# Mathematics 640 Homework 3

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### Exercise 1

$$\mathbf{y}|\boldsymbol{\mu},\boldsymbol{\Sigma} \sim N(\boldsymbol{\mu},\boldsymbol{\Sigma})$$

$$\mu \sim N(\mu_0, \Delta_0)$$

Note: we should focus solely on the exponent of the multi-normal distribution.

Since the numerator and the formula within the integral of the denominator are identical please assume the steps approached are done for both.

$$\begin{split} &\exp\left\{-\frac{1}{2}*\sum_{i=1}^{n}\left(y_{i}-\boldsymbol{\mu}\right)^{T}*\boldsymbol{\Sigma}^{-1}*\left(y_{i}-\boldsymbol{\mu}\right)\right\}*\exp\left\{-\frac{1}{2}\left(-\boldsymbol{\mu}_{0}\right)^{T}\boldsymbol{\triangle}_{0}^{-1}\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)\right\} = \\ &\exp\left\{-\frac{1}{2}*\sum_{i=1}^{n}y_{i}^{T}\boldsymbol{\Sigma}^{-1}y_{i}-\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}y_{i}-\boldsymbol{y}_{i}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}\right\}*\exp\left\{-\frac{1}{2}\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)^{T}\boldsymbol{\triangle}_{0}^{-1}\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)\right\} = \\ &\exp\left\{-\frac{1}{2}*\left[\sum_{i=1}^{n}y_{i}^{T}\boldsymbol{\Sigma}^{-1}y_{i}-\sum_{i=1}^{n}\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}y_{i}-\sum_{i=1}^{n}y_{i}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\sum_{i=1}^{n}\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)^{T}\boldsymbol{\triangle}_{0}^{-1}\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)\right]\right\} = \\ &\exp\left\{-\frac{1}{2}*\left[\sum_{i=1}^{n}y_{i}^{T}\boldsymbol{\Sigma}^{-1}y_{i}-\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}n\bar{y}-n\bar{y}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+n\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)^{T}\boldsymbol{\triangle}_{0}^{-1}\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{0}\right)\right]\right\} \end{split}$$

Since  $\sum_{i=1}^n y_i^T \mathbf{\Sigma}^{-1} y_i$  does not contain  $\mu$ , it can be factored out

$$exp\left\{-\frac{1}{2}*\left(-\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}n\bar{y}-n\bar{y}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+n\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\boldsymbol{\mu}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}_{0}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}_{0}+\boldsymbol{\mu}_{0}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}_{0}\right)\right\}=0$$

$$exp\left\{-\frac{1}{2}*\left(-\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}n\bar{y}-n\bar{y}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+n\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\boldsymbol{\mu}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}_{0}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}_{0}+\boldsymbol{\mu}_{0}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}_{0}\right)\right\}$$

Since  $\mu_0^T \Delta_0^{-1} \mu_0$  does not contain  $\mu$ , it can be factored out

$$exp\left\{-\frac{1}{2}*\left(-\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}n\bar{y}-n\bar{y}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+n\boldsymbol{\mu}^{T}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}+\boldsymbol{\mu}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}_{0}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}^{T}\boldsymbol{\Delta}_{0}^{-1}\boldsymbol{\mu}_{0}\right)\right\}=0$$

$$exp\left\{-\frac{1}{2}*\left(\left(-n\bar{y}^T\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}_0^T\boldsymbol{\Delta}_0^{-1}\boldsymbol{\mu}\right)+\left(\boldsymbol{\mu}^T\boldsymbol{\Delta}_0^{-1}\boldsymbol{\mu}+\boldsymbol{\mu}^Tn\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}\right)+\left(-\boldsymbol{\mu}^T\boldsymbol{\Delta}_0^{-1}\boldsymbol{\mu}_0-\boldsymbol{\mu}^T\boldsymbol{\Sigma}^{-1}n\bar{y}\right)\right)\right\}=0$$

$$exp\left\{-\frac{1}{2}*\left(-\left(\bar{y}^{T}n\Sigma^{-1}+\mu_{0}^{T}\Delta_{0}^{-1}\right)\mu+\mu^{T}\left(\Delta_{0}^{-1}+n\Sigma^{-1}\right)\mu-\mu^{T}\left(\Delta_{0}^{-1}\mu_{0}+n\Sigma^{-1}\bar{y}\right)\right)\right\}$$

