

Lecture 7: Dynamic Panel Data Models

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EESP-FGV

Econometrics 2

- Recommended Reading: Bond (2002)
- Problem Set 6 - Deadline: June 20th at 5:00 pm

1. Motivation
2. Autoregressive Models
3. Multivariate Dynamic Models
4. Frontier in Panel Data Research and Dynamic Models

Motivation

Motivation

We want to estimate dynamic models.

- Relationship between a treatment in the past and an outcome today.
- Relationship between an endogenous variable in the past and its value today.

To estimate this type of model, we need a time dimension. However, time series data may be insufficient.

- Aggregate time series data may obscure microeconomic relationships.
 - Robinson [1950]: In the U.S. in 1930, immigrants were less likely to be literate than native citizens, but states with more immigrants had higher literacy rates.

Motivation

Panel data → microeconomic heterogeneity: firms, households or individuals.

Examples:

- Lecture 1 - Income Dynamics
- Household consumption - Euler Equations
- Adjustment cost models for firm's factor demands
- Economic growth
- Production functions with serially correlated productivity shocks
- Company investment rates

Motivation

Goal:

- Estimation of single equation, autoregressive-distributed lag models from panels with a large number of cross-section units, each observed for a small number of time periods.
- Estimation methods that do not require the time dimension to become large in order to obtain consistent parameter estimates.
- Focus on **micro panel data** on individuals or firms.

Identification depends on limited serial correlation in the error term.

- It relies on somewhat hard-to-justify assumptions.

Autoregressive Models

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Autoregressive Models

We start with a simple $AR(1)$ model:

$$Y_{i,t} = \alpha \cdot Y_{i,t-1} + \eta_i + \nu_{i,t},$$

where $|\alpha| < 1$, $i \in \{1, 2, \dots, N\}$ and $t \in \{2, 3, \dots, T\}$.

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Cross-sectional asymptotics: N goes to infinite, while T is fixed.

Autoregressive Models

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OLS estimator is inconsistent.

Autoregressive Models

“Are we in the woods without a dog?”



Autoregressive Models

“Are we in the woods without a dog?”



“If we don't have a dog, we hunt with a cat.”

Autoregressive Models

Within-groups estimator:

- Run OLS in the demeaned model

$$\tilde{Y}_{i,t} = \alpha \cdot \tilde{Y}_{i,t-1} + \tilde{\nu}_{i,t},$$

where $\tilde{Y}_{i,t} := Y_{i,t} - \frac{\sum_{\tau=2}^T Y_{i,\tau}}{T-1}$, $\tilde{Y}_{i,t-1} := Y_{i,t-1} - \frac{\sum_{\tau=1}^{T-1} Y_{i,\tau}}{T-1}$ and

$$\tilde{\nu}_{i,t} := \nu_{i,t} - \frac{\sum_{\tau=2}^T \nu_{i,\tau}}{T-1}.$$

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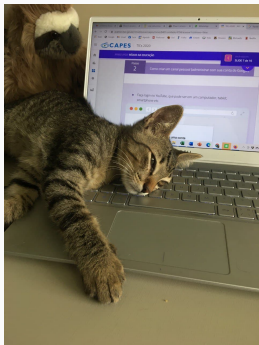
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Within-groups estimator is inconsistent.

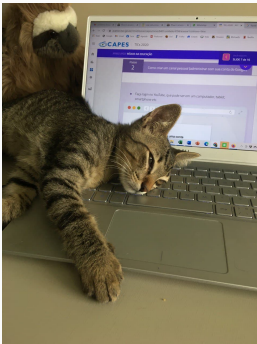
Autoregressive Models

“It was a house cat!”



Autoregressive Models

“It was a house cat!”



Let's try something different.

First-differences estimator:

- Run OLS in the first-differences model

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First-differences estimator is inconsistent.

Autoregressive Models

“It was another house cat! A sleepy one!”



Autoregressive Models

“It was another house cat! A sleepy one!”



We will need a dog!

Anderson-Hsiao estimator!



Anderson-Hsiao estimator:

- Run 2SLS in the first-differences model

$$\Delta Y_{i,t} = \alpha \cdot \Delta Y_{i,t-1} + \Delta \nu_{i,t}.$$

- Instrument is $Y_{i,t-2}$.

