Building Bayesian Influence Ontologies
Literature Review

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1   Introduction

The need to detect and recover influence structures arises in a number of applications, including plagiarism detection [Merlo 2007], comparison between genomes [Koonin 1999], and even automatic detection of similarity between pieces of music [Slaney *et al.* 2008].

One of the simplest statistical measures for quantifying influence between random variables X_i and $X_j \in \mathcal{X}$ is the correlation matrix $Corr(i,j)$. One such measure is the Pearson correlation coefficient, developed by Pearson [1895], whose form is

$$Corr(i,j) = \frac{cov(X_i, X_j)}{\sigma_{X_i} \sigma_{X_j}}$$

where $cov(X_i, X_j)$ denotes the covariance between X_i and X_j, and where σ_{X_k} denotes the standard deviation of X_k. However, this measure is only equipped to measure influence in the case of a linear relationship, and even more complicated forms such as the Spearman rank correlation coefficient, developed by Spearman [1904], can only measure influence in the case of monotonically related functions. Worse still, while these measures are able to detect correlation, they are unable to recover the underlying influence structure [Damghani *et al.* 2012]. We therefore require a method that can recover the "ground truth" structure which produced the data.

The remainder of this chapter will focus on Bayesian methods in the recovering of influence structures. Section 2 provides a brief introduction to Bayesian networks and discusses some of their properties. In section 3, we discuss methods by which Bayesian networks are constructed given a set of observations. We briefly discuss constraint-based methods in section 3.1 before discussing score-based methods in greater depth in section 3.2. In section 4, we discuss parameter estimation techniques in the case of partially observed data, leading up to a description of the Expectation Maximisation algorithm. Finally, methods for the construction of networks of independently learned models are discussed in section 5, with the Kullback–Leibler divergence metric for measuring similarity between distributions discussed in section 5.1.
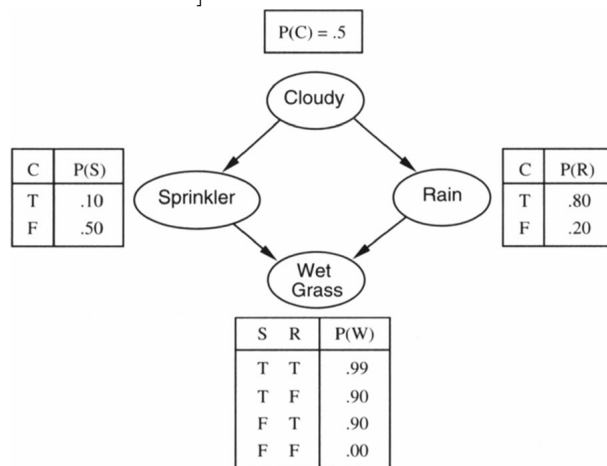
2   Bayesian Networks

When considering a joint probability distribution across n random variables, classical probability states that the number of parameters needed to represent the distribution grows exponentially in n [Koller and Friedman 2009]. Even in the simple case of binary variables, we would still need

$2^n - 1$ parameters to describe the distribution. This is clearly unfeasible for practical applications, in which the number of random variables can grow very large.

Bayesian networks, originally developed by Pearl [1988], present a way of reducing the number of parameters needed to represent a joint distribution. A Bayesian network is a directed acyclic graph (DAG) whose nodes represent random variables and whose edges represent influence of one variable on another. This structure can also be thought of as a representation of the conditional independencies between the random variables [Koller and Friedman 2009]. Indeed, it is through the exploitation of these independency assumptions that a Bayesian network can more compactly represent a joint distribution.

Figure 1: A famous example of a Bayesian network, showing how a complete representation of any random variable X requires considering only those variables who are parents of X in the graphical representation [Norvig and Russell 1994].

An important notion in Bayesian networks is that of d-separation, first presented by Pearl [1986], which is used to find the set $\mathcal{I}(\mathcal{G})$ of conditional independencies in the graph $\mathcal{G}$. $\mathcal{I}(\mathcal{G})$ is used as the basis for an equivalence relation, I-equivalence, for which any two I-equivalent graphs represent the same independency assumptions [Verma and Pearl 1991]. An important developement by Pearl [1986] is that any I-equivalence class can be represented as a partially directed acyclic graph (PDAG) in which undirected edges represent edges that can be oriented any way and still result in a graph belonging to the same class.