Since both  $\Delta_0^{-1}$  and  $\Sigma^{-1}$  are symmetric  $\left(\Delta_0^{-1}\right)^T = \Delta_0^{-1}$  and  $\left(\Sigma^{-1}\right)^T = \Sigma^{-1}$ 

$$exp\left\{-\frac{1}{2}*\left(-\left(\triangle_0^{-1}\boldsymbol{\mu}_0+n\boldsymbol{\Sigma}^{-1}\bar{\boldsymbol{y}}\right)^T\boldsymbol{\mu}+\boldsymbol{\mu}^T\left(\triangle_0^{-1}+n\boldsymbol{\Sigma}^{-1}\right)\boldsymbol{\mu}-\boldsymbol{\mu}^T\left(\triangle_0^{-1}\boldsymbol{\mu}_0+n\boldsymbol{\Sigma}^{-1}\bar{\boldsymbol{y}}\right)\right)\right\}$$

Let 
$$X = \Delta_0^{-1} \boldsymbol{\mu}_0 + n \boldsymbol{\Sigma}^{-1} \bar{y}$$
 and  $Y = \Delta_0^{-1} + n \boldsymbol{\Sigma}^{-1}$ 

note: Y is basically  $\triangle_n^{-1}$  and  $u_n = Y^{-1}X$  so the goal is to manipulate in the multivariate normal.

$$\exp\left\{ -\frac{1}{2}*\left( -\left( X\right) ^{T}\boldsymbol{\mu}+\boldsymbol{\mu}^{T}\left( Y\right) \boldsymbol{\mu}-\boldsymbol{\mu}^{T}\left( X\right) \right)\right\}$$

multiply the numerator and denominator by  $\exp\left\{-\frac{1}{2}*X^TY^{-1}X\right\}$ 

$$\exp\left\{-\frac{1}{2}\left(-\boldsymbol{\mu}\left(\boldsymbol{X}\right)^{T}-\boldsymbol{\mu}^{T}\left(\boldsymbol{X}\right)+\boldsymbol{\mu}^{T}\left(\boldsymbol{Y}\right)\boldsymbol{\mu}\right)\right\}*\exp\left\{-\frac{1}{2}*\boldsymbol{X}^{T}\boldsymbol{Y}^{-1}\boldsymbol{X}\right\}=$$

$$\exp\left\{-\frac{1}{2}\left(-\boldsymbol{\mu}\left(\boldsymbol{X}\right)^{T}-\boldsymbol{\mu}^{T}\left(\boldsymbol{X}\right)+\boldsymbol{\mu}^{T}\left(\boldsymbol{Y}\right)\boldsymbol{\mu}+\boldsymbol{X}^{T}\boldsymbol{Y}^{-1}\boldsymbol{X}\right)\right\}=$$

$$exp\left\{ -\frac{1}{2}\left(\boldsymbol{\mu}^{T}\left(Y\right)\boldsymbol{\mu}-\boldsymbol{\mu}\left(X\right)^{T}-\boldsymbol{\mu}^{T}\left(X\right)+X^{T}Y^{-1}X\right)\right\}$$

 $I = Y^{-1}Y = YY^{-1}$  because Y is symmetric and invertible

$$exp\left\{ -\frac{1}{2}\left(\boldsymbol{\mu}^{T}\left(Y\right)\boldsymbol{\mu}-\boldsymbol{\mu}Y\left(Y^{-1}X\right)^{T}-\boldsymbol{\mu}^{T}Y\left(Y^{-1}X\right)+\left(Y^{-1}X\right)^{T}Y\left(Y^{-1}X\right)\right)\right\}$$

