

IPS / Translational Bioinformatics

# Composite Biomarker Analysis of Patients with Urothelial Carcinoma

April 10, 2024

Abraham Apfel

Translational Bioinformatics / Bioinformatics Methodology

 Bristol Myers Squibb™

# Key Research Question

- Estimating associations between 13 biomarkers and Disease-Free Survival (DFS) in Urothelial Carcinoma (UC) patients treated with adjuvant therapy
  - Do the estimated associations depend on whether:
    - the patient received Treatment or Placebo?
    - we adjust for other biomarkers?

# Acknowledgements

Scott Chasalow

Jun Li

Justin David

Bhakti Dwivedi

# Our Cohort

- Phase 3 Randomized Clinical Trial of 699 patients
  - 2 treatment arms
- Whole Exome Sequencing (WES)
  - N = 458
  - Used to derive TMB
  - Identified genes with most mutations
- RNA-Seq
  - N = 323
  - Genome-wide single gene and gene signature analyses
- **Composite Analysis**
  - **Used literature and stratification variables to select a subset of 13 biomarkers for which to take deeper dive and run composite analysis.**

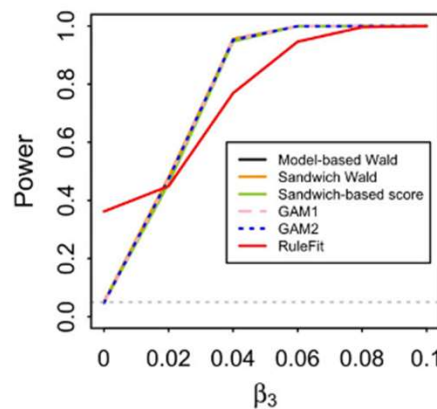
# Statistical Considerations (1)

Challenge	Solution
<b>Straight-line assumption can yield misleading results when assessing interactions among variables<sup>1</sup></b>	<b>Transform each continuous variable into 2 basis functions to allow for cubic splines in model</b>
Previous studies imply different associations between PDL1_IHC (2 measures) = 0 and PDL1_IHC > 0. Not sufficiently accounted for by splines	Force model to allow for point of discontinuity at PDL1_IHC = 0 by including extra terms in model
Train model to focus on discovering interactions between variables and treatment	Parameterize treatment as -1/+1 instead of standard 0/1 “dummy” coding <sup>2</sup>

## Straight Line Assumption Leads to High False Discovery Rate in Many Non-Linear Scenarios

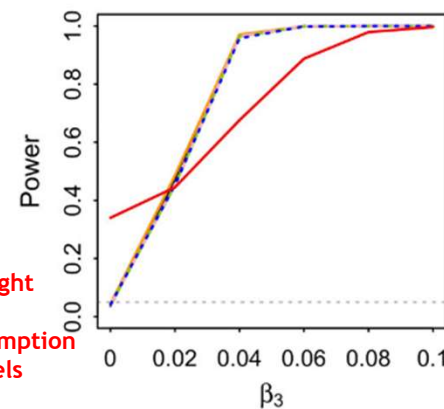
TRUE RELATIONSHIP  
Linear

## $X_1$ and $X_2$ Independent



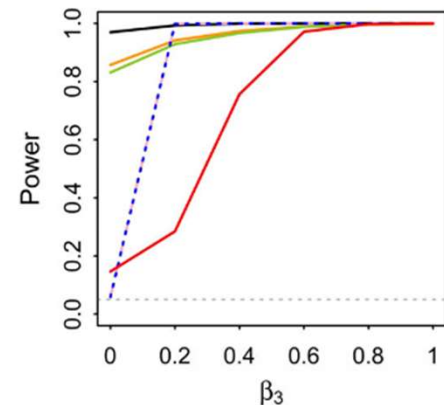
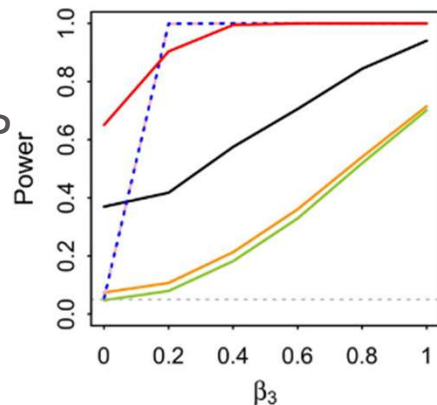
## Straight Line Assumption Models

## $X_1$ and $X_2$ Correlated



## TRUE RELATIONSHIP

### Quadratic



- $N = 500$
- $\beta_3$  is coefficient for  $X_1 * X_2$ , the interaction term
- “RuleFit” is method proposed by Friedman, et al for discovering interactions
- Intercept (Power for  $\beta_3 = 0$ ) is Type I error

- NOTE: Results were similar for Log and Exponential relationship with outcome
- For Binary outcome, even when predictors were independent there is high Type I Error

- “RuleFit” is method proposed by Friedman, et al for discovering interactions

- Intercept (Power for  $\beta_3 = 0$ ) is Type I error

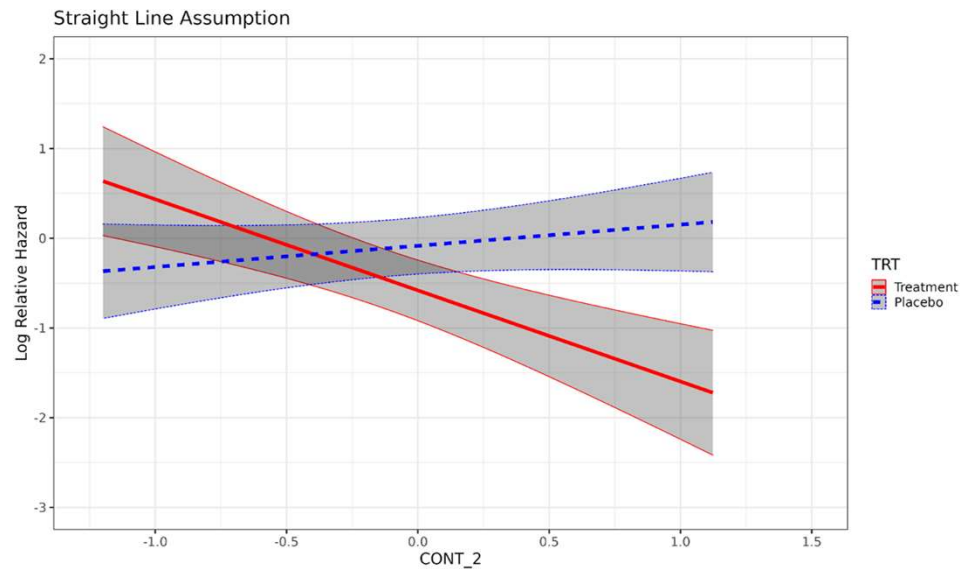
- NOTE: Results were similar for Log and Exponential relationship with outcome
- For Binary outcome, even when predictors were independent there is high Type I Error

- For Binary outcome, even when predictors were independent there is high Type I Error

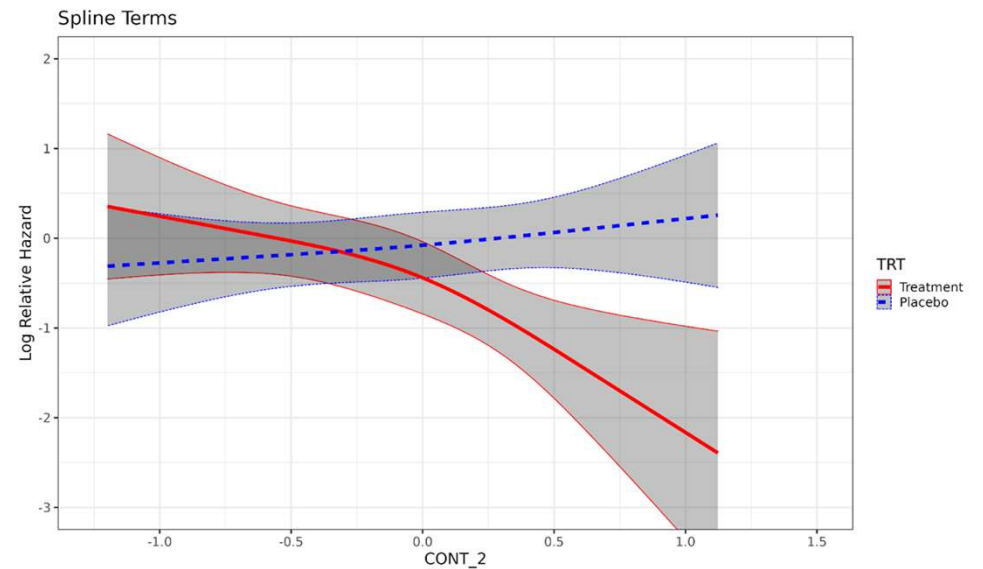
# Model with One Continuous Biomarker at a Time