3 Structure Learning

The manual construction of networks is generally unfeasible for a large number of variables [Koller and Friedman 2009]. Fortunately, strategies exist to learn model structures from data $\mathcal{D}$.

3.1 Constraint-Based Structure Learning

One approach to the construction of model structures is the constraint-based approach, in which dependencies between variables are first queried and then, based on these dependencies, a PDAG is constructed [Koller and Friedman 2009]. This strategy can be traced back to Verma and Pearl [1991].

However, this approach is generally not preferred, as failure in the individual independence queries can lead to the construction of a network which poorly matches the data [Koller and Friedman 2009].

3.2 Score-Based Structure Learning

A more popular approach to the problem is score-based structure learning, in which entire networks are constructed and then evaluated and modified based on some scoring metric [Koller and Friedman 2009]. Two areas of interest in this approach are the choice of scoring function and the method of structure search.

3.2.1 Scoring Function

One possible scoring function would be the maximum likelihood function (most often in its logarithm form), finding graph $\mathcal{G}$ that maximises

$$score_L(\mathcal{G} : \mathcal{D}) = l(\hat{\theta}_{\mathcal{G}} : \mathcal{D})$$

which decomposes to

$$score_L(\mathcal{G} : \mathcal{D}) = M \sum_{i=1}^{n} [\mathbb{I}_{\hat{P}}(X_i; Pa_{X_i}^{\mathcal{G}}) - H_{\hat{P}}(X_i)]$$

for number of variables n, number of samples M, mutual information $\mathbb{I}_{\hat{P}}$ and entropy $H_{\hat{P}}$ [Koller and Friedman 2009]. However, this score always prefers a more connected network, and is thus prone to overfitting.

Other scores designed to balance fit to data with network complexity are the Akaike Information Criterion (AIC), proposed by Akaike [1998]

$$score_{AIC}(\mathcal{G} : \mathcal{D}) = l(\hat{\theta}_{\mathcal{G}} : \mathcal{D}) - Dim(\mathcal{G})$$

and the Bayesian Information Criterion (BIC), proposed by Schwarz [1978]

$$score_{BIC}(\mathcal{G} : \mathcal{D}) = l(\hat{\theta}_{\mathcal{G}} : \mathcal{D}) - \frac{logM}{2}Dim(\mathcal{G})$$

where $Dim(\mathcal{G})$ denotes the dimension of $\mathcal{G}$. In particular, Schwarz [1978] shows that the BIC is an assymptotic approximation of the Bayesian score under the assumptions of independent, identically distributed observations with a density function of the form

$$f(x, \theta) = exp(\theta \cdot y(x) - b(\theta))$$

where y is the sufficient statistic, and where it is also assumed that the penalty for guessing an incorrect model is fixed. The $\frac{logM}{2}$ term in the BIC ensures that, as M grows, more consideration is placed in models of greater complexity [Koller and Friedman 2009].

3.2.2 Structure Search

The problem of structure search is to find the graph $\mathcal{G}$ that maximises the chosen scoring function for the given data $\mathcal{D}$. In general, this problem is NP-hard for a graph whose variables have at most $d \geq 2$ parents [Chickering 1996]. Fortunately, there exist heuristic algorithms which can find

assist in this regard. Some of the earliest of these algorithms include the K2 algorithm of Cooper and Herskovits [1992], which relied on a predetermined ordering of variables, and the local search algorithms proposed by Heckerman *et al.* [1995].

These algorithms define a search space of graphs, where each graph can be transformed into another by a set of operators [Koller and Friedman 2009]. These operators commonly include edge addition, edge deletion and edge reversal.

