# Random Restrictions on Boolean Functions with Small Influences

Ronen Eldan, Avi Wigderson, Pei Wu Oct., 2022

# Some background

# Influences

 $f: \{0,1\}^n \to \{0,1\}.$ 

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### Definition (influence and max influence)

$$\operatorname{Inf}_{i}[f] = \Pr_{x}[f(x) \neq f(x \oplus e_{i})],$$

$$\operatorname{MaxInf}[f] = \max_{i \in [n]} \operatorname{Inf}_{i}(f).$$

$$|.\mathbf{MAJ}_{\mathbf{n}}: x \mapsto \mathbf{I}_{\{|x| > n/2\}}.$$

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Fact. MaxInf(MAJ<sub>n</sub>) = 
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#### **Proof:**

$$\operatorname{Inf}_{i}(f) = \operatorname{Pr}\left[ \left| |x| - \frac{n}{2} \right| \le 1 \right] \cdot \Theta(1).$$

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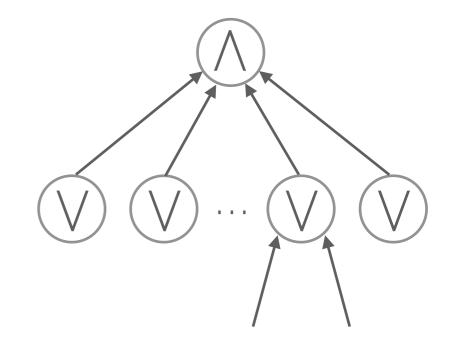
#### **Proof:**

$$\operatorname{Inf}_{i}(f) = \operatorname{Pr}\left[\left| |x| - \frac{n}{2} \right| \le 1 \right] \cdot \Theta(1).$$

e.g. 01010101010, any 0 is sensitive

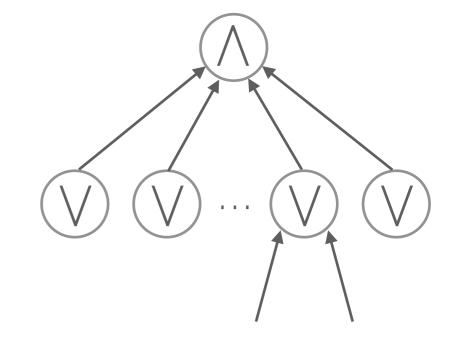
2. TRIBES<sub>n</sub>(x) = 
$$\bigwedge_{i=1}^{S} \bigvee_{j=1}^{w} (x_{i,j})$$
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$$s = n/w, w \approx \log n - \log \log n$$
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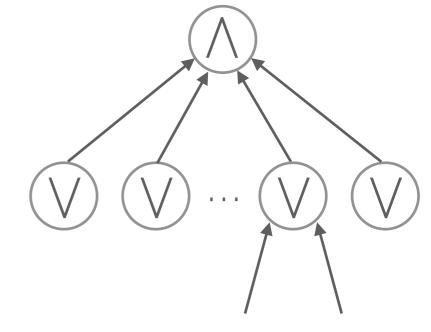
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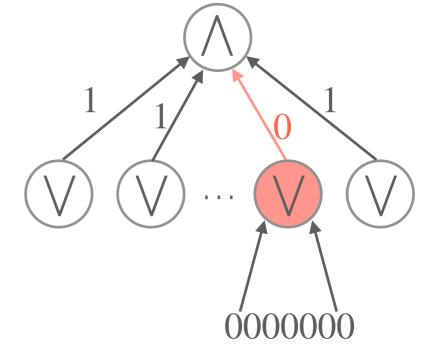
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Suffices to check the case when TRIBES(x) = 0. With  $\Theta(1)$  probability, exactly

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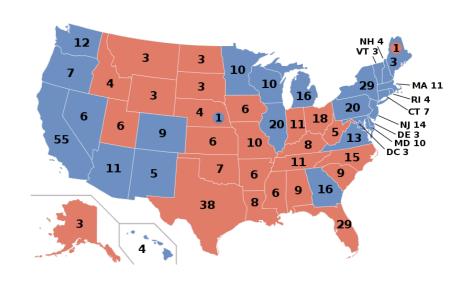




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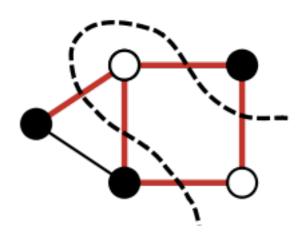
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- 3. Invariance principle

$$\sum X_i \approx Gaussian$$

# Random restrictions

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0	*	1	0	*	1	0	*	0	1			
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 $f: \{0,1\}^n \to \{0,1\}$ .



p-random restriction  $f|_{R_p}$ : fix p random bits

### Theorem (Eldan, Wigderson, W.)

For any  $f: \{0,1\}^n \to \{0,1\}$ , with  $\operatorname{MaxInf}(f) = \tau = o(1)$ ,  $\Omega(1)$  variance. Then for alive probability

$$ho = \tilde{\Omega}\left(\frac{1}{\log 1/ au}\right)$$
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Proof. by Hastad's Switching Lemma

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#### **Remarks:**

- 1. Optimality (w.r.t.  $\rho$ ): the Tribes function.
- 2. Optimality (w.r.t. Var): the Majority function.

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- Koehler-Lifshitz-Minzer-Mossel have a different approach '22

### Corollary.

For any balanced function f, with  $\tau = o(1)$  max influence, then

$$\Pr[\mathsf{bs}_f(x) \ge \tilde{\Omega}(\log(1/\tau))] = 1 - o(1).$$

Sensitivity  $s_f(x)$ : number of sensitive bits

$$s_f(x) := |\{i : f(x) \neq f(x \oplus e_i)\}|.$$

Block sensitivity  $\mathbf{bs}_f(x)$ : max number of disjoint sensitive blocks

$$bs_f(x) := max \mid \{disjoint S_1, S_2, ..., S_k \subseteq [n] : f(x) \neq f(x \oplus 1_{S_i})\} \mid .$$

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Remark: By KKL inequality,

$$\mathbf{E}[\mathrm{bs}_f(x)] \ge \mathbf{E}[\mathrm{s}_f(x)] \ge \Omega(\log(1/\tau)).$$

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Partition [n] into  $M = \tilde{O}(\log(1/\tau))$  random blocks,

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#### Random input *x*

Partition [n] into  $M = \tilde{O}(\log(1/\tau))$  random blocks, Any M-1 blocks, induces a random restriction.