Let 
$$\Delta_n^{-1} = Y = \Sigma^{-1}n + \Delta_0^{-1} = \Delta_0^{-1} + n\Sigma^{-1}$$
 and  $\mu_n = Y^{-1}X = (\Delta_0^{-1} + n\Sigma^{-1}) * (\Sigma^{-1}ny + \Delta_0^{-1}\mu_0)$ 

$$exp\left\{-\frac{1}{2}\left(\boldsymbol{\mu}^{T}\triangle_{n}^{-1}\boldsymbol{\mu}-\boldsymbol{\mu}\triangle_{n}^{-1}\mu_{n}^{T}-\boldsymbol{\mu}^{T}\triangle_{n}^{-1}\mu_{n}+\mu_{n}^{T}\triangle_{n}^{-1}\mu_{n}\right)\right\}=$$

$$exp\left\{-\frac{1}{2}\left(\left(\boldsymbol{\mu}^{T}-\boldsymbol{\mu}_{n}^{T}\right)\triangle_{n}^{-1}\left(\boldsymbol{\mu}-\boldsymbol{\mu}_{n}\right)\right)\right\}=$$

$$exp\left\{-\frac{1}{2}\left(\left(\boldsymbol{\mu}-\mu_{n}\right)^{T}\triangle_{n}^{-1}\left(\boldsymbol{\mu}-\mu_{n}\right)\right)\right\}$$
 Based on this kernel, it can be concluded that

$$\mu|Y \sim N(\mu_n, \triangle_n)$$
 where

$$\triangle_n^{-1} = \boldsymbol{\Sigma}^{-1} n + \triangle_0^{-1} = \triangle_0^{-1} + n \boldsymbol{\Sigma}^{-1} \text{ and } \mu_n = \left(\triangle_0^{-1} + n \boldsymbol{\Sigma}^{-1}\right) * \left(\boldsymbol{\Sigma}^{-1} n y + \triangle_0^{-1} \boldsymbol{\mu}_0\right)$$

### Exercise 2

#### Part A

Under  $H_0$ : the following are true

$$\pi_1 = \pi_2$$

$$X_j|\pi_1 \sim Binomial(n_j, \pi_1)$$
 where  $j = 1, 2$ 

$$\pi | H_0 \sim Beta(a,b)$$

under the assumption that  $x_1$  and  $x_2$  are independent:

$$p(x_1, x_2 | \pi_1, \pi_2) = p(x_1 | \pi_1) * p(x_2 | \pi_2) \left[ \begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \pi_1^{x_1} * (1 - \pi_1)^{n_1 - x_1} \right] * \left[ \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \pi_2^{x_2} * (1 - \pi_2)^{n_2 - x_2} \right]$$

since  $\pi_1 = \pi_2$  under  $H_0$ 

$$p(x_1, x_2 | \pi) = p(x_1 | \pi) * p(x_2 | \pi) \left[ \left( \begin{array}{c} n_1 \\ x_1 \end{array} \right) \pi^{x_1} * (1 - \pi)^{n_1 - x_1} \right] * \left[ \left( \begin{array}{c} n_2 \\ x_2 \end{array} \right) \pi^{x_2} * (1 - \pi)^{n_2 - x_2} \right] = 0$$

$$\binom{n_1}{x_1} * \binom{n_2}{x_2} \pi^{x_1+x_2} * (1-\pi)^{n_1-x_1+n_2-x_2}$$
 Since  $\pi|H_0 \sim Beta(a,b)$  then

$$p(x_1, x_2|H_0) = \int_{\pi} p(x_1, x_2|\pi) * p(\pi|H_0)d\pi =$$

$$\int \left(\begin{array}{c} n_1 \\ x_1 \end{array}\right) \left(\begin{array}{c} n_2 \\ x_2 \end{array}\right) \pi^{x_1 + x_2} (1 - \pi)^{n_1 - x_1 + n_2 - x_2} \left[\frac{1}{B(a, b)} \pi^{a - 1} (1 - \pi)^{b - 1}\right] d\pi =$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \frac{1}{B(a,b)} \int \pi^{x_1+x_2+a-1} (1-\pi)^{n_1-x_1+n_2-x_2+b-1} d\pi =$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \frac{B(x_1 + x_2 + a, n_1 - x_1 + n_2 - x_2 + b)}{B(a,b)} \int \frac{1}{B(x_1 + x_2 + a, n_1 - x_1 + n_2 - x_2 + b)} \pi^{x_1 + x_2 + a - 1} (1 - \pi)^{n_1 - x_1 + n_2 - x_2 + b - 1} d\pi = 0$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \frac{B(x_1+x_2+a,n_1-x_1+n_2-x_2+b)}{B(a,b)}$$