A search procedure is then required to traverse the search space and select an optimal graph. A common choice is the greedy hill-climbing algorithm, which applies only the operations which maximise the score [Koller and Friedman 2009]. This technique is prone to local maxima and the plateaus in score caused by I-equivalent graphs. Methods which work around this problem include the tabu search, proposed by Glover [1986], which keeps track of recent operations and does not allow them to be reversed until a certain number of iterations has passed, and the method of random restarts, which restarts the search with random initial conditions [Koller and Friedman 2009].

4   Partially Observed Data

The methods discussed in section 3 require data that are fully observed, in that every datum assigns a value to each variable [Koller and Friedman 2009]. In practice, this is not always possible. Sometimes, these latent or hidden variables are simply impractical to measure. In other cases, these variables represent factors that cannot be measured quantitatively.

Unfortunately, the introduction of hidden variables complicates the structure learning process by introducing a number of terms to the maximum likelihood estimate, with the number of terms growing exponentially as more variables are hidden [Koller and Friedman 2009]. Furthermore, the introduction of latent variables results in a likelihood function which is not locally decomposable, and this renders the techniques discussed in section 3 useless in general. Thus different methods must be applied for both the learning of parameters and the learning of structure. For the sake of relevance, we will discuss only the problem of parameter estimation, as the problem of structure learning is beyond the scope of this paper.

Here we discuss two popular methods for parameter estimation. The first is gradient ascent, in which parameters are chosen by iteratively moving in the direction of the gradient of the likelihood function [Koller and Friedman 2009]. Early applications of gradient methods to the likelihood function include the work of Thiesson [1995] and Binder *et al.* [1997]. In these algorithms, the gradient of the likelihood function is given as

$$\frac{\partial l(\boldsymbol{\theta} : \mathcal{D})}{\partial P(x|\boldsymbol{u})} = \frac{1}{P(x|\boldsymbol{u})} \sum_{m=1}^{M} P(x, \boldsymbol{u}|\boldsymbol{o}[m], \boldsymbol{\theta})$$

where $\boldsymbol{\theta}$ are the parameters, $\mathcal{D} = \{\boldsymbol{o}[1], ..., \boldsymbol{o}[M]\}$ is the set of partially observed data, $x \in Val(X)$ for a variable $X \in \mathcal{X}$ and $\boldsymbol{u} \in Val(Pa_X)$ [Koller and Friedman 2009].

The second method to be discussed is Expectation Maximisation (EM), introduced by Dempster *et al.* [1977] as a generalisation of several earlier methods such as those presented by Baum *et al.* [1970] and Orchard and Woodbury [1972]. EM consists of two steps: Expectation (E-step) and Maximisation (M-step), which are repeated until convergence, starting from some initial $\boldsymbol{\theta}^0$ [Koller and Friedman 2009].

In the E-step, we compute the expected sufficient statistic for each $x \in X$ and $u \in U$

$$\bar{M}_{\boldsymbol{\theta}^t}[x, \boldsymbol{u}] = \sum_{m=1}^{M} P(x, \boldsymbol{u}|\boldsymbol{o}[m], \boldsymbol{\theta}^t)$$

where $\boldsymbol{\theta}^t$ denotes the value of $\boldsymbol{\theta}$ at iteration t.

In the M-step, the $\bar{M}$ is treated as the observed sufficient statistic and is used to calculate $\boldsymbol{\theta}^{t+1}$ using maximum likelihood estimation.

An important property of EM is that it is guaranteed to improve $l(\boldsymbol{\theta^t} : \mathcal{D})$ monotonically as t increases [Koller and Friedman 2009].

5 Influence between Models

We now turn to the specific problem of tracking influence between models. A number of different approaches have been developed to suit different models. For example, Pan *et al.* [2005] use Jeffrey's rule [Pearl 1990] to propagate beliefs on variables from one Bayesian network to another and thus map concepts between ontologies.

However, the methods that are most relevant to this paper are those presented by Ajoodha and Rosman [2017]. In this work, the authors track influence between a set of naïve Bayes models (NBMs). They do so by first partitioning the observable data into k sets and learning each of the k NBMs independently through the use of EM. They then compute the score of the overarching network (see figure 5) using the BIC, relearn model parameters for the new independence assertions using EM, use the search operators discussed in section 3.2.2 to try and improve the network's score, and then repeat this process until no more improvement can be made to the score. This process is thus a greedy hill-climbing, although in their final implementation, the authors also made use of tabu lists and random restarts to avoid the pitfalls discussed in section 3.2.2.