# Proof of the main result

 $f: \{-1,1\}^n \to \{0,1\}$  (by multilinear extension  $f: \mathbb{R}^n \to \mathbb{R}$ )

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- I. Random permutation  $\pi$ ,
- 2.  $X(0) = 0^n$ ,
- 3.  $X_i(t) = X_i(t-1)$ , for  $i \neq \pi(t)$ ;  $X_{\pi(t)}(t) \sim \{-1,1\}$ , for t = 1,2,...,n

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$$X(0)$$
= 0 0 0 0 0 0 0 0 ......  
 $X(1)$ = 0 0 0 0 0 1 0 0 ......  
 $X(2)$ = 0 0 -1 0 0 1 0 0 ......

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### Uniform process

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#### Observation.

- i. X(t) induces a random restriction  $f|_{R(t)}$ ,
- ii. X(n) is a uniformly random element from  $\{-1,1\}^n$ ,
- iii.  $f(X(t)) = \mathbf{E}_{z \in \{-1,1\}^{n-t}}[f|_{R(t)}(z)]$  is a martingale.

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e.g. 
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### Conditioned process

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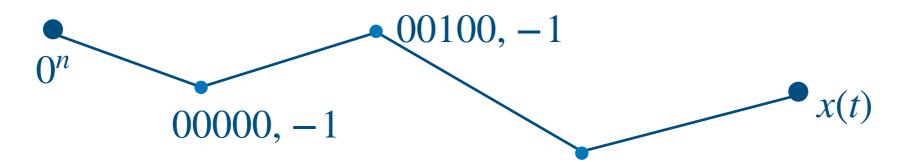
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**Fact.** Fix some some path  $x(0), x(1), ..., x(t) \in \{-1,0,1\}^n$ ,

$$\frac{\Pr[\forall i \in [t], Y(t) = x(t)]}{\Pr[\forall i \in [t], X(t) = x(t)]} = \frac{f(x(t))}{f(0)}.$$



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**Pf.** This = 
$$\prod_{i=1}^{t} 2 \frac{f(x(i-1) + x_{\pi(i)} e_{\pi(i)})}{2f(x(i-1))} = \frac{f(x(t))}{f(0)}.$$





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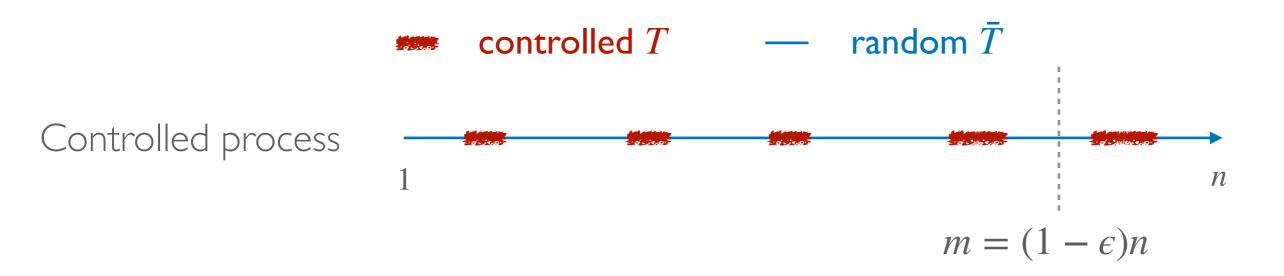
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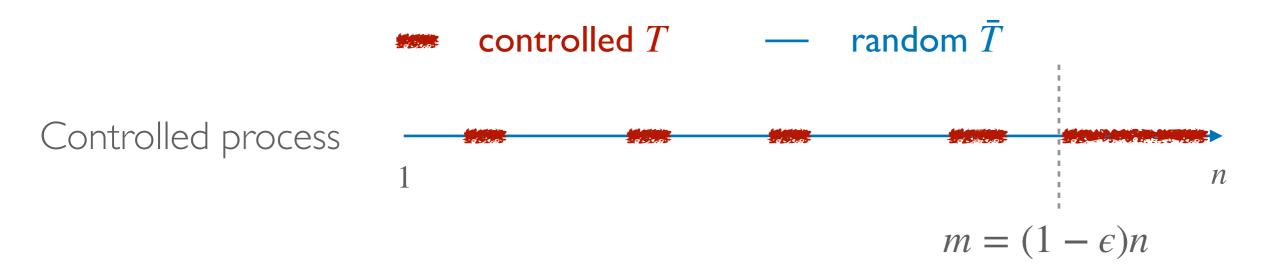
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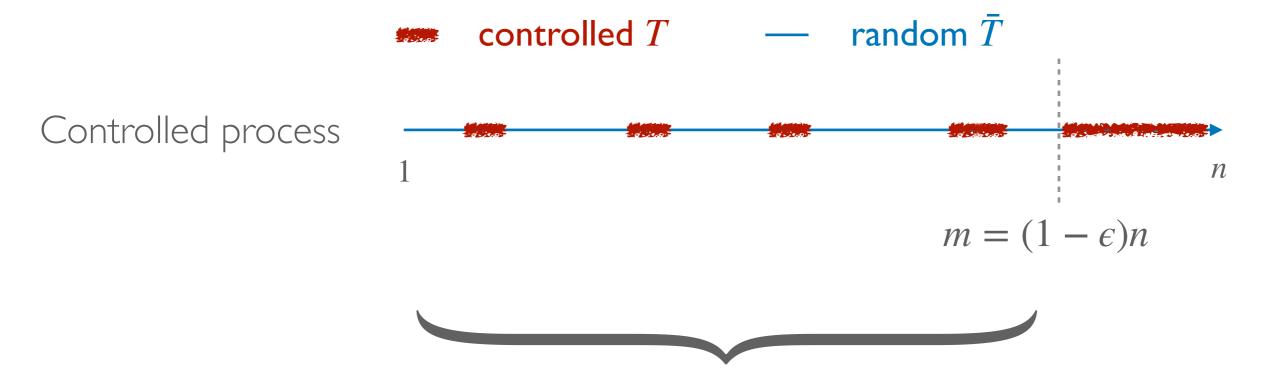
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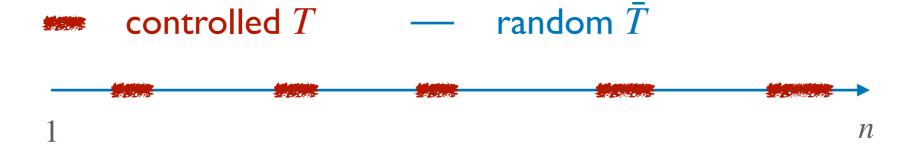


the blue variables = random restriction

### Controlled process



### Controlled process



$$T \subseteq [n]$$
, a random  $(1 - \epsilon)$ -set

(Random coordinate  $t \notin T$ )

$$Pr[Y_{\pi(t)}(t) = \pm 1] = 0.5$$

(Controlled coordinate  $t \in T$ )

$$\Pr[Y_{\pi(t)}(t) = \pm 1] = \frac{1}{2} \pm \frac{\partial_{\pi(t)} f(Y(t-1))}{2\epsilon f(Y(t-1))}.$$

