Lecture 8: Causality in Time Series and Panel Data Models

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EESP-FGV

Econometrics 2

Administrative

- Optional Reading: Rambachan and Shephard [2021], Bojinov et al. [2021] and Heckman et al. [2016]
- Even though there are only optional readings, you should read my lecture notes carefully.

Outline

1. Introduction

2. When Do Common Time Series Estimands Have a Causal Meaning?

3. Dynamic Treatment Effects

- Hamilton's textbook is quite old.
- It predates two important changes in economic research.
 - 1. Micro-founded Macroeconomics
 - 2. Credibility Revolution

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- It predates two important changes in economic research.
 - 1. Micro-founded Macroeconomics
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- This lecture is an attempt to connect the classic time series and panel data models with the modern literature on causality.

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 - 2. Local Projections.

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- We will focus on two time series estimands that are closely connected to the ones we saw during the course.
 - 1. Impulse Response Functions.
 - 2. Local Projections.
- This paper is very readable and we will go through some of their proofs.

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 Panel Experiments.
- It illustrates that even experiments can have interesting econometric problems.
 - For more problems like this, check Susan Athey's interview on *Mixtape: The Podcast*: https://open.spotify.com/episode/226ISyTNqr7jTwaCafcX1S?si= 01173248d1724962.

- Their approach to uncertainty is completely different from what you were exposed to in Econometrics 1 and 2.
 - They focus on design-based uncertainty instead of sampling uncertainty.
 - Consequently, they do not rely on asymptotic approximations. They derive inferential methods based on a finite population perspective.

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 - They focus on design-based uncertainty instead of sampling uncertainty.
 - Consequently, they do not rely on asymptotic approximations. They derive inferential methods based on a finite population perspective.
- Since a lot of things are new for you here, I will just emphasize two of their main points.
 - Inverse Probability Weighting Estimators work.
 - Fixed effects models may be problematic.

• If you are interested in this type of work, we can talk more about relevant readings.

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- It can still identify many interesting economic objects related to structural models but it does not require strong parametric assumptions.
- I always learn something new when I read this paper.
- Since it is a lot to digest, I will focus on two topics.
 - Defining causal parameters based on a flexible economic model.
 - Classic IV estimands are hard to interpret.

When Do Common Time Series

Meaning?

Estimands Have a Causal

Outline

Introduction

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- Single unit (e.g., macroeconomy or market) observed over time.
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- Single unit (e.g., macroeconomy or market) observed over time.
- At each time period $t \ge 1$, the unit receives a vector of assignments W_t and an associated vector of outcomes Y_t are generated.
- The outcomes are causally related to the assignments through a potential outcome process.

Definition: Potential Outcome Process

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Intuition: Marvel's Multiverse (*Loki* and *What if?* are particularly useful to understand potential outcomes.)

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• For each time period $t \ge 1$ and assignment paths $\{w_s\}_{s \ge 1}$ and $\{w_s'\}_{s \ge 1}$, $Y_t\left(\{w_s\}_{s \ge 1}\right) - Y_t\left(\{w_s'\}_{s \ge 1}\right)$ is a random variable.

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- We will focus on moments of this random variable.

Advantages:

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- No restriction on the extent to which past assignments may causally affect outcomes.
- Most leading econometric models used to study dynamic causal effects in time series settings (e.g., Structural VAR) can be cast as special cases of the direct potential outcome system.
- The potential outcome system provides a flexible, nonparametric foundation upon which to analyze dynamic causal effects in time series settings.

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 - Impulse Response Functions and Local Projections identify, respectively, a dynamic average treatment effect and a weighted average of marginal treatment effects if the assignments are independent of the potential outcome process and over time.
- Case2: The researcher only observes Y_t .
 - Researchers can recover familiar model-based approaches by introducing functional form restrictions on the potential outcome process.

There is a single unit. At each time period $t \geq 1$, the unit receives a d_w -dimensional assignment $\{W_t\}_{t\geq 1}$. Associated with this assignment process, we observe a d_y -dimensional outcome $\{Y_t\}_{t\geq 1}$.

