

2) Binomial distribution $\sim \text{Bin}(n, \pi)$:

$$f(y) = \binom{n}{y} \left(\frac{\pi}{1-\pi}\right)^y (1-\pi)^n$$

Taking log on both sides.

$$f(y) = \exp \left\{ \log \binom{n}{y} + y \log \left(\frac{\pi}{1-\pi} \right) + n \log(1-\pi) \right\} \quad \text{--- (1)}$$

$$\text{Let } \log \left(\frac{\pi}{1-\pi} \right) = \theta \Rightarrow \frac{\pi}{1-\pi} = e^\theta \Rightarrow \pi = \frac{e^\theta}{1+e^\theta}$$

$$1-\pi = 1 - \frac{e^\theta}{1+e^\theta} = \frac{1}{1+e^\theta}$$

~~comparing~~

$$\Rightarrow f(y) = \exp \left\{ y\theta - n \log(1+e^\theta) + \log \binom{n}{y} \right\} \quad \text{--- (2)}$$

comparing (2) &

$$f_\eta(y; \theta, \phi) = \exp \left\{ y \frac{\theta}{\phi} - \frac{b(\theta)}{\phi} + c(y, \phi) \right\} \quad \text{--- (3)}$$

$$\text{we get, } \theta = \log \left(\frac{\pi}{1-\pi} \right), \quad \phi = 1,$$

$$b(\theta) = n \log(1+e^\theta)$$

$$E(Y) = \mu = b'(\theta)$$

$$= \frac{ne^{\theta}}{1+e^{\theta}} = n\pi \quad \text{assumed}$$

$$\text{var}(Y) = \phi b''(\theta)$$

$$= \frac{1}{n} \frac{\partial}{\partial \theta} \frac{ne^{\theta}}{1+e^{\theta}} = \frac{ne^{\theta}}{(1+e^{\theta})^2} = n\pi(1-\pi)$$

To show canonical link:

To show: $g(\mu) = \theta$

$$\frac{ne^{\theta}}{1+e^{\theta}} = \mu \Rightarrow \frac{\mu}{n} = \frac{e^{\theta}}{1+e^{\theta}} = \pi$$

$$\Rightarrow \theta = \log\left(\frac{\pi}{1-\pi}\right) = \log\left(\frac{\mu/n}{1-\mu/n}\right)$$

\Rightarrow Logit fn.

\therefore Logit function is the canonical link for binomial distribution.

3) Poisson distribution $\text{Pois}(\lambda)$

$$f(y) = \frac{1}{y!} \lambda^y e^{-\lambda}$$

Taking log obs

$$f(y) = \exp \{ y \log \lambda - \lambda - \log(y!) \} \quad \text{--- (1)}$$

To show:

~~Comparing~~

$$f(y; \theta, \phi) = \exp \left\{ y \frac{\theta - b(\theta)}{\phi} + c(y, \phi) \right\} \quad \text{--- (2)}$$

Comparing (1) + (2)

$$\text{Let } \log \lambda = \theta \Rightarrow \lambda = e^\theta \quad \phi = 1 \quad c(y, \phi) = -\log(y!)$$

$$b(\theta) = e^\theta$$

$$E(y) = \phi b'(\theta) = \frac{d}{d\theta} e^\theta = \lambda = \mu$$

$$\text{Var}(y) = \phi b''(\theta) = \frac{d}{d\theta} e^\theta = e^\theta = \lambda$$

Canonical link : $\eta(y) = \theta \Rightarrow$ To show.

$$\lambda = e^\theta \Rightarrow \log \lambda = \theta$$

$\therefore \log$ is the canonical link for Poisson distribution.

$$5) f(y, \beta) = \binom{y+m-1}{m-1} \beta^m (1-\beta)^y.$$

$$= f(y) = \exp \left[\log \binom{y+m-1}{m-1} + m \log \beta + y \log (1-\beta) \right]$$

$$\cancel{= f(y)} = \text{let } \theta = \log (1-\beta)$$

$$1-\beta = e^\theta \Rightarrow \beta = 1 - e^\theta$$

$$\Rightarrow \log \beta = \log (1 - e^\theta)$$

$$= \exp \left[y \log \theta + m \log (1 - e^\theta) + \log \binom{y+m-1}{m-1} \right]$$

$$\phi = 1 \quad b(\theta) = -m \log (1 - e^\theta)$$

$$c(y, \phi) = \log \binom{y+m-1}{m-1}$$

$$E[y] = \frac{+m e^\theta}{1 - e^\theta} \quad \phi = \frac{m(1-\beta)}{\beta}$$

$$\text{Var}(y) = \phi \frac{\partial^2 b(\theta)}{\partial \theta^2} = \frac{m e^\theta}{(1 - e^\theta)^2} = \frac{r(1-\beta)}{\beta^2}$$

