



# MAXIMUM LIKELIHOOD ESTIMATION OF REPLICATED LINEAR FUNCTIONAL RELATIONSHIP MODEL

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### Introduction



#### • Errors-in-variables model :

- functional relationship model (LFRM),
- structural relationship model and
- ultrastructural relationship model.
- •Measurement error can occur in many disciplines such as in econometrics, environmental sciences, engineering, manufacturing, and many others (Buonaccorsi, 2010; Doganaksoy and Van Meer, 2015; Hu and Wansbeek, 2017).

#### Linear functional relationship model

- \* the variable *X* is fixed or deterministic
- \*can be divided into unreplicated and replicated LFRM with certain recommendations



## **Objectives**



ODISCUSS the parameter estimates as well as the asymptotic covariance in replicated linear functional relationship model (LFRM).

### Emphasize on a balanced replicated LFRM

- \*Review the maximum likelihood estimation for the balanced replicated model
- Derive the asymptotic closed-form of the variance-covariance matrix
- Investigate the precision of the estimated parameters and its variance-covariance matrix by simulation study

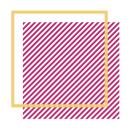


### **Literature Review**



Findings	Authors
the unidentifiability problem; the assumption on the ratio of error variance in unreplicated LFRM	Lindley, 1947; Kendall and Stuart, 1979
parameter estimation in the errors-in-variables model	Lindley, 1947; Villegas, 1961; Kendall and Stuart, 1979; Fuller, 1987; Buonaccorsi, 2010
maximum likelihood estimation method in estimating the parameters in both linear and circular models	Barnett, 1970; Chan and Mak, 1979; Hussin et al., 2005; Mokhtar et al., 2017
derived the asymptotic variance-covariance matrix in the errors-in-variables model	Hussin, 2005; Hussin et al., 2010; Mamun et al., 2013

Motivation: The balanced replicated LFRM can be used to overcome the inconsistencies i.e the unidentifiability problem and also the assumption on the ratio of error variance in unreplicated LFRM.



# Methodology



- Replicated Linear Functional Relationship Model
- Maximum Likelihood Estimation of the Model
- ■Variance-covariance Matrix of the Model
- **□**Simulation Study



# Replicated Linear Functional Relationship Model



 $\circ$  A linear relationship between  $X_i$  and  $Y_i$  are given by

$$Y_i = \alpha + \beta X_i$$

- $\alpha \implies \text{intercept parameter}$
- $\beta \implies$  slope parameter
- o  $x_{ij}$  and  $y_{ij}$  are subject to random errors  $\delta_{ij}$  and  $\varepsilon_{ij}$

$$x_{ij} = X_i + \delta_{ij} \longrightarrow \delta_{ij} \sim N(0, \sigma^2)$$

$$y_{ij} = Y_i + \varepsilon_{ik} \longrightarrow \varepsilon_{ik} \sim N(0, \tau^2)$$

for 
$$i = 1, 2, \dots, p$$
,  $j = 1, 2, \dots, m$  and  $k = 1, 2, \dots, m$ 

The replicated model is balanced and equal.



# Maximum Likelihood Estimation of 20 International Control of the Model



- Maximum Likelihood Estimation method which involves an iterative technique.
- The log likelihood function can be expressed as

$$\log L(\alpha, \beta, \sigma^{2}, \tau^{2}, X_{1}, \dots, X_{p}) = constant$$

$$-\frac{1}{2} \sum_{i=1}^{p} m(\log \sigma^{2} + \log \tau^{2}) - \frac{1}{2} \left\{ \sum_{i=1}^{p} \sum_{j=1}^{m} \frac{(x_{ij} - X_{i})^{2}}{\sigma^{2}} + \sum_{i=1}^{p} \sum_{j=1}^{m} \frac{(y_{ij} - \alpha - \beta X_{i})^{2}}{\tau^{2}} \right\}$$



