000 0.00000000 0.00000000 0.00000000

```

```
## -0.05 0.10967742 0.09032258 0.03539823 0.05769231
## -0.18 0.05161290 0.05161290 0.05309735 0.03846154
## -0.25 0.03225806 0.01935484 0.01769912 0.05769231
```

5.2(d) Conditional expectation of Y given X for all values of X=x

$$E[Y|X = x] = \sum_y y P(Y|X = x) \forall x \in X$$

```
y = as.numeric(rownames(cond_dist))
cond_exp = numeric(ncol(income))
for (i in (1:ncol(income))) {
  cond_exp[i] = sum(y * cond_dist[i])
}
x = as.numeric(colnames(cond_dist))

# conditional expectation of Y given X=x for all values of
# X=x
condition_exp = cbind(X = x, `P(Y|X=x)` = cond_exp)
condition_exp
```

```
##           X      P(Y|X=x)
## [1,]  0.5 -0.01121951
## [2,]  1.5  0.06462366
## [3,]  2.5  0.04849462
## [4,]  3.5  0.09841463
## [5,]  4.5  0.07946903
## [6,]  5.5  0.08262136
## [7,]  6.7  0.11167742
## [8,]  8.8  0.12909677
## [9,] 12.5  0.15371681
## [10,] 17.5 0.16134615
```

Expectation of Y, $E[Y] = \sum_y y P(Y = y) = 0.0987$

```
# Expectation of Y, E[Y]
mean_y = sum(y * PMF_Y)
mean_y
```

```
## [1] 0.0987
```

$$E[E[Y|X]] = \sum_x E[Y|X = x] P(X = x) = 0.0987$$

```
# Expectation of Y, E[E[Y|X]]
exp_y = sum(cond_exp * PMF_X)
exp_y
```

```
## [1] 0.0987
```

Hence it's proved that $E[E[Y|X]] = E[Y] = 0.0987$

5.2(e) Best Linear predictor (BLP) of Y given X

BLP of Y given X is $Y = \alpha + \beta X$

$$\alpha = E[Y] - \frac{Cov(X,Y)}{var(X)} E[X] = 0.0432$$

$$\beta = \frac{Cov(X,Y)}{var(X)} = 0.0086$$

BLP of Y = 0.0432 + 0.0086X

```

# Expectation of X
x = as.numeric(colnames(cond_dist))
mean_x = sum(x * PMF_X)
mean_x

## [1] 6.4795

# variance of X
exp_xsq = sum(x^2 * PMF_X)
var_x = exp_xsq - (mean_x^2)
var_x

## [1] 17.76773

# Cov(X,Y) = E[XY] - E[X]E[Y]
exp_xy = 0
for (i in (1:ncol(income))) {
  for (j in (1:nrow(income))) {
    exp_xy = exp_xy + x[i] * y[j] * income[j, i]
  }
}
cov_xy = exp_xy - (mean_x * mean_y)
cov_xy

## [1] 0.1520084

# intercept alpha
alpha = mean_y - (cov_xy * mean_x/var_x)
alpha

## [1] 0.0432659

# slope beta
beta = cov_xy/var_x
beta

## [1] 0.008555305

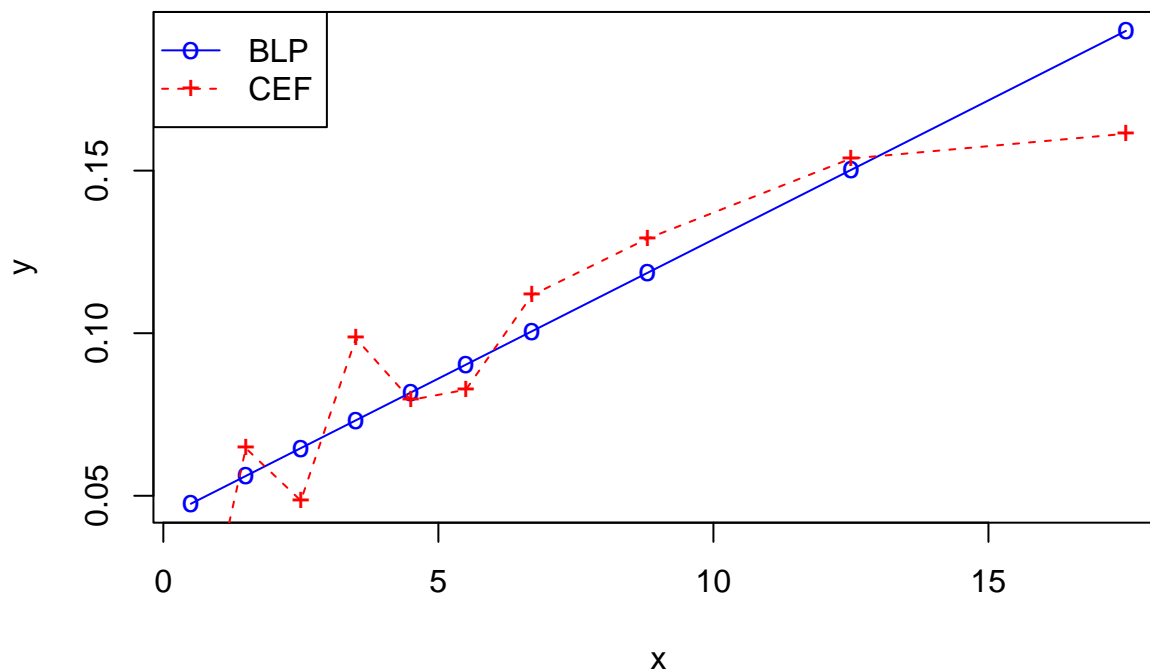
# Y from Best Linear Predictor
y_blp = alpha + beta * x
blp = cbind(X = x, Y_blp = y_blp)
blp

