

# From Statistics to Data Mining

Master 1
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#### Introduction

#### Definition

- o the principal components of a collection of points in a real p-space that are a sequence of p direction vectors, where the  $i^{\text{th}}$  vector is the direction of a line that best fits the data while being orthogonal to the first i-1 vectors
- a best-fitting line is defined as one that minimizes the average squared distance from the points to the line
- these directions constitute an orthonormal basis in which different individual dimensions of the data are linearly uncorrelated
- → PCA: process of computing the principal components and using them to perform a change of basis on the data

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- Introduction
- > Uses
- principal component analysis (PCA), also known as the Karhunen-Love transform, is widely used for:
- ☐ dimensionality reduction:
  - it projects data points living in a d-dimensional space onto a M-dimensional subspace, where M < d
- o if M = 2, PCA allows **data visualization** while preserving the variance of the original data
- ☐ feature extraction: it generates new uncorrelated (i.e., without redundancies) meaningful features

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Introduction

#### > Example

- main characteristics of the planets of the solar system:
- distance to the sun (in UA)
- diameter (in km)
- density (in g / cm<sup>3</sup>)

	Distance	Diamètre	Densité
Mercure	0,387	4 878	5,42
Vénus	0,723	12 104	5,25
Terre	1,000	12 756	5,52
Mars	1,524	6 787	3,94
Jupiter	5,203	142 800	1,31
Saturne	9,539	120 660	0,69
Uranus	19,180	51 118	1,29
Neptune	30,060	49 528	1,64
Pluton	39,530	2 300	2,03

 since a 3D-plot is not always very readable, can we find a 2D-plot of the data such that close points in that new space mean similar planets in the original 3D-space?

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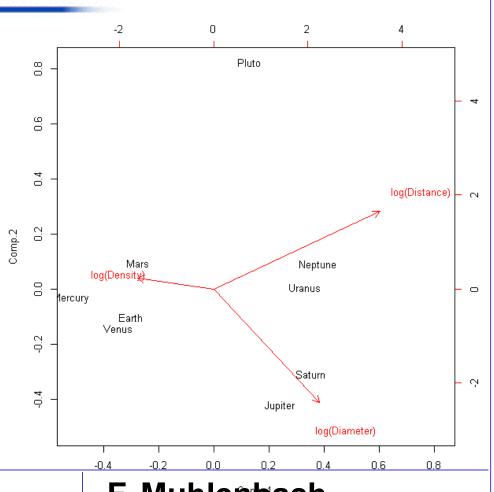




Introduction

#### > Example

 possible solution with PCA: reduction of dimensionality



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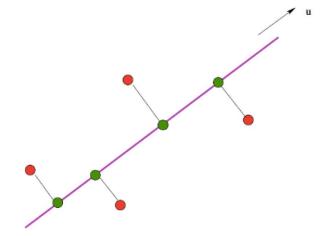




Goal of PCA

#### > Example

- o the goal of PCA is to linearly project the data  $x_i \in \mathbb{R}^d$  onto a space having dimensionality M < d such that close points in that new M-space mean similar examples in the original d-space
- $\circ$  here, d=2 and M=1
- we have to define the direction of this space using a 2-dimensional vector u



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- Maximization of the variance of the projected data
- o let us suppose that the training data are zero mean (that is,  $\forall i$ ,  $x_i$  is changed into  $x_i \leftarrow x_i \bar{x}$ )
- $\circ$  PCA seeks a new space of size M < d by applying a linear transformation  $\mathbf{U}^{\mathrm{T}}$  on the original data
- the new representation of a training data  $x_i$ , denoted by  $t_i$ , is computed as follows:  $t_i = \mathbf{U}^T x_i$
- o where  $\mathbf{U} = (\mathbf{u}_1, \dots, \mathbf{u}_M)$  is a  $(d \times M)$ -matrix of new bases and  $\mathbf{u}_i \in \mathbb{R}^d$
- o we impose that  $\mathbf{U}^{\mathrm{T}}\mathbf{U} = I$ , that is  $\mathbf{U}$  is orthogonal, meaning:
- $\square$  every new feature  $\mathbf{u}_i$  is linearly independent from the others,
- $\square \forall j, \mathbf{u}_i^{\mathrm{T}} \mathbf{u}_i = 1$
- $\circ$  note that each  $t_i$  is a linear combination of the original features