Assumption 1 (Assignment and Potential Outcome)

The assignment process $\{W_t\}_{t\geq 1}$ satisfies $W_t\in\mathcal{W}:=X_{k=1}^{d_w}\mathcal{W}_k\subseteq\mathcal{R}^{d_w}$. The potential outcome process is, for any deterministic sequence $\{w_s\}_{s\geq 1}$ with $w_s\in\mathcal{W}$ for all $s\geq 1,$ $\Big\{Y_t\left(\{w_s\}_{s\geq 1}\right)\Big\}_{t\geq 1}$, where the time-t potential outcome satisfies $Y_t\left(\{w_s\}_{s\geq 1}\right)\in\mathcal{Y}\subseteq\mathbb{R}^{d_y}$.

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For each $t \geq 1$ and all deterministic $\{w_t\}_{t \geq 1}$ and $\{w_t'\}_{t \geq 1}$ with $w_t, w_t' \in \mathcal{W}$,

$$Y_t\left(\mathbf{w_{1:t}}, \{\mathbf{w_s}\}_{s \geq t+1}\right) = Y_t\left(\mathbf{w_{1:t}}, \{\mathbf{w_s'}\}_{s \geq t+1}\right).$$

Under Assumption 2, we drop references to the future assignments in the potential outcome process and write $\left\{Y_t\left(\left\{w_s\right\}_{s\geq 1}\right)\right\}_{t\geq 1}=\left\{Y_t\left(w_{1:t}\right)\right\}_{t\geq 1}$.

Assumption 3 (Output)

The output is $\{W_t, Y_t\}_{t\geq 1} = \{W_t, Y_t(W_{1:t})\}_{t\geq 1}$. The $\{Y_t\}_{t\geq 1}$ is called the outcome process.

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The Fundamental Problem of Causal Inference is that we only observe the potential outcome evaluated at the observed assignment path.

Assumption 4 (Sequentially Probabilistic Assignment Process)

The assignment process satisfies $0<\mathbb{P}\left[\left.W_{t}=w\right|\left.W_{1:t-1},\,Y_{1:t-1}\right]<1$ for any $w\in\mathcal{W}.$

For each time period $t \geq 1$ and assignment paths $w_{1:t}$ and $w'_{1:t}$,

$$Y_t\left(w_{1:t}\right) - Y_t\left(w_{1:t}'\right)$$

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- We focus on causal parameters that average over these dynamic causal effects along various underlying assignment paths.

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- Enormous class of dynamic causal effects.
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- Shorthand Notation: For $t \ge 1$, $h \ge 0$ and any $w \in \mathcal{W}$, write the time-(t+h) potential outcome at the assignment process $(W_{1:t-1}, w, W_{t+1:t+h})$ as

$$Y_{t+h}(w) = Y_{t+h}(W_{1:t-1}, w, W_{t+1:t+h}).$$

Definition 1 (Impulse Causal Effect)

For $t \geq 1$, $h \geq 0$ and any $w, w' \in \mathcal{W}$, the time-t, h-period ahead impulse causal effect is

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Note: The time-t, h-period ahead impulse causal effect is a random variable.

Intuition: It measures the ceteris paribus causal effect of intervening to switch the time-t assignment from w' to w on the h-period ahead outcomes holding all else fixed along the assignment process.

Definition 2 (Average Treatment Effect)

For $t \geq 1$, $h \geq 0$ and any $w, w' \in \mathcal{W}$, the time-t, h-period ahead average treatment effect is

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Intuition: It measures the unconditional expectation of the impulse causal effect $Y_{t+h}(w) - Y_{t+h}(w')$.

We further define analogous versions of the dynamic causal effects for a particular scalar assignment. For any $w_k \in \mathcal{W}_k$, define

$$Y_{t+h}\left(w_{k}\right) = Y_{t+h}\left(W_{1:t-1}, W_{1:k-1,t}, w_{k}, W_{k+1:d_{W},t}, W_{t+1:t+h}\right).$$

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The corresponding time-t, h—period ahead impulse causal effect and average treatment effect for the k-th assignment are, respectively,

$$Y_{t+h}(w_k) - Y_{t+h}(w'_k)$$

and

$$\mathbb{E}\left[Y_{t+h}\left(w_{k}\right)-Y_{t+h}\left(w_{k}'\right)\right].$$

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Definition 3 (Marginal Impulse Causal Effect and Average Marginal Treatment Effect)

The time-t, h-period ahead Marginal Impulse Causal Effect and Average Marginal Treatment Effect are, respectively,

$$Y'_{t+h}(w_k) = \frac{\partial Y_{t+h}(w_k)}{\partial w_k}$$

and

$$\mathbb{E}\left[Y'_{t+h}\left(w_{k}\right)\right].$$

From now on, we assume that the outcome Y_t is univariate. There is no loss in generality in doing so because the more general case is covered by running the analysis equation by equation.

