CS719: Data Stream Processing

Lecture 14,15, part of 16 Heavy Hitters

> IIT Kanpur Jan-Apr 2010

Outline

Problem Definition and Introduction

 ℓ_1 point query and heavy-hitters: COUNT-MIN sketch COUNT-MIN-sketch: ApproxHH₁ by domain iteration COUNT-MIN sketch: Approx HH by Group Testing

ℓ₂ Point Query and Heavy Hitters: the

COUNTSKETCH structure

COUNTSKETCH: Structure and Simple Analysis COUNTSKETCH and COUNT-MINSketch: Residual F₂ based bounds

Comparing ℓ_1 and ℓ_2 point query estimators

Sparse Approximation



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\text{COUNT-MIN-sketch: ApproxHH}_1 \] by domain iteration
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\$\ell_2\$ Point Query and Heavy Hitters: the

COUNTSKETCH structure

COUNTSKETCH: Structure and Simple Analysis COUNTSKETCH and COUNT-MINSketch: Residual F_2 based bounds

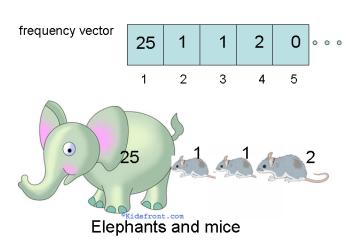
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Heavy Hitters: Illustration

Heavy Hitters are items with large absolute frequencies (Elephants) stream:(1,10)(2,1)(3,1)(4,2)(1,10)...



Heavy Hitter Problem: Find the elephants

Applications

- Among the most popular applications of data streaming.
 - 1. Find the IP-addresses that send the most traffic.
 - Find source-IP, dest-IP pairs that send the most traffic to each other.
 - 3. Find the most visited web sites.

Heavy Hitters: Definition



• ℓ_p heavy hitters with parameter ϕ :

$$HH_{p}^{\phi}(f) = \left\{ i \in [n] : |f_{i}|^{p} \geq \phi \sum_{i \in [n]} |f_{i}|^{p} \right\} .$$

- ℓ_p -heavy hitters problem. First Attempt: Given ϕ , find the set HH^{ϕ}_p over stream.
- ▶ We want to achieve this in low space, close to $O(\frac{1}{\phi})$.
- Finding HH_p^φ exactly requires $\Omega(n)$ space. [Proof later lectures.]
- ► Settle for Approximate Heavy Hitters.



Approximate Heavy Hitters: Definition



- ► Recall: $HH_p^{\phi}(f) = \{i \in [n] : |f_i|^p \ge \phi \sum_{i \in [n]} |f_i|^p \}$.
- ▶ Approximate heavy hitters: ApproxHH_p^{ϕ,ϕ'}. Two parameters ϕ and ϕ' with $\phi' < \phi$.
- Return set S such that
 - 1. *S* includes HH_p^{ϕ} . Meaning: Do not miss *i* with $|f_i|^p > \phi F_p$.
 - 2. *S* is included in $HH_p^{\phi'}$. Do not include *i* with $|f_i|^p < \phi' F_p$.

Simple Fact



▶ Since $HH^{\phi}_{p}(f) = \left\{i \in [n] : |f_i|^p \ge \phi \sum_{i \in [n]} |f_i|^p\right\}$.

$$\left|\mathsf{HH}^{\phi}_{p}\right| \leq \left\lfloor \frac{1}{\phi} \right\rfloor \ .$$

- ▶ Let $S = \text{Approximate Heavy Hitters } (\phi, \phi').$
- ▶ Since $S \subseteq HH_p^{\phi'}$, number of approximate HH (ϕ, ϕ') is at most $\lfloor 1/\phi' \rfloor$.

ℓ_p Point Query/Estimating Frequencies

▶ Point query: Estimate frequency of any item i. Cannot be done exactly in o(n) space. Allow bounded error:

$$\hat{f}_i^p = f_i^p \pm \phi F_p, \quad \forall i \in [n]$$
.



Frequency Vector



Estimated frequencies



An obvious question

- Which is better: PtQuery₁(φ) or PtQuery₂(φ)?
- i.e., which is smaller: ϕF_1 or $(\phi F_2)^{1/2}$?

Plan for this topic

- ▶ ℓ_1 heavy-hitters.
 - 1. Count-Min Algorithm for ℓ_1 point query.
 - 2. Application to ℓ_1 heavy-hitters.
- ▶ ℓ₂ heavy-hitters.
 - 1. Countsketch algorithm for ℓ_2 point query.
 - 2. Application to ℓ_2 heavy-hitters.
- Brief description of a few other algorithms.

References

- COUNT-MIN algorithm: Cormode, Muthukrishnan. An improved data stream summary: the count-min sketch and its applications. J. Algorithms 55(1): 58-75 (2005).
- \(\ell_1\)-HH: Cormode, Muthukrishnan: What's New: Finding Significant Differences in Network Data Streams. IEEE INFOCOM 2004.
- ► COUNTSKETCH algorithm: Charikar, Chen, Farach-Colton. Finding frequent items in data streams. Theor. Comput. Sci. 312(1): 3-15 (2004).