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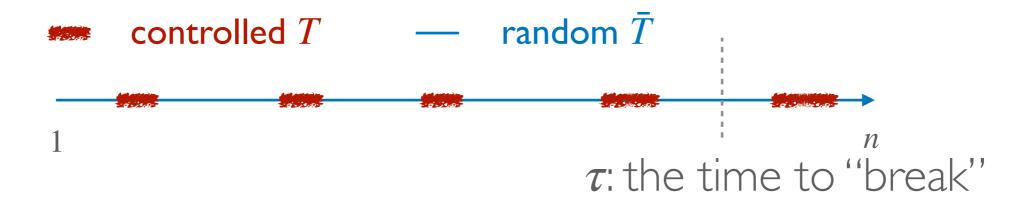
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 $\tau$ : the time to "break"

 $T \subseteq [n]$ , a random  $(1 - \epsilon)$ -set (Random coordinate  $t \notin T$ )  $\Pr[Y_{\pi(t)}(t) = \pm 1] = 0.5$  in particular, the first time t when  $\max_{i} |\partial_{i} f(Y(t-1))| > \epsilon \delta,$  $f(Y(t-1)) < \delta$ 

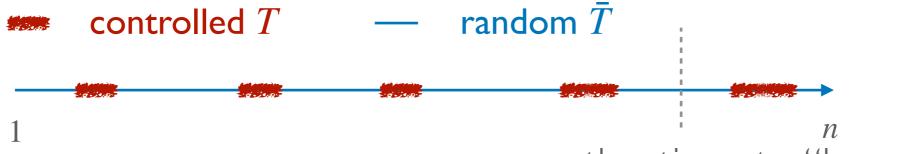
(Controlled coordinate  $t \in T$ )

$$\Pr[Y_{\pi(t)}(t) = \pm 1] = \frac{1}{2} \pm \frac{\partial_{\pi(t)} f(Y(t-1))}{2\epsilon f(Y(t-1))}.$$

(After  $\tau$ )

$$\Pr[Y_{\pi(t)}(t) = \pm 1] = \frac{1}{2} \pm \frac{\partial_{\pi(t)} f(Y(t-1))}{2f(Y(t-1))}.$$

### Controlled process



 $\tau$ : the time to "break"

 $T \subseteq [n]$ , a random  $(1 - \epsilon)$ -set (Random coordinate  $t \notin T$ )  $\Pr[Y_{\pi(t)}(t) = \pm 1] = 0.5$ 

in particular, the first time 
$$t$$
 when 
$$\max_{i} |\partial_{i} f(Y(t-1))| > \epsilon \delta,$$
$$f(Y(t-1)) < \delta$$

(Controlled coordinate  $t \in T$ )

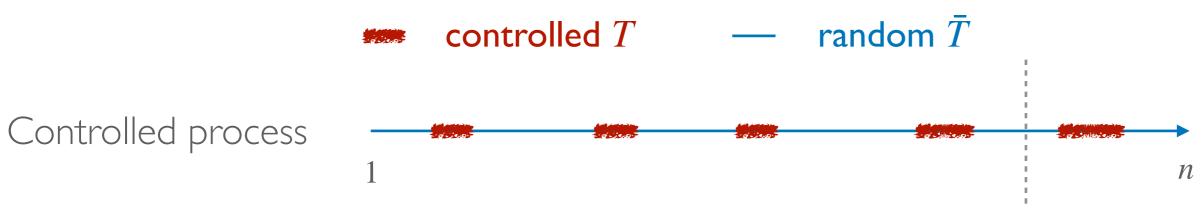
$$\Pr[Y_{\pi(t)}(t) = \pm 1] = \frac{1}{2} \pm \frac{\partial_{\pi(t)} f(Y(t-1))}{2\epsilon f(Y(t-1))}.$$

Goal accomplished, as long as  $\tau > m!$ 

(After  $\tau$ )

$$\Pr[Y_{\pi(t)}(t) = \pm 1] = \frac{1}{2} \pm \frac{\partial_{\pi(t)} f(Y(t-1))}{2f(Y(t-1))}.$$

## Analysis



τ: the time to "break"

in particular, the first time t when  $\max_{i} |\partial_{i} f(Y(t-1))| > \epsilon \delta,$  $f(Y(t-1)) < \delta$ 

$$\tau = \min\{\tau_1, \tau_2\},\$$

$$\tau_1 = \min_t \{f(Y(t-1)) < \delta\},\$$

$$\tau_2 = \min_t \{\max_i |\partial_i f(Y(t-1))| > \epsilon \delta\}$$

## Analysis

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**Lemma I.** w.h.p.  $\tau_1 > (1 - \epsilon)n$ 

# Analysis

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**Lemma I.** w.h.p.  $\tau_1 > (1 - \epsilon)n$ 

**Lemma 2.** w.h.p.  $\tau_2 > (1 - \epsilon)n$ 

# Mean remains large

$$\tau_1 = \min_{t} \{ f(Y(t-1)) < \delta \}$$

**Lemma I.** w.h.p.  $\tau_1 > (1 - \epsilon)n$ 

**Proof.** 
$$\Pr_{Y}[\tau_{1} \leq (1 - \epsilon)n] \leq \Pr_{X}[\tau_{1} \leq (1 - \epsilon)n] \cdot \frac{\delta}{f(0)}$$
 
$$\leq \frac{\delta}{f(0)}.$$



## Partial derivatives remain small

$$\tau_2 = \min_{t} \{ \max_{i} |\partial_i f(Y(t-1))| > \epsilon \delta \}$$

**Lemma 2.** w.h.p.  $\tau_2 > (1 - \epsilon)n$ 

**Proof.** 
$$\Pr_{Y}[\tau_2 \leq (1-\epsilon)n] \leq \Pr_{X}[\tau_2 \leq (1-\epsilon)n] \cdot \frac{1}{f(0)}$$
.

## Partial derivatives remain small

$$\tau_2 = \min_{t} \{ \max_{i} |\partial_i f(Y(t-1))| > \epsilon \delta \}$$

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$$\tau_2 > (1 - \epsilon)n$$
**Proof.**  $\Pr_Y[\tau_2 \le (1 - \epsilon)n] \le \Pr_X[\tau_2 \le (1 - \epsilon)n] \cdot \frac{1}{f(0)}$ .