### Maximum Likelihood Estimation of 29101 the Model



There are (p+4) parameters to be estimated and can be obtained by differentiating the log-likelihood function with respect to  $\hat{\alpha}$ ,  $\hat{\beta}$ ,  $\hat{\sigma}^2$ ,  $\hat{\tau}^2$  and  $\hat{X}_i$ :

$$\widehat{X}_{i} = \frac{1}{\widehat{\Delta}} \left\{ \frac{m \bar{x}_{i.}}{\widehat{\sigma}^{2}} + \frac{m \widehat{\beta}}{\widehat{\tau}^{2}} \left( \bar{y}_{i.} - \widehat{\alpha} \right) \right\}, \quad \widehat{\sigma}^{2} = \frac{\sum \sum \left( x_{ij} - \widehat{X}_{i} \right)^{2}}{\sum m}, \quad \widehat{\tau}^{2} = \frac{\sum \sum \left( y_{ij} - \widehat{\alpha} - \widehat{\beta} \widehat{X}_{i} \right)^{2}}{\sum m},$$

$$\hat{\alpha} = \frac{\sum m(\bar{y}_{i.} - \hat{\beta}\hat{X}_{i})}{\sum m} \text{ , } \hat{\beta} = \frac{\sum m\hat{X}_{i}(\bar{y}_{i.} - \hat{\alpha})}{\sum m\hat{X}_{i}^{2}} \text{ where } \bar{x}_{i.}, \bar{y}_{i.} \text{ are sample means for each}$$
 group and 
$$\hat{\Delta}_{i} = \frac{m}{\hat{\sigma}^{2}} + \frac{m\hat{\beta}^{2}}{\hat{\tau}^{2}}.$$

• The parameters can be solved iteratively by using unreplicated LFRM



# Variance-covariance Matrix of the Model



- By inverting the estimated Fisher information matrix for balanced replicated LFRM.
- The second derivative for the log-likelihood function is obtained followed by their negatives expected values.
- •The estimated Fisher information matrix, F, for  $\hat{X}_1, \dots, \hat{X}_p$ ,  $\hat{\sigma}^2$ ,  $\hat{\tau}^2$ ,  $\hat{\alpha}$  and  $\hat{\beta}$ :

$$F = \begin{bmatrix} B & 0 & E \\ 0 & C & 0 \\ E^T & 0 & D \end{bmatrix}$$

where B, C and D are a square matrix with sizes p, 2 and 2 respectively and E is a  $p \times 2$  matrix.

•The asymptotic covariance matrix of our interest,  $\hat{\sigma}^2$ ,  $\hat{\tau}^2$ ,  $\hat{\alpha}$  and  $\hat{\beta}$  is the bottom right minor of order  $4 \times p$  of the inverse of matrix F.



Model

# Variance-covariance Matrix of the



oFrom the theory of partitioned matrices, (Graybill, 1961), this is given by,

$$\widehat{Var} \begin{bmatrix} \widehat{\sigma}^2 \\ \widehat{\tau}^2 \\ \widehat{\alpha} \\ \widehat{\beta} \end{bmatrix} = \begin{bmatrix} C^{-1} & 0 \\ 0 & (D - E^T B^{-1} E)^{-1} \end{bmatrix}$$

•The asymptotic covariance matrix for  $\hat{\sigma}^2$ ,  $\hat{\tau}^2$ ,  $\hat{\alpha}$  and  $\hat{\beta}$  is given by

$$M = \begin{bmatrix} a_{11} & 0 & 0 & 0 \\ 0 & a_{22} & 0 & 0 \\ 0 & 0 & a_{33} & a_{34} \\ 0 & 0 & a_{43} & a_{44} \end{bmatrix}$$

where  $a_{11}=2\sigma^4/n$ ,  $a_{22}=2\tau^4/n$ ,  $a_{33}=Q\sum_{i=1}^p X_i^2$ ,  $a_{34}=-Q\sum_{i=1}^p X_i$ ,  $a_{44}=Qp$  and

$$a_{43} = -Q \sum_{i=1}^{p} X_i$$
 and  $Q = \frac{m\tau^2 + m\beta^2 \sigma^2}{m^2 \left\{ p \sum_{i=1}^{p} X_i^2 - \left(\sum_{i=1}^{p} X_i\right)^2 \right\}}$ . Other elements in matrix  $M$  is 0.



# **Simulation Study**



- •To evaluate the performance of the parameters.
- Fixed the true value of  $\alpha = 0$  and different true values of  $\beta$ ,  $\sigma^2$  and  $\tau^2$ .
- The sample size, n are 50,100 and 180 with p-subgroups of 5, 10, and 12.
- •The values of  $x_{ij}$  and  $y_{ij}$  are divided into p-subgroups with m elements such that  $p \times m = n$ .
- The parameters of interest can be solved iteratively.
- •The variance-covariance matrix of the parameters can be obtained.
- The steps are repeated for 5000 simulations.
- The performance of the estimated parameters is measured by estimated bias, mean square error, and standard deviation.