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Point Query and Heavy Hitters: the COUNTSKETCH structure COUNTSKETCH: Structure and Simple Analysis COUNTSKETCH and COUNT-MINSketch: Residual F₂ based bounds

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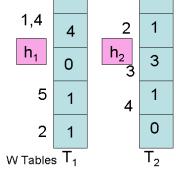


Count-Min Sketch: Basic algorithm

- T_1, T_2, \ldots, T_w
- ► Each table *T_j*:

 1. *B* buckets.
 - 2. hash fn. $h_j:[n] \rightarrow [B]$.
 - 3. $h_j \in_R$ pair-wise indep. family.
- h_j's independent.
 ▶ UPDATE(i, v):
- $T_j[h_j(i)]$ += v, $j = 1, 2, \dots, w$.
- ► ESTIMATE(i):

All non-negative frequencies
$$\hat{f}_i = \min_{i=1}^{w} T_i[h_i(i)]$$



General frequencies $\hat{f}_i = \text{median}_{i=1}^W T_i[h_i(i)]$

1,5

3

3

5

 h_{w}

4

000

3

Simple Analysis: Non-negative frequencies

▶ Fix table index $j \in \{1, 2, ..., w\}$. Let $T = T_j$ and $h = h_j$.

$$T[h(i)] = f_i + \sum_{h(k) = h(i), i \neq k} f_k .$$
So, $E[T[h(i)] - f_i] = \sum_{k \neq i} f_k \Pr[h(k) = h(i)] = \frac{F_1 - f_i}{B} .$

- ▶ By Markov's ineq. $\Pr\left[T[h(i)] f_i > \frac{2F_1}{B}\right] < \frac{1}{2}$.
- Consider all w tables.

$$\Pr\left[\min_{j=1}^{w}(T_{j}[h_{j}(i)]-f_{i})>\frac{2F_{1}}{B}\right]=\Pr\left[\forall j\ T_{j}[h_{j}(i)]-f_{i}>\frac{2F_{1}}{B}\right]<\left(\frac{1}{2}\right)^{2}$$

$$f_i \leq \min_{j=1}^{w} T_j[h_j(i)] \leq f_i + \frac{2F_1}{B}$$
, with prob. $1 - 1/2^w$.

COUNT-MIN sketch: point query, fi's non-negative

- $\blacktriangleright \hat{f}_i = \min_{j=1}^w T_j[h_j(i)].$
- ► For non-negative frequency vector:

$$f_i \le \hat{f}_i \le f_i + \frac{2F_1}{B}$$
, with prob. $1 - 1/2^w$.

- ► To solve point query s.t. $\hat{f}_i \in f_i \pm \phi F_1$ with prob. 1δ set $B = \frac{1}{2\phi}$ and $w = \log \frac{1}{\delta}$.
- ▶ Space requirement: $Bw = \frac{1}{\phi} \log \frac{1}{\delta}$ counters of size $\log(mn)$, where, $m = \max_i f_i$.

Analysis: $f \in \mathbb{Z}^n$

▶ Fix table $T = T_i$. Now,

$$T[h(i)] - f_i = \sum_{\substack{i \neq k \\ h_j(i) = h_j(k)}} f_k = \sum_{\substack{i \neq k}} f_k x_k$$

where, x_k 's are indicator variables, $x_k = 1$ if h(k) = h(i) and 0 otherwise.

► So,

$$\mathsf{E}\Big[\big|T[h(i)] - f_i\big|\Big] = \mathsf{E}\Big[\big|\sum_{i \neq k} f_k x_k\big|\Big] \le \mathsf{E}\Big[\sum_{i \neq k} |f_k| x_k\Big]$$
$$= \sum_{i \neq k} \frac{f_k}{B} = \frac{F_1 - |f_i|}{B} .$$

By Markov's inequality,

$$\Pr\left[\left|T[h(i)]-f_i\right|>\frac{4F_1}{B}\right]\leq \frac{1}{4}.$$

COUNT-MIN-Sketch: Point query analysis

- ▶ jth hash table gives good estimate $\equiv |T_j[h_j(i)] f_i| > \frac{4F_1}{B}$ with prob. at least 3/4.
- ▶ By classical boosting argument (Chernoff bound), using $w = O(\log \frac{1}{\delta})$ tables

$$\Pr\left[\left|\left(\operatorname{median}_{j=1}^{w}T_{j}[h_{j}(i)]\right)-f_{i}\right|>\frac{4F_{1}}{B}\right]<\delta$$
.

- ► Recall $\hat{f}_i = \text{median}_{j=1}^w T_j[h_j(i)].$
- ▶ Therefore, $|\hat{f}_i f_i| < \frac{4F_1}{B}$ with prob. 1δ .
- Solves point query problem with parameter ϕ for general frequency vectors. Choose $B = \frac{4}{\phi}$, $w = O(\log \frac{1}{\delta})$.

COUNT-MIN-Sketch: Space and Update time

- ► Each counter uses $\log(mn)$ bits. There are Bw counters. So space is $O(\frac{1}{\phi}\log(\frac{1}{\delta}))$ counters.
- Each item is inserted in the right bucket in each table.
- ▶ Since h_j 's are from pair-wise indep. family, $h_j(i)$ is computed in using O(1) + and · operations over a finite field of size O(n).
- ► Time for inserting/updating a table is *O*(1).
- ▶ Update time is $O(w) = O(\log(1/\delta))$.
- ▶ Number of random bits: $O(\log n)$ random bits per hash table.