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Definition 4 (Impulse Response Function)

For $h \geq 0$ and deterministic $w_k, w_k' \in \mathcal{W}$, the impulse response function is defined by

$$IRF_{k,t,h}\left(w_{k},w_{k}'\right) := \mathbb{E}\left[\left.Y_{t+h}\right|W_{k,t} = w_{k}\right] - \mathbb{E}\left[\left.Y_{t+h}\right|W_{k,t} = w_{k}'\right].$$

Definitions

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Definition 5 (Local Projection)

For $h \ge 0$, the Local Projection estimand is defined by

$$LP_{k,t,h} := \frac{\operatorname{Cov}(Y_{t+h}, W_{k,t})}{\operatorname{Var}(W_{k,t})}.$$

Theorem 6

Consider some
$$k = 1, ..., d_w, t \ge 1$$
, $h \ge 0$, fix $w_k, w_k' \in \mathcal{W}_k$ and that
$$\mathbb{E}\left[|Y_{t+h}(w_k) - Y_{t+h}(w_k')|\right] < \infty. \text{ Then, under Assumptions 1-4,}$$

$$IRF_{k,t,h}\left(w_k, w_k'\right) = \mathbb{E}\left[Y_{t+h}(w_k) - Y_{t+h}(w_k')\right] + \Delta_{k,t,h}\left(w_k, w_k'\right),$$

AverageTreatmentEffect

SelectionBias

Theorem 6

Consider some $k=1,\ldots,d_w,t\geq 1$, $h\geq 0$, fix $w_k,w_k'\in \mathcal{W}_k$ and that $\mathbb{E}\left[|Y_{t+h}\left(w_k\right)-Y_{t+h}\left(w_k'\right)|\right]<\infty$. Then, under Assumptions 1-4,

$$IRF_{k,t,h}\left(w_{k},w_{k}'\right) = \underbrace{\mathbb{E}\left[Y_{t+h}\left(w_{k}\right) - Y_{t+h}\left(w_{k}'\right)\right]}_{AverageTreatmentEffect} + \underbrace{\Delta_{k,t,h}\left(w_{k},w_{k}'\right)}_{SelectionBias},$$

where

$$\Delta_{k,t,h}(w_k,w_k') := \frac{\mathsf{Cov}(Y_{t+h}(w_k),1\{W_{k,t}=w_k\})}{\mathbb{E}[1\{W_{k,t}=w_k\}]} - \frac{\mathsf{Cov}(Y_{t+h}(w_k'),1\{W_{k,t}=w_k'\})}{\mathbb{E}[1\{W_{k,t}=w_k'\}]}.$$

Proof: Note that

$$\mathbb{E}[Y_{t+h}|W_{k,t} = w_k] = \frac{\mathbb{E}[Y_{t+h} \cdot 1\{W_{k,t} = w_k\}]}{\mathbb{E}[1\{W_{k,t} = w_k\}]}$$

$$= \frac{\mathbb{E}[Y_{t+h}(w_k) \cdot 1\{W_{k,t} = w_k\}]}{\mathbb{E}[1\{W_{k,t} = w_k\}]}$$

$$= \frac{\mathbb{E}[Y_{t+h}(w_k)] \cdot \mathbb{E}[1\{W_{k,t} = w_k\}] + \text{Cov}(Y_{t+h}(w_k), 1\{W_{k,t} = w_k\})}{\mathbb{E}[1\{W_{k,t} = w_k\}]}$$

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The result is then immediate by (i) applying the same calculation to $\mathbb{E}\left[\left.Y_{t+h}\right|\,W_{k,t}=w_k'\right]$ and (ii) taking the difference.

Our last theorem states that the impulse response function is equal to the average treatment effect if and only if the selection bias term

$$\Delta_{k,t,h}\left(w_{k},w_{k}'\right):=\frac{\mathsf{Cov}\left(Y_{t+h}\left(w_{k}\right),1\left\{W_{k,t}=w_{k}\right\}\right)}{\mathbb{E}\left[1\left\{W_{k,t}=w_{k}\right\}\right]}-\frac{\mathsf{Cov}\left(Y_{t+h}\left(w_{k}'\right),1\left\{W_{k,t}=w_{k}'\right\}\right)}{\mathbb{E}\left[1\left\{W_{k,t}=w_{k}'\right\}\right]}$$

is equal to zero.