“Simple” Model:  $\text{Surv(DFS, Censor)} \sim \text{Biomarker} * \text{TRT} + \text{Stratification\_Variables}$

## Straight Line Assumption



## Restricted Cubic Spline



# Statistical Considerations (1)

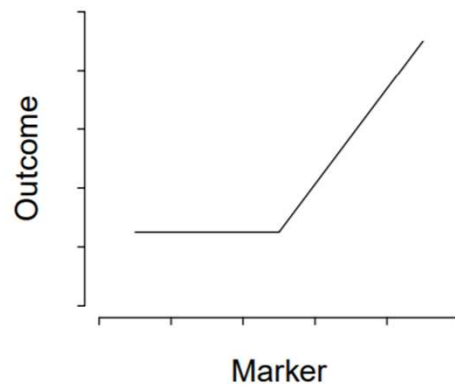
Challenge	Solution
Straight-line assumption can yield misleading results when assessing interactions among variables <sup>1</sup>	Transform each continuous variable into 2 basis functions to allow for cubic splines in model
<b>Previous studies imply different associations between PDL1_IHC (2 measures) = 0 and PDL1_IHC &gt; 0. Not sufficiently accounted for by splines</b>	<b>Force model to allow for point of discontinuity at PDL1_IHC = 0 by including extra terms in model</b>
Train model to focus on discovering interactions between variables and treatment	Parameterize treatment as -1/+1 instead of standard 0/1 “dummy” coding <sup>2</sup>

# Implications of Dichotomizing a Continuous Variable

## Continuous (Can be modeled via Splines)

### Can Occur in Biology

Not Handled by Dichotomization



## Dichotomized

### Unlikely to Occur

Assumed in Much of Biomarker Research

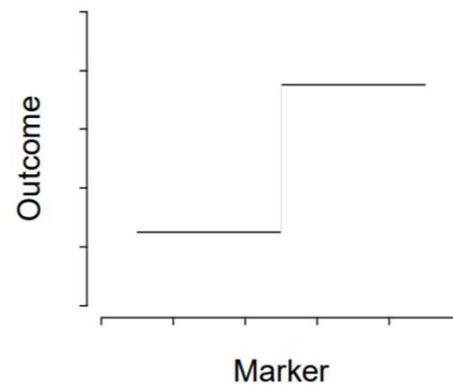


Figure 18.2: Two kinds of thresholds. The pattern on the left represents a discontinuity in the first derivative (slope) of the function relating a marker to outcome. On the right there is a lowest-order discontinuity.

- Prior studies suggest that each PDL1 measure has point of discontinuity at 0
- Still treat as continuous (i.e., do not assume a flat line) for values greater than 0

Figure copied from:

<http://hbiostat.org/doc/bbr.pdf> (Frank Harrell)



# Statistical Considerations (1)

## Challenge

Straight-line assumption can yield misleading results when assessing interactions among variables<sup>1</sup>

Previous studies imply different associations between PDL1\_IHC (2 measures) = 0 and PDL1\_IHC > 0. Not sufficiently accounted for by splines

**Train model to focus on discovering interactions between variables and treatment**

## Solution

Transform each continuous variable into 2 basis functions to allow for cubic splines in model

Force model to allow for point of discontinuity at PDL1\_IHC = 0 by including extra terms in model

**Parameterize treatment as -1/+1 instead of standard 0/1 “dummy” coding<sup>2</sup>**

## Statistical Considerations (2)

Challenge	Solution
Small sample size (N = 323) relative to number of predictors (13 + Treatment) and terms (2 basis functions and interaction for each continuous predictor)	Penalized Regression
Maintain hierarchy while allowing for multiple basis functions (for splines) and interaction with treatment in penalized setting	Use Group Lasso <sup>3</sup> : Each group consists of all terms associated with a given variable
Different scaling of categorical and continuous predictors makes it difficult to infer relative importance of variables in penalized setting	Scale each continuous variable by dividing by 2*Standard Deviation instead of typical scaling of 1*Standard Deviation <sup>4</sup>
Appropriate visualizations to assess effect size when allowing for non-straight-line effects	Make partial effects plots and coefficient trajectory plots

# Intuition Behind Scaling by $2 \cdot SD_4$

- Each continuous variable will have  $SD = 0.5$
- Binary variable,  $X$ , with  $P(X = 1) = 0.5$  has  $SD = 0.5$ 
  - Thus by dividing continuous predictors by  $2 \cdot SD$ , they will be on same scale as categorical predictor with  $P(X = 1) = 0.5$
- Even categorical predictor with skewed distribution, e.g.  $\Pr(X = 1) = 0.3$ , has  $SD = 0.45$ 
  - Thus, unless categorical predictor has “extremely skewed” distribution, dividing by  $2 \cdot SD$  is sufficient to put continuous variables on “similar” scale
    - In our dataset, most skewed variable had  $\Pr(X = 1) = 0.21$ , with  $SD = 0.4$

4. Gelman, A. (2008). Scaling regression inputs by dividing by two standard deviations. *Statistics in medicine*, 27(15), 2865-2873.

# Composite Model

Scaled by 2\*SD

2<sup>nd</sup> basis function to allow for nonlinear relationship with outcome

Penalty goes on entire group at a time (typically 4 terms: 2 basis functions + 2 interaction terms)

**Composite Model:**  $\text{Surv(DFS, Censor)} \sim \{[(\text{Scaled\_Biomarker1} + \text{Scaled\_Biomarker1}')*\text{TRT}] + \text{Penalty}\} + \{[(\text{I(PDL1} == 0) + \text{Scaled\_PDL1} + \text{Scaled\_PDL1}')*\text{TRT}] + \text{Penalty}\} + \dots$

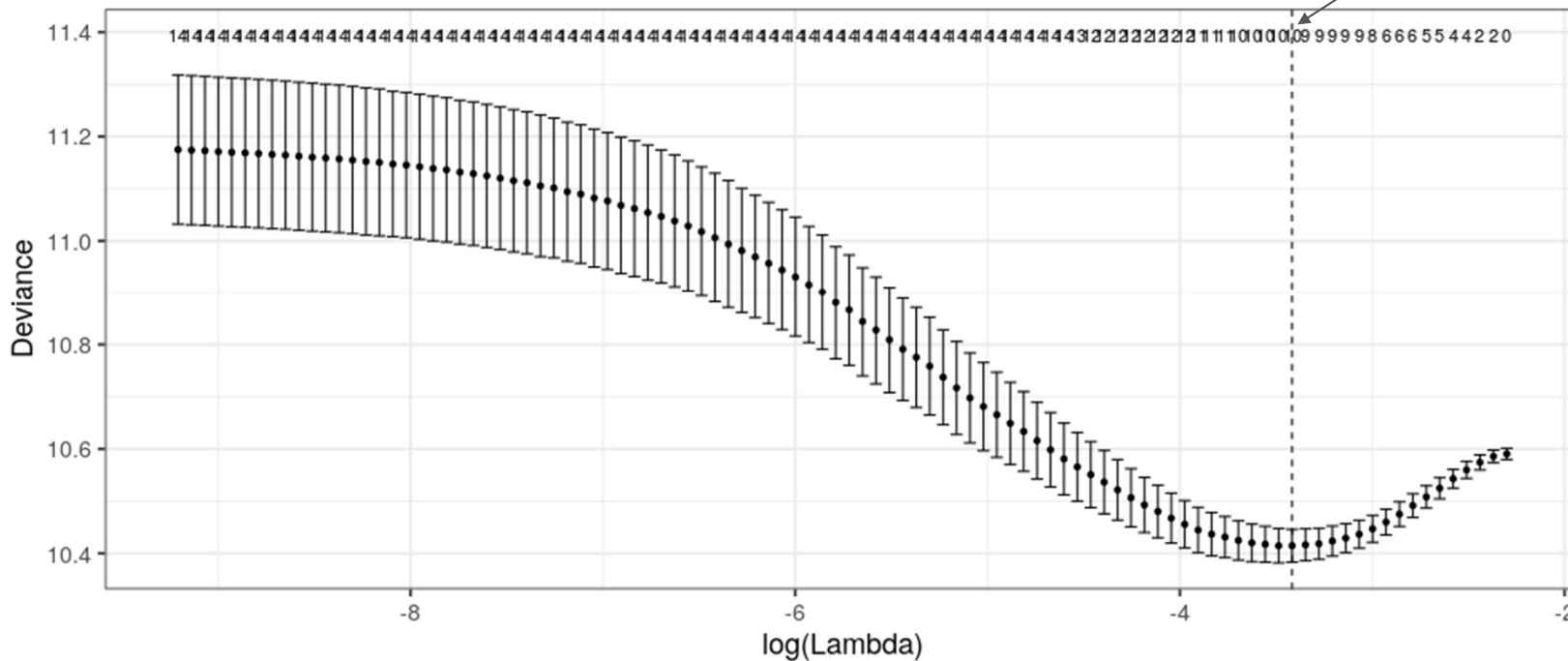
For PDL1, we include an indicator function to allow for discontinuity at 0.