COUNT-MIN-Sketch: ℓ_1 ApproxHH

Simple (but inelegant) approach:

- 1. Estimate F_1 as \hat{F}_1 correct to within $1 \pm \frac{1}{8}$ with high prob.
- 2. Keep $B = \lceil \frac{12}{\phi} \rceil$ buckets per table, and $w = O(\log(n/\delta))$ buckets.
- 3. Iterate over domain [n] and obtain \hat{f}_i for each i.
- 4. Return *i* with $\hat{f}_i \geq \frac{2\phi}{3}\hat{F}_1$.

COUNT-MIN sketch: ApproxHH₁ $^{\phi,\phi/3}$ by domain iteration

Assumption for simplicity: $\hat{F}_1 = F_1$. ¹

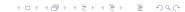
- 1. $|\hat{f}_i f_i| \leq \frac{\phi F_1}{3}$ with prob. $1 \frac{\delta}{n}$.
- 2. If $|f_i| \ge \phi F_1$, then,

$$\hat{f}_i \ge f_i - \frac{\phi F_1}{3} \ge \frac{2\phi F_1}{3}$$
 with prob. $1 - \frac{\delta}{n}$

3. If $|f_i| < \frac{\phi F_1}{3}$, then,

$$\hat{f}_i < f_i + \frac{\phi F_1}{3} \le \frac{\phi F_1}{3} + \frac{\phi F_1}{3} = \frac{2\phi F_1}{3}$$
.

4. All *n* inferences hold jointly with prob. $\geq 1 - \frac{n\delta}{n} = 1 - \delta$.



¹Otherwise, constants change a little.

COUNT-MIN sketch: ApproxHH₁^{\phi,\phi'}

- A possible design:
 - 1. $f_i > \phi F_1$ should imply

$$\hat{f}_i > \frac{(\phi + \phi')F_1}{2} = \phi F_1 - \frac{(\phi - \phi')F_1}{2}$$
.

2. $f_i < \phi' F_1$ should imply

$$\hat{f}_i < \frac{(\phi + \phi')F_1}{2} = \phi' F_1 + \frac{(\phi - \phi')F_1}{2} \ .$$

- Suffices to solve point query with parameter $\frac{\phi \phi'}{2}$ with prob. $1 \frac{\delta}{n}$.
- ▶ Space: $O(\frac{1}{\phi \phi'}(\log \frac{n}{\delta}))$ counters.

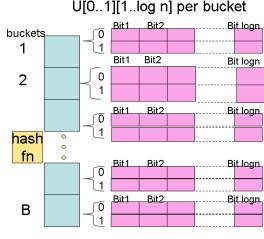
COUNT-MIN sketch: domain iteration

- ► Solves approximate ApproxHH₁^{$\phi,\phi/3$} with prob. 1- δ .
- Resources used
 - 1. Space: $O(\frac{1}{\phi \phi'} \log(\frac{n}{\delta}))$ counters.
 - 2. Update time: $O(\log(\frac{n}{\delta}))$.
 - 3. Discovery time of approximate HH: O(n)—the problem of this approach!
 - 4. Implicit: F_1 has to be estimated.

HH: COUNT-MIN sketch with Group testing

- ► Each bucket T_j[b] now a 2 × log n array U.
- ▶ UPDATE(*i*, *v*):

```
for each table T_j {
Go to bucket h_j(i)
for each bit position I {
    if Ith bit of X is 1 then
    increment U[1][I] by V
else
    increment U[0][I] by V.
```



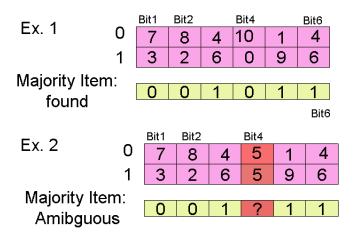
Discovering Candidate Approximate HH: Basic Idea

- ▶ Heavy Hitter threshold ϕF_1 .
- ► For simplicity, let F₁ be known exactly. ²
- ► Use *U*[2][log *n*] array to find one possible candidate.
- ► Candidate Item: abs. value of frequency is at least $(\phi + \phi')F_1/2$.
- If multiple possible candidate, then no inference is made from this array.

 $^{^2}F_1$ is estimated using within accuracy of 1 \pm 1/8 with prob. 1 - δ /2. Replace F_1 by \hat{F}_1 .

Phase I: Discovering Candidate Approx. HH

Assume: Non-negative frequencies



Candidate HH: an umambiguous majority item.

Analysis

- Suffices to show that every heavy hitter is a majority item in some table bucket, with high probability.
- ▶ Let *i* be a HH₁^{ϕ} item, so $|f_i| > \phi F_1$.
- ▶ Prior analysis: $\left|T_{j}[h_{j}(i)] |f_{i}|\right| < \frac{4(F_{1} |f_{i}|)}{B}$ with prob. 3/4.
- ▶ So *i* is majority item in bucket $h_i(i)$ with prob. 3/4 if

$$\frac{4(F_1-|f_i|)}{B}<\frac{\phi F_1}{2}.$$

- ▶ Suffices if $B > \lceil \frac{8}{\phi} \rceil + 1$.
- ▶ Prob. that *i* is not a majority item in any of its buckets is at most $\frac{1}{4^w}$.