### Simulation Results and Discussion



Table 1. Results for  $\alpha = 0$  and  $\beta = 1$  with different sets of  $(\sigma^2, \tau^2)$  with n is the sample size

		(0.8,1)				(1,0.8)				
Statistics	Sample size, n	â	ĝ	$\hat{\sigma}^2$	$\hat{ au}^2$	$\hat{lpha}$	ĝ	$\hat{\sigma}^2$	$\hat{ au}^2$	
	50	0.0117	0.0024	0.0550	0.0659	0.0137	0.0027	0.0618	0.0582	
Estimated Bias	100	0.0088	0.0014	0.0499	0.0529	0.0089	0.0015	0.0527	0.0500	
	180	0.0010	0.0004	0.0315	0.0351	0.0004	0.0004	0.0329	0.0332	
Mean	50	0.1995	0.0045	0.0279	0.0422	0.1997	0.0045	0.0430	0.0275	
Square Error	100	0.0881	0.0023	0.0153	0.0227	0.0877	0.0023	0.0229	0.0151	
	180	0.0466	0.0012	0.0081	0.0123	0.0464	0.0012	0.0122	0.0081	
	50	0.4299	0.0648	0.1490	0.1868	0.4306	0.0649	0.1876	0.1484	
Standard Deviation	100	0.2812	0.0453	0.1061	0.1339	0.2814	0.0453	0.1340	0.1061	
	180	0.2090	0.0341	0.0810	0.1017	0.2091	0.0341	0.1019	0.0808	

- oThe estimated bias for estimated parameters,  $\hat{\alpha}$ ,  $\hat{\beta}$ ,  $\hat{\sigma}^2$  and  $\hat{\tau}^2$  become smaller and are approximately close to 0 when the sample size is increased from 50 to 180.
- This shows the unbiasedness of the parameters.
- The mean square error and the standard deviation also show similar trends
- The estimated values of parameters are consistent.



## Simulation Results and Discussion



Table 2. Results for  $\alpha = 0$  and  $\beta = 1.2$  with different sets of  $(\sigma^2, \tau^2)$  with n is the sample size

		(0.8,1)				(1,1)				
Statistics	Sample size, n	â	$\hat{eta}$	$\hat{\sigma}^2$	$\hat{ au}^2$	â	β̂	$\hat{\sigma}^2$	$\hat{ au}^2$	
	50	0.0146	0.0029	0.0504	0.0715	0.0177	0.0034	0.0597	0.0749	
Estimated Bias	100	0.0104	0.0017	0.0436	0.0608	0.0115	0.0019	0.0498	0.0655	
	180	0.0018	0.0006	0.0272	0.0403	0.0017	0.0006	0.0308	0.0435	
Mean	50	0.2383	0.0054	0.0276	0.0428	0.2707	0.0061	0.0428	0.0432	
Square Error	100	0.1049	0.0027	0.0148	0.0234	0.1188	0.0031	0.0227	0.0239	
	180	0.0556	0.0015	0.0078	0.0126	0.0629	0.0017	0.0121	0.0129	
	50	0.4704	0.0709	0.1499	0.1857	0.5017	0.0756	0.1881	0.1850	
Standard Deviation	100	0.3076	0.0496	0.1070	0.1328	0.3278	0.0528	0.1344	0.1322	
	180	0.2286	0.0373	0.0815	0.1012	0.2436	0.0397	0.1022	0.1008	

- oWhen the sample size is increased, the value of the estimated bias, the mean square error and the standard deviation also decreased.
- These results clearly show that the estimated values of parameters are unbiased and consistent.



### Conclusion



- The estimated parameters,  $\hat{\sigma}^2$ ,  $\hat{\tau}^2$ ,  $\hat{\alpha}$  and  $\hat{\beta}$  can be obtained iteratively using the maximum likelihood estimation method.
- •The variance-covariance matrix can be obtained using the Fisher information matrix and partitioned matrix.
- The results from simulation study suggest that the estimated parameters unbiased and consistent.





# THANK YOU

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