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is equal to zero.

These covariace terms depend on how the assignment $W_{k,t}$ covaries with the potential outcome $Y_{t+h}(w_k) = Y_{t+h}(W_{1:t-1}, w_k, W_{-k,t}, W_{t+1:t+h})$. Consequently, the selection bias term depends on how the assingment $W_{k,t}$ relates to past assignments, other contemporaneous assignments, future assignments and the potential outcome process.

By placing further restrictions on the assignment process, we arrive at sufficient conditions for $\Delta_{k,t,h}(w_k,w_k')=0$. In particular, we impose that the assignment $W_{k,t}$ is randomized in the sense that it is independent of all other assignments and the time-(t+h) potential outcomes.

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We formalize this intuition in the next slide.

Theorem 7

Under Assumptions 1-4, if $Cov(Y_{t+h}(w_k), 1\{W_{k,t} = w_k\}) = 0$ and $Cov(Y_{t+h}(w_k'), 1\{W_{k,t} = w_k'\}) = 0$, then $\Delta_{k,t,h}(w_k, w_k') = 0$. Moreover, this condition is satisfied if

$$W_{k,t} \perp Y_{t+h}(w_k)$$
, and $W_{k,t} \perp Y_{t+h}(w'_k)$,

which is in turn implied by

$$W_{k,t} \perp \{Y_{t+h}(w_k) : w_k \in \mathcal{W}_k\},$$

which is in turn implied by

$$W_{k,t} \perp \left(W_{1:t-1}, W_{1:k-1,t}, W_{k+1:d_W,t}, W_{t+1:t+h}, \left\{Y_{t+h}\left(w_{1:t+h}\right) : w_{1:t+h} \in \mathcal{W}^{t+h}\right\}\right).$$

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- 3. Independence: $W_{k,t} \perp \{Y_{t+h}(w_k) : w_k \in \mathcal{W}_k\}$.

Then,

$$LP_{k,t,h} = \frac{\int_{\mathcal{W}_k} \mathbb{E}\left[Y'_{t+h}\left(w_k\right)\right] \mathbb{E}\left[G_t\left(w_k\right)\right] dw_k}{\int_{\mathcal{W}_k} \mathbb{E}\left[G_t\left(w_k\right)\right] dw_k},$$

Theorem 8

Under Assumptions 1-4, further assume that:

- 1. The support of $W_{k,t}$ is a closed interval, $\mathcal{W}_k := [\underline{w}_k, \bar{w}_k] \subset \mathbb{R}$.
- 2. Differentiability: $Y_{t+h}(w_k)$ is continuously differentiable in w_k , as is $\mathbb{E}\left[Y'_{t+h}(w_k)\right]$.
- 3. Independence: $W_{k,t} \perp \{Y_{t+h}(w_k) : w_k \in \mathcal{W}_k\}$.

Then,

$$LP_{k,t,h} = \frac{\int_{\mathcal{W}_k} \mathbb{E}\left[Y'_{t+h}\left(w_k\right)\right] \mathbb{E}\left[G_t\left(w_k\right)\right] dw_k}{\int_{\mathcal{W}_k} \mathbb{E}\left[G_t\left(w_k\right)\right] dw_k},$$

where $G_t(w_k) = 1\{w_k \leq W_{k,t}\} (W_{k,t} - \mathbb{E}[W_{k,t}])$, noting $\mathbb{E}[G_t(w_k)] \geq 0$.