Width of majority structure

- ▶ Height of each hash table can be $\lceil \frac{8}{\phi} \rceil + 1$ or larger.
- ▶ Width w:
 - 1. Prob. that a heavy hitter is a majority item in some bucket is $1 4^{-w}$.
 - 2. There are at most $\lceil \frac{1}{\phi} \rceil$ heavy hitters.
 - 3. Prob. that each heavy hitter is a majority item in some bucket is $1 \frac{4^{-w}}{\phi}$.
 - 4. To ensure that all heavy hitters are detected with a prob. of 1 $-\delta$, set

$$\frac{4^{-w}}{\phi} < \delta$$
 or, $w > \frac{1}{2} \log \frac{1}{\phi \delta}$.

▶ Number of candidate heavy hitters at most $Bw = O\left(\frac{1}{\phi}\log\frac{1}{\phi\delta}\right)$.



Phase 2: Verifying Candidate Heavy Hitters

- ▶ Discovery phase gives at most $K = O\left(\frac{1}{\phi}\log\frac{1}{\phi\delta}\right)$ candidate heavy hitters.
- ▶ We now use a Count-Min sketch structure to solve the $PtQuery_1(\phi, \phi')$ for each of the K candidate heavy hitters.
- ▶ Items that cross the threshold $\hat{f}_i \ge \frac{(\phi + \phi')F_1}{2}$ are returned, rest are rejected.

Resources Consumed: ApproxHH $_1^{\phi,\phi'}$

- ▶ Space required by PtQuery₁(ϕ, ϕ') data structure is $O\left(\frac{1}{\phi \phi'} \log \frac{K}{\delta}\right)$ counters.
- ▶ Discovering the heavy hitters requires time $O(K \log \frac{K}{\delta})$.
- ▶ Updating the structures requires time $O\left(\log \frac{1}{\phi \delta}\right)$.
- ► Final Guarantee: ApproxHH₁^{ϕ,ϕ'} is solved with prob. 1 -2δ .

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ℓ₂ heavy hitters: the Countsketch structure

- ► COUNTSKETCH structure:
 - 1. w tables T_1, \ldots, T_w .
 - 2. $h_j : [n] \rightarrow [B]$ corresponding to T_i .
 - 3. *h_j* randomly chosen from a pair-wise indep. family.
 - 4. h_1, \ldots, h_w are independently chosen.
 - 5. Sketch fn. $\xi_j : [n] \to \{-1, +1\}$ corresponding to T_j , 4-wise independent.

6.

$$T_j[b] = \sum_{i:h_j(i)=b} f_i \xi_j(i)$$

Each bucket keeps AMS sketch of sub-stream mapping to it

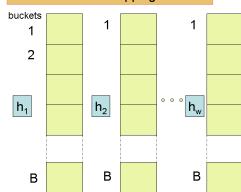


Table T₁

Table T₂

Table T_w

Countsketch Structure

$$b = 1, \ldots, B, j = 1, 2, \ldots, vv$$
.

COUNTSKETCH structure

Each bucket keeps AMS sketch of sub-stream mapping to it

- ► UPDATE(i, v): for j = 1 to w { $T_j[h_j(i)] += v \cdot \xi_j(i)$ }
- ► ESTIMATE(i):

$$\hat{f}_i = \mathsf{median}_{j=1}^w T_j[h_j(i)] \cdot \xi_j(i)$$

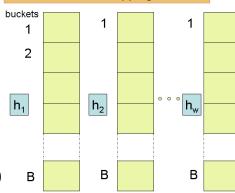


Table T₁ Table T₂

Table T_w

Countsketch Structure

Frequency recovery: Basic idea

- ▶ Let $X = \sum_{k \in [n]} f_k \xi(k)$ be AMS sketch.
- ▶ Consider $X \cdot \xi(i)$.

$$X \cdot \xi(i) = f_i(\xi(i))^2 + \sum_{\substack{k \in [n] \\ k \neq i}} f_k \xi(k) \xi(i) .$$

▶ By linearity of expect. and by pair-wise indep. of $\xi(\cdot)$, $E[\xi(k)\xi(i)] = E[\xi(k)]E[\xi(i)] = 0 \cdot 0 = 0$, we have

$$E[X \cdot \xi(i)] = f_i$$
.

Variance calculation.

$$Var[X \cdot \xi(i)] = E[(X \cdot \xi(i))^{2}] - f_{i}^{2} = F_{2} - f_{i}^{2}$$
.



COUNTSKETCH: Simple Analysis

- ▶ Useful fact: h and ξ use independent random bits.
- ▶ Fix table index *j*. Let $T = T_j$, $h = h_j$, $\xi_j = \xi$.
- ► Fix arbitrary item i.
- ▶ Let $X_i = T[h(i)]$ be sketch of sub-stream mapping to bucket i.
- ▶ Then, $E_{\xi}[X_i \cdot \xi(i)] = f_i$, from previous slide. This is independent of h.
- And

$$\operatorname{Var}_{\xi} \left[X_i \cdot \xi(i) \right] = \sum_{\substack{k: h(k) = h(i) \\ k \neq i}} f_k^2$$

which is a random function of h (and not of ξ).

COUNTSKETCH: Simple Analysis

► Define event Lowvar(i) as

$$\operatorname{Var}_{\xi}\left[X_{i}\cdot\xi(i)\right]\leq 5\operatorname{Var}_{h,\xi}\left[X_{i}\cdot\xi(i)\right]$$

or equivalently

$$\sum_{\substack{k: h(k) = h(i) \\ k \neq i}} f_k^2 \le \frac{5(F_2 - f_i^2)}{B} .$$

▶ By Markov's inequality, $Pr_h[LOWVAR(i)] \ge 4/5$.