- Maximization of the variance of the projected data
- o if  $x_j \in \mathbb{R}^d$ , then the PCA can generate a maximum of d new components, i.e.,  $\mathbf{U} = (\mathbf{u}_1, \cdots, \mathbf{u}_d)$  is a  $(d \times d)$ -matrix of new bases
- o if the linear transformation  ${\bf U}$  is composed with d new bases, then it is possible to perfectly rebuild the data of the initial space (i.e., it is a bijection) but if  ${\bf U}$  is composed only with M new bases, with M < d, then some information is lost in the projection in the M-dimension space and the reconstruction of the data in the initial space is not perfect
- therefore the objective of the PCA is to minimize this reconstruction error in order to keep as much information as possible from the original space despite the dim. reduction

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- Maximization of the variance of the projected data
- o let  $\widehat{x}_i = \mathbf{U}t_i$  be the reconstruction of the original vector  $x_i$  using the transformation  $\mathbf{U}$
- o the objective of PCA is to optimize **U** s.t. the mean square error  $J(\mathbf{U})$  between  $x_i$  and  $\hat{x}_i$  is as small as possible:

$$\min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} (x_i - \hat{x}_i)^2$$

$$\Leftrightarrow \min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} (x_i - \mathbf{U} \mathbf{U}^\mathsf{T} x_i) (x_i - \mathbf{U} \mathbf{U}^\mathsf{T} x_i)$$

$$\Leftrightarrow \min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} (x_i^{\mathsf{T}} x_i - 2x_i^{\mathsf{T}} \mathbf{U} \mathbf{U}^{\mathsf{T}} x_i + x_i^{\mathsf{T}} \mathbf{U} \mathbf{U}^{\mathsf{T}} \mathbf{U} \mathbf{U}^{\mathsf{T}} x_i)$$

$$\Leftrightarrow \min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} (x_i^{\mathsf{T}} x_i - x_i^{\mathsf{T}} \mathbf{U} \mathbf{U}^{\mathsf{T}} x_i)$$
 because  $\mathbf{U}^{\mathsf{T}} \mathbf{U} = 1$ 

$$\Leftrightarrow \min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} x_{i}^{\mathsf{T}} x_{i} - \frac{1}{n} \sum_{i} x_{i}^{\mathsf{T}} \mathbf{U} \mathbf{U}^{\mathsf{T}} x_{i}$$

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- Maximization of the variance of the projected data
- $\triangleright$  Optimization of  $U \rightarrow$  minimization of J(U) (conclusion):
- $\circ \min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} (x_i \hat{x}_i)^2 = \min_{\mathbf{U}} \frac{1}{n} \sum_{i} x_i^{\mathsf{T}} x_i \frac{1}{n} \sum_{i} x_i^{\mathsf{T}} \mathbf{U} \mathbf{U}^{\mathsf{T}} x_i$
- $\Leftrightarrow \min_{\mathbf{U}} J(\mathbf{U}) = \min_{\mathbf{U}} \operatorname{Tr}(\Sigma) \operatorname{Tr}(\mathbf{U}^{\mathrm{T}} \Sigma \mathbf{U})$
- where Σ is the covariance matrix of the original data and U<sup>T</sup>ΣU
  is covariance in the new space
- since Tr(Σ) does not depend on U, minimizing J(U) boils down to maximizing U<sup>T</sup>ΣU, that is,

$$\max_{\mathbf{U}} \mathbf{U}^{\mathrm{T}} \mathbf{\Sigma} \mathbf{U}$$

s. t. 
$$\forall j = 1, \dots, M, \mathbf{u}_i^T \mathbf{u}_i = 1$$





- Maximization of the variance of the projected data
- ightharpoonup Minimization of  $J(\mathbf{U}) \rightarrow$  optimization problem:  $\max_{\mathbf{U}} \mathbf{U}^{\mathrm{T}} \Sigma \mathbf{U}$

s. t. 
$$\forall j = 1, \dots, M, \mathbf{u}_i^T \mathbf{u}_i = 1$$

- o introducing Lagrange multipliers (denoted by the feature vector  $\lambda = (\lambda_1, \dots, \lambda_M)$ ), we get the unconstrained maximization problem:  $\max_{\mathbf{U}} \mathbf{U}^T \Sigma \mathbf{U} + \lambda (1 \mathbf{U}^T \mathbf{U})$
- let us consider the first component u₁ of the new space
- o find u₁ requires to solve:

$$\frac{\partial \mathbf{U}^{\mathrm{T}} \mathbf{\Sigma} \mathbf{U} + \lambda (1 - \mathbf{U}^{\mathrm{T}} \mathbf{U})}{\partial \mathbf{u}_{1}} = 0$$

