# Imprecise Linear Discriminant Analysis based on robust Bayesian inference

27èmes rencontres francophones sur la logique floue et ses applications

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# **Overview Imprecise Linear Discriminant Analysis Classifica tion**

- Classification
  - Motivation
  - Decision Making
  - Discriminant Analysis
- Imprecise Classification
  - Imprecise Linear discriminant analysis
  - Cautious Decision
- Evaluation
  - Cautious accuracy measure
  - Experiments
- Conclusions





#### **Overview**

#### Classification

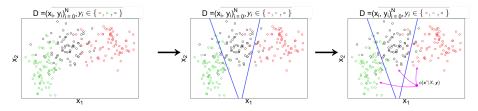
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### **Classification - Outline (Example)**

- Data training  $D = \{x_i, y_i\}_{i=0}^N$  such that :
  - (Input)  $x_i \in \mathcal{X}$  are regressors or features (often  $x_i \in \mathbb{R}^p$ ).
  - (Output)  $y_i \in \mathcal{K}$  is a response category variable, with  $\mathcal{K} = \{ m_1, ..., m_K \}$



#### **Objective**

Given training data  $D = \{x_i, y_i\}_{i=0}^N$ , we need to learn a classification rule :  $\phi : \mathcal{X} \to \mathcal{Y}$  in order to predict a new observation  $\phi(\mathbf{x}^*)$ 

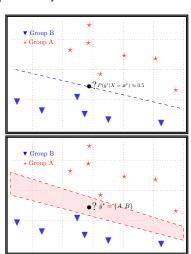




# **Motivation**What is the bigger problem in (precise) Classification?

 Precise models can produce many mistakes for hard to predict unlabeled instances.

 One way to recognize such instances and avoid making such mistakes too often → Making a cautious decision.





#### **Decision Making in Statistic**

1. In Statistic : An optimal model for classification under 1/0 loss function is Bayes Classifier :

$$\phi(\mathbf{x}^*) := \underset{m_k \in \mathcal{K}}{\operatorname{arg\,max}} P(y = m_k | X = \mathbf{x}^*)$$
 (1)

2. Preference ordering:

## Definition (Preference ordering [3, pp. 47])

Let P a conditional probability distribution,  $m_a$  is preferred to  $m_b$ :

$$m_a > m_b \iff P(y = m_a | X = \boldsymbol{x}^*) > P(y = m_b | X = \boldsymbol{x}^*)$$

We then take the *maximal element* of the complete order > :

$$m_{i_K} > m_{i_{K-1}} > .... > m_{i_1} \iff P(y = m_{i_K} | \boldsymbol{x}^*) > .... > P(y = m_{i_1} | \boldsymbol{x}^*)$$





#### **Decision Making in Statistic**

$$m_{i_K} > m_{i_{K-1}} > \dots > m_{i_1} \iff$$

$$P(y = m_{i_K} | \mathbf{x}^*)$$

$$V$$

$$P(y = m_{i_K} | \mathbf{x}^*)$$

$$V$$

$$\dots$$

$$V$$

$$P(y = m_{i_1} | \mathbf{x}^*)$$

$$m_{i_1}$$

• How can we estimate the cond. probability distribution  $P_{Y|X}$ ?



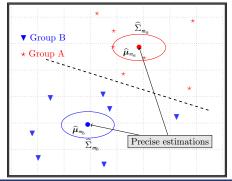


#### (Precise) Linear Discriminant Analysis

Applying Baye's rules to  $P(Y = m_a | X = \mathbf{x}^*)$ :

$$P(y = m_k | X = \mathbf{x}^*) = \frac{P(X = \mathbf{x}^* | y = m_k) P(y = m_k)}{\sum_{m_l \in \mathcal{X}} P(X = \mathbf{x}^* | y = m_l) P(y = m_l)}$$

Normality  $P_{X|Y=m_k} \sim \mathcal{N}(\mu_k, \Sigma_k)$  and precise marginal  $\pi_k := P_{Y=m_k}$ .





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### **Imprecise Linear Discriminant Analysis**

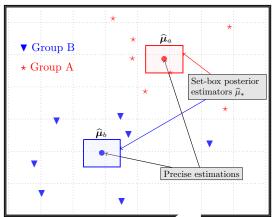
**Objective :** Making imprecise the parameter mean  $\mu_k$  of each Gaussian distribution family  $\mathscr{G}_k := P_{X|Y=m_k} \sim \mathcal{N}(\mu_k, \widehat{\Sigma})$ 