COUNTSKETCH: Simple Analysis

•

$$\mathsf{E}_{h}\left[\mathsf{Var}_{\xi}\left[X_{i}\cdot\xi(i)\right]\right] = \mathsf{E}_{h}\left[\sum_{\substack{k:h(k)=h(i)\\k\neq i}} f_{k}^{2}\right]$$
$$= \sum_{\substack{k\neq i}} f_{k}^{2} \cdot \mathsf{Pr}\left[h(k) = h(i)\right] = \frac{F_{2} - f_{i}^{2}}{B} \ .$$

Also

$$\operatorname{Var}_{h,\xi}\left[X_i \cdot \xi(i) \mid \operatorname{LOWVAR}(i)\right] = \operatorname{E}_h\left[\operatorname{Var}_{\xi}\left[X_i \cdot \xi(i)\right] \mid \operatorname{LOWVAR}(i)\right] \\ \leq 5(F_2 - f_i^2)/B \ .$$

COUNTSKETCH: Simple Analysis

► By Chebychev's inequality

$$\begin{aligned} & \mathsf{Pr}_{h,\xi} \bigg[\big| X_i \cdot \xi(i) - f_i \big| > \bigg(\frac{25(F_2 - f_i)}{B} \bigg)^{1/2} \bigg| \mathsf{LOWVAR}(i) \bigg] \\ & \leq \frac{\mathsf{Var}_{h,\xi} \left[X_i \cdot \xi(i) \mid \mathsf{LOWVAR}(i) \right]}{25(F_2 - f_i^2)/B} \\ & \leq \frac{5(F_2 - f_i^2)/B}{25(F_2 - f_i^2)/B} \\ & = \frac{1}{5} \ . \end{aligned}$$

Point query estimator

▶ Unconditioning w.r.t. event LOWVAR(i)

$$\Pr\left[\left|X_i \cdot \xi(i) - f_i\right| \le \left(\frac{25F_2}{B}\right)^{1/2}\right] \ge 1 - \frac{1}{5} - \frac{1}{5} = \frac{3}{5} .$$

- ▶ Let $X_{j,i} = T_j[h_j(i)]$.
- $\hat{f}_i = \text{median}_{j=1}^w X_{j,i} \cdot \xi_j(i).$
- ► Then, by independence of h_j 's and ξ_j 's, by standard boosting argument we have

$$\Pr\left[\left|\hat{f}_i - f_i\right| \le \left(\frac{25F_2}{B}\right)^{1/2}\right] > 1 - \delta.$$

- ▶ PtQuery₂(ϕ) : $\hat{f}_i \in f_i \pm (\phi F_2)^{1/2}$.
- ► COUNTSKETCH solves PtQuery₂(ϕ). Set $B = \lceil \frac{25}{\phi} \rceil$.



COUNTSKETCH Point Query Estimator

- ▶ Basic guarantee: Let $B = \lceil \frac{25}{\phi} \rceil$. Then, $|\hat{f}_i f_i| < (\phi F_2)^{1/2}$ with prob. 1δ .
- ► Space required: $O\left(\frac{1}{\phi}\log\frac{1}{\delta}\right)$.
- ► Random bits: $2 \log n$ random bits for each h_j and $4 \log n$ random bits for each ξ_j . Total = $O(\log n \cdot \log \frac{1}{\delta})$.
- ▶ Update time: $O(\log \frac{1}{\delta})$.
- ► Solving point query: time $O(\log \frac{1}{\delta})$.

Point query: (slightly) better analysis

Order items in non-increasing order of absolute frequencies

$$|f_{\mathcal{S}_1}| \geq |f_{\mathcal{S}_2}| \geq \ldots \geq |f_{\mathcal{S}_n}|$$
.

- ▶ A Top-k item is one of $s_1, s_2, ..., s_k$, item with one of the top-k frequencies.
- ▶ Hash table *T* with *pairwise independent* hash fn *h*.
- ▶ Say that item *i* collides with item *j* if h(i) = h(j) and $i \neq j$.
- ► Event NoCollision(i) $\equiv i$ does not collide with any of the Top-k items.
- ▶ Property: $Pr[NoCollision(i)] \ge 1 k/B$.

No Collision with any Top-k items

- ► Fix item i.
- ▶ Let j be a fixed Top-k item, $j \neq i$. Then,

$$Pr[i \text{ collides with } j] = Pr[h(i) = h(j)] = \frac{1}{B}$$
.

► There are *k* TOP-*k* items. So by union bound

$$\Pr[i \text{ collides with one of the Top-}k \text{ items}] \leq \frac{k}{B}$$
.

► So, $Pr[NoCollision(i)] \ge 1 - \frac{k}{B}$, for any fixed *i*.



Consequence of NoCollision(i)

- ▶ Leads to smaller error in inference made from bucket *h*(*i*).
- Notion of residual moments.
- ▶ Define

$$F_p^{\text{res}}(k) = \sum_{i=k+1}^n |f_{s_i}|^p$$
.