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- Maximization of the variance of the projected data
- > Derivatives of matrices and vectors
- o let  $v \in \mathbb{R}^d$  a vector and M a  $d \times d$  matrix:

$$\frac{\partial v^{\mathrm{T}} M v}{\partial v} = (M + M^{\mathrm{T}}) v$$

o if M is symmetric, then  $M=M^T$  and

$$\frac{\partial v^{\mathrm{T}} M v}{\partial v} = 2 M v$$

applying the previous on

$$\frac{\partial \mathbf{U}^{\mathrm{T}} \mathbf{\Sigma} \mathbf{U} + \lambda (1 - \mathbf{U}^{\mathrm{T}} \mathbf{U})}{\partial \mathbf{u}_{1}} = 0$$

we get  $\Sigma \mathbf{u}_1 = \lambda_1 \mathbf{u}_1$ 

 $\circ$  which says that  $\mathbf{u}_1$  must be an eigenvector of  $\Sigma$ 

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- Maximization of the variance of the projected data
- o for maximizing  $\mathbf{U}^{T}\Sigma\mathbf{U}$ , we have the constraint  $\Sigma\mathbf{u}_{1}=\lambda_{1}\mathbf{u}_{1}$
- o if we left-multiply by  $\mathbf{u}_1^T$  and make use of  $\mathbf{u}_1^T \mathbf{u}_1 = 1$ , we see that the variance is given by

$$\mathbf{u}_1^{\mathrm{T}} \Sigma \mathbf{u}_1 = \lambda_1$$

- o and the variance will be maximum when we set  $\mathbf{u}_1$  equal to the eigenvector having the largest eigenvalue  $\lambda_1$
- o this eigenvector is known as the first principal component
- o **conclusion**: constraining  $\mathbf{U}^{\mathrm{T}}\mathbf{U} = I$  means that we restrict the optimization problem to find an orthogonal matrix  $\mathbf{U}$
- o therefore, we get the same result  $\mathbf{U}^T \Sigma \mathbf{U} = \lambda$  as that of which would have been obtained with a diagonalizable PD matrix

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- Properties of the components
- $\circ$  the eigenvalues of  $\Sigma$  are always positive because is  $\Sigma$  PSD
- the number of components is equal to the number of non zero eigenvalues
- o the total variance of the original data is  $V = \text{Tr}(\Sigma)$  because the diagonal elements of Σ contain the variances
- o we deduce that:

$$V = \operatorname{Tr}(\Sigma) = \operatorname{Tr}(\mathbf{U}\lambda\mathbf{U}^{-1}) = \operatorname{Tr}(\mathbf{U}^{-1}\mathbf{U}\lambda) = \operatorname{Tr}(\lambda) = \lambda_1 + \lambda_2 \cdots + \lambda_d$$

when we project the data on a two-dimensional plane corresponding to the eigenvectors  $\mathbf{u}_1$  and  $\mathbf{u}_2$  associated with the two largest eigenvalues  $\lambda_1$ ,  $\lambda_2$ , we get a new covariance matrix  $\mathbf{U}\Sigma\mathbf{U}^T$  whose total variance  $\hat{V} = \text{Tr}(\mathbf{U}\Sigma\mathbf{U}^T) = \lambda_1 + \lambda_2$ 

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- Properties of the components
- o projection of data onto a 2-dimensional plane space  $\rightarrow$  covariance matrix  $\mathbf{U}\Sigma\mathbf{U}^{\mathrm{T}}$  with variance  $\hat{V} = \mathrm{Tr}(\mathbf{U}\Sigma\mathbf{U}^{\mathrm{T}}) = \lambda_1 + \lambda_2$
- o therefore, we can compute the ratio of variance "explained" by the projected data:  $\frac{\lambda_1 + \lambda_2}{\lambda_1 + \lambda_2 + \cdots + \lambda_{d-1} + \lambda_d}$
- the higher the ratio, the better the projection
- interpretation of the results of the PCA → it depends on:
- ☐ quality of the representation on the main planes
- ☐ choice of size (number of axes to be used)
- ☐ "internal" interpretation (correlations between variables, place and importance of individuals, size effect, etc.)
- ☐ "external" interpretation (variables and additional individuals)





- Algorithmic complexity of PCA
- $\circ$  PCA involves evaluating the mean  $\bar{x}$  and the covariance matrix  $\Sigma$  of the data set and then finding the M eigenvectors of  $\Sigma$  corresponding to the M largest eigenvalues:
- $\Box$  the computational cost of computing the full eigenvector decomposition for a matrix of size  $d \times d$  is  $\mathcal{O}(d^3)$
- $\Box$  however, if we are only interested in the the projection onto the first M principal components, efficient techniques exist, such as the *power method* that scale like  $\mathcal{O}(Md^2)$ , or alternatively we can make use of the EM algorithm