### Part B

Under  $H_1$ : the following are true

$$\pi_1 \neq \pi_2$$

 $X_j|\pi_1 \sim Binomial(n_j, \pi_1)$  where j = 1, 2

$$\pi | H_1 \sim Beta(a,b)$$
 where  $j=1,2$ 

under the assumption that  $x_1$  and  $x_2$  are independent:

$$p(x_1, x_2 | \pi_1, \pi_2) = p(x_1 | \pi_1) * p(x_2 | \pi_2) = \left[ \left( \begin{array}{c} n_1 \\ x_1 \end{array} \right) \pi_1^{x_1} * (1 - \pi_1)^{n_1 - x_1} \right] * \left[ \left( \begin{array}{c} n_2 \\ x_2 \end{array} \right) \pi_2^{x_2} * (1 - \pi_2)^{n_2 - x_2} \right]$$

$$p(x_1,x_2|H_1) = \int_{\pi_1} \int_{\pi_2} p(x_1,x_2|\pi_1,\pi_2) * p(\pi_1|H_1) * p(\pi_2|H_1) d\pi_2 d\pi_1 =$$

$$\int_{\pi_1} \int_{\pi_2} \left[ \left( \begin{array}{c} n_1 \\ x_1 \end{array} \right) \pi_1^{x_1} * (1 - \pi_1)^{n_1 - x_1} \right] * \left[ \left( \begin{array}{c} n_2 \\ x_2 \end{array} \right) \pi_2^{x_2} * (1 - \pi_2)^{n_2 - x_2} \right] *$$

$$\left(\frac{1}{B(a,b)}\pi_1^{a-1}(1-\pi_1)^{b-1}\right) * \left(\frac{1}{B(a,b)}\pi_2^{a-1}(1-\pi_2)^{b-1}\right) d\pi_2 d\pi_1 =$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \left(\frac{1}{B(a,b)}\right)^2 \int_{\pi_1} \pi_1^{x_1} * (1-\pi_1)^{n_1-x_1} * \left(\pi_1^{a-1}(1-\pi_1)^{b-1}\right) *$$

$$\int_{\pi_2} \left[ \pi_2^{x_2} * (1 - \pi_2)^{n_2 - x_2} \right] * \left( \pi_2^{a - 1} (1 - \pi_2)^{b - 1} \right) d\pi_2 d\pi_1 =$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \left(\frac{1}{B(a,b)}\right)^2 \int_{\pi_1} \pi_1^{x_1+a-1} * (1-\pi_1)^{n_1-x_1+b-1} \int_{\pi_2} \pi_2^{x_2+a-1} * (1-\pi_2)^{n_2-x_2+b-1} d\pi_2 d\pi_1 = 0$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \left(\frac{1}{B(a,b)}\right)^2 B(x_2+a,n_2-x_2+b) * B(x_1+a,n_1-x_1+b) \int_{\pi_1} \frac{1}{B(x_1+a,n_1-x_1+b)} \pi_1^{x_1+a-1} * \frac{1}{B(x_1+a,n_1-x_1+b$$

$$(1-\pi_1)^{n_1-x_1+b-1} \int_{\pi_2} \frac{1}{B(x_2+a,n_2-x_2+b)} \pi_2^{x_2+a-1} * (1-\pi_2)^{n_2-x_2+b-1} d\pi_2 d\pi_1 = 0$$

$$\begin{pmatrix} n_1 \\ x_1 \end{pmatrix} \begin{pmatrix} n_2 \\ x_2 \end{pmatrix} \left(\frac{1}{B(a,b)}\right)^2 B(x_2+a, n_2-x_2+b) * B(x_1+a, n_1-x_1+b)$$