► So

$$F_1^{\text{res}}(k) = \sum_{i=k+1}^n |f_{s_i}| \text{ and } F_2^{\text{res}}(k) = \sum_{i=k+1}^n |f_{s_i}|^2.$$

COUNTSKETCH PtQuery₂(ϕ)

- ► Condition analysis on LOWVAR(i) and NoCollision(i).
- ▶ Set B = 8k. Then, $\Pr[NoCollision(i)] \ge \frac{7}{8}$, $k = \lceil 1/\phi \rceil$.
- ► From previous calculation

$$\operatorname{Var}_{\xi}\left[T[h(i)] \cdot \xi(i)\right] = \sum_{k: h(k) = h(i), k \neq i} f_k^2.$$

► Assume NoCollision(i) holds. Then,

$$\operatorname{Var}_{\xi} \left[T[h(i)] \cdot \xi(i) \right] = \sum_{\substack{k: h(k) = h(i), k \neq i \\ k \notin \operatorname{TOP} - \mathbf{k}}} f_k^2 .$$

So

$$\mathsf{E}_h\left[\mathsf{Var}_\xi\left[T[h(i)]\cdot\xi(i)
ight]\left|\mathsf{NoCollision}(i)
ight] \leq rac{F_2^{\mathsf{res}}(k)}{B}$$
 .

Analysis contd.

► By Markov's inequality

$$\Pr_{h}\left[\operatorname{Var}_{\xi}\left[T[h(i)]\cdot \xi(i)
ight] > rac{8F_{2}^{\operatorname{res}}}{B} \middle| \operatorname{NoCollision}(i)
ight] \leq rac{1}{8} \ .$$

► Change constant in defn of LOWVAR(i):

$$\mathsf{LOWVAR}(i) \equiv \mathsf{Var}_{\xi}\left[T[h(i)] \cdot \xi(i)\right] < \frac{8F_2^{\mathsf{res}}(k)}{B} \ .$$



Analysis

By Chebychev's inequality,

$$\begin{aligned} \Pr_{h,\xi} \bigg\{ \big| T[h(i)] \cdot \xi(i) - f_i \big| &> \bigg(\frac{64 F_2^{\text{res}}(k)}{B} \bigg)^{1/2} \bigg| \text{NoCollision}(i) \land \text{Lowvar}(i) \bigg\} \\ &\leq \\ \frac{\text{Var}_{h,\xi} \left[\big| T[h(i)] \cdot \xi(i) \big| \quad \text{NoCollision}(i) \land \text{Lowvar}(i) \right]}{64 F_2^{\text{res}}(k) / B} \\ &\leq \\ \frac{8 F_2^{\text{res}}(k) / B}{64 F_2^{\text{res}}(k) / B} = \frac{1}{8} \end{aligned} .$$

Analysis contd.

Already shown (assuming B = 8k)

$$\Pr_{h,\xi} \left\{ \left| T[h(i)] \cdot \xi(i) - f_i \right| \le \left(\frac{64F_2^{\text{res}}(k)}{8k} \right)^{1/2} \middle| \text{NoCollision}(i) \land \text{Lowvar}(i) \right\} \\ \ge \frac{7}{8} .$$

 $\Pr[\mathsf{NoCollision}(i)] \ge \frac{7}{8}, \quad \Pr[\mathsf{Lowvar}(i) \mid \mathsf{NoCollision}(i)] \ge \frac{7}{8}$

From basic probability:

$$\Pr\left[E\right] = \Pr\left[E \mid A \wedge B\right] \cdot \Pr\left[B \middle| A\right] \cdot \Pr\left[A\right] \ .$$

Therefore,

$$\Pr_{h,\xi} \left\{ \left| T[h(i)] \cdot \xi(i) - f_i \right| > \left(\frac{64F_2^{\text{res}}(k)}{8k} \right)^{1/2} \right\} \ge \frac{7}{8} \cdot \frac{7}{8} \cdot \frac{7}{8} > 2/3 \ .$$

Point Query: Detailed Analysis

- ▶ We have $\Pr_{h,\xi} \left\{ \left| T[h(i)] \cdot \xi(i) f_i \right| > \left(\frac{8F_2^{\text{res}}(k)}{k} \right)^{1/2} \right\} \geq \frac{2}{3}$.
- ▶ Keep $w = O(\log \frac{1}{\delta})$ independent tables $T_1, T_2, ..., T_w$.
- ▶ Let $\hat{f}_i = \text{median}_{j=1}^w T_j[h_j(i)] \cdot \xi_j(i)$.
- ► Then,

$$\Pr\left\{\left|\hat{f}_i - f_i\right| > \left(\frac{8F_2^{\text{res}}(k)}{k}\right)^{1/2}\right\} \geq 1 - \delta$$
.

No change in algorithm, Space increases by a factor of 2, update time does not change. Only analysis is stronger.

COUNT-MIN Sketch: Detailed Analysis

- ▶ Same idea can be applied. Choose B = 4k.
- Final guarantee obtained as

$$\Pr\Big\{ \left| \hat{f}_i - f_i \right| > \left(\frac{F_1^{\text{res}}(k)}{4k} \right) \} \ge 1 - \delta .$$

No change in point query estimator. Space increases by factor of 2. No change in update time.

Outline

Problem Definition and Introduction

ℓ₁ point query and heavy-hitters: COUNT-MIN sketch COUNT-MIN-sketch: ApproxHH₁ by domain iteration COUNT-MIN sketch: Approx HH by Group Testing

ℓ₂ Point Query and Heavy Hitters: the COUNTSKETCH structure

COUNTSKETCH: Structure and Simple Analysis COUNTSKETCH and COUNT-MINsketch: Residual F_2 based bounds

Comparing ℓ_1 and ℓ_2 point query estimators

Sparse Approximation



Which is better?