# Switch to continuous process

### Continuous uniform process

```
Sample \eta(i) \in [0,1], i = 1,2,...,n,
Each Z_i(t) is 0 until t = \eta(i), set Z_i(t) \sim \{-1,1\}.
```

# Switch to continuous process

### Continuous uniform process

Sample 
$$\eta(i) \in [0,1], i = 1,2,...,n,$$
  
Each  $Z_i(t)$  is  $0$  until  $t = \eta(i)$ , set  $Z_i(t) \sim \{-1,1\}$ .

$$au_3 = \min_t \{ \max_i | \partial_i f(Z(t)) | > \epsilon \beta \}$$
  
**Lemma 3.** w.h.p.  $au_3 > (1 - \epsilon)$ 

$$\tau_3 = \min_t \{ \max_i | \partial_i f(Z(t)) | > \epsilon \delta \}$$
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**Lemma 3.** w.h.p.  $\tau_3 > (1 - \epsilon)$ 

$$\Pr\left[\sup_{0\leq s\leq t}\beta(s)\geq\theta\right]\leq\Pr\left[\sup_{0\leq s\leq t}\sum_{i=1}^{n}|\partial_{i}f(Z(s))|^{2+\eta}\geq\theta^{2+\eta}\right]$$

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$$\leq\theta^{-2-\eta}\sum_{i}\mathbf{E}[|\partial_{i}f(Z(t))|^{2+\eta}]$$

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Proof.

$$\Pr\left[\sup_{0\leq s\leq t}\beta(s)\geq\theta\right]\leq\Pr\left[\sup_{0\leq s\leq t}\sum_{i=1}^{n}|\partial_{i}f(Z(s))|^{2+\eta}\geq\theta^{2+\eta}\right]$$
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 $\partial_i f(\mathbf{Z}(t))$  is a martingale.

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$$\leq\theta^{-2-\eta}\sum_{i}(\mathbf{E}[\partial_{i}f(Z(T))^{2}])^{1+\eta/2}$$

A hypercontractivity inequality for random restrictions

# Theorem (Hypercontractive inequality for random restriction).

For any multilinear function  $f: [-1,1]^n \to \mathbb{R}$ , and  $0 \le t \le T \le 1$ . Then, for  $\eta \le T - t$ ,  $\mathbf{E}[|f(Z(t))|^{2+\eta}]^{\frac{1}{2+\eta}} \le \mathbf{E}[f(Z(T))^2]^{\frac{1}{2}},$ 

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For any multilinear function 
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$$\mathbf{E}[|f(Z(t))|^{2+\eta}]^{\frac{1}{2+\eta}} \le \mathbf{E}[f(Z(T))^2]^{\frac{1}{2}},$$

c.f. the standard HC inequality 
$$\mathbf{E}[(\mathbf{T}_{\epsilon(\eta)}f(x)^{2+\eta}])^{\frac{1}{2+\eta}} \leq \mathbf{E}[f(x)^2]^{\frac{1}{2}}$$

$$\tau_3 = \min_{t} \{ \max_{i} | \partial_i f(Z(t)) \} > \epsilon \delta \}$$

**Lemma 3.** w.h.p.  $\tau_3 > (1 - \epsilon)$ 

$$\Pr\left[\sup_{0\leq s\leq t}\beta(s)\geq\theta\right]\leq\Pr\left[\sup_{0\leq s\leq t}\sum_{i=1}^{n}|\partial_{i}f(Z(s))|^{2+\eta}\geq\theta^{2+\eta}\right]$$

$$\leq\theta^{-2-\eta}\sum_{i}\mathbf{E}[|\partial_{i}f(Z(t))|^{2+\eta}]$$

$$\leq\theta^{-2-\eta}\sum_{i}(\mathbf{E}[\partial_{i}f(Z(T))^{2}])^{1+\eta/2}$$

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### Proof.

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$$\leq\theta^{-2-\eta}\sum_{i}\inf_{(T,t)}2\mathbf{E}[\partial_{i}f(Z(T))^{2}]$$

**Proposition.**  $E[\partial_i f(Z(t))^2] \leq Inf_i(f)$ 

$$\tau_3 = \min_{t} \{ \max_{i} | \partial_i f(Z(t)) \} > \epsilon \delta \}$$

**Lemma 3.** w.h.p.  $\tau_3 > (1 - \epsilon)$ 

$$\Pr\left[\sup_{0\leq s\leq t}\beta(s)\geq\theta\right]\leq\Pr\left[\sup_{0\leq s\leq t}\sum_{i=1}^{n}|\partial_{i}f(Z(s))|^{2+\epsilon}\geq\theta^{2+\eta}\right]$$

