Consistency of ELBO maximization for model selection

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Outline of the talk

- Tempered Variational Bayes
 - Tempered posteriors
 - Variational Bayes
- 2 Model Selection
 - Framework
 - ELBO criterion
- Consistency result
 - Main result
 - Application

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Notations

Assume that we observe X_1, \ldots, X_n i.i.d from P^0 in a model $\mathcal{M}_K = \{P_\theta, \theta \in \Theta_K\}$ associated with a likelihood L_n . We define a prior Π_K on Θ_K .

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The posterior

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The tempered posterior - $0 < \alpha < 1$

$$\pi_{n,\alpha}^{K}(\mathrm{d}\theta) \propto [L_{n}(\theta)]^{\alpha} \Pi_{K}(\mathrm{d}\theta).$$

Various reasons to use a tempered posterior

Easier to sample from



G. Behrens, N. Friel & M. Hurn. (2012). Tuning tempered transitions. $\it Statistics \ and \ Computing.$

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Robust to model misspecification



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Theoretical analysis easier



A. Bhattacharya, D. Pati & Y. Yang (2016). Bayesian fractional posteriors. *Preprint arxiv*:1611.01125.

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$$\begin{split} \tilde{\pi}_{n,\alpha}^K &= \arg\min_{\rho \in \mathcal{F}_K} \mathcal{K}(\rho, \pi_{n,\alpha}^K) \\ &= \arg\max_{\rho \in \mathcal{F}_K} \left\{ \alpha \int \ell_n(\theta) \rho(\mathrm{d}\theta) - \mathcal{K}(\rho, \Pi_K) \right\}. \end{split}$$

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Examples:

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parametric approximation

$$\mathcal{F}_{K} = \left\{ \mathcal{N}(\mu, \Sigma) : \mu \in \mathbb{R}^{d}, \Sigma \in \mathcal{S}_{d}^{+} \right\}.$$

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• mean-field approximation, $\Theta = \Theta_1 \times \Theta_2$ and

$$\mathcal{F}_{K} = \{ \rho : \rho(\mathrm{d}\theta) = \rho_{1}(\mathrm{d}\theta_{1}) \times \rho_{2}(\mathrm{d}\theta_{2}) \}.$$

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Several models

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$$\pi = \sum_{K \in \mathbb{N}^*} \pi_K \Pi_K$$

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Tempered posteriors & their VBs

$$\pi_{n,\alpha}^{K}(d\theta_{K}) \propto L_{n}(\theta_{K})^{\alpha}\Pi_{K}(d\theta_{K})$$

and

$$\tilde{\pi}_{n,\alpha}^{K} = \arg\min_{\rho_{K} \in \mathcal{F}_{K}} \mathcal{K}(\rho_{K}, \pi_{n,\alpha}^{K})$$

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ELBO criterion

ELBO maximization program

$$\tilde{\pi}_{n,\alpha}^{K} = \arg\max_{\rho_{K} \in \mathcal{F}_{K}} \left\{ \alpha \int \ell_{n}(\theta_{K}) \rho_{K}(d\theta_{K}) - \mathcal{K}(\rho_{K}, \Pi_{K}) \right\}$$

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ELBO

$$\mathcal{L}(K) = \alpha \int \ell_n(\theta_K) \tilde{\pi}_{n,\alpha}^K(d\theta_K) - \mathcal{K}(\tilde{\pi}_{n,\alpha}^K, \Pi_K)$$

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Model selection criterion

$$\hat{\mathcal{K}} = \arg\max_{\mathcal{K} \geq 1} \left\{ \mathcal{L}(\mathcal{K}) - \log\left(\frac{1}{\pi_{\mathcal{K}}}\right) \right\}$$

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Technical condition for posterior concentration

If there is a true model $(\exists K_0, \exists \theta^0 \in \Theta_{K_0}, P_{\theta^0} = P^0)$:

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Prior mass condition for concentration of tempered posteriors

The rate (r_n) is such that

$$\Pi_{K_0}[\mathcal{B}(r_n)] \geq e^{-nr_n}$$

where
$$\mathcal{B}(r) = \{\theta \in \Theta_{K_0} : \mathcal{K}(P_{\theta^0}, P_{\theta}) \leq r\}.$$

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Prior mass condition for concentration of Variational Bayes

The rate (r_n) is such that there exists $\rho_{n,K_0} \in \mathcal{F}_{K_0}$ such that

$$\int \mathcal{K}(P_{\theta^0}, P_{\theta}) \rho_{n, K_0}(\mathrm{d}\theta) \leq r_n, \ \text{and} \ \mathcal{K}(\rho_{n, K_0}, \Pi_{K_0}) \leq n r_n.$$

Consistency of the true approximation



P. Alquier & J. Ridgway (2017). Concentration of Tempered Posteriors and of their Variational Approximations. *Preprint arxiv* :1706.09293.