##           X      Y_blp
## [1,]  0.5 0.04754355
## [2,]  1.5 0.05609886
## [3,]  2.5 0.06465416
## [4,]  3.5 0.07320947
## [5,]  4.5 0.08176477
## [6,]  5.5 0.09032008
## [7,]  6.7 0.10058644
## [8,]  8.8 0.11855259
## [9,] 12.5 0.15020721
## [10,] 17.5 0.19298374

```

5.2(f) Plot of BLP and CEF(=E[Y|X])

```
# Plotting BLP and CEF together plot x vs BLP in blue line
plot(x, y_blp, type = "o", col = "blue", pch = "o", ylab = "y",
     xlab = "x", lty = 1)
# adding x vs CEF in red line to the previous plot
points(x, cond_exp, col = "red", pch = "+")
lines(x, cond_exp, col = "red", lty = 2)
legend(x = "topleft", legend = c("BLP", "CEF"), col = c("blue",
  "red"), pch = c("o", "+"), lty = c(1, 2), ncol = 1)
```



Problem 6

6(a)

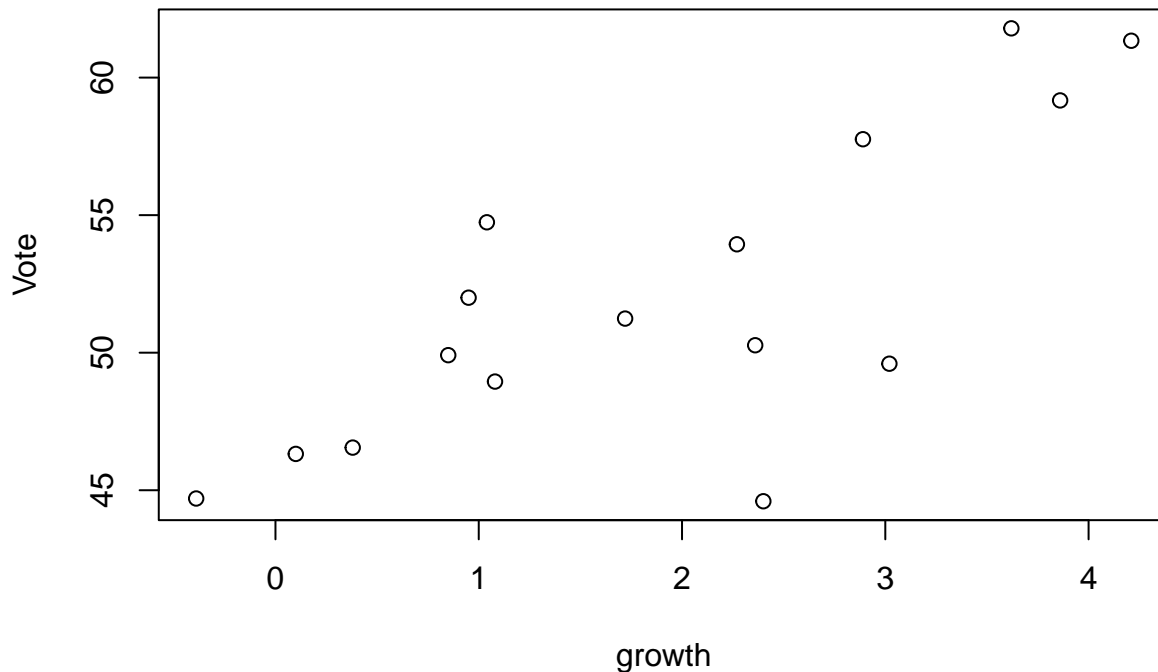
Yes, there could be a positive correlation between income growth and incumbent vote%. As the income increases the confidence in the incumbent party's economic policy might get stronger and hence they might get higher share of vote.

6(b) Scatter plot of vote Vs growth