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$$\leq\theta^{-2-\eta}\frac{\operatorname{MaxInf}(f)^{\eta/2}}{1-T}$$

$$\tau_3 = \min_{t} \{ \max_{i} | \partial_i f(Z(t)) | > \epsilon \delta \}$$

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$$\leq\theta^{-2-\eta}\frac{\operatorname{MaxInf}(f)^{\eta/2}}{1-T}$$

$$\Pr{o\leq s\leq t}$$

$$=\frac{1}{1-T}$$

$$=\frac{1-T}{1-T}$$

$$\tau_3 = \min_{t} \{ \max_{i} | \partial_i f(Z(t)) | > \epsilon \delta \}$$

**Lemma 3.** w.h.p.  $\tau_3 > (1 - \epsilon)$ 

$$\Pr\left[\sup_{0 \le s \le t} \beta(s) \ge \theta\right] \le \Pr\left[\sup_{0 \le s \le t} \sum_{i=1}^{n} |\partial_{i} f(Z(s))|^{2+\epsilon} \ge \theta^{2+\eta}\right]$$

$$\le \theta^{-2-\eta} \sum_{i} \mathbf{E}[|\partial_{i} f(Z(t))|^{2+\eta}]$$

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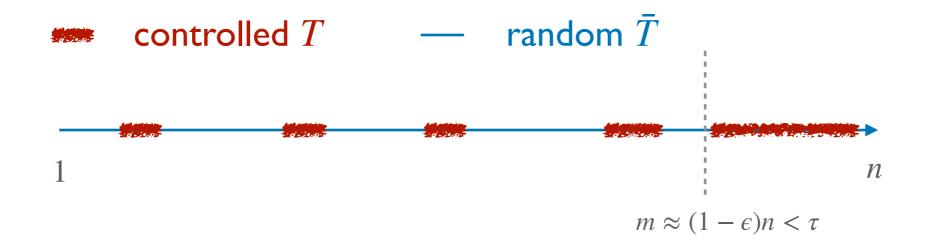
$$\le \theta^{-2-\eta} \sum_{i} \inf_{i} (f)^{\eta/2} \mathbf{E}[\partial_{i} f(Z(T))^{2}]$$

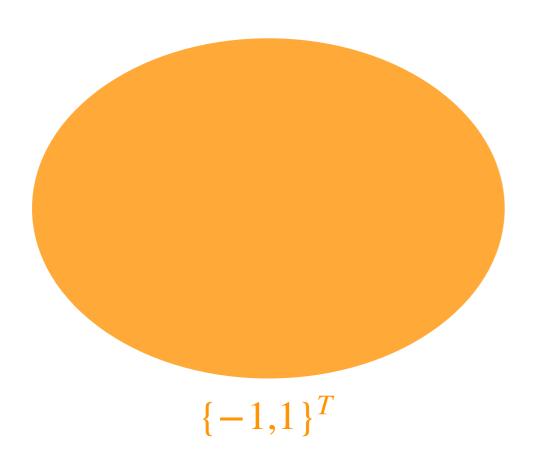
$$\le \theta^{-2-\eta} \frac{\operatorname{MaxInf}(f)^{\eta/2}}{1-T} \le 2\theta^{-3+t/2} \frac{\operatorname{MaxInf}(f)^{(1-t)/4}}{1-t}$$

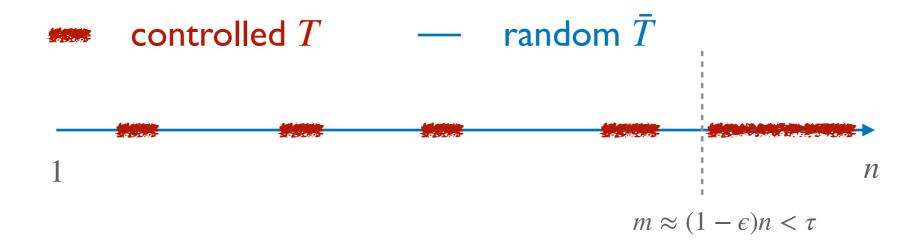
$$\operatorname{set} T = (1+t)/2, \eta = T-t$$