$$g(u) = \eta = 0$$

$$E(y) = \frac{m e^0}{1 - e^0} = \frac{m}{e^{-0} - 1}$$

$$g^{-1}(\theta) = \frac{e^{-\theta} - 1}{m}$$

$$m g^{-1}(\theta) + 1 = e^{-\theta}$$

$$\log(m g^{-1}(\theta) + 1) = e^{-\theta}$$

$$\log\left(\frac{1}{m g^{-1}(\theta) + 1}\right) = +\theta$$

$$\log\left(\frac{1}{\frac{m}{E y} + 1}\right) = \theta$$

$$\log\left[\frac{E y}{m + E y}\right] = \theta$$

\therefore canonical link is log

c) ~~Gamma~~ Gamma $(\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-\beta y}$ where $\alpha, \beta > 0$.

$$f(y) = \ln \{ \alpha \log \beta - \log(\Gamma(\alpha)) + (\alpha-1) \log y - \beta y \} \quad (1)$$

To show :

$$f(y; \phi; \theta) = \ln \left\{ \frac{y^\theta - b(\theta)}{\phi} + c(y, \phi) \right\} \quad (2)$$

$$= \ln \left\{ \frac{y(-\beta/\alpha) - (-\log \beta)}{1/\alpha} + (\alpha-1) \log y - \log(\Gamma(\alpha)) \right\}$$

let $\theta = -\beta/\alpha$ $\phi = 1/\alpha$ $\alpha = -\beta/\phi$

$$-\beta = \alpha \theta \Rightarrow \beta = -\theta/\phi$$

~~$\log \beta = -\log \theta - \log \phi$~~

~~$b(\theta) = -\log \theta$~~

$$= \ln \left\{ \frac{y \theta - [-\log \theta]}{\phi} - \log(\phi)/\phi + \left(\frac{1}{\phi} - 1 \right) \log y - \log \Gamma(1/\phi) \right\} \quad (3)$$

comparing (2) + (3).

$$b(\theta) = -\log(-\theta) \quad \phi = \frac{1}{\alpha} \quad \theta = -\beta/\alpha$$

~~Canonical link $g(\mu) = \theta$~~

$$EY = b'(\theta) = -\frac{1}{\theta} = \mu \quad (4)$$

$$\text{Var } Y = \phi b''(\theta) = \phi / \theta^2$$

To show:

Canonical link: $g(\mu) = \theta$

From (4), we get $\theta = -\frac{1}{\mu} = g(\mu)$

\therefore The canonical link function is negative inverse function

i) $\text{Exp}(\lambda)$

$$f(y; \lambda) = \lambda e^{-\lambda y}$$

$$f(y) = \text{exp} \{ \log \lambda - \lambda y \} \quad \text{--- (1)}$$

Comparing with

$$f(y; \theta, \phi) = \text{exp} \{ \underbrace{\eta y - b(\theta)}_{\phi} + c(y, \phi) \} \quad \text{--- (2)}$$

$$\theta = -\lambda \quad \text{--- (3)} \Rightarrow \lambda = -\theta$$

$$f(y) = \text{exp} \{ \underbrace{\theta y - (-\log(-\theta))}_{\phi} \}$$

$$\text{we get } b(\theta) = -\log(-\theta) \quad \phi = 1, \quad \theta = -\lambda$$

$$EY = b'(\theta) = \frac{-1}{-\theta} = \frac{1}{\lambda} = \mu \quad \text{--- (4)}$$

$$\text{Var } Y = \phi b''(\theta) = \frac{1}{\theta^2} = \frac{1}{\lambda^2}$$

To show: $g(\mu) = \theta$

$$\text{From (4)} \quad \mu = \frac{1}{\lambda} \Rightarrow \theta = \frac{-1}{\mu}$$

\therefore it is an inverse function

4) Chi-squared distribution.

$$f(y; k) = \frac{1}{\Gamma(\frac{k}{2}) 2^{k/2}} y^{k/2-1} e^{-y/2}$$

$$f(y) = \exp \left\{ -\log(\Gamma(\frac{k}{2})) + \frac{k}{2} \log \frac{1}{2} + (\frac{k}{2} - 1) \log y - y/2 \right\}$$

$$f(y) = \exp \left\{ -y/2 + \frac{k}{2} \log \frac{1}{2} + (\frac{k}{2} - 1) \log y - \log(\Gamma(\frac{k}{2})) \right\}$$

$$= \exp \left\{ \underbrace{\frac{-y/2}{k/2} + \log \frac{1}{2} + (\frac{k}{2} - 1) \log y - \log(\Gamma(\frac{k}{2}))}_{\frac{k}{2}} \right\}$$

$$= \exp \left\{ \underbrace{-\frac{y}{k} - \log 2 + (\frac{k}{2} - 1) \log y - \log(\Gamma(\frac{k}{2}))}_{\frac{k}{2}} \right\}$$

