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Rayleigh distribution

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Not to be confused with [Rayleigh mixture distribution](#).

In [probability theory](#) and [statistics](#), the **Rayleigh distribution** /ˈreɪli/ is a [continuous probability distribution](#) for positive-valued [random variables](#).

A Rayleigh distribution is often observed when the overall magnitude of a vector is related to its directional [components](#). One example where the Rayleigh distribution naturally arises is when wind velocity is analyzed into its orthogonal 2-dimensional vector components. Assuming that the magnitudes of each component are [uncorrelated](#), [normally distributed](#) with equal [variance](#), and zero [mean](#), then the overall wind speed ([vector](#) magnitude) will be characterized by a Rayleigh distribution. A second example of the distribution arises in the case of random complex numbers whose real and imaginary components are i.i.d. (independently and identically distributed) [Gaussian](#) with equal variance and zero mean. In that case, the absolute value of the complex number is Rayleigh-distributed.

The distribution is named after [Lord Rayleigh](#).^[*citation needed*]

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Definition [edit]

The [probability density function](#) of the Rayleigh distribution is^[1]

$$f(x;\sigma) = \frac{x}{\sigma^2}e^{-x^2/(2\sigma^2)}, \quad x \geq 0,$$

where $\sigma > 0$, is the [scale parameter](#) of the distribution. The [cumulative distribution function](#) is^[1]

$$F(x) = 1 - e^{-x^2/(2\sigma^2)}$$

for $x \in [0, \infty]$.

Relation to random vector lengths [edit]

Consider the two-dimensional vector $\mathbf{Y} = (U, V)$ which has components that are Gaussian-distributed and independent.

Then $f_U(u;\sigma) = \frac{e^{-u^2/2\sigma^2}}{\sqrt{2\pi\sigma^2}}$, and similarly for $f_V(v;\sigma)$.

Let x be the length of \mathbf{Y} . It is distributed as

$$f(x;\sigma) = \frac{1}{2\pi\sigma^2} \int_{-\infty}^{\infty} du \int_{-\infty}^{\infty} dv \, e^{-u^2/2\sigma^2} e^{-v^2/2\sigma^2} \delta(x - \sqrt{u^2 + v^2}).$$

By transforming to the [polar coordinate system](#) one has

$$f(x;\sigma) = \frac{1}{2\pi\sigma^2} \int_0^{2\pi} d\phi \int_0^{\infty} dr \, \delta(r - x) r e^{-r^2/2\sigma^2} = \frac{x}{\sigma^2} e^{-x^2/2\sigma^2},$$

which is the Rayleigh distribution. It is straightforward to generalize to vectors of dimension other than 2. There are also generalizations when the components have unequal variance or correlations.

Properties [edit]

The raw [moments](#) are given by:

$$\mu_k = \sigma^k 2^{\frac{k}{2}} \Gamma\left(1 + \frac{k}{2}\right)$$

where $\Gamma(z)$ is the [Gamma function](#).

The [mean](#) and [variance](#) of a Rayleigh [random variable](#) may be expressed as:

$$\mu(X) = \sigma \sqrt{\frac{\pi}{2}} \approx 1.253\sigma$$

and

$$\mathrm{var}(X) = \frac{4 - \pi}{2} \sigma^2 \approx 0.429\sigma^2$$

The mode is σ and the maximum pdf is

$$f_{\max} = f(\sigma;\sigma) = \frac{1}{\sigma} e^{-\frac{1}{2}} \approx \frac{1}{\sigma} 0.606$$

The [skewness](#) is given by:

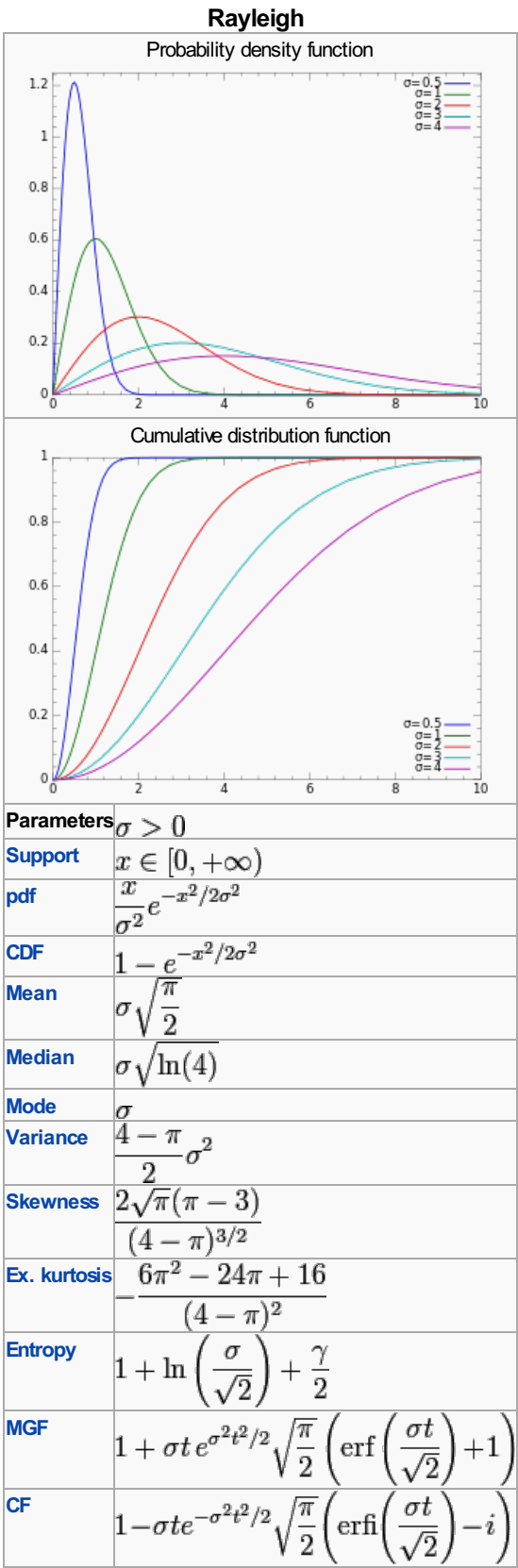
$$\gamma_1 = \frac{2\sqrt{\pi}(\pi - 3)}{(4 - \pi)^{\frac{3}{2}}} \approx 0.631$$

The excess [kurtosis](#) is given by:

$$\gamma_2 = -\frac{6\pi^2 - 24\pi + 16}{(4 - \pi)^2} \approx 0.245$$

The [characteristic function](#) is given by:

$$\varphi(t) = 1 - \sigma t e^{-\frac{1}{2}\sigma^2 t^2} \sqrt{\frac{\pi}{2}} \left[\operatorname{erfi}\left(\frac{\sigma t}{\sqrt{2}}\right) - i \right]$$



where erfi(*z*) is the imaginary error function. The moment generating function is given by

M
(
t
)
=
1
+
σ
t

e

1

2

σ

2

t

2

√

π

2

⌈
erf
⁡
(

σ
t

√
2

)
+
1
⌋

{\displaystyle M(t)=1+\sigma te^{\frac {1}{2}\sigma ^{2}t^{2}}{\sqrt {\frac {\pi }{2}}}\left[\mathrm {erf} \left({\frac {\sigma t}{\sqrt {2}}}\right)+1\right]}

where erf(*z*) is the error function.

Differential entropy [edit]

The differential entropy is given by^[*citation needed*]

H
=
1
+
ln
⁡
(

σ

√
2

)
+

γ

2

{\displaystyle H=1+\ln \left({\frac {\sigma }{\sqrt {2}}}\right)+{\frac {\gamma }{2}}}

where γ is the Euler–Mascheroni constant.

Differential equation [edit]

⌈

σ

2

x

f
′
(
x
)
+
f
(
x
)
(

x

2

−

σ

2

)
=
0
,
f
(
1
)
=

e

−

1

2
σ

2

x

σ

2

⌋

{\displaystyle \left\{\sigma ^{2}xf'(x)+f(x)\left(x^{2}-\sigma ^{2}\right)=0,f(1)={\frac {e^{-{\frac {1}{2\sigma ^{2}}x}}}{\sigma ^{2}}}\right\}}

Parameter estimation [edit]

Given a sample of *N* independent and identically distributed Rayleigh random variables *x*_{*i*} with parameter σ,

σ

2

≈

1

2
N

∑

i
=
1

N

x

i

2

{\displaystyle {\widehat {\sigma ^{2}}}\approx {\frac {1}{2N}}\sum _{i=1}^{N}x_{i}^{2}}

 is an unbiased maximum likelihood estimate.