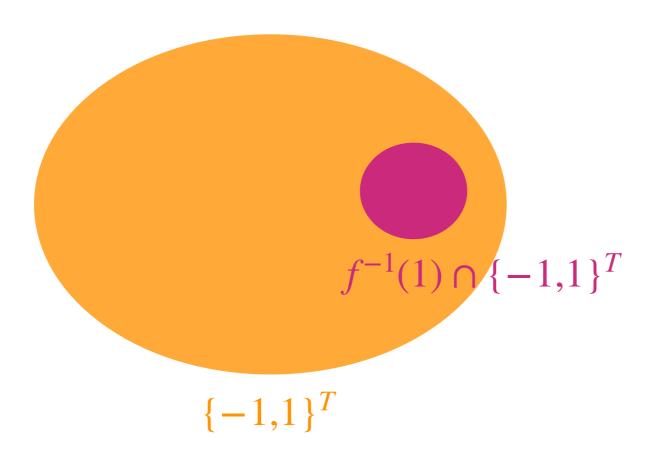
### Bound the variance

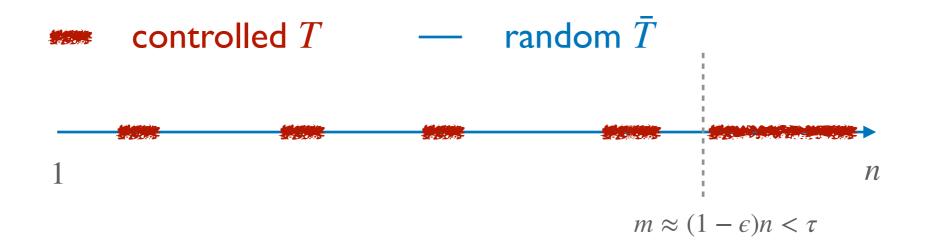


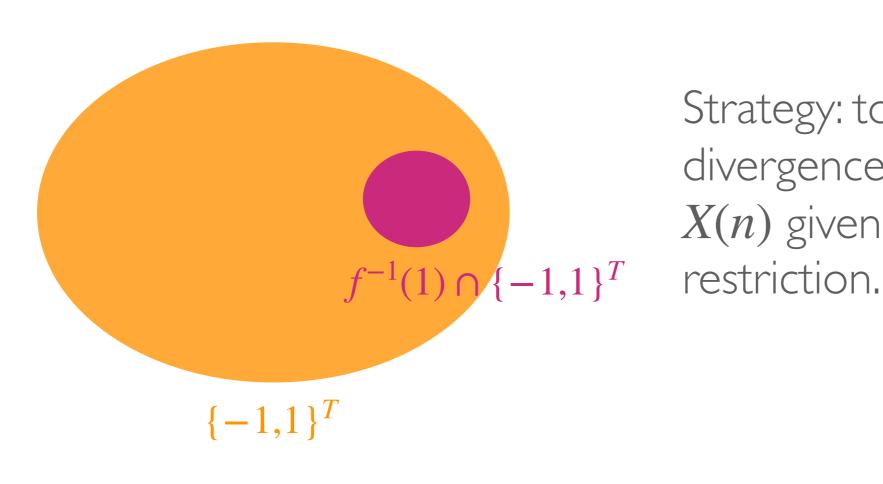




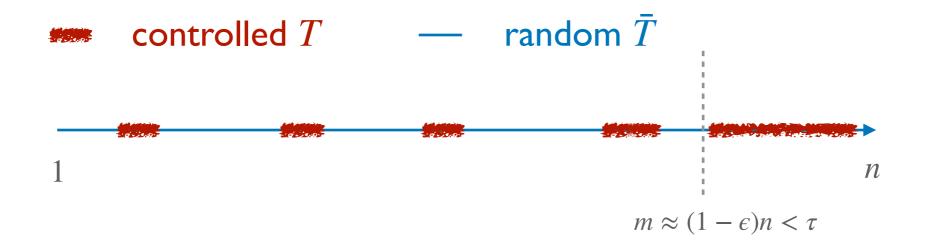


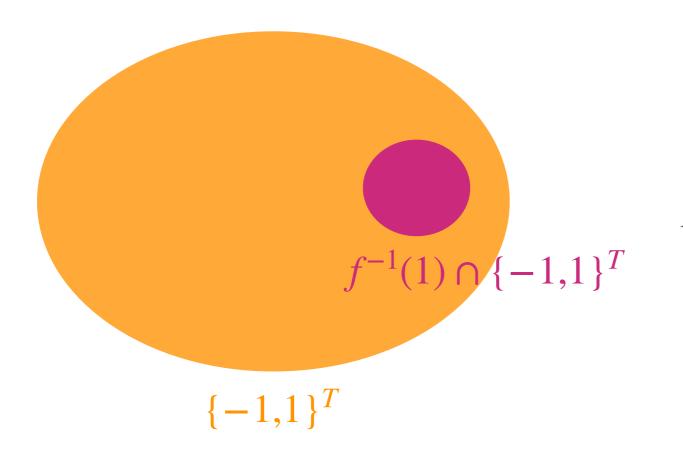






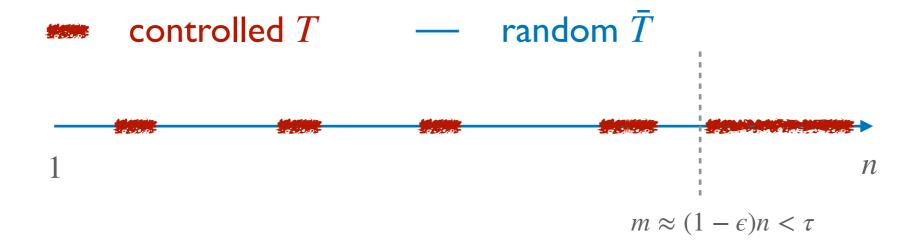
Strategy: to bound the KL-divergence between Y(n) and X(n) given  $\pi$ , T, and the restriction.

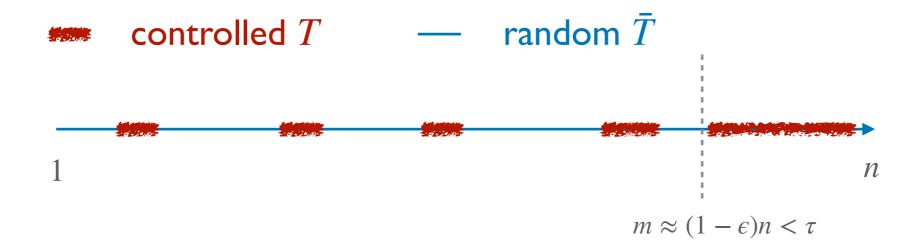


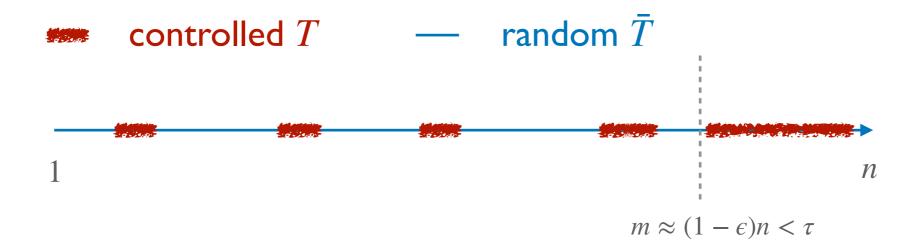


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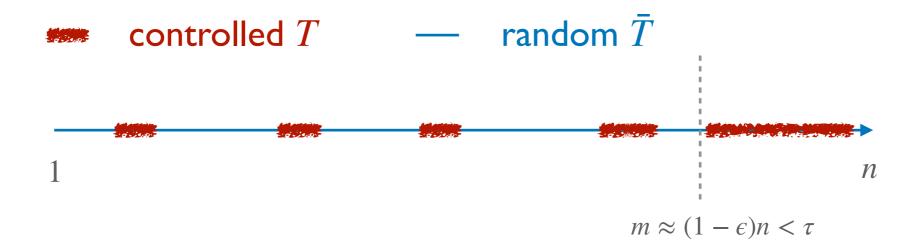
$$|T| - \log |f^{-1}(1) \cap \{-1,1\}^T|$$



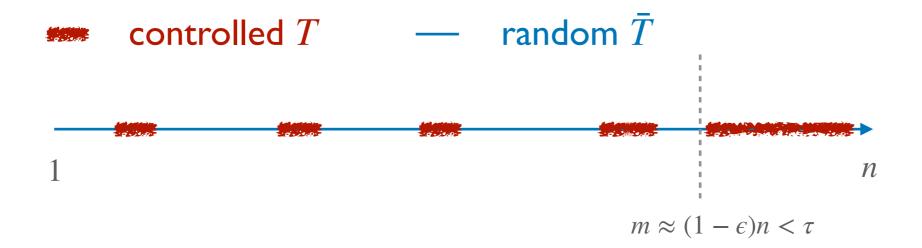




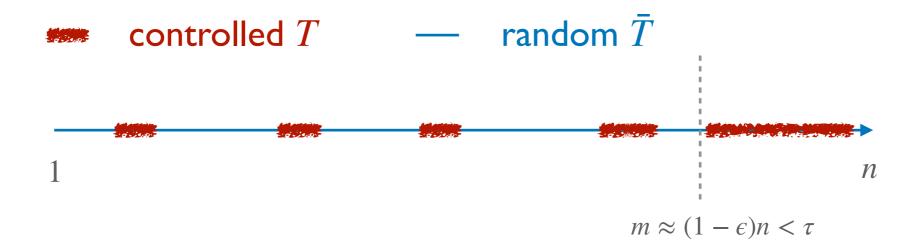
$$\mathbf{E}_{\mathcal{G}_m}[\mathrm{KL}(Y(n)\mid \mathcal{G}_m \| X(n)\mid \mathcal{G}_m)]$$