**Proposition:** Using a set of posterior distribution  $\mathcal{P}$  ([4, eq 17]).

$$\mathcal{P}_{\mu_k} = \left\{ \begin{aligned} &\mu_k \left| \ell \propto \mathcal{N} \left( \frac{\ell + n_k \overline{\mathbf{x}}_{n_k}}{n_k}, \frac{1}{n_k} \widehat{\boldsymbol{\Sigma}} \right) \right. \\ &\overline{\mathbf{x}}_k = \frac{1}{n_k} \sum_{i=1}^{n_k} x_{i,k}, \ell \in \mathbb{L} \end{aligned} \right.$$

where  $\mathbb{L}$  is a hypercube.

$$\mathbb{L} = \left\{ \begin{aligned} \ell \in \mathbb{R}^d : \ell_i \in [-c_i, c_i], \\ c_i > 0, i = \{1, ..., d\} \end{aligned} \right\}$$





#### **Decision Making in Imprecise Probabilities**

### **Definition (Partial Ordering by Maximality [1])**

Let  $\mathscr{P}$  a set of probabilities then  $m_a$  is preferred to  $m_b$  if the cost of exchanging  $m_a$  with  $m_b$  if and only if :

$$m_a \succ_M m_b \iff \inf_{P \in \mathscr{P}} \frac{P(y = m_a | X = \mathbf{x}^*)}{P(y = m_b | X = \mathbf{x}^*)} > 1$$
 (2)

• This definition give us a partial order  $\succ_M$ , such as the maximal element of partial order is the cautious decision :

$$Y_M = \{m_a \in \mathcal{K} \mid \not\exists m_b : m_a \succ_M m_b\}$$

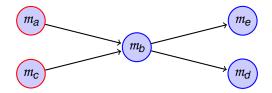




#### **Decision Making in Imprecise Probabilities**

For instance, if we have  $\mathcal{K} = \{m_a, m_b, m_c, m_d, m_e\}$ , we could have a set of comparisons from last definition :

$$\{m_a \succ_M m_b, m_c \succ_M m_b, m_a \succ \prec_M m_c, m_b \succ_M m_d, m_b \succ_M m_d, m_d \succ \prec_M m_e\}$$



Where cautious decision is a set-value categories :  $Y_M = \{m_a, m_c\}$ 



 We take to equation (2) of the criterion of maximality and applying Bayes' rule:

$$\inf_{P_{X|y} \in \mathscr{P}_{X|y}} \frac{P(\boldsymbol{x}^*|y = m_a)P(y = m_a)}{P(\boldsymbol{x}^*|y = m_b)P(y = m_b)} > 1$$

• Assuming an precise estimation for marginal :  $\pi_k := \widehat{P}_{Y=m_k}$  of LDA.

$$\frac{\widehat{P}(y=m_a)}{\widehat{P}(y=m_b)} \inf_{P_{X|y} \in \mathscr{P}_{X|y}} \frac{P(\mathbf{x}^*|y=m_a)}{P(\mathbf{x}^*|y=m_b)} > 1$$

• Given that normality assumption  $P_{X|Y=m_k} \sim \mathcal{N}(\mu_k, \widehat{\Sigma})$  of LDA, each component of numerator/denominator are independent, then :

$$\frac{\widehat{P}(y=m_a)}{\widehat{P}(y=m_b)} \quad \frac{\inf_{P_{X|y=m_a} \in \mathscr{D}_{X|y=m_b}} P(\boldsymbol{x}^* | y=m_a)}{\sup_{P_{X|y=m_b} \in \mathscr{D}_{X|y=m_b}} P(\boldsymbol{x}^* | y=m_b)} > 1$$





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$$\frac{\widehat{P}(y=m_a)}{\widehat{P}(y=m_b)} \quad \frac{\inf_{P_{X|y=m_a} \in \mathscr{P}_{X|y=m_a}} P(\boldsymbol{x}^*|y=m_a)}{\sup_{P_{X|y=m_b} \in \mathscr{P}_{X|y=m_b}} P(\boldsymbol{x}^*|y=m_b)} > 1$$



 We then have two optimization: box-constrained quadratic problem (BQP) and non-convex BQP (NBQP)

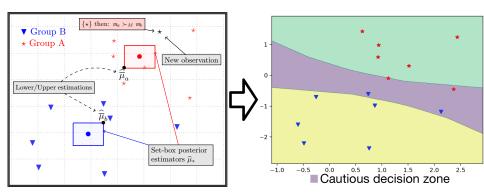
$$\inf_{P \in \mathscr{P}_{\mu_a}} P(\boldsymbol{x}^* | y = m_a) \iff \underline{\mu}_a = \underset{\mu_a \in \mathscr{P}_{\mu_a}}{\operatorname{arg \, min}} - \frac{1}{2} (\boldsymbol{x}^* - \mu_a)^T \widehat{\Sigma}^{-1} (\boldsymbol{x}^* - \mu_a) \text{ (NBQP)}$$

$$\sup_{P \in \mathscr{P}_{\mu_b}} P(\boldsymbol{x}^* | y = m_b) \iff \overline{\mu}_b = \underset{\mu_b \in \mathscr{P}_{\mu_a}}{\operatorname{arg \, max}} - \frac{1}{2} (\boldsymbol{x}^* - \mu_b)^T \widehat{\Sigma}^{-1} (\boldsymbol{x}^* - \mu_b) \text{ (BQP)}$$

• First problem non-convex BQP  $\rightarrow$  solved through Branch and Bound method.