Adjusted for all 13 biomarkers in the same model

# Selecting lambda (degree of penalization)

CV Curve for Median Errors Across 50 Repeats  
grpsurv model

Number of groups with non-zero  
coefficients in model (out of 14 groups)



We chose the  
Lambda with  
minimum Deviance:  
 $\text{Lambda} = 0.0328$   
( $\log(\text{lambda}) = -3.42$ )

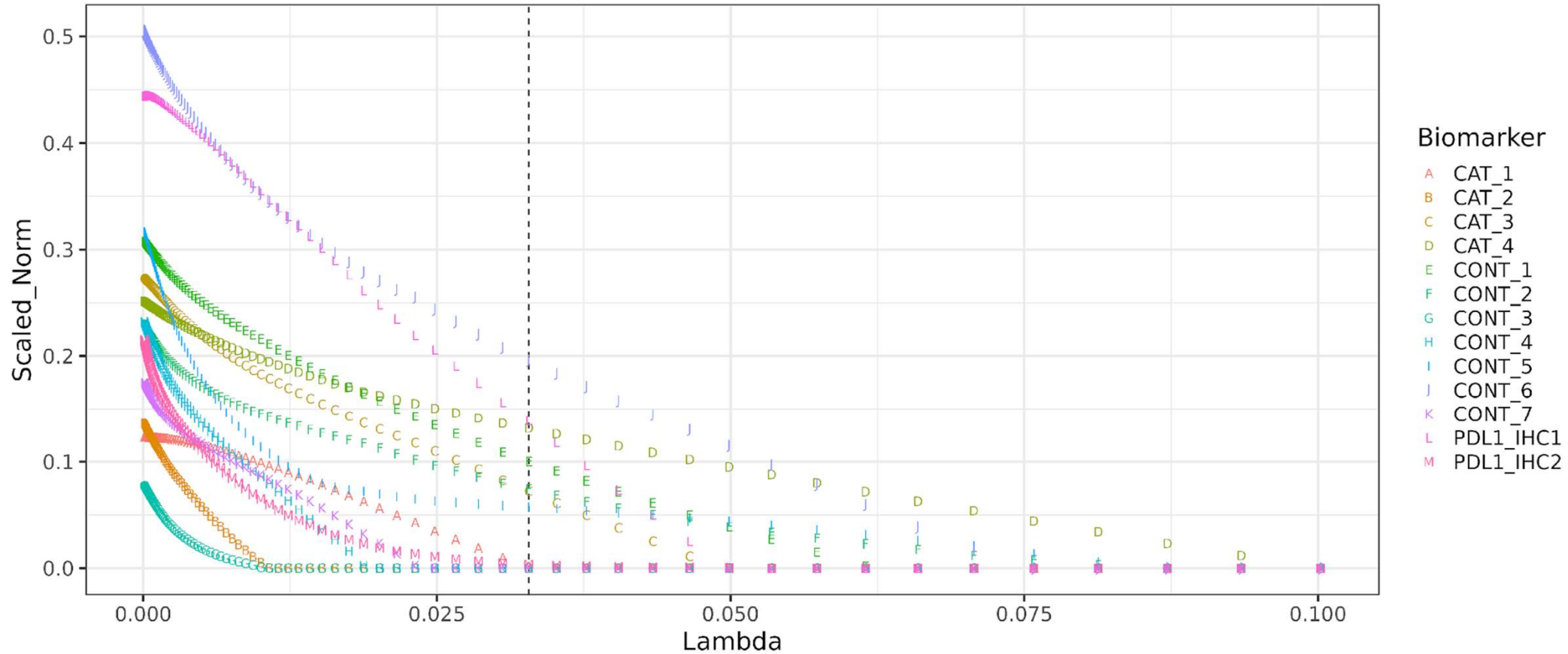
## Statistical Considerations (2)

Challenge	Solution
Small sample size ( $N = 323$ ) relative to number of predictors ( $13 + \text{Treatment}$ ) and terms (2 basis functions and interaction for each continuous predictor)	Penalized Regression
Maintain hierarchy while allowing for multiple basis functions (for splines) and interaction with treatment in penalized setting	Use Group Lasso <sup>3</sup> : Each group consists of all terms associated with any given variable
Different scaling of categorical and continuous predictors makes it difficult to infer relative importance of variables in penalized setting	Scale each continuous variable by dividing by $2 \times \text{Standard Deviation}$ instead of typical scaling of $1 \times \text{Standard Deviation}$ <sup>4</sup>
<b>Appropriate visualizations to assess effect size when allowing for non-straight-line effects</b>	<b>Make partial effects plots and coefficient trajectory plots</b>

# Trajectories of Group Coefficient Norms

Scaled Variables by 2\*Standard Deviation

Selected Lambda = 0.0328

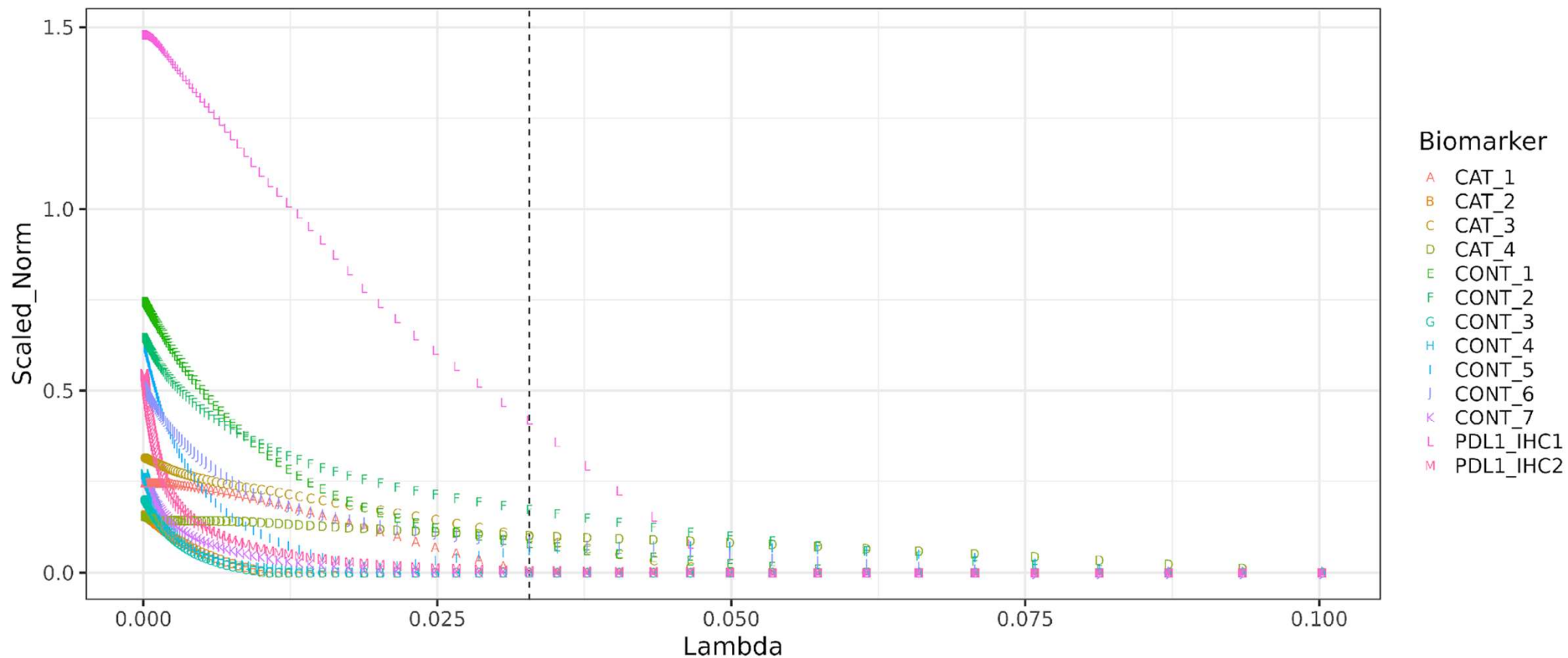


$\text{Scaled\_Norm} = \sqrt{(\sum B^2)/k}$  , where B = Coefficient and k = Number of coefficients for a given predictor