$$\begin{split} \mathbf{E}_{\mathcal{G}_{m}}[\mathrm{KL}(Y(n) \mid \mathcal{G}_{m} \parallel X(n) \mid \mathcal{G}_{m})] \\ & \leq \sum_{t=1}^{m} \mathbf{E}[\mathbb{I}_{\{t \in T\}} \mathrm{KL}((Y_{\pi(t)}(t) \mid \mathcal{G}_{m}, Y(t-1)) \parallel (X_{\pi(t)}(t) \mid \mathcal{G}_{m}, Y(t-1)))] \\ & + \mathbf{E}[\mathrm{KL}((Y(n) \mid \mathcal{G}_{m}, Y(m)) \parallel (X(n) \mid \mathcal{G}_{m}, Y(m)))] \end{split}$$



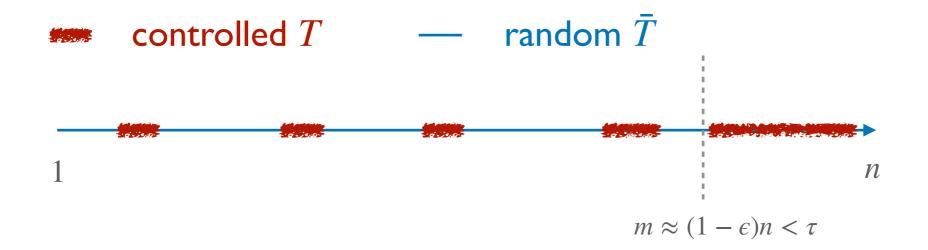
$$\begin{split} \mathbf{E}_{\mathcal{G}_{m}}[\mathrm{KL}(Y(n) \mid \mathcal{G}_{m} \parallel X(n) \mid \mathcal{G}_{m})] \\ & \leq \sum_{t=1}^{m} \mathbf{E}[\mathbb{I}_{\{t \in T\}} \mathrm{KL}((Y_{\pi(t)}(t) \mid \mathcal{G}_{m}, Y(t-1)) \parallel (X_{\pi(t)}(t) \mid \mathcal{G}_{m}, Y(t-1)))] \\ & + \mathbf{E}[\mathrm{KL}((Y(n) \mid \mathcal{G}_{m}, Y(m)) \parallel (X(n) \mid \mathcal{G}_{m}, Y(m)))] \\ & \leq \log \frac{1}{\delta} \end{split}$$



 $\mathcal{G}_m = (\pi, T, x(m)|_{\bar{T}})$ : all the information we know about coordinates come before time m on  $\bar{T}$ .

t=1

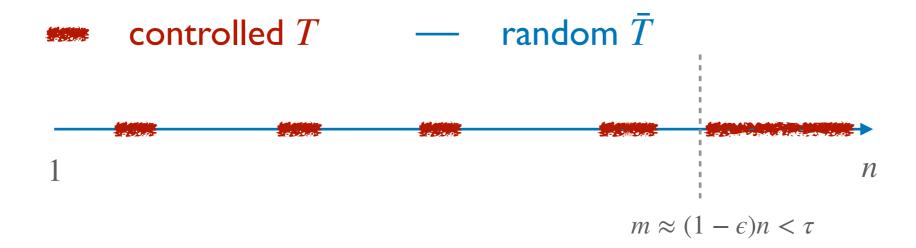
$$\begin{split} \mathbf{E}_{\mathcal{G}_{m}}[\mathrm{KL}(Y(n) \mid \mathcal{G}_{m} \parallel X(n) \mid \mathcal{G}_{m})] \\ \leq \sum_{t=T}^{m} \mathbf{E}[\mathbb{I}_{\{t \in T\}} \mathrm{KL}((Y_{\pi(t)}(t) \mid \mathcal{G}_{m}, Y(t-1)) \parallel (X_{\pi(t)}(t) \mid \mathcal{G}_{m}, Y(t-1)))] + \log \frac{1}{\delta} \end{split}$$



$$\mathbf{E}_{\mathcal{G}_m}[\mathrm{KL}(Y(n)\mid \mathcal{G}_m \parallel X(n)\mid \mathcal{G}_m)]$$

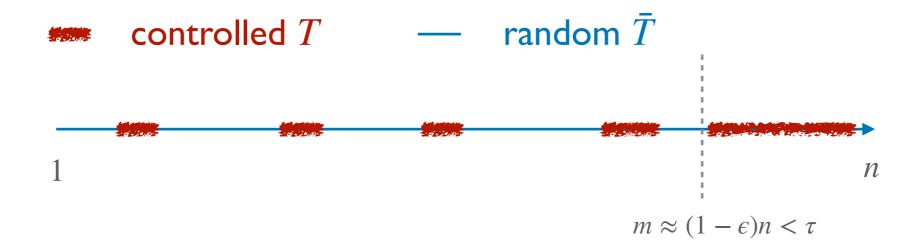
$$\leq \sum_{t=1}^{m} \mathbf{E}[\mathbb{I}_{\{t \in T\}} \mathrm{KL}((Y_{\pi(t)}(t) \mid \mathcal{G}_m, Y(t-1)) \parallel (X_{\pi(t)}(t) \mid \mathcal{G}_m, Y(t-1)))] + \log \frac{1}{\delta}$$

$$\leq \sum_{t=1}^{m} \frac{\epsilon}{n-t+1} \sum_{i:Y:(t-1)=0} \left( \frac{\partial_{i} f(Y(t-1))}{2\epsilon f(Y(t-1))} \right)^{2} + \log \frac{1}{\delta}$$



$$\mathbf{E}_{\mathcal{G}_m}[\mathrm{KL}(Y(n)\mid\mathcal{G}_m\parallel X(n)\mid\mathcal{G}_m)]$$

$$\leq \sum_{t=1}^{m} \frac{\epsilon}{n-t+1} \sum_{i:Y_i(t-1)=0} \left( \frac{\partial_i f(Y(t-1))}{2\epsilon f(Y(t-1))} \right)^2 + \log \frac{1}{\delta}$$