Comparing with

$$f(y; \phi, \theta) = \exp \left\{ \frac{y\theta - b(\theta)}{\phi} + c(y, \phi) \right\}$$

we get $\phi = \frac{k}{2}$, $b(\theta) = \log 2$, $\theta = -\frac{1}{2}$,

$Y \sim \text{Bernoulli}$

Problem 3) $b(y_i, \pi) = \pi^{y_i} (1-\pi)^{1-y_i}$

~~$L(Y, \pi) = \prod_{i=1}^n b(y_i, \pi) = \prod_{i=1}^n \pi^{y_i} (1-\pi)^{1-y_i}$~~
log-likelihood:

$$l(Y, \pi) = \sum_{i=1}^n [y_i \log \pi + (1-y_i) \log(1-\pi)]$$

- (1)

$$S(\pi) = \frac{\partial l(Y, \pi)}{\partial \pi} = \sum_{i=1}^n \left[\frac{y_i}{\pi} - \frac{1-y_i}{1-\pi} \right]$$

$$= \frac{\sum_{i=1}^n y_i}{\pi} - \frac{\sum_{i=1}^n (1-y_i)}{1-\pi}$$

$$= \frac{\sum_{i=1}^n y_i}{\pi} - n - \frac{\sum_{i=1}^n y_i}{1-\pi}$$

Assume $\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i = \frac{n\bar{y}}{\pi} - n - \frac{n\bar{y}}{1-\pi}$

$$\frac{n\bar{y}(1-\pi) - n(1-\bar{y})\pi}{\pi(1-\pi)} = \frac{n\bar{y} - n\pi}{\pi(1-\pi)} \quad \text{--- (2)}$$

For MLE, put (2) = 0
 $n\bar{y} = n\pi \Rightarrow \hat{\pi}_{MLE} = \bar{y}$

$$I(\pi) = E \left(- \frac{\partial^2 l(y, \pi)}{\partial \pi^2} \right)$$

$$= E \left[\frac{-n\pi(1-\pi) - n(\bar{y} - \pi)(1-2\pi)}{\pi^2(1-\pi^2)} \right]$$

$$= E \left[\frac{n\pi(1-\pi) + n(\bar{y} - \pi)(1-2\pi)}{\pi^2(1-\pi^2)} \right]$$

$$= \frac{n\pi(1-\pi) + n(1-2\pi)E(\bar{y} - \pi)}{\pi^2(1-\pi)^2}$$

$$= \frac{n\pi(1-\pi)}{\pi^2(1-\pi)^2}$$

Fishing statistic :-

$$= \frac{n}{\pi(1-\pi)} \quad (3)$$

$$\text{Wald : } T_{SW} = (\hat{\pi}_{MLE} - \pi_0) \cdot I(\hat{\pi}_{MLE}) (\hat{\pi}_{MLE} - \pi_0)$$

$$= (\bar{y} - \pi_0)^2 \frac{n}{\bar{y}(1-\bar{y})} = n \frac{(\bar{y} - \pi_0)^2}{\bar{y}(1-\bar{y})}$$

$$\text{Score : } T_S = S(\pi_0) \cdot I^{-1}(\pi_0) \cdot S(\pi_0)$$

$$= \left[\frac{n\bar{y} - n\pi_0}{\pi_0(1-\pi_0)} \right]^2 \frac{\pi_0(1-\pi_0)}{n}$$

$$= n(\bar{y} - \pi_0)^2 / \pi_0(1-\pi_0)$$

$$LR : TS_{LR} = 2 [l(y, \hat{\pi}_{MLE}) - l(y, \pi_0)]$$

$$l(y, \bar{y}) = \sum_{i=1}^n [y_i \log \bar{y} + (1-y_i) \log (1-\bar{y})]$$

$$l(y, \pi_0) = \sum_{i=1}^n \left[y_i \log \frac{\bar{y}}{\pi_0} + (1-y_i) \log \frac{1-\bar{y}}{1-\pi_0} \right]$$

$$TS_{LR} = 2 \sum_{i=1}^n \left[y_i \log \frac{\bar{y}}{\pi_0} + (1-y_i) \log \frac{1-\bar{y}}{1-\pi_0} \right]$$

$$= 2 \left[n\bar{y} \log \frac{\bar{y}}{\pi_0} + (n-n\bar{y}) \log \frac{1-\bar{y}}{1-\pi_0} \right]$$