# Trajectories of Interaction Coefficient Norms

Interaction Coefficients

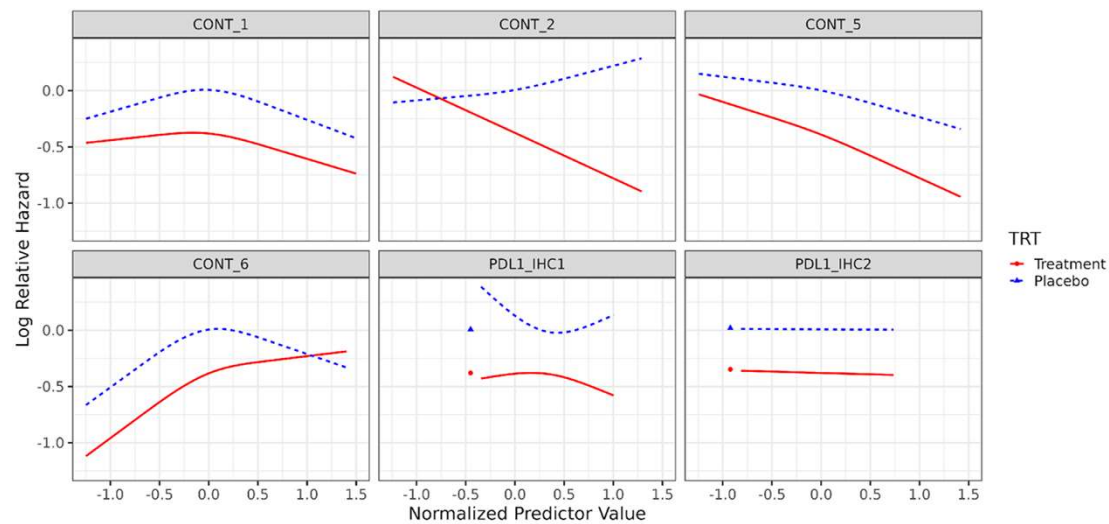
Selected Lambda = 0.0328



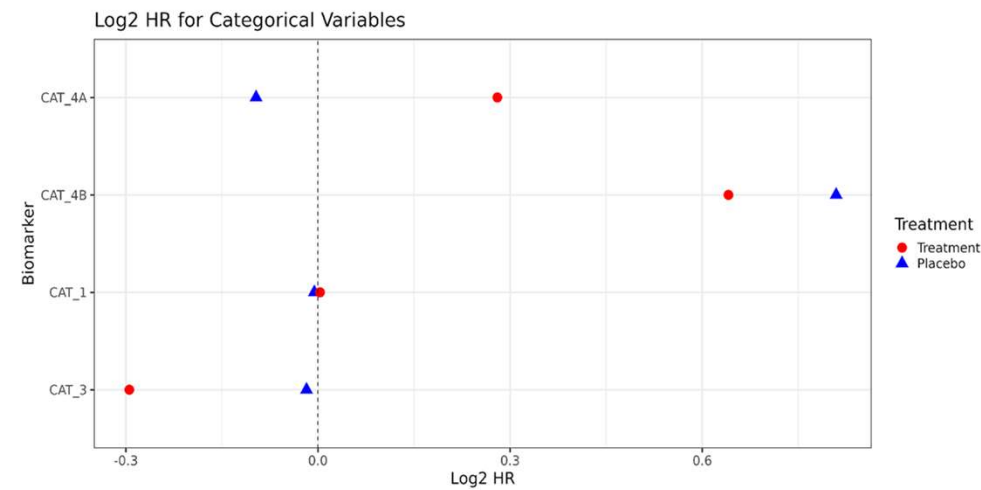


# Results from Model: Partial Effects Plots

## Continuous Variables

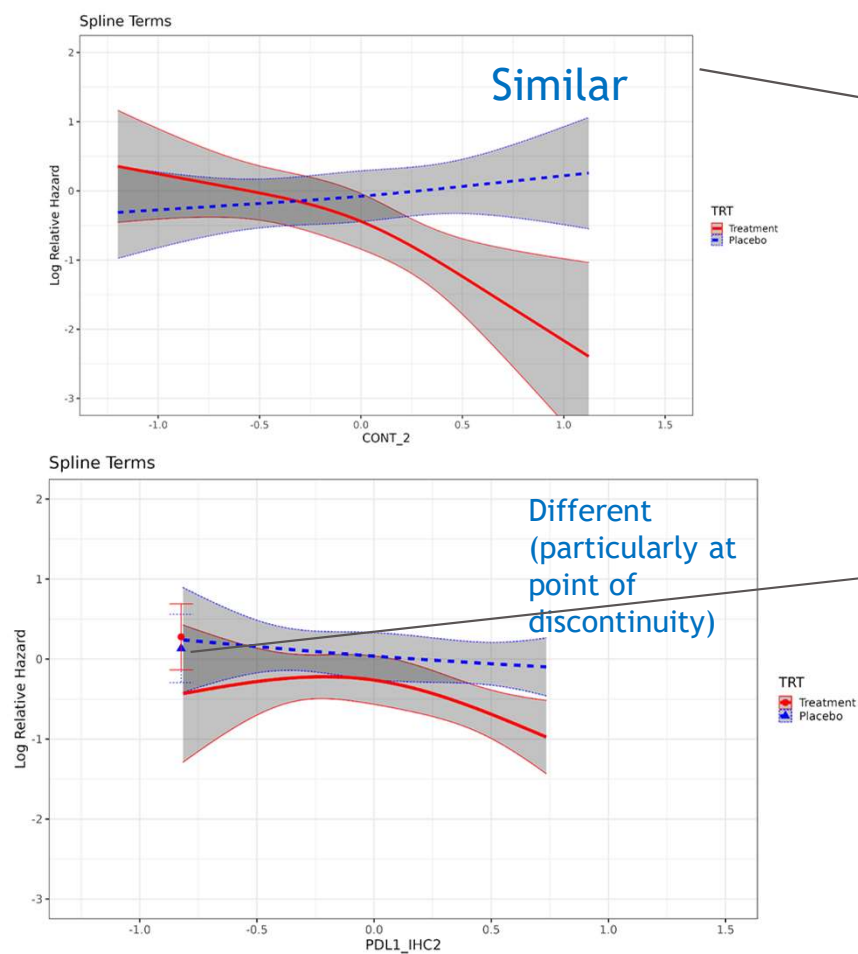


## Categorical Variables

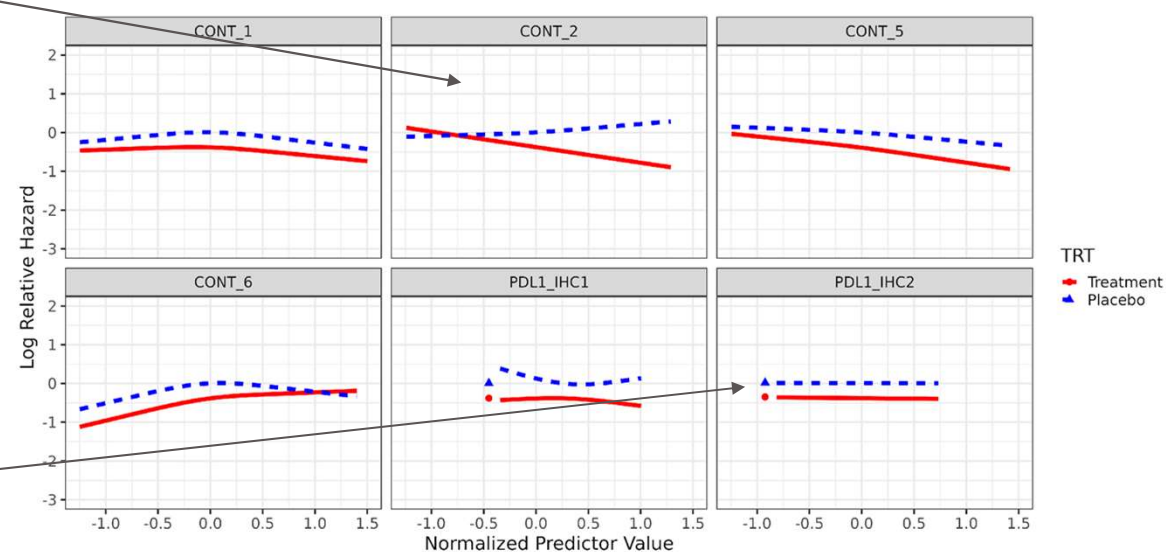


# Composite Model More Impactful for Some Biomarkers Than Others

## One-At-A-Time Model

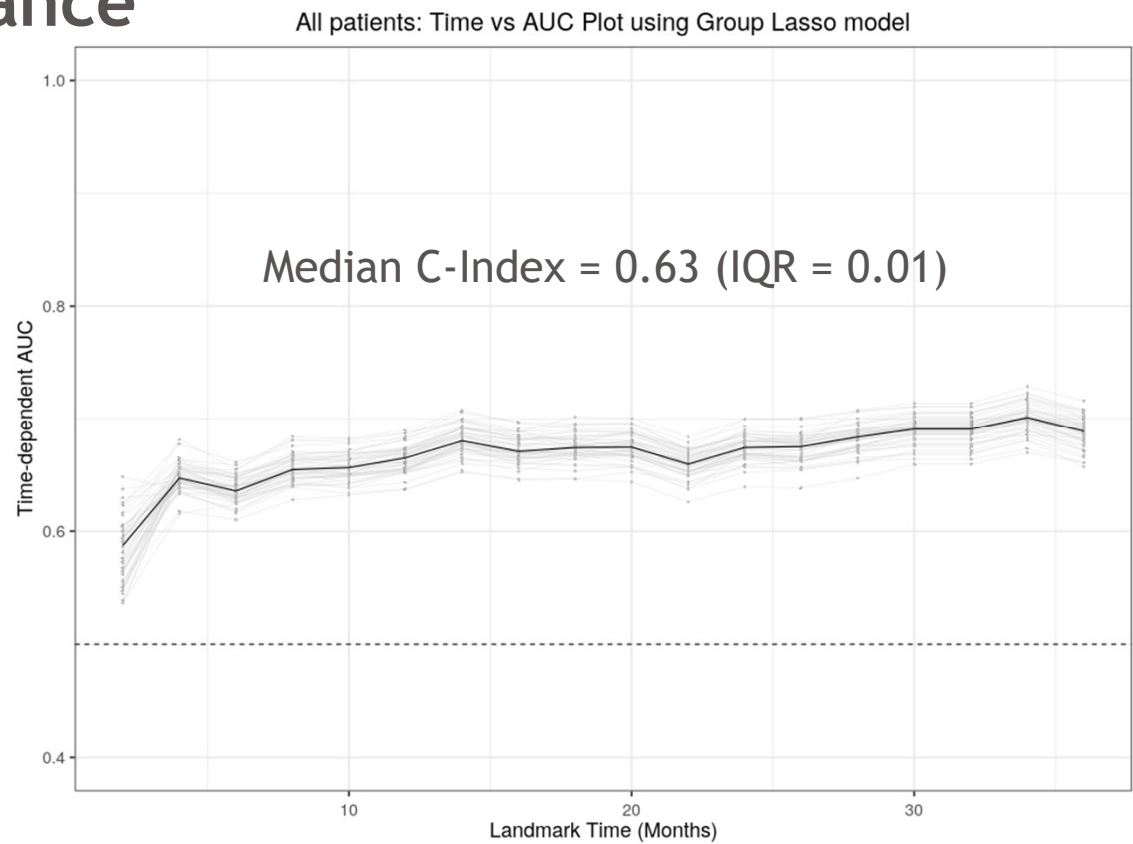


## Composite Model



# Model Performance

Note: Performance Metrics are from 50 repeats of 10-fold CV



