

Homework 5

Exercise 1

1, 2, 3 are all aperiodical because $\gcd(1) = \gcd(2) = \gcd(3) = 1$

Exercise 2

States 1, 2 are periodic with $\gcd(1) = \gcd(2) = 2$ so the period for both stages is also 2.

Exercise 3

$$f(x|\theta) = \frac{1}{\theta + \frac{1}{2} - (\theta - \frac{1}{2})} = 1 \quad (1)$$

From $f(x|\theta)$, we know that $L(\theta|x) = \prod_{i=1}^n 1^n = 1$. This means that there is no unique value for θ

However $X_{(1)} > \theta - \frac{1}{2}$ and $X_{(n)} < \theta + \frac{1}{2}$

From Pitman 319, $f_{\min}(x) = n(1 - F(x))^{n-1}f(x)$ and $f_{\max}(x) = n(F(x))^{n-1}f(x)$

$$\begin{aligned} X_{(1)} &= \min(X_1 \dots X_n) \sim n(1/2 + \theta - x)^{n-1} \\ X_{(n)} &= \max(X_1 \dots X_n) \sim n(1/2 - \theta + x)^{n-1} \end{aligned} \quad (2)$$

The MLE of θ should occur when $X_{(1)} = X