Proof: First, note that

$$Y_{t+h} = Y_{t+h}(W_{k,t})$$

$$= Y_{t+h}(\underline{w}_k) + \int_{\underline{w}_k}^{W_{k,t}} \frac{\partial Y_{t+h}(\tilde{w}_k)}{\partial \tilde{w}_k} d\tilde{w}_k$$

by the Fundamental Theorem of Calculus

$$=Y_{t+h}(\underline{w}_k)+\int_{w_k}^{\overline{w}_k}\frac{\partial Y_{t+h}(\widetilde{w}_k)}{\partial \widetilde{w}_k}\cdot 1\{\widetilde{w}_k\leq W_{k,t}\}\ d\widetilde{w}_k. \tag{1}$$

Moreover, observe that

$$\begin{aligned} \mathsf{Cov}\left(Y_{t+h}, W_{k,t}\right) &= \mathbb{E}\left[Y_{t+h}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] \\ &= \mathbb{E}\left[Y_{t+h}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] - 0 \\ &= \mathbb{E}\left[Y_{t+h}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] - \mathbb{E}\left[Y_{t+h}\left(\underline{w}_{k}\right)\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] \\ &= \mathsf{because} \ W_{k,t} \perp \left\{Y_{t+h}\left(w_{k}\right) : w_{k} \in \mathcal{W}_{k}\right\} \\ &= \mathbb{E}\left[\left(Y_{t+h} - Y_{t+h}\left(\underline{w}_{k}\right)\right)\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] \\ &= \mathbb{E}\left[\left(\int_{\underline{w}_{k}}^{\overline{w}_{k}} \frac{\partial Y_{t+h}\left(\widetilde{w}_{k}\right)}{\partial \widetilde{w}_{k}} 1\left\{\widetilde{w}_{k} \leq W_{k,t}\right\} d\widetilde{w}_{k}\right)\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] \\ &= \mathsf{by} \ \mathsf{Equation} \ (1). \end{aligned}$$

Moreover, observe that

$$\mathsf{Cov}\left(Y_{t+h}, W_{k,t}\right) = \mathbb{E}\left[\left(\int_{\underline{w}_{k}}^{\overline{w}_{k}} \frac{\partial Y_{t+h}\left(\widetilde{w}_{k}\right)}{\partial \widetilde{w}_{k}} 1\left\{\widetilde{w}_{k} \leq W_{k,t}\right\} d\widetilde{w}_{k}\right) \left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right]$$

$$= \int_{\underline{w}_{k}}^{\overline{w}_{k}} \mathbb{E}\left[\frac{\partial Y_{t+h}\left(\widetilde{w}_{k}\right)}{\partial \widetilde{w}_{k}} 1\left\{\widetilde{w}_{k} \leq W_{k,t}\right\} \left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] d\widetilde{w}_{k}$$

because our differentiability assumption ensures

that we can apply Fubini's Theorem

$$= \int_{\underline{w}_{k}}^{\overline{w}_{k}} \mathbb{E}\left[\frac{\partial Y_{t+h}\left(\tilde{w}_{k}\right)}{\partial \tilde{w}_{k}}\right] \mathbb{E}\left[1\left\{\tilde{w}_{k} \leq W_{k,t}\right\}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right] d\tilde{w}_{k}$$
because $W_{k,t} \perp \left\{Y_{t+h}\left(w_{k}\right) : w_{k} \in \mathcal{W}_{k}\right\}$.

Analogously, we have that

$$\operatorname{Var}\left(W_{k,t}
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$$\mathbb{E}\left[1\left\{W_{k,t} \geq \tilde{w}_{k}\right\}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)\right]$$

$$= \mathbb{E}\left[1\left\{W_{k,t} \geq \tilde{w}_{k}\right\} W_{k,t}\right] - \mathbb{E}\left[1\left\{W_{k,t} \geq \tilde{w}_{k}\right\}\right] \mathbb{E}\left[W_{k,t}\right]$$

$$= \left(\mathbb{E}\left[W_{k,t} \mid W_{k,t} \geq \tilde{w}_{k}\right] - \mathbb{E}\left[W_{k,t}\right]\right) \mathbb{P}\left(W_{k,t} \geq \tilde{w}_{k}\right) \geq 0$$

since $\mathbb{E}[W_{k,t} \mid W_{k,t} \geq \tilde{w}_k] \geq \mathbb{E}[W_{k,t}]$ for $\tilde{w}_k \in [\underline{w}_k, \bar{w}_k]$.

According to our last theorem, we know that

$$LP_{k,t,h} = \frac{\int_{\mathcal{W}_k} \mathbb{E}\left[Y'_{t+h}\left(w_k\right)\right] \mathbb{E}\left[G_t\left(w_k\right)\right] dw_k}{\int_{\mathcal{W}_k} \mathbb{E}\left[G_t\left(w_k\right)\right] dw_k}.$$

Intuitively, the local projection estimand is a weighted average of average marginal treatment effects of $W_{k,t}$ on Y_{t+h} , where the weights are non-negative and sum to one.