Figure 2: A simple example of an influence network between a set of independently learned NBMs, where each latent variable L_i is learned from a set of observations $\{O_1^i, ..., O_K^i\}$ and is related to other latent variables via the high-level influence network [Ajoodha and Rosman 2017]
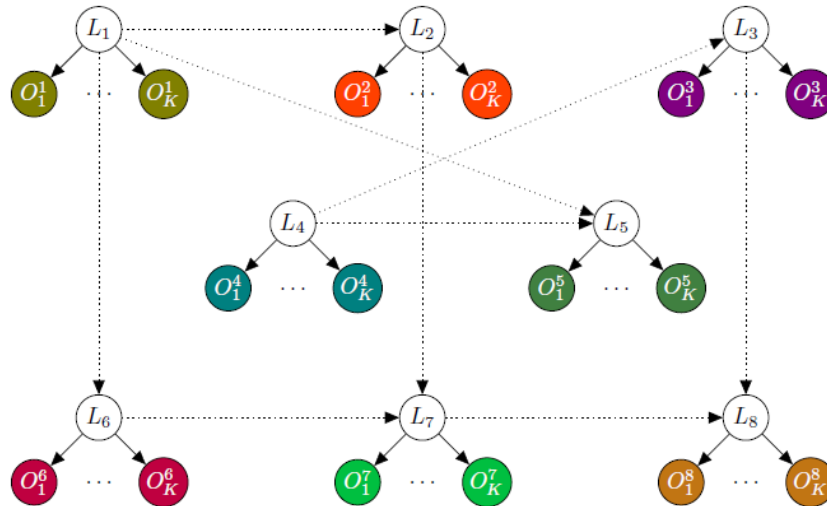
Interesting extensions to this work include the work of Ajoodha and Rosman [2018], who extend the above process to temporal models. In essence, they replace NBMs with hidden Markov models (HMMs) in order to model stochastic processes. They then use very similar techniques to construct a delayed dynamic influence network that models the high-level influence between the HMMs.

5.1 Model Evaluation

When constructing an influence network, it is important to measure how well one's learned model recovers the "ground truth" structure. A useful measure for evaluating this is the Kullback–Leibler divergence, also known as relative entropy, developed by Kullback and Leibler [1951], which sees use in both Ajoodha and Rosman [2017] and Ajoodha and Rosman [2018] discussed above. The relative entropy between two models, P_1 and P_2, over a shared variable space $\mathcal{X}$ is given by

$$\mathbb{D}(P_1||P_2) = \mathbb{E}_{P_1}\left(log\frac{P_1(\mathcal{X})}{P_2(\mathcal{X})}\right) = \sum_{x \in \mathcal{X}} P_1(x)log\frac{P_1(x)}{P_2(x)}$$

in the discrete case [Koller and Friedman 2009]. There are two important observations to make about the measure. Firstly, $\mathbb{D}(P_1||P_2) \geq 0$ in all cases and, secondly, because probabilities must sum to 1, $\mathbb{D}(P_1||P_2) = 0$ if and only if $P_1(x) = P_2(x) \ \forall x \in \mathcal{X}$ [Koller and Friedman 2009]. Thus the closer the relative entropy between a reconstructed model and the underlying distribution gets to 0, the more confidently we can say that our model has recovered the "ground truth" structure.

6 Conclusion

In this chapter we have discussed Bayesian networks and their advantages for compactly representing joint probability distributions by making use of the underlying independence assumptions between random variables. We discussed methods by which network structures could be learnt from fully-observed data, focusing particularly on score-based structure learning. We then presented the problem of learning Bayesian networks from partially observed data, and discussed methods of parameter estimation in this case. Finally, we discussed methods by which the influence between models can be recovered, and the Kullback–Leibler divergence metric which can be used to evaluate how well a learned model has recovered the underlying structure of the data.

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