▶ We have shown estimators for PtQuery $_1^{\phi}$ and PtQuery $_2^{\phi}$ as follows.

$$\begin{split} & \mathsf{PtQuery}_1: \ |\hat{f}_i - f_i| \leq \frac{F_1^{\mathsf{res}}(k)}{k} \\ & \mathsf{PtQuery}_2: \ |\hat{f}_i - f_i| \leq \left(\frac{F_2^{\mathsf{res}}(k)}{k}\right)^{1/2} \ . \end{split}$$

▶ Which is more accurate (more than just constant factors)?

Comparing different norms: Two results

▶

$$\left(\frac{F_2^{\rm res}(k)}{k}\right)^{1/2} \le \frac{F_1}{k} .$$

Many examples where, $F_1^{\text{res}}(k) \approx F_1$.

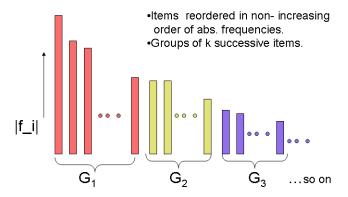
► Follows from inequality:

$$F_q^{\mathsf{res}}(k) \leq rac{F_p^{q/p}}{k^{q/p-1}}, \quad q \geq p$$
 .

A version of Holder's inequality:

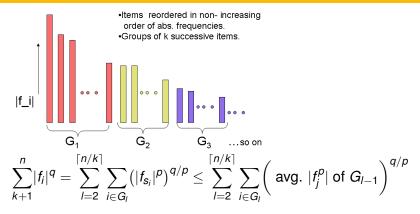
$$\left(\frac{F_p}{F_0}\right)^{1/p} \leq \left(\frac{F_q}{F_0}\right)^{1/q}, \quad q \geq p$$
.

Proof of inequality 1

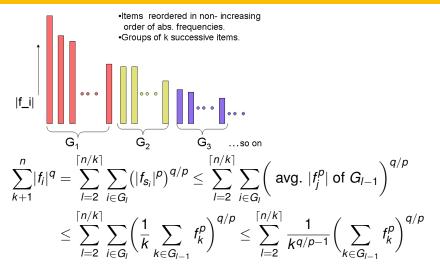


- Arrange items 1 to n in non-increasing order of absolute values of frequency: $|f_{s_1}| \ge |f_{s_2}| \ge ... \ge |f_{s_n}|$.
- Items are grouped k at a time (in non-increasing order), G₁, G₂, ..., G⌊n/k⌋ groups.
- ▶ Last group may have less than *k* items.

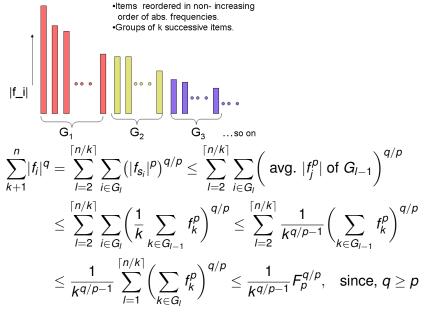
Proof of CS inequality



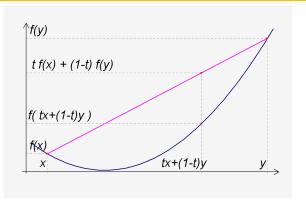
Proof of CS inequality



Proof of CS inequality



Jensen's inequality and convex functions 3



▶ A function ϕ is convex in an interval I of \mathbb{R} if for any $x_1, x_2 \in I$ and 0 < t < 1,

$$\phi(tx_1 + (1-t)x_2) \le t\phi(x_1) + (1-t)\phi(x_2), \quad t \in [0,1]$$
.

▶ Convex function $\phi(X)$ of a random variable X. ϕ is convex in an interval that contains the support of X.



Jensen's inequality

- ▶ Let ϕ be a convex function of a random variable X. Then, $E[\phi(X)] \ge \phi(E[X])$.
- ► Proof:

$$\begin{split} \mathsf{E}\left[\phi(X)\right] &= \sum_{x \in \mathsf{Supp}(X)} \mathsf{Pr}\left[X = x\right] \phi(x) \\ &\geq \phi\bigg(\sum_{x \in \mathsf{Supp}(X)} \mathsf{Pr}\left[X = x\right]\bigg), \quad \text{by convexity of } \phi \\ &= \phi\big(\mathsf{E}\left[X\right]\big) \ . \end{split}$$

Application of Jensen's inequality

- ▶ Let f be an n-dimensional vector. 4
- ▶ Random experiment: choose a random element $i \in_R [n]$. Let $X = |f_i|^p$.
- ▶ $E[X] = \frac{F_p}{n}$.
- ▶ Consider $\phi(X) = X^{q/p}$, where, $q \ge p$. ϕ is convex.
- By Jensen's inequality,

$$\begin{split} & \mathsf{E}\left[X^{q/p}\right] \geq \left(\mathsf{E}\left[X\right]\right)^{q/p} \\ & \text{or, } \frac{F_q}{n} \geq \left(\frac{F_p}{n}\right)^{q/p} \\ & \text{or, } \left(\frac{F_q}{n}\right)^{1/q} \geq \left(\frac{F_p}{n}\right)^{q/p} \;. \end{split}$$



 $^{^4}n$ can be replaced by F_0 .