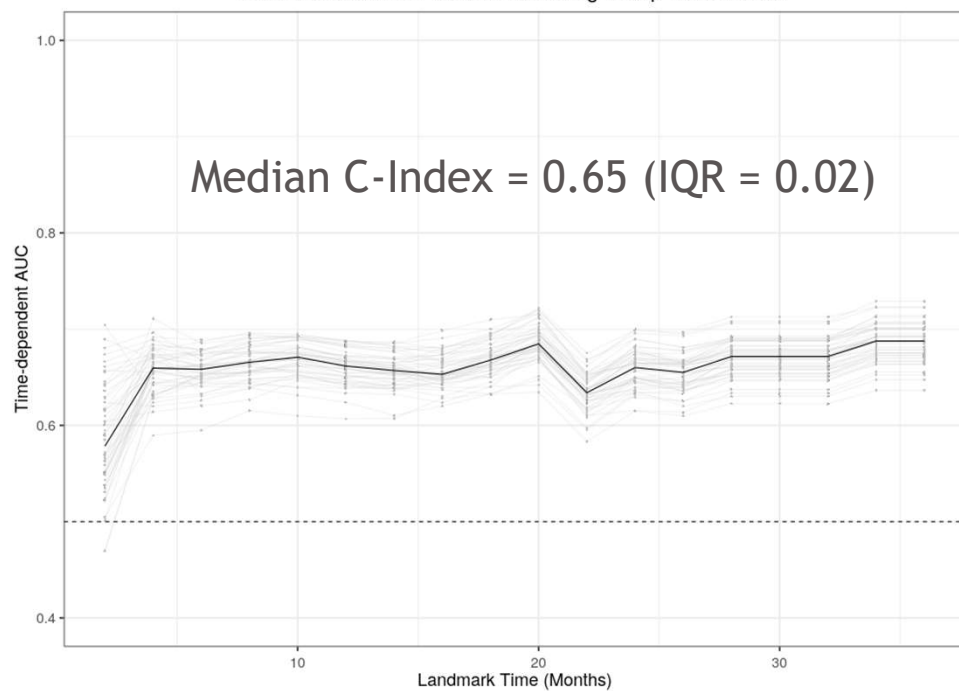
## Distribution of DFS

10%	25%	50%	75%	90%
2.6	3.4	13.7	25.8	36.9

# Model Performance, by Arm

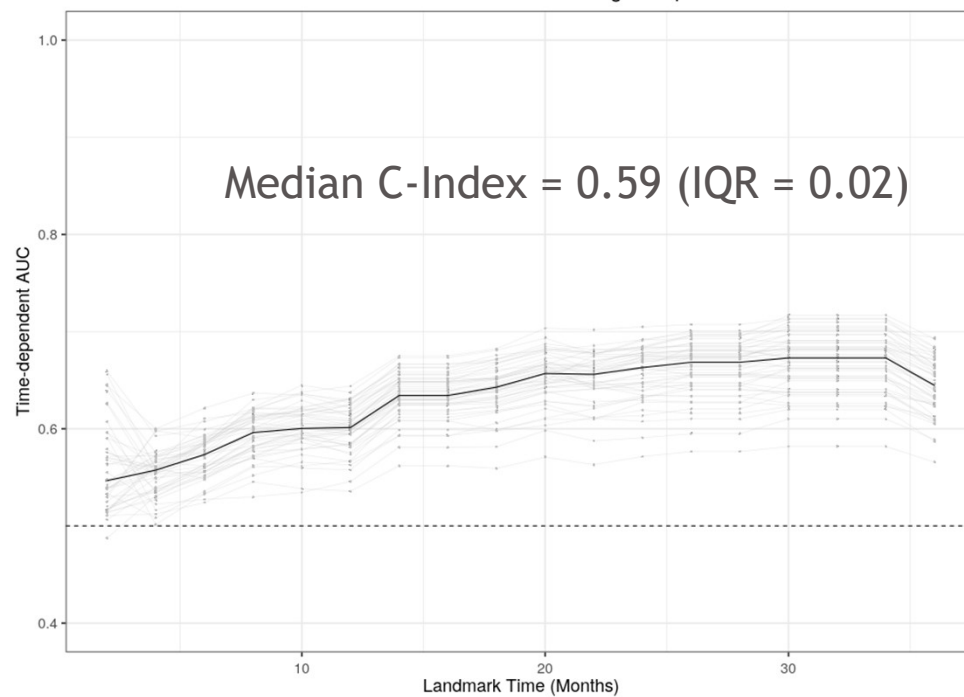
## Treatment Subjects

Nivo Patients: Time vs AUC Plot using Group Lasso model



## Placebo Subjects

Placebo Patients: Time vs AUC Plot using Group Lasso model



Note: Performance Metrics are from 50 repeats of 10-fold CV

# Statistical Considerations (1)

## Challenge

Straight-line assumption can yield misleading results when assessing interactions among variables<sup>1</sup>