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- Extra Assumption - Predetermined Initial Conditions:

$$\text{Cov}(Y_{i,1}, \nu_{i,t}) = 0 \text{ for } t \in \{2, 3, \dots, T\}$$

Autoregressive Models

Anderson-Hsiao estimator: For $t \in \{3, 4, \dots, T\}$, we have that

$$\begin{aligned} \text{Cov}(\Delta \nu_{i,t}, Y_{i,t-2}) &= \text{Cov}(\nu_{i,t} - \nu_{i,t-1}, Y_{i,t-2}) \\ &= \text{Cov}(\nu_{i,t}, Y_{i,t-2}) - \text{Cov}(\nu_{i,t-1}, Y_{i,t-2}) \\ &= \text{Cov}\left(\nu_{i,t}, \alpha^{t-3} \cdot Y_{i,1} + \left\{ \sum_{\tau=0}^{t-4} \alpha^{\tau} \cdot \nu_{i,t-2-\tau} \right\}\right) \\ &\quad - \text{Cov}\left(\nu_{i,t-1}, \alpha^{t-3} \cdot Y_{i,1} + \left\{ \sum_{\tau=0}^{t-4} \alpha^{\tau} \cdot \nu_{i,t-2-\tau} \right\}\right) \\ &= 0 \end{aligned}$$

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Anderson-Hsiao estimator is large- N consistent if $T \geq 3$.

We can do better if $T > 3$.

Arellano-Bond Estimator

- Period $t = 3$: $Y_{i,1}$ is the only valid instrument.
- Period $t = 4$: $Y_{i,1}$ and $Y_{i,2}$ are valid instruments.
- \vdots
- Period $t = T$: $Y_{i,1}, Y_{i,2}, \dots, Y_{i,T-2}$ are valid instruments.

Arellano-Bond Estimator ($T > 3$):

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Autoregressive Models

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- Efficient GMM estimator based on $\mathbb{E}[Z_i' \Delta \nu_i] = 0$, where

$$Z_i := \begin{bmatrix} Y_{i,1} & 0 & 0 & \cdots & 0 & \cdots & 0 \\ 0 & Y_{i,1} & Y_{i,2} & \cdots & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & & \vdots & & \vdots \\ 0 & 0 & 0 & \cdots & Y_{i,1} & \cdots & Y_{i,T-2} \end{bmatrix}$$

$$\Delta \nu_i := (\Delta \nu_{i,3}, \Delta \nu_{i,4}, \dots, \Delta \nu_{i,T})'$$

Autoregressive Models

Arellano-Bond Estimator ($T > 3$):

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 - Older values of $Y_{i,t-j}$ are valid instruments.

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 - Older values of $Y_{i,t-j}$ are valid instruments.
- Drawbacks:
 - Large number of instruments \Rightarrow Poor Small Sample Performance and Weak Instruments.
 - Too few AR terms \Rightarrow Arellano-Bond estimator is inconsistent.
 - Too many AR terms \Rightarrow Weak Instruments

Empirical Example: Company Investment Rates (Bond, Klemm, Newton-Smith, Syed and Vlieghe, 2002)

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- $I_{i,t}$: gross investment expenditures
- $K_{i,t}$: net capital stock at replacement cost
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- η_i : firm-specific depreciation rates

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Table 1. Alternative estimates of the $AR(1)$ specification for company investment rate

Dependent variable: $(I/K)_t$					
	OLS levels	Within groups	2SLS DIF	GMM DIF	GMM DIF
$(I/K)_{t-1}$	0.2669 (.0185)	-0.0094 (.0181)	0.1626 (.0362)	0.1593 (.0327)	0.1560 (.0318)
Sargan Instruments			(p-value)	.36	.43
			$(I/K)_{t-2}$	$(I/K)_{t-2}$	$(I/K)_{t-2}$
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- Within-Groups is likely to be biased downwards. (See PSet 6)

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- 2SLS and GMM: between OLS and Within-Groups

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- 2SLS and GMM: between OLS and Within-Groups
- 2SLS and GMM: close to each other
- Sargan Test: do not reject
- Many IVs: Column (4) = 23, Column (5) = 78

Multivariate Dynamic Models

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Now, we analyze a slightly more complicated model:

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- $X_{i,t}$: current and lagged values of additional explanatory variables.
- No need to model the time series $X_{i,t}$.
- $X_{i,t}$ can be correlate with η_i .

Multivariate Dynamic Models

First-differences model:

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 $\Rightarrow X_{i,t-1}$ is a valid instrument too.

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- $X_{i,t}$ is strictly exogenous: $\text{Cov}(X_{i,t}, \nu_{i,\tau}) = 0$ for any $\tau \in \{1, 2, \dots, T\}$.
 $\Rightarrow X_i' := (X_{i,1}, \dots, X_{i,T})'$ is a valid instrument.

Frontier in Panel Data Research and Dynamic Models

1. Motivation
2. Autoregressive Models
3. Multivariate Dynamic Models
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\Rightarrow Previous models are infrequently used today.

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- Arellano and Bonhomme [2017]: Survey paper on Non-linear Panel Data Models

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- Both authors rely on control function arguments
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- Shrinkage methods increases precision by allowing for some bias.
- Navigating this trade-off: Minimize $MSE \rightarrow$ Kwon [2021].

Choice-based Treatment Effect Models:

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- Heckman et al. [2016]:
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- Han [2021]:
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- Ben-Michael et al. [2021], Cattaneo et al. [2023]:
 - Synthetic Controls with Staggered Treatment Adoption.

