

<u>Sampling distribution</u> <u>Conditional Distribution</u> $f_{Y X} = \frac{f_{XY}(x, y)}{f_X(x)}$ If independent: $f_{XY}(x, y) = f_X(x) f_Y(y)$ so $f_{Y X}(y x) = f_Y(y)$ <u>Law of Large Number</u> For LLN means Sample (x_1, x_2, \dots, x_n) $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$ is sample mean as $n \rightarrow \infty, \bar{x} \rightarrow E(X)$ <u>LLN mediums</u> m is sample mean As $n \rightarrow \infty, m \rightarrow \mu$ <u>Central Limit Theorem</u> Let $\bar{x}_n = \frac{1}{n} \sum_{i=1}^n x_i$ (mean of first n obs.) $Var(X) = \sigma^2$ Then: $\sqrt{n}(\bar{x}_n - E(X)) \sim N(0, \sigma^2)$ Let $S_n = \sum_{i=1}^n x_i$, Then $\frac{S_n - nE(X)}{\sqrt{n}} \sim N(0, \sigma^2)$ SD is more often used to refer to the individual observations, whereas SE is the SD of the sampling distribution <u>Neyman-Pearson Lemma</u> Of all tests of H_0 vs. H_a with given sig. level α , the likelihood ratio test has the largest power $P(T > C H_a)$ <u>Likelihood function</u> $L(\theta \dots)$ $L(\theta_a \dots)$ <u>Type I error (Drug example)</u> Drug is effective when it's not <u>Type II error (Drug example)</u> Drug is claimed ineffective when it is $P(\text{Type I error}) = P(\text{Reject } H_0 H_0 = T)$ $B = P(\text{Type II error}) = P(\text{Don't reject } H_0 H_a = T)$ $1 - B = 1 - P(\text{Reject } H_0 H_a = T)$ (power) <u>Rejection Region</u> Region in which the probability of the null hypothesis is true is lower than the defined significance level	<u>Permutative test</u> <u>comparing means</u> observed mean comparison used sampling to see varying differences checked to see how many sample difference are greater than observations - that percentage is the p-value <u>goodness of fit</u> Apply chi-square to a bunch of samples, then follow same logic as above <u>two-way tables</u> one sample isn't change other sample's order is randomized - check difference in mean follow remain steps <u>Bootstrapping</u> 1) Draw a resample of size n (size of pop.) with replacement Compute statistic (mean/median/etc.) 2) Repeat at least 10,000 times 3) Map dist. of Dist - To find SE of bootstrap, compute sd of Distribution - bias is mean of Dist. - mean(pop) two-sample sample each sample independent of the other small sample size with mean that the a mean varies	<u>confidence Interval for mean</u> If σ known, then use z score - If σ is unknown, then calculate sample SD (s) and use t-score When looking at Diff. of means $\bar{X} - \bar{Y} \sim N(\mu_1 - \mu_2, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2})$ use t-statistic if σ_1, σ_2 unknown $T = \frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{\sqrt{s_1^2/n_1 + s_2^2/n_2}}$ with df: $(s_1^2/n_1 + s_2^2/n_2)$ $(s_1^2/n_1)/(n_1-1) + (s_2^2/n_2)/(n_2-1)$ <u>T-Confidence Interval</u> $(\bar{X} - \bar{Y}) \pm q \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$ where q is T-score at $(1-\alpha/2)$ with above df <u>One-sided (I for mean)</u> q is $(1-\alpha)$ quantile of t-distribution $n-1$ df $\bar{X} - q \cdot \frac{s}{\sqrt{n}}$ (lower bound) upper bound ∞ so <u>upper confidence bound</u> $[\bar{X} - q \cdot \frac{s}{\sqrt{n}}, \infty]$ lower confidence bound $(-\infty, \bar{X} + q \cdot \frac{s}{\sqrt{n}}]$ <u>Confidence Interval for prop.</u> use z-score for q : the $(1-\alpha/2)$ quantile $CI = \hat{p} \pm q^z(zn) - q \sqrt{\hat{p}(1-\hat{p})/n + q^2/n}$ $1 + q^2/n$ (use prop. test in R) <u>CI for Difference in prop.</u> $\bar{P}_1 - \bar{P}_2 \pm q \sqrt{\frac{\bar{P}_1(1-\bar{P}_1)}{n_1} + \frac{\bar{P}_2(1-\bar{P}_2)}{n_2}}$ $\bar{P}_i = (x_i + 1)/(n_i + 2)$ $\bar{P}_i = n_i + 2$ bootstrap + Confidence Int. - bootstrap method - quantile for CI	<u>Correlation coefficient</u> $P(x, y) = \text{cov}(x, y) / (\sigma_x \sigma_y)$ <u>Simple Linear Regression</u> $y = \alpha + \beta x$ $\beta = r \frac{s_y}{s_x}$ $\alpha = \bar{y} - \beta \bar{x}$ <u>residuals</u> $r_i = y_i - (\alpha + \beta x_i)$ <u>Fitted Values</u> $\hat{y} = \alpha + \beta x$, \hat{y} is the fitted value aka predicted value.
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