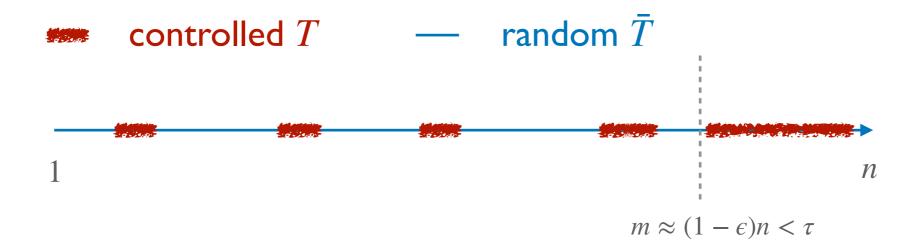
 $\mathcal{G}_m = (\pi, T, x(m)|_{\bar{T}})$ : all the information we know about coordinates come before time m on  $\bar{T}$ .

$$\mathbf{E}_{\mathcal{G}_m}[\mathrm{KL}(Y(n)\mid \mathcal{G}_m \parallel X(n)\mid \mathcal{G}_m)]$$

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### Theorem ([Talagrand 96])

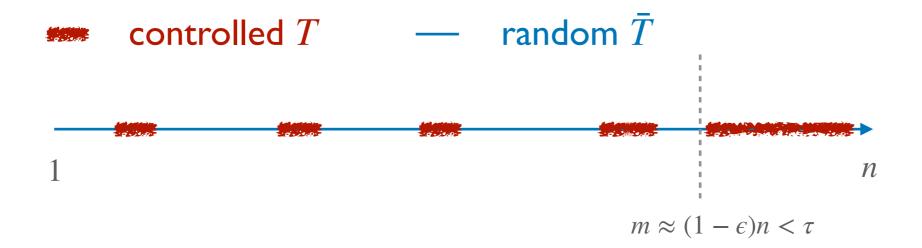
For any 
$$g: \{-1,1\}^n \to [0,1]$$
, we have  $\sum \partial_i g(0)^2 \le Cg(0)^2 \log \frac{e}{g(0)}$ .



$$\mathbf{E}_{\mathcal{G}_m}[\mathrm{KL}(Y(n)\mid \mathcal{G}_m \parallel X(n)\mid \mathcal{G}_m)]$$

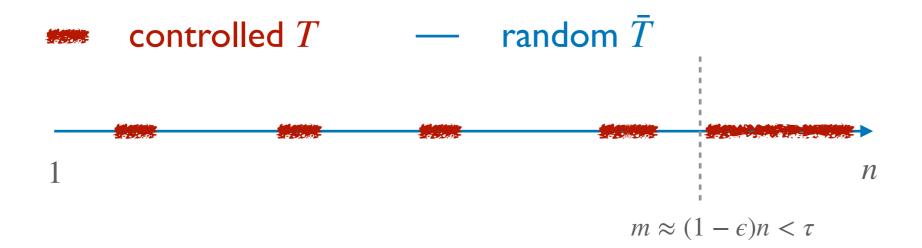
$$\leq \sum_{t=1}^{m} \frac{\epsilon}{n-t+1} \sum_{i:Y_{i}(t-1)=0} \left( \frac{\partial_{i} f(Y(t-1))}{2\epsilon f(Y(t-1))} \right)^{2} + \log \frac{1}{\delta}$$

$$\leq \sum_{t=1}^{m} \frac{C}{4\epsilon(n-t+1)} \log \frac{e}{f(Y(t-1))} + \log \frac{1}{\delta}$$

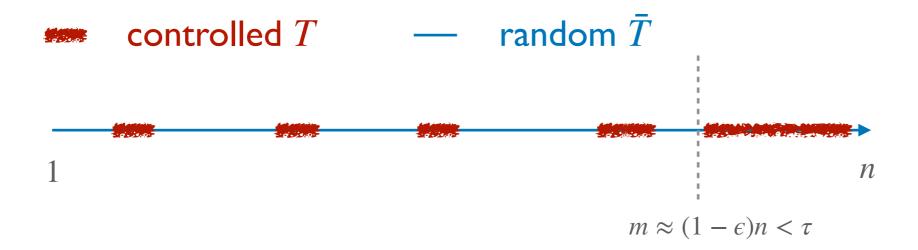


$$\mathbf{E}_{\mathcal{G}_{m}}[\mathrm{KL}(Y(n) \mid \mathcal{G}_{m} \mid X(n) \mid \mathcal{G}_{m})]$$

$$\leq \sum_{t=1}^{m} \frac{C}{4\epsilon(n-t+1)} \log \frac{e}{f(Y(t-1))} + \log \frac{1}{\delta}$$



$$\begin{split} \mathbf{E}_{\mathcal{G}_{m}}[\mathrm{KL}(Y(n) \mid \mathcal{G}_{m} \parallel X(n) \mid \mathcal{G}_{m})] \\ &\leq \sum_{t=1}^{m} \frac{C}{4\epsilon(n-t+1)} \log \frac{e}{f(Y(t-1))} + \log \frac{1}{\delta} \\ &\leq O\left(\frac{1}{\epsilon} \log \frac{1}{\delta}\right) \cdot \sum_{t=1}^{m} \frac{1}{n-t+1} \\ &= O\left(\frac{1}{\epsilon} \log \frac{1}{\delta} \log \frac{n-m}{n}\right). \end{split}$$



 $\mathcal{G}_m = (\pi, T, x(m)|_{\bar{T}})$ : all the information we know about coordinates come before time m on  $\bar{T}$ .

$$\begin{split} \mathbf{E}_{\mathcal{G}_{m}}[\mathrm{KL}(Y(n) \mid \mathcal{G}_{m} || X(n) \mid \mathcal{G}_{m})] \\ &\leq \sum_{t=1}^{m} \frac{C}{4\epsilon(n-t+1)} \log \frac{e}{f(Y(t-1))} + \log \frac{1}{\delta} \\ &\leq O\left(\frac{1}{\epsilon} \log \frac{1}{\delta}\right) \cdot \sum_{t=1}^{m} \frac{1}{n-t+1} \\ &= O\left(\frac{1}{\epsilon} \log \frac{1}{\delta} \log \frac{n-m}{n}\right). \end{split}$$

Thank you!