Previous studies imply different associations between PDL1\_IHC (2 measures) = 0 and PDL1\_IHC > 0. Not sufficiently accounted for by splines

Train model to focus on discovering interactions between variables and treatment

## Solution

Transform each continuous variable into 2 basis functions to allow for cubic splines in model

Force model to allow for point of discontinuity at PDL1\_IHC = 0 by including extra terms in model

Parameterize treatment as -1/+1 instead of standard 0/1 “dummy” coding<sup>2</sup>

## Statistical Considerations (2)

### Challenge

### Solution

Small sample size ( $N = 323$ ) relative to number of predictors ( $13 + \text{Treatment}$ ) and terms (2 basis functions and interaction for each continuous predictor)

Penalized Regression

Maintain hierarchy while allowing for multiple basis functions (for splines) and interaction with treatment in penalized setting

Use Group Lasso<sup>3</sup>: Each group consists of all terms associated with any given variable

Different scaling of categorical and continuous predictors makes it difficult to infer relative importance of variables in penalized setting

Scale each continuous variable by dividing by  $2 \times \text{Standard Deviation}$  instead of typical scaling of  $1 \times \text{Standard Deviation}$ <sup>4</sup>

Appropriate visualizations to assess effect size when allowing for non-straight-line effects

Make partial effects plots and coefficient trajectory plots

## Conclusions: Clinical

- In our composite model:
  - The effect of PDL1\_IHC1 and CONT\_2 on DFS differed greatly between the Treatment and Placebo arms
  - CONT\_1, CONT\_5, CONT\_6, and CAT\_4 were possibly prognostic: they had strong associations with DFS but NOT strong evidence that the associations differed between treatment arms

## Conclusions: Methods

- The results for some biomarkers changed greatly between the “one continuous biomarker at a time” model and the composite model
  - Most notably, CONT\_3 and PDL1\_IHC2 were found to have minimal association with DFS in the composite model but each appeared to have strong association in the one-at-a-time models
- The detailed modeling choices had important influences on our conclusions
  - Relationship of PDL1\_IHC1 and PDL1\_IHC2 to DFS changed greatly after allowing for discontinuity at 0
  - Appropriate scaling changed the relative importance of several variables
  - The straight-line assumption appeared unreasonable for several variables



# References

1. Zhang, M., Yu, Y., Wang, S., Salvatore, M., G. Fritsche, L., He, Z., & Mukherjee, B. (2020). Interaction analysis under misspecification of main effects: Some common mistakes and simple solutions. *Statistics in Medicine*, 39(11), 1675-1694.
2. Tian, L., Alizadeh, A. A., Gentles, A. J., & Tibshirani, R. (2014). A simple method for estimating interactions between a treatment and a large number of covariates. *Journal of the American Statistical Association*, 109(508), 1517-1532.
3. Yuan, M., & Lin, Y. (2006). Model selection and estimation in regression with grouped variables. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 68(1), 49-67.
4. Gelman, A. (2008). Scaling regression inputs by dividing by two standard deviations. *Statistics in medicine*, 27(15), 2865-2873.
5. Huang J, Breheny P, and Ma S. (2012). A selective review of group selection in high dimensional models. *Statistical Science*, **27**: 481-499. doi: [10.1214/12-sts392](https://doi.org/10.1214/12-sts392)

# Backup

# Methods - 1

- Parameterized Treatment as +1/-1 (instead of standard 0/1 “dummy coding”) to train model to focus on identifying interactions with treatment<sup>2</sup>
- Applied Grouped lasso-regularized Cox PH regression<sup>5</sup>
  - Imputed missing values via single imputation (small amount of missing)
  - Allowed continuous variables to have non-linear relationship with outcome via cubic splines (via “rcs” function from “rms” package)
  - Scaled continuous variables to have mean = 0 and std = 0.5 to achieve similar scaling as categorical variables<sup>4</sup>
  - Each measure for PDL1 was modeled using  $\text{Ind}(x = 0)$  to allow for a discontinuity between 0 and 1, together with cubic splines for the  $x > 0$  portion of the domain, and interaction with treatment for both parts.
  - Treatment was constrained to have a non-zero coefficient

<sup>2</sup> Tian L, Alizadeh AA, Gentles AJ, Tibshirani R. A Simple Method for Estimating Interactions between a Treatment and a Large Number of Covariates. *J Am Stat Assoc.* 2014;109(508):1517-1532. doi:10.1080/01621459.2014.951443

<sup>4</sup> Gelman A. Scaling regression inputs by dividing by two standard deviations. *Stat Med.* 2008 Jul 10;27(15):2865-73. doi: 10.1002/sim.3107. PMID: 17960576.

<sup>5</sup> Huang J, Breheny P, and Ma S. (2012). A selective review of group selection in high dimensional models. *Statistical Science*, **27**: 481-499. doi: [10.1214/12-sts392](https://doi.org/10.1214/12-sts392)

## Methods - 2

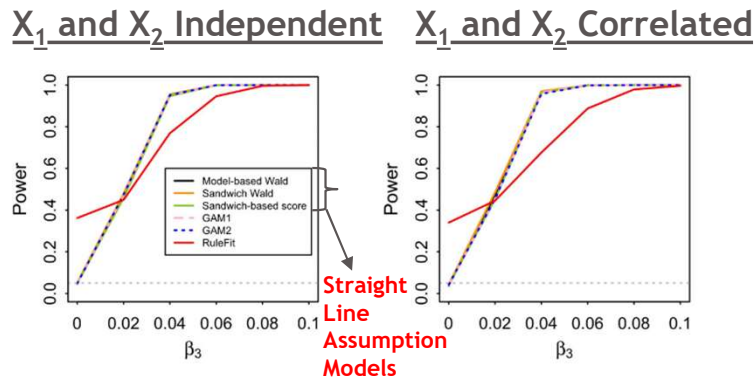
- Chose the tuning parameter ( $\lambda$ ) value that had minimum median Deviance from 50 repeated iterations of 5-fold cross validation (CV) across the full dataset (inner loop of CV)
- Model Performance was assessed via C-index and plots of AUC (of time-dependent ROC curve) vs Landmark Time, estimated from 50 repeats of 10-fold CV (outer loop)
- Key Packages:
  - grpreg
  - SurvivalROC

# Bioinformatics Details for RNA-Seq

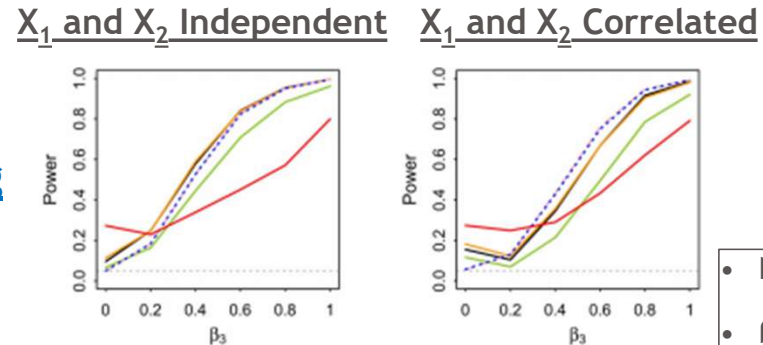
- TMM-normalized CPM
- Filtered ~35,000 genes with  $\log_2(\text{CPM}) < 0.5$ 
  - ~24,000 genes remained in analysis
- Gene Set Enrichment Analysis (GSEA)
  - Genes were ranked by  $\log_{\text{HR}}$  (IQR) for expression in placebo arm -  $\log_{\text{HR}}$  (IQR) for expression in treatment arm
  - Looked for patterns of enrichment based on above criterion to identify interesting pathways
  - GSEA used 50 hallmark gene sets from Molecular Signature Data Base (MSigDB)