$\mathsf{Theorem}$

If there is a true model $(\exists K_0, \exists \theta^0 \in \Theta_{K_0}, P_{\theta^0} = P^0)$, then under the prior mass condition, for any $\alpha \in (0,1)$,

$$\mathbb{E}\bigg[\int D_{\alpha}(P_{\theta},P^{0})\tilde{\pi}_{n,\alpha}^{K_{0}}(d\theta)\bigg] \leq \frac{1+\alpha}{1-\alpha}r_{n}.$$

Consistency of the selected approximation

Theorem

If there is a true model $(\exists K_0, \exists \theta^0 \in \Theta_{K_0}, P_{\theta^0} = P^0)$, then under the prior mass condition, for any $\alpha \in (0,1)$,

$$\mathbb{E}\bigg[\int D_{\alpha}(P_{\theta},P^{0})\tilde{\pi}_{n,\alpha}^{\hat{K}}(d\theta)\bigg] \leq \frac{1+\alpha}{1-\alpha}r_{n} + \frac{\log(\frac{1}{\pi_{K_{0}}})}{n(1-\alpha)}.$$

Robustness to misspecification

Theorem

For any $\alpha \in (0,1)$, for any K, for any $\rho_K \in \mathcal{F}_K$,

$$\begin{split} &\mathbb{E}\bigg[\int D_{\alpha}(P_{\theta},P^{0})\tilde{\pi}_{n,\alpha}^{\hat{K}}(d\theta)\bigg] \\ &\leq \frac{\alpha}{1-\alpha}\int \mathcal{K}(P^{0},P_{\theta_{K}})\rho_{n,K}(d\theta_{K}) + \frac{\mathcal{K}(\rho_{K},\Pi_{K})}{n(1-\alpha)} + \frac{\log(\frac{1}{\pi_{K}})}{n(1-\alpha)}. \end{split}$$

Example: Univariate Gaussian mixture models

The true distribution P^0 is such that $\mathbb{E}|X| < +\infty$. Let L > 1, $\pi_K = 2^{-K}$, $\Pi_K = \mathcal{D}_K(\alpha_1, \dots, \alpha_K) \bigotimes \mathcal{N}(0, \mathcal{V}^2)^{\otimes n}$ and

$$r_{n,K} = \left\lceil \frac{8K \log(nK)}{n} \bigvee \left(\frac{8K \log(nV)}{n} + \frac{8KL^2}{nV^2} \right) \right\rceil + \frac{K \log(2)}{n(1-\alpha)}.$$

Theorem

For any $\alpha \in (0,1)$,

$$\begin{split} \mathbb{E}\bigg[\int &D_{\alpha}\big(P_{\theta},P^{0}\big)\tilde{\pi}_{n,\alpha}^{\hat{K}}(d\theta)\bigg] \\ &\leq \inf_{K\geq 0}\bigg\{\frac{\alpha}{1-\alpha}\inf_{\theta^{*}\in\mathcal{S}_{K}\times[-L,L]^{K}}\mathcal{K}(P^{0},P_{\theta^{*}})+\frac{1+\alpha}{1-\alpha}r_{n,K}\bigg\}. \end{split}$$

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Gaussian mixtures :
$$heta = (p, (m_1, \sigma_1^2), ..., (m_K, \sigma_K^2))$$

$$\mathbb{E}\bigg[\int D_{\alpha}(P_{\theta}, P^{0})\tilde{\pi}_{n,\alpha}^{\hat{K}}(d\theta)\bigg] = \mathcal{O}\bigg(\frac{K_{0}\log(nK_{0})}{n}\bigg)$$

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Probabilistic PCA : $\theta \in \mathbb{R}^{d \times K}$

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Main result Application

Thank you!