σ
^
≈

1

2
N

∑

i
=
1

N

x

i

2

{\displaystyle {\hat {\sigma }}\approx {\sqrt {\frac {1}{2N}\sum _{i=1}^{N}x_{i}^{2}}}}

 is a biased estimator that can be corrected via the formula

σ
=

σ
^

Γ
(
N
)

√
N

Γ
(
N
+

1
2

)

=

σ
^

4

N

N
!
(
N
−
1
)
!

√
N

[2]

{\displaystyle \sigma ={\hat {\sigma }}{\frac {\Gamma (N){\sqrt {N}}}{\Gamma (N+{\frac {1}{2}})}}={\hat {\sigma }}{\frac {4^{N}N!(N-1)!{\sqrt {N}}}{(2N)!{\sqrt {\pi }}}}[2]}

Confidence intervals [edit]

To find the (1 − *α*) confidence interval, first find the two numbers *χ*₁², *χ*₂² where:

Pr
(

χ

2

(
2
N
)
≤

χ

1

2

)
=
α
/
2
,

Pr
(

χ

2

(
2
N
)
≤

χ

2

2

)
=
1
−
α
/
2

{\displaystyle Pr(\chi ^{2}(2N)\leq \chi _{1}^{2})=\alpha /2,\quad Pr(\chi ^{2}(2N)\leq \chi _{2}^{2})=1-\alpha /2}

then

N

x
¯

2

χ

2

2

≤

σ
^

2

≤

N

x
¯

2

χ

1

2

[3]

{\displaystyle {\frac {N{\overline {x}}^{2}}{\chi _{2}^{2}}}\leq {\hat {\sigma }}^{2}\leq {\frac {N{\overline {x}}^{2}}{\chi _{1}^{2}}}[3]}

Generating random variates [edit]

Given a random variate *U* drawn from the uniform distribution in the interval (0, 1), then the variate

X
=
σ

√
−
2
ln
⁡
(
U
)

{\displaystyle X=\sigma {\sqrt {-2\ln(U)}}}

has a Rayleigh distribution with parameter σ. This is obtained by applying the inverse transform sampling-method.

Related distributions [edit]

- R* ∼ Rayleigh(σ) is Rayleigh distributed if *R* = √*X*² + *Y*², where *X* ∼ *N*(0,σ²) and *Y* ∼ *N*(0,σ²) are independent normal random variables.^[4] (This gives motivation to the use of the symbol "sigma" in the above parameterization of the Rayleigh density.)
- The chi distribution with *v* = 2 is equivalent to Rayleigh Distribution with σ = 1. I.e., if *R* ∼ Rayleigh(1), then *R*² has a chi-squared distribution with parameter *N*, degrees of freedom, equal to two (*N* = 2)

[
Q
=

R

2

]
∼

χ

2

(
N
)
.

{\displaystyle [Q=R^{2}]\sim \chi ^{2}(N)\, .}

- If *R* ∼ Rayleigh(σ), then

∑

i
=
1

N

R

i

2

{\displaystyle \sum _{i=1}^{N}R_{i}^{2}}

 has a gamma distribution with parameters *N* and 2σ²

[
Y
=

∑

i
=
1

N

R

i

2

]
∼
Γ
(
N
,
2

σ

2

)
.

{\displaystyle \left[Y=\sum _{i=1}^{N}R_{i}^{2}\right]\sim \Gamma (N,2\sigma ^{2}).}

- The Rice distribution is a generalization of the Rayleigh distribution.
- The Weibull distribution is a generalization of the Rayleigh distribution. In this instance, parameter σ is related to the Weibull scale parameter λ: λ = σ√2.
- The Maxwell–Boltzmann distribution describes the magnitude of a normal vector in three dimensions.
- If *X* has an exponential distribution *X* ∼ Exponential(λ), then *Y* = √2*Xσ*²λ ∼ Rayleigh(σ).

Applications [edit]

An application of the estimation of σ can be found in magnetic resonance imaging (MRI). As MRI images are recorded as complex images but most often viewed as magnitude images, the background data is Rayleigh distributed. Hence, the above formula can be used to estimate the noise variance in an MRI image from background data.^[5]

See also [edit]

- Normal distribution
- Rayleigh fading
- Rayleigh mixture distribution



This article includes a list of references, but **its sources remain unclear because it has insufficient inline citations**. Please help to **improve** this article by introducing more precise citations. *(April 2013)*

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Categories: Continuous distributions | Exponential family distributions | Probability distributions

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