# Simulations for Continuous Outcome: Straight Line Assumption Leads to High False Discovery Rate in Many Non-Linear Scenarios

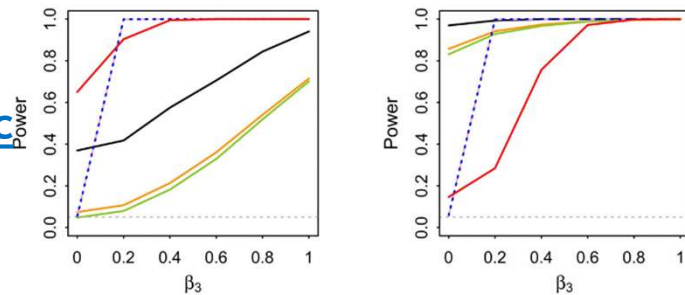
Linear



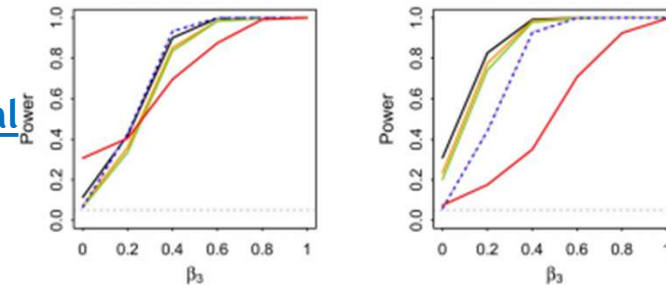
Log



Quadratic



Exponential

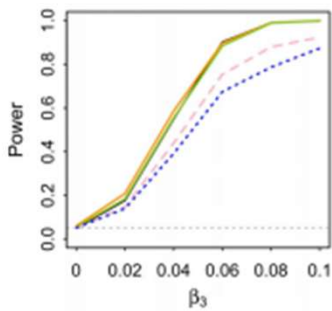


- $N = 500$
- $\beta_3$  is coefficient for  $X_1 \times X_2$ , the interaction term
- “RuleFit” is method proposed by Friedman, et al for discovering interactions
- Intercept (Power for  $\beta_3 = 0$ ) is Type I error

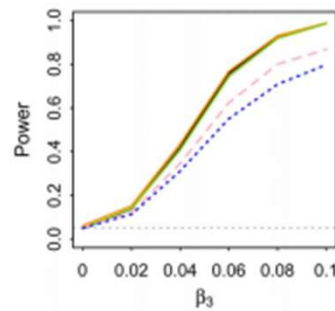
# Simulations for Binary Outcome: Straight Line Assumption Leads to High False Discovery Rate in ALL Simulated Non-Linear Scenarios

Linear

$X_1$  and  $X_2$  Independent

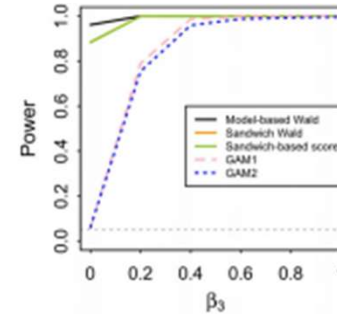


$X_1$  and  $X_2$  Correlated

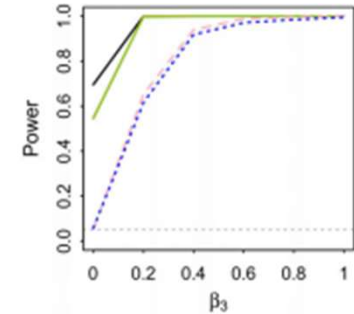


Log

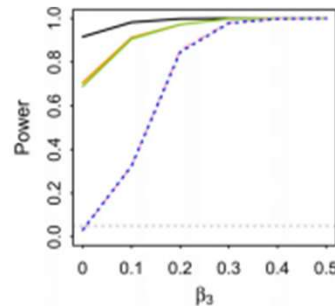
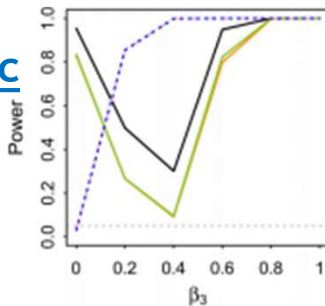
$X_1$  and  $X_2$  Independent



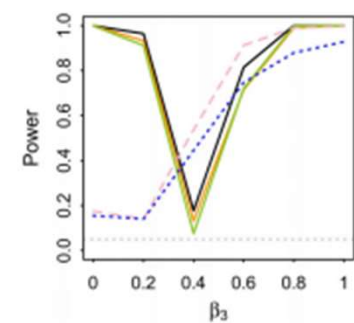
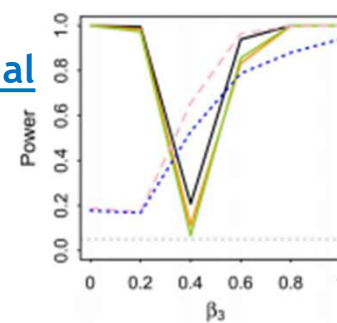
$X_1$  and  $X_2$  Correlated



Quadratic



Exponential



- $N = 2000$
- $\beta_3$  is coefficient for  $X_1 \times X_2$ , the interaction term
- Intercept (Power for  $\beta_3 = 0$ ) is Type I error

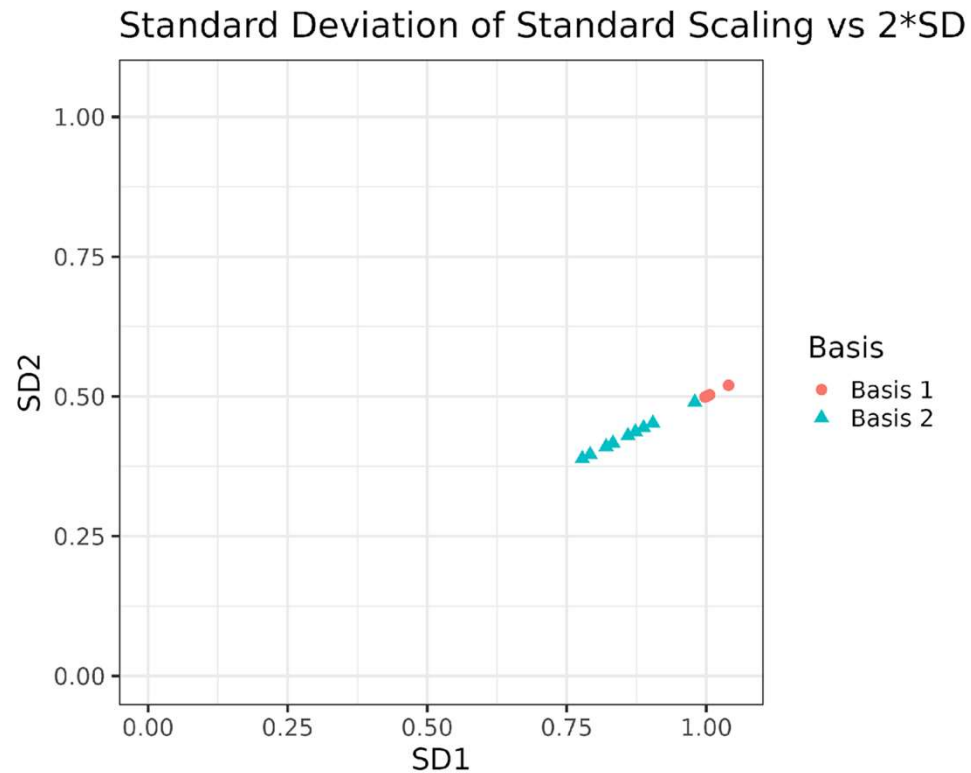
# Parameterization of Treatment

- Parameterize treatment as +1/-1 instead of standard dummy coding
- Objective is to estimate  $\widehat{\Delta z}$ , the causal treatment effect for patients with covariates  $\mathbf{z}$ 
  - $\Delta z = E(Y^{(1)} - Y^{(-1)} | \mathbf{Z} = \mathbf{z})$

2. Tian, L., Alizadeh, A. A., Gentles, A. J., & Tibshirani, R. (2014). A simple method for estimating interactions between a treatment and a large number of covariates. *Journal of the American Statistical Association*, 109(508), 1517-1532.



