

STAT 231 Online Spring 2020

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1 Week 7

1.1 Test of Hypothesis Part I

Definition

A **test statistic** or **discrepancy measure** is a function of the data $D = g(Y)$ that is constructed to measure the degree of "agreement" between the data Y and the null hypothesis H_0 .

Summary (Steps of a Statistical Test of Hypothesis)

1. Assume that the null hypothesis H_0 will be tested using data Y
2. Adopt a test statistic or discrepancy measure $D(Y)$ for which large value of D are less consistent with H_0 . Let $d = D(y)$ be the corresponding observed value of D
3. Calculate

$$\begin{aligned} p\text{-value} &= P(D \geq d; \text{assuming } H_0 \text{ is true}) \\ &= P(D \geq d; H_0) \end{aligned}$$

4. Draw a conclusion based on the p -value

1.2 Test of Hypothesis Part II

Summary (Hypothesis Test for Binomial Model)

1. Test $H_0 : \theta = \theta_0$ using $Y \sim \text{Bi}(n, \theta)$
2. Test statistic $D(Y) = |Y - n\theta_0|$. Let $d = |y - n\theta_0|$ be the observed value of D
3. Calculate

$$\begin{aligned} p\text{-value} &= P(D \geq d; \theta = \theta_0) \\ &= P(|Y - n\theta_0| \geq d) \end{aligned}$$

where $Y \sim \text{Bi}(n, \theta)$

If n is large, we can approximate

$$\begin{aligned} p\text{-value} &\approx P\left(|Z| \geq \frac{d}{\sqrt{n\theta_0(1-\theta_0)}}\right) \quad \text{where } Z \sim N(0, 1) \\ &= 2 \left[1 - P\left(Z \leq \frac{d}{\sqrt{n\theta_0(1-\theta_0)}}\right)\right] \end{aligned}$$

4. Draw a conclusion

Summary (Hypothesis Test for Gaussian Model)

When the **variance** is unknown

1. 123

When the **mean** is unknown

2 Week 8

Summary (Likelihood Ratio Test of Hypothesis)

1. To test $H_0: \theta = \theta_0$

2. Likelihood ratio test statistic $\Lambda(\theta_0) = -2 \log \left[\frac{L(\theta_0)}{L(\hat{\theta})} \right],$

with the observed value $\lambda(\theta_0) = -2 \log \left[\frac{L(\theta_0)}{L(\hat{\theta})} \right]$

3. Calculate

$$\begin{aligned} p\text{-value} &\approx P(W \leq \lambda(\theta_0)) \quad \text{where } W \sim \chi^2(1) \\ &= 2[1 - P(Z \leq \sqrt{\lambda(\theta_0)})] \quad \text{where } Z \sim G(0, 1) \end{aligned}$$

4. Draw a conclusion

3 Week 9

3.1 Gaussian Response Models Part I

Confidence Interval for beta β

A 100p% Confidence Interval for β is given by

$$\hat{\beta} \pm a \frac{s_e}{\sqrt{S_{xx}}}$$

where $P(T \leq a) = \frac{(1+p)}{2}$ and $T \sim t(n-2)$

3.2 Gaussian Response Models Part II

Summary (Hypothesis testing of No Relationship for beta β)

1. Test $H_0 : \beta = 0$

2. Test statistic $D = \frac{|\tilde{\beta} - 0|}{s_e/\sqrt{S_{xx}}}$, with observed value $d = \frac{|\hat{\beta} - 0|}{s_e/\sqrt{S_{xx}}}$

3. Calculate

$$\begin{aligned} p\text{-value} &= P\left(|T| \geq \frac{|\hat{\beta} - 0|}{s_e/\sqrt{S_{xx}}}\right) \\ &= 2 \left[1 - P\left(T \leq \frac{|\hat{\beta} - 0|}{s_e/\sqrt{S_{xx}}}\right) \right] \end{aligned}$$

where $T \sim t(n-2)$

4. Draw Conclusion

Confidence Interval for Mean $\mu(x) = \alpha + \beta x$

A 100p% Confidence Interval for $\mu(x) = \alpha + \beta x$ is given by

$$\begin{aligned} \hat{\mu}(x) \pm as_e \sqrt{\frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}} \\ = \hat{\alpha} + \hat{\beta}x \pm as_e \sqrt{\frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}} \end{aligned}$$

where $P(T \leq a) = \frac{(1+p)}{2}$ and $T \sim t(n-2)$

4 Week 10

4.1 Gaussian Response Models Part III

Prediction Interval For Future Response Y

A 100p% Prediction Interval for a future response Y is given by

$$\begin{aligned} \hat{\mu}(x) \pm as_e \sqrt{1 + \frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}} \\ = \hat{\alpha} + \hat{\beta}x \pm as_e \sqrt{1 + \frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}} \end{aligned}$$

where $P(T \leq a) = \frac{(1+p)}{2}$ and $T \sim t(n-2)$

Definition

For the simple linear regression model, let

$$\hat{\mu}_i = \hat{\alpha} + \hat{\beta}x_i$$

and let

$$\hat{r}_i = y_i - \hat{\mu}_i \quad i = 1, 2, \dots, n$$

The \hat{r}_i 's are called **residuals** since \hat{r}_i represents what is "left" after the model has been "fitted" to the data