2) i) $\pi_0 = 0.1$, $n = 10$, $\pi = 0.3$

$$TS_w = \frac{10(0.3 - 0.1)^2}{0.3(1-0.3)} = 1.9 \quad \checkmark$$

$$TS_s = \frac{10(0.3 - 0.1)^2}{0.1(1-0.1)} = 4.44 \quad \times$$

$$TS_{LR} = 2 \left[10 \cdot 0.3 \log \frac{0.3}{0.1} + 10(1-0.3) \log \frac{1-0.3}{1-0.1} \right] \\ = 3.07 \quad \checkmark$$

$$ii) \quad \pi_0 = 0.3, \quad \pi = 0.3, \quad n = 10.$$

$$TS_w = \frac{10(0.3 - 0.3)^2}{0.3(1-0.3)} = 0. \quad \checkmark$$

$$TS_s = \frac{10(0.3 - 0.3)^2}{0.3(1-0.3)} = 0. \quad \checkmark$$

$$TS_{LP} = 0 = 2 \left[10 \cdot 0.3 \log \frac{0.3}{0.3} + 10(1-0.3) \log \frac{1-0.3}{1-0.3} \right] = 0. \quad \checkmark$$

$$iii) \quad \pi_0 = 0.5$$

$$TS_w = \frac{10(0.3 - 0.5)^2}{0.3(1-0.3)} = 1.9 \quad \checkmark$$

$$TS_s = \frac{10(0.3 - 0.5)^2}{0.5(1-0.5)} = 1.6 \quad \checkmark$$

$$TS_{LP} = 2 \left[10 \cdot 0.3 \log \frac{0.3}{0.5} + 10(1-0.3) \log \frac{1-0.3}{1-0.5} \right] \\ = 1.65 \quad \checkmark$$

3) Not all of them.

Problem 2: $y_i \sim \text{Bin}(m, \pi_i)$.

$$f(y) = \binom{m}{y_i} \log \pi_i^{y_i} (1 - \pi_i)^{m - y_i}$$

$$\ell(\beta, \pi_i) = \sum_{i=1}^n \left\{ y_i \log \pi_i + (m - y_i) \log (1 - \pi_i) \right. \\ \left. + \log \binom{m}{y_i} \right\}$$

$$\Rightarrow \log \frac{\pi_i}{1 - \pi_i} = x_i \beta = m_i$$

For a ~~the~~ saturated model, all π_i 's are different.
So $\beta = [\pi_1, \dots, \pi_N]^T$

$$\text{MLE} \Rightarrow \hat{\pi}_i = \frac{e^{x_i \hat{\beta}}}{1 + e^{x_i \hat{\beta}}} = \frac{y_i}{m_i}$$

$$\ell(\text{binom}, y) = \sum \left[y_i \log \left(\frac{y_i}{m_i} \right) - y_i \log \left(\frac{m_i - y_i}{m_i} \right) \right. \\ \left. + m_i \log \left(\frac{m_i - y_i}{m_i} \right) + \log \binom{m_i}{y_i} \right]$$

For any other model

$$\text{let } \hat{y}_i = m_i \hat{\pi}_i$$

$$l(b; y) = \sum \left[y_i \log \left(\frac{\hat{y}_i}{m} \right) - y_i \log \left(\frac{m_i - \hat{y}_i}{m} \right) + m_i \log \left(\frac{m_i - \hat{y}_i}{m} \right) + \log(m_i) \right]$$

$$D = 2 \left[l(b_{max}; y) - l(b; y) \right]$$

$$= 2 \sum_{i=1}^n \left[y_i \log \left(\frac{y_i}{\hat{y}_i} \right) + (m_i - y_i) \log \left(\frac{m_i - y_i}{m_i - \hat{y}_i} \right) \right]$$

Since $\hat{y}_i = m_i \hat{\pi}_i$

$$= 2 \sum_{i=1}^n \left[y_i \log \left(\frac{y_i}{m_i \hat{\pi}_i} \right) + (m_i - y_i) \log \left(\frac{m_i - y_i}{m_i - m_i \hat{\pi}_i} \right) \right]$$

where $\hat{\pi}_i = \frac{e^{x_i \hat{\beta}}}{1 + e^{x_i \hat{\beta}}}$

deviance residual

$$d_i = \text{sign}(y_i - m_i \hat{\pi}_i) \sqrt{2 \left[y_i \log \left(\frac{y_i}{m_i \hat{\pi}_i} \right) + (m_i - y_i) \log \left(\frac{1 - y_i/m_i}{1 - \hat{\pi}_i} \right) \right]}$$

where $\hat{\pi}_i = \frac{e^{x_i^T \beta}}{1 + e^{x_i^T \beta}}$

Pearson's statistic:

$$\begin{aligned} G &= \sum_{i=1}^n (y_i - \hat{\mu}_i)^2 / v(\hat{\mu}_i) \\ &= \sum_{i=1}^n \left(\frac{y_i - m_i \hat{\pi}_i}{\sqrt{m_i \hat{\pi}_i (1 - \hat{\pi}_i)}} \right)^2 \end{aligned}$$

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