# Scaling Comparison: Standard Deviation

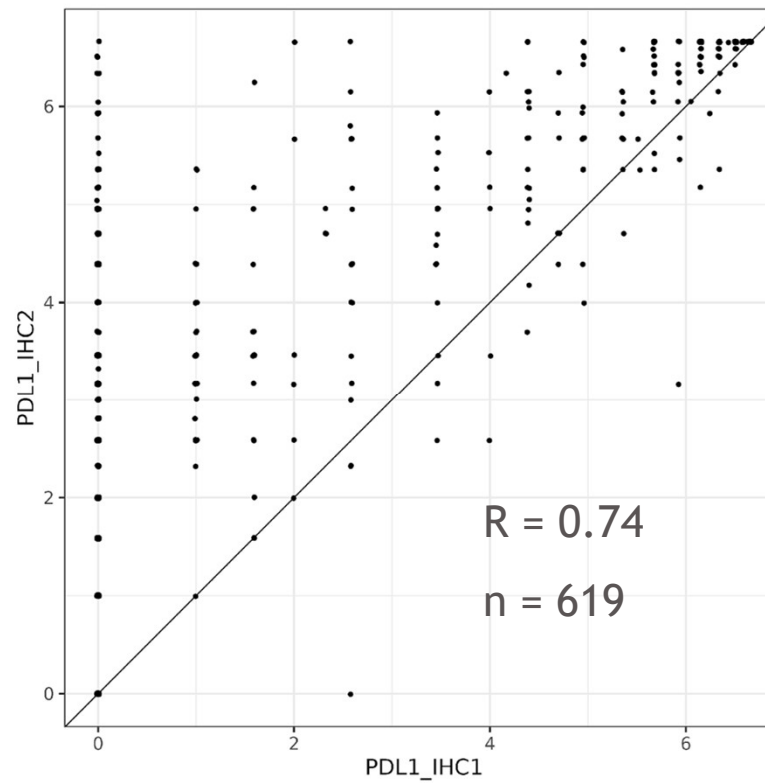


Notes:

1. Used `rms::rcs()` to create splines because `ns()` has unclear scaling method.
2. We 1<sup>st</sup> scaled variables before creating basis function

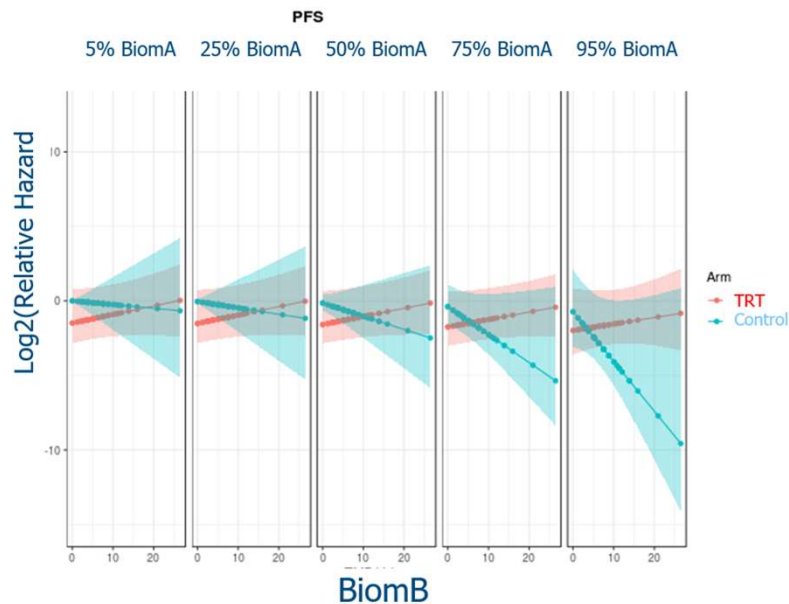
## PDL1\_IHC1 and PDL1\_IHC2 Correlation?

(All observations)



# Why Multivariable Models?

1. Controls for confounders  
— Fewer spurious findings
2. Can increase precision
3. Interactions



## Google searches for 'my cat scratched me'

correlates with

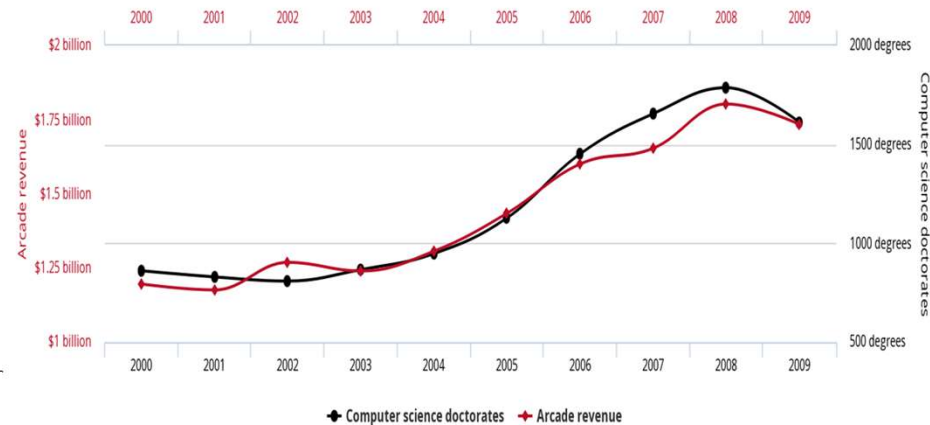
## The Coca-Cola Company's stock price (KO)



## Total revenue generated by arcades

correlates with

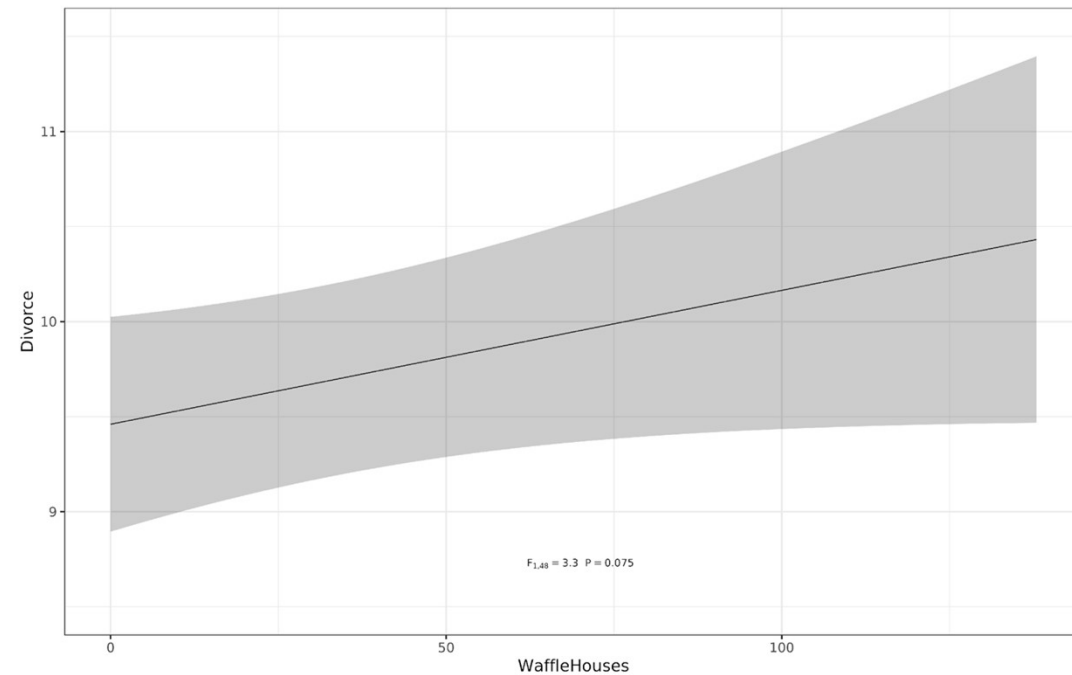
## Computer science doctorates awarded in the US



# WaffleHouses → High divorce rates?



Divorce Rate ~ Waffle Houses per Million People



# Well, after adjusting for location...

Divorce Rate ~ Waffle Houses per Million + South