#### Part C

$$p(H_0|x_1,x_2) = \frac{p(H_0)*p(x_1,x_2|H_0)}{p(x_1,x_2)} = \frac{p(H_0)*p(x_1,x_2|H_0)}{p(x_1,x_2|H_0)*p(H_0) + p(x_1,x_2|H_1)*p(H_1)} = \frac{\omega*p(x_1,x_2|H_0)}{p(x_1,x_2|H_0)*\omega + p(x_1,x_2|H_1)*(1-\omega)}$$

#### Part D

$$BF = \frac{p(H_1|x_1,x_2)/p(H_0|x_1,x_2)}{p(H_1)/p(H_2)} = \frac{\frac{(1-\omega)*p(x_1,x_2|H_1)}{p(x_1,x_2|H_0)*\omega+p(x_1,x_2|H_1)*(1-\omega)/\frac{p(x_1,x_2|H_0)*\omega+p(x_1,x_2|H_1)*(1-\omega)}{p(x_1,x_2|H_1)*(1-\omega)/\omega}}{\frac{(1-\omega)*p(x_1,x_2|H_1)/\omega*p(x_1,x_2|H_0)}{(1-\omega)/\omega}} = \frac{\frac{(1-\omega)*p(x_1,x_2|H_1)}{p(x_1,x_2|H_0)}}{\frac{(1-\omega)}{p(x_1,x_2|H_0)}} = \frac{\frac{p(x_1,x_2|H_1)}{p(x_1,x_2|H_0)}}{\frac{(1-\omega)}{p(x_1,x_2|H_0)}} = \frac{\frac{n_1}{n_1} \binom{n_2}{n_2} \binom{1}{B(a,b)}}{\binom{n_2}{x_2} \binom{1}{B(a,b)}}^{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}}{\binom{n_1}{x_1} \binom{n_2}{x_2} \binom{1}{B(a,b)}} = \frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{n_1-x_1+n_2-x_2+b)}} * \binom{1}{B(a,b)}$$

#### Part E

$$log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+n_2-x_2+b)}*\left(\frac{1}{B(a,b)}\right)\right) = log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+n_2-x_2+b)}\right) - log(B(a,b)) = log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+n_2-x_2+b)}\right) - log(B(a,b)) = log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+n_2-x_2+b)}\right) - log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+b)}\right) - log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+b)}\right) - log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+b)}\right) - log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+b)}\right) - log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+b)}\right) - log\left(\frac{B(x_2+a,n_2-x_2+b)*B(x_1+a,n_1-x_1+b)}{B(x_1+x_2+a,n_1-x_1+b)}\right)$$

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\begin{array}{l} logB = function\,(a\,,b) \{ \\ lgamma\,(a) + lgamma\,(b) - lgamma\,(a + b) \\ 1 = 46 \\ n1 = 143 \\ x2 = 30 \\ n2 = 151 \\ omega = 0.05 \\ a = 2 \\ b = 8 \\ exp\,(\,logB\,(\,x2 + a\,,\,n2 - x2 + b) + logB\,(\,x1 + a\,,\,n1 - x1 + b) - logB\,(\,x1 + x2 + a\,,\,n1 - x1 + n2 - x2 + b) - logB\,(\,a\,,\,b\,)) \\ 4.534823 \end{array}
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Under the Kass & Raftery – scale of evidence in favor of  $H_1$ , the interpretation is postive since  $3 < 4.534823 \le 20$ 

# Part F

prop. test (
$$x=c(x1, x2), n=c(n1, n2)$$
)

2-sample test for equality of proportions with

 $\begin{array}{c} \text{continuity correction} \\ \text{data:} \quad c(\text{x1, x2}) \text{ out of } c(\text{n1, n2}) \\ \text{X-squared} = 5.1733, \quad \text{df} = 1, \text{ p-value} = 0.02294 \\ \text{alternative hypothesis: two.sided} \\ 95 \text{ percent confidence interval:} \\ 0.01663715 \quad 0.22936850 \quad \text{sample estimates:} \\ \text{prop 1} \quad \text{prop 2} \quad 0.3216783 \quad 0.1986755 \\ \end{array}$ 

The p-value is significant, which leads us to reject  $H_0$  in favor of  $H_1$