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Intuitively, the local projection estimand is a weighted average of average marginal treatment effects of $W_{k,t}$ on Y_{t+h} , where the weights are non-negative and sum to one.

Thus, if the assignment $W_{k,t}$ is an exogenous shock, the local projection estimand has a nonparametric causal interpretation.

Food for Thought: Are the weights,

$$\frac{1\left\{\tilde{w}_{k} \leq W_{k,t}\right\}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right)}{\int_{\mathcal{W}_{k}} 1\left\{\check{w}_{k} \leq W_{k,t}\right\}\left(W_{k,t} - \mathbb{E}\left[W_{k,t}\right]\right) d\check{w}_{k}},$$

policy-relevant? Do the weights capture a economically meaningful parameter?

Interpreting Local Projections

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policy-relevant? Do the weights capture a economically meaningful parameter?

In some contexts, linearity may have relevant costs.

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- Researchers introduce parametric models to study the dynamic causal effects of unobservable structural shocks.
- These structural shocks must be inferred from the outcomes.
- Rambachan and Shephard [2021] illustrate that this approach can be nested in the direct potential outcome system framework.

Focusing on the one lag model with no intercept for brevity, the Structural VAR approach assumes that the potential outcome process satisfies

$$A_0 Y_t (w_{1:t}) = w_t + A_1 Y_{t-1} (w_{1:t-1}),$$

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In this case, the h-period ahead average marginal treatment effect is

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- $(A_0^{-1}A_1)$ can be determined from the dynamics of the observable outcomes if this process is stationary.
- A_0 and $Var(W_t)$ cannot be separately identified from the observable outcomes. So further structural assumptions are needed.

Outline

1. Introduction

2. When Do Common Time Series Estimands Have a Causal Meaning

• Heckman et al. [2016] develops models for estimating and interpreting treatment effects arising from both ordered and unordered multi-stage decision problems.

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 effects arising from both ordered and unordered multi-stage decision problems.
 (We will focus on the ordered case.)
- Identification is secured through instrumental variable and conditional independence assumptions.
- They decompose treatment effects into direct effects and continuation values associated with moving to the next stage of a decision problem.
- IV estimand does not capture an economically interpretable parameter in dynamic discrete choice models unless policy variables are instruments.

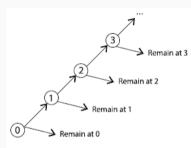


Fig. 1. An ordered multi-stage dynamic decision model.

- This figure illustrates one ordered decision problem.
- The stages could correspond to a sequence of educational choices.
- All agents start at stage 0. Some transit to 1, while other stay at 0 forever, and some of those who go to 1 stop there while others go on, etc.
- At each stage, agents update their information sets and decide whether to transit to the next stage.
- Associated with each stage is a set of potential outcomes.

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- Heckman et al. [2016] do not take a position on the precise content of the information sets and preferences governing agent choices at different nodes or the exact decision rules used.
- Agents could make irrational choices and their choices could be governed by behavioral anomalies.
- Thus, the model does not rule out dynamic inconsistency (e.g., ex-post regret).

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- ullet We denote each potential outcome by Y^k_s for $k\in\mathcal{K},s\in\mathcal{J}$.
- The observed outcome Y^k is given by

$$Y^k = \sum_{s \in \mathcal{J}} D_s \cdot Y_s^k.$$

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- ullet Endowments heta are not directly observed by the econometrician but are proxied by a measurement system.
 - ullet If we interpret heta as cognitive skill, we can proxy it with test scores.

Outcomes are also represented by a separable model:

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Assumptions

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 θ captures all sources of dependence across transitions apart from X and Z. Along with the observed variables, it generates dependence between choices and outcomes.

Formally, we assume that, conditional on X,

$$u_j \perp \!\!\!\!\perp \nu_l \quad \forall l \neq j \quad l, j \in \{0, \dots, \bar{s} - 1\}$$
 $\omega_s^k \perp \!\!\!\perp \omega_{s'}^k \quad \forall s \neq s' \quad \forall k \text{ and } s, s' \in \mathcal{J}$
 $\omega_s \perp \!\!\!\!\perp \nu \quad \forall s \in \mathcal{J}$
 $\theta \perp \!\!\!\!\perp Z$
 $(\omega_s, \nu) \perp \!\!\!\!\perp (\theta, Z) \quad \forall s \in \mathcal{J}$

Heckman et al. [2016] define treatment effects that take into account the direct effect of transiting to the next node in a decision tree plus the benefits associates with the options opened up by the additional choices made possible by such transitions.