```
election = read.delim("/Users/anuram/Library/Mobile Documents/com~apple~CloudDocs/MS Stats/Winter 2023/
  sep = " ")

# Scatter plot of vote vs growth
plot(election$growth, election$vote, main = "Scatter plot of Vote Vs Growth",
     xlab = "growth", ylab = "Vote")
```

Scatter plot of Vote Vs Growth



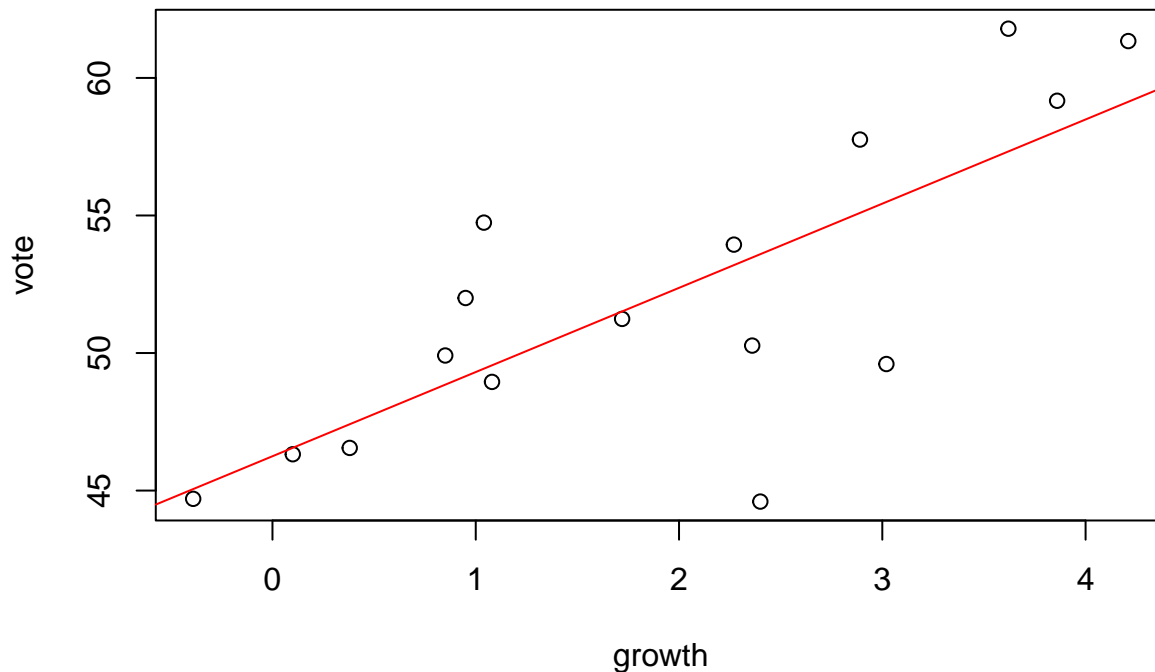
Simple linear regression model of Y on X Intercept: 46.25 Slope: 3.06

```
fit = lm(election$vote ~ election$growth)
summary(fit)
```

```
##
## Call:
## lm(formula = election$vote ~ election$growth)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -8.9929 -0.6674  0.2556  2.3225  5.3094
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   46.2476     1.6219  28.514 8.41e-14 ***
## election$growth  3.0605     0.6963   4.396 0.00061 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.763 on 14 degrees of freedom
## Multiple R-squared:  0.5798, Adjusted R-squared:  0.5498
## F-statistic: 19.32 on 1 and 14 DF,  p-value: 0.00061
```

The regression line seems to predict the trend of the data (ie, positive correlation between vote and growth), but adjusted R-squared is 0.55.

```
# plot with regression line
plot(election$vote ~ election$growth, ylab = "vote", xlab = "growth")
abline(fit, col = "red")
```



6(c) Regression model summary

The predicted regression model is $Vote = 46.25 + 3.06 * Growth$

From the regression line, it can be interpreted that as growth in income increases by 1% the incumbent party's vote percentage increases by 3.06%.

The estimated regression coefficients are : Intercept: 46.25 This means that when there is no income growth in the previous years, the incumbent party's vote percentage on average would be 46.25%

Slope:3.06 Since the slope is positive, it indicates a positive linear relationship between income growth and incumbent party's vote %. Since slope is defined as $\frac{\text{change in } y}{\text{change in } x}$, if income grows by 1% in the previous years, it can be predicted that the incumbent party's vote percentage would increase by 3.06%.

6(d) Prediction when average income growth is 2%

Given: if average income growth is 2%, from the regression model the incumbent party's vote percentage is 52.37%.

$$\text{Vote\%} = 46.25 + 3.06 * 2 = 52.37$$

The actual vote% for the incumbent party was 51.1% in 2016 when the income growth was 2%. Hence the linear regression model is similar to the actual vote% observed.