Thank you!

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References

- A. Abadie and J. Gardeazabal. The Economic Costs of Conflict: A Case Study of the Basque Country. *American Economic Review*, 93(1):113–132, 2003.
- A. Abadie, A. Diamond, and J. Hainmueller. Synthetic Control Methods for Comparative Case Studies: Estimating the Effect of California's Tobacco Control Program. *Journal of the American Statistical Association*, 105(490):493–505, 2010.
- A. Abadie, A. Diamond, and J. Hainmueller. Comparative Politics and the Synthetic Control Method. *American Journal of Political Science*, 59(2):495–510, 2015.
- M. Arellano and S. Bonhomme. Nonlinear Panel Data Methods for Dynamic Heterogeneous Agent Models. *Annual Review of Economics*, 9:pp. 471–496, 2017.

- E. Ben-Michael, A. Feller, and J. Rothstein. Synthetic Controls with Staggered Adoption. *Journal of the Royal Statistical Society - Series B*, Jan. 2021. Available at <https://arxiv.org/abs/1912.03290>.
- B. Callaway and P. H. Sant'anna. Difference-in-Differences with Multiple Time Periods. *Journal of Econometrics*, 225(2):pp. 200–230, Mar. 2021. Available at https://papers.ssrn.com/sol3/papers.cfm?abstract_id=3148250.
- B. Callaway, A. Goodman-Bacon, and P. H. Sant'Anna. Difference-in-Differences with a Continuous Treatment. Available at <https://arxiv.org/abs/2107.02637>., July 2021.
- M. D. Cattaneo, Y. Feng, F. Palomba, and R. Titiunik. Uncertainty Quantification in Synthetic Controls with Staggered Treatment Adoption. Available at https://mdcattaneo.github.io/papers/Cattaneo-Feng-Palomba-Titiunik_2024_wp.pdf., 2023.

- L. Davezies, X. DHaultfoeuille, and L. Laage. Identification and Estimation of Average Marginal Effects in Fixed Effects Logit Models. Available at <https://arxiv.org/abs/2105.00879>., May 2021.
- B. Ferman, C. Pinto, and V. Possebom. Cherry Picking with Synthetic Controls. *Journal of Policy Analysis and Management*, 39(2):pp. 510–532, 2020.
- A. F. Galvao. Quantile Regression for Dynamic Panel Data with Fixed Effects. *Journal of Econometrics*, 164(1):pp. 142–157, 2011. Available at <https://doi.org/10.1016/j.jeconom.2011.02.016>.
- W. Y. Gao and M. Li. Robust Semiparametric Estimation in Panel Multinomial Choice Models. Available at https://www.waynegao.com/uploads/8/1/4/6/81465138/g1_210327.pdf., Mar. 2021.

- A. Goodman-Bacon. Difference-in-Differences with Variation in Treatment Timing. *Journal of Econometrics*, 225(2):pp. 254–277, 2021.
- S. Han. Identification in Nonparametric Models for Dynamic Treatment Effects. *Journal of Econometrics*, 225, 2021.
- J. J. Heckman, J. E. Humphries, and G. Veramendi. Dynamic Treatment Effects. *Journal of Econometrics*, 191(2):pp. 276–292, 2016.
- S. Kwon. Optimal Shrinkage Estimation of Fixed Effects in Linear Panel Data Models. Available at https://soonwookwon.github.io/files/Soonwoo_Kwon_JMP.pdf., Jan. 2021.
- L. Laage. A Correlated Random Coefficient Panel Model with Time-Varying Endogeneity. Available at <https://arxiv.org/abs/2003.09367>., Mar. 2020.

- M. Li. A Time-Varying Endogenous Random Coefficient Model with an Application to Production Functions. Available at https://uc6c705c8f483c94e338df344781.d1.dropboxusercontent.com/cd/0/inline2/BfB6ZQh8j0BkLg17Cy2aZgLPf4o3x6Wj-EX5ddq_82dp8wn4Iki13tJYlcb8VRelF8u9Y2H6wlz4JUMN1IV17_qRGgrdAApEfXjfAiE05bRgeIgT03T1ZKjIEGn209tXrg1y0hxdFMrhAtpVKv3MaXxc1aqRsDBamYbsAJp5v-8G7kS7d8f2AM0fJULA8NkzHTT3PVKSYNhXjvXPHLrdX_LRYPRYIWLjfdwCSKsI5gy7lGCrYRszfySIT0Xys5L86H1qINt1flu00b8_yNy9ZNnxvH5hQKxnl4VQul24v3pPGc480Id_4JhPIL2PCF0kQAEDxxbN5i8vU4QTyP2-t8RuB0dNmisSD4/file#, Oct. 2021.
- W. S. Robinson. Ecological Correlations and the Behavior of Individuals. *American Sociological Review*, 15(3):pp. 351–357, 1950.