Definition

Define the **standardized residuals** as

$$\hat{r}_i = \frac{\hat{r}_i}{s_e} = \frac{y_i - \hat{\mu}_i}{s_e} \quad i = 1, 2, \dots, n$$

4.2 Comparing Means of Two Populations

Definition

Pooled estimate of variance is defined as

$$\begin{aligned} s_p^2 &= \frac{1}{n_1 + n_2 - 2} \left[\sum_{i=1}^{n_1} (y_{1i} - \bar{y}_1)^2 + \sum_{i=1}^{n_2} (y_{2i} - \bar{y}_2)^2 \right] \\ &= \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2} \end{aligned}$$

C.I. for Difference in Means $\mu_1 - \mu_2$

A 100p% C.I. for $\mu_1 - \mu_2$ is given by

$$\bar{y}_1 - \bar{y}_2 \pm as_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

where $P(T \leq a) = \frac{(1+p)}{2}$ and $T \sim t(n_1 + n_2 - 2)$

Summary (Hypothesis testing of Difference in Means $\mu_1 - \mu_2$)

1. Test $H_0 : \mu_1 - \mu_2 = 0$

2. Test statistic $D = \frac{|\bar{Y}_1 - \bar{Y}_2 - 0|}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ and observed value $d = \frac{|\bar{y}_1 - \bar{y}_2 - 0|}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$

3. Calculate

$$\begin{aligned} p\text{-value} &= P(|T| \geq d) \\ &= 2[1 - P(T \leq d)] \end{aligned}$$

where $T \sim t(n_1 + n_2 - 2)$

4. Draw a conclusion

C.I. for σ

A 100p% Confidence Interval for σ is given by

5 Week 11

5.1 Gaussian Response Models Part IV

Summary (Hypothesis testing of Difference in Means $\mu = \mu_1 - \mu_2$ with-in-pair)

1. Test $H_0: \mu = 0$ using $Y_i = Y_{1i} - Y_{2i} \sim G(\mu_1 - \mu_2, \sigma)$, $i = 1, 2, \dots, n$, independently

2. Test statistic $D = \frac{|\bar{Y} - 0|}{S/\sqrt{n}}$, with observed value $d = \frac{|\bar{y} - 0|}{s/\sqrt{n}}$

3. Calculate

$$p\text{-value} = 2[1 - P(T \leq d)]$$

where $T \sim t(n-1)$

4. Draw a conclusion

5.2 Multinomial Models and Goodness of Fit

Summary (Hypothesis testing of θ)

1. Test $H_0: \theta = \theta_0 = \left(\frac{1}{k}, \frac{1}{k}, \dots, \frac{1}{k}\right)$ using $f(y_1, y_2, \dots, y_k; \theta_1, \theta_2, \dots, \theta_k)$

2. Likelihood ratio test statistic $\Lambda = 2 \sum_{j=1}^k Y_j \log \left(\frac{Y_j}{E_j} \right)$

and observed value $\lambda = 2 \sum_{j=1}^k y_j \log \left(\frac{y_j}{e_j} \right)$

3. Calculate

$$p\text{-value} = P(\Lambda \geq \lambda; H_0) \approx P(W \geq \lambda)$$

where $W \sim \chi^2(k-1-p)$

4. Draw a conclusion

6 Week 12

6.1 Two-Way Tables and Tests of Hypothesis

Definition

An example of **Two-way table**...

Let θ_{ij} be the probability a randomly selected individual is combined type (A_i, B_j) and note that $\sum_{i=1}^a \sum_{j=1}^b \theta_{ij} = 1$. The $a \times b$ frequencies $(Y_{11}, Y_{12}, \dots, Y_{ab})$ follow a Multinomial distribution with $k = ab$ classes.

Summary (Hypothesis testing of Dependence of A and B class)

1. Test $H_0: \theta_{ij} = \alpha_i \beta_j$

Note that the expected frequencies under the hypothesis H_0 are given by $e_{ij} = \frac{r_i c_j}{n}$

2. Likelihood ratio statistic $\Lambda = 2 \sum_{i=1}^a \sum_{j=1}^b Y_{ij} \log \left(\frac{Y_{ij}}{E_{ij}} \right)$,

with observed value $\lambda = 2 \sum_{i=1}^a \sum_{j=1}^b y_{ij} \log \left(\frac{y_{ij}}{e_{ij}} \right)$

3. Calculate

$$p\text{-value} \approx P(W \geq \lambda)$$

where $W \sim \chi^2((a-1) \times (b-1))$

Note:

- (a) If $((a-1) \times (b-1)) = 1$, then

$$p\text{-value} = 2[1 - P(Z \leq \sqrt{\lambda})], \quad \text{where } Z \sim G(0, 1)$$

- (b) If $((a-1) \times (b-1)) = 2$, then

$$p\text{-value} = P(W \geq \lambda) = e^{-\lambda/2}, \quad \text{where } W \sim \chi^2(2) = \text{Exponential}(2)$$

4. Draw a conclusion

6.2 Cause and Effect

Definition

Let be y a response variate and let be x an explanatory variate associated with units in a population or process. Then, if all other factors that affect are held constant, let us change x (or observe different values of x) and see if y changes. If it does, we say that x has a **causal effect** on y .

Definition

x has a **causal effect** on Y if, when all other factors that affect are held constant, a change in x (or observing different values of x) induces a change in a property of the distribution of Y .