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Sparse Approximation



Sparse Approximation: Basic Idea

- ▶ x: n-dimensional frequency vector.
- Wish to summarize x using another vector x' that has "low complexity" k.

$$\|\mathbf{x} - \mathbf{x}'\|_{p} \le (1 + \alpha) \mathsf{Err}_{p}^{k}$$

where,

$$\operatorname{Err}_p^k = \min_{x'' \text{ with complexity } k} ||x - x''||_p$$

and α is a parameter.

▶ Notion of complexity: $\|\cdot\|_0$ or number of non-zero coordinates . That is, find x' s.t.

$$||x - x'||_p \le (1 + \alpha) \min_{||x''||_0 \le k} ||x - x''||_p$$
.

Definition

- ▶ A vector is said to be k-sparse if it has at most k non-zero coordinates, i.e., $||x||_0 \le k$.
- ℓ_1 -sparse approximation problem: Find x' such that

$$||x - x'||_1 \le (1 + \phi) \min_{x'' \text{ k-sparse}} ||x - x''||_1$$
.

▶ ℓ_2 -sparse approximation problem: Find x' such that

$$||x - x'||_2 \le (1 + \phi) \min_{x'' \text{ k-sparse}} ||x - x''||_2$$
.

ℓ_1 sparse approximation

- ▶ Min value of $||x x''||_p$ for k-sparse x is attained for x^* set to the top-k values of x (and 0's elsewhere).
- ▶ Let $x_{s_1} \ge x_{s_2} \ge ... \ge x_{s_n}$.
- So

$$\min_{x^{\prime\prime}\text{ k-sparse}} \lVert x - x^{\prime\prime} \rVert_p = \sum_{j=k+1}^n \lvert x_{s_j} \rvert^p, \quad p \geq 0.$$

- ► From PtQuery₁ estimator: $|\hat{x}_i x_i| \le \frac{\sum_{j=k+1}^n |x_{s_j}|}{4k}$ with prob. $1 2^{-\Omega(w)}$.
- ℓ_1 -sparse approximation: x' = top-k of the $|\hat{x}|_i$'s.

ℓ_1 Sparse Approximation

- \triangleright x': top-k by estimated values, x^* : actual top-k values.
- ▶ Arrange x in descending order: $|x_{s_1}| \ge |x_{s_2}| ... \ge |x_{s_k}|$.
- ▶ Keep PtQuery₁ $(\lceil \frac{\phi}{k} \rceil)$ data structure. So,

$$|x_i^* - x_i| \le \frac{\phi}{4k} \sum_{j=\lfloor k/\phi \rfloor} |x_{s_j}|$$
.

▶ Arrange x' in descending order: $|\hat{x}|_{t_1} \ge |\hat{x}|_{t_2} \ldots \ge |\hat{x}|_{t_k}$. So,

$$||x^* - x'|| \le \sum_{j=1}^k |x_{s_j} - \hat{x}_{t_j}| \le k \cdot \frac{\phi}{4k} \sum_{j=\lfloor k/\phi \rfloor} |x_{s_j}| \le \frac{\phi}{4} ||x - x^*||$$

ℓ_1 sparse approximation

- ▶ Min value of $||x x''||_p$ for k-sparse x is attained for x^* set to the top-k values of x (and 0's elsewhere).
- ▶ Let $x_{s_1} \ge x_{s_2} \ge ... \ge x_{s_n}$.
- ► So

$$\min_{x^{\prime\prime}\text{ k-sparse}} \lVert x - x^{\prime\prime} \rVert_p = \sum_{j=k+1}^n \lvert x_{s_j} \rvert^p, \quad p \geq 0.$$

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- ℓ_1 -sparse approximation: x' = top-k of the $|\hat{x}_i|$'s.

Estimating $F_1^{res}(k)$

- ▶ Use PtQuery₁(ϕ/k).
- ▶ Obtain top-k estimated frequencies: $\hat{x}_{t_1} \ge \hat{x}_{t_2} \ge ... \ge \hat{x}_{t_k}$. Denote vector as x'.
- Skimming: Remove contribution of estimated frequencies.

$$X = X - \sum_{i=1}^k \hat{x}_{t_i} \xi_{t_i}, \quad ext{ for all sketches } X.$$

- Apply to Cauchy sketches.
- ▶ After skimming, effective frequency is \bar{x}_i :

$$X = \sum_{i \in [n]} \bar{x}_i \xi(i) .$$

$$\bar{x}_i = \begin{cases} x_i & \text{if } i \text{ is not among top-} k \text{ estimated freq.} \\ x_i - \hat{x}_i & \text{otherwise.} \end{cases}$$

Estimating residual ℓ_1 norm

$$\begin{split} \|\bar{x}\|_1 &= \sum_{j=1}^k |x_{t_j} - \hat{x_{t_j}}| + \|x - x'\|_1 \\ &\leq \frac{\phi}{4} \|x - x^*\|_1 + \big(1 + \frac{\phi}{4}\big) \|x - x^*\|_1 \text{ from previous slide} \\ &\leq \bigg(1 + \frac{\phi}{2}\bigg) \|x - x^*\| \ . \end{split}$$

Guarantee is probabilistic: above result holds with high probability for any x.
Note: Argument is a variation on Chandan Saha's original argument [CS719 course: 2004 Sem-I].