The person-specific treatment effect T_j^k for outcome $k \in \mathcal{K}$ for an individual selected from the population with $Q_j=1$ with characteristics $X=x,Z=z,\theta=\overline{\theta}$ is the difference between the individual's outcomes under two possible actions:

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$$T_{j}^{k} [Y_{k} | X = x, Z = z, \theta = \overline{\theta}]$$

$$:= (Y_{k} | X = x, Z = z, \theta = \overline{\theta}, Q_{j} = 1, \text{Fix } D_{j} = 0)$$

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where

• $(Y_k|X=x,Z=z,\theta=\overline{\theta},Q_j=1, \text{Fix }D_j=0)$ is the value of Y_k for a person with characteristics $X=x,Z=z,\theta=\overline{\theta}$ from the population that attains at least node j and for whom we fix $D_j=0$ so that the agent is forced to go on to the next node.

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- $(Y_k|X=x,Z=z,\theta=\overline{\theta},Q_j=1, \text{Fix }D_j=0)$ is the value of Y_k for a person with characteristics $X=x,Z=z,\theta=\overline{\theta}$ from the population that attains at least node j and for whom we fix $D_j=0$ so that the agent is forced to go on to the next node.
- $(Y_k|X=x,Z=z,\theta=\overline{\theta},Q_j=1, \text{Fix }D_j=1)$ is defined for the same individual but forces the person with these characteristics not to transit to the next node.

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We have that the **Total Effect** of fixing $D_j = 0$ on Y^k satisfies

$$T_j^k = DE_j^k + C_{j+1}^k.$$

The associated population level average treatment effect conditional on $Q_i=1$ is

$$ATE_{j}^{k} := \int \mathbb{E}\left(T_{j}^{k}\left[\left.Y_{k}\right|X=x,Z=z,oldsymbol{ heta}=\overline{oldsymbol{ heta}}
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The population continuation value at j+1 conditional on $\mathit{Q}_{j}=1$ is

$$ACV_{j+1}^{k} = \mathbb{E}_{X,Z,\theta} \left[\sum_{l=j+1}^{\overline{s}-1} \left\{ \begin{array}{c} \mathbb{E} \left[\left. Y_{l+1}^{k} - Y_{l}^{k} \right| X, Z, \theta, Q_{l+1} = 1, \text{Fix } Q_{j+1} = 1 \right] \\ \cdot \mathbb{P} \left[\left. Q_{l+1} = 1 \right| \text{Fix } Q_{j+1} = 1, X, Z, \theta, Q_{j} = 1 \right] \end{array} \right\} \middle| Q_{j} = 1 \right].$$

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• $\mathbb{E}\left[|Y_j - Y_{j-1}||S(z_2) \ge j > S(z_1)\right]$ is the mean gain of going from j-1 to j for a set of people who pass through j even though they may end up far above j and start far below j.

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• $\mathbb{E}\left[Y_j - Y_{j-1} \middle| S(z_2) \ge j > S(z_1)\right]$ is the mean gain of going from j-1 to j for a set of people who pass through j even though they may end up far above j and start far below j. Is this an interesting margin of choice?

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 What about different policies? Building a new school? Decreasing the tuition price?

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 It does not disentangles the direct effect and the continuation value. Wald Estimand is hard to interpret.

Thank you!

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References

- Bojinov, A. Rambachan, and N. Shephard. Panel Experiments and Dynamic Causal Effects: A Finite Population Perspective. Quantitative Economics, 12:pp, 1171–1196, 2021.
- J. J. Heckman, J. E. Humphries, and G. Veramendi. Dynamic Treatment Effects. *Journal of Econometrics*, 191(2):pp. 276–292, 2016.
- A. Rambachan and N. Shephard. When Do Common Time Series Estimands Hace Nonparametric Causal Meaning? Working Paper. Available at https://asheshrambachan.github.io/assets/files/arns_commontimeseries_causal.pdf, Oct. 2021.