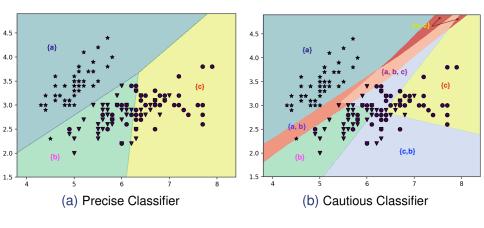


#### Cautious decision zone (example with 2 class)





#### **Another Example with 3 class** {*a*, *b*, *c*}





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#### **Utility-discount accuracy measure**

**Example:** Given a ground-truth category  $\{m_a\}$  and a binary cautious prediction  $\{m_a, m_b\}$ , how should we reward cautious prediction?

Case 1: If we reward it 0.5:

Cautious accuracy measure Experiments

Reward = 
$$\frac{1}{|\{m_a, m_b\}|}$$

 $\implies$  Equivalent to randomly pick out  $m_b$  or  $m_a$ .



- → too low reward confusing cautiousness and randomness.



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- → too low reward confusing cautiousness and randomness.
- Case 2: If we reward it with 1, then classifier always returning all classes is the best  $\rightarrow$  too high reward to cautiousness, no penalty for non-informativeness.
- → Actual reward should be between those two cases, and depend on how much we are cautiousness seeking.



### Utility-discount accuracy measure

Cautious accuracy measure Experiments

Zaffalon et al in [2] proposes an utility-discounted accuracy measure:

$$u(y, \widehat{Y}_M) = \begin{cases} 0 & \text{if } y \notin \widehat{Y}_M \\ \frac{\alpha}{|\widehat{Y}_M|} - \frac{1-\alpha}{|\widehat{Y}_M|^2} & \text{else} \end{cases}$$

Where  $u(\cdot,\cdot)$  rewards to be more informative in cautious decision:  $u_{65}$  with a gain of 0.65 and  $u_{80}$  with a gain of 0.8.

### **Experiments**

Cautious accuracy measure Experiments

Cross-Validation with 60% training data and 40% validation data, it is repeated with 10 resampling.

#	Name	# Obs.	# Regr.	# Classes
а	iris	150	4	3
b	seeds	210	7	3
С	glass	214	9	6

TABLE – Data sets used in the experiments

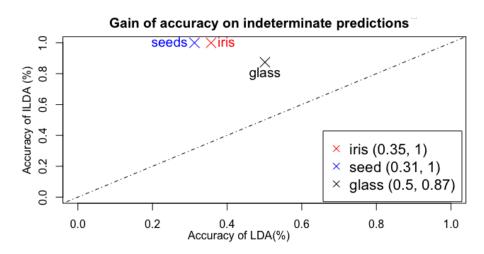
	DLA	IDLA		Inference
#	DLA	<i>u</i> <sub>65</sub>	<i>u</i> <sub>80</sub>	time
а	0.961	0.969	0.975	0.56 sec.
b	0.959	0.959	0.962	1.50 sec.
С	0.594	0.589	0.642	8.66 sec

TABLE – Average utility-discounted accuracies



### **Experiments**

Cautious accuracy measure Experiments





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# **Conclusions and Perspectives Imprecise Linear Discriminant Classification Analysis**

- Increasing in imprecision on the estimators has allowed us to be more cautious in doubt and to improve the prediction of classification.
- More experiments with all imprecise components.
- Creation of new imprecise statistic models for a sensibility analysis and a more (cautious) robust prediction.



#### References



Matthias CM TROFFAES. "Decision making under uncertainty using imprecise probabilities". In: International Journal of Approximate Reasoning 45.1 (2007), p. 17-29.



Marco ZAFFALON, Giorgio CORANI et Denis MAUÁ. "Evaluating credal classifiers by utility-discounted predictive accuracy". In: International Journal of Approximate Reasoning 53.8 (2012), p. 1282-1301.



James O BERGER. Statistical decision theory and Bayesian analysis. Springer Science & Business Media, 2013.



Alessio BENAVOLI et Marco ZAFFALON. "Prior near ignorance for inferences in the k-parameter exponential family". In: Statistics 49.5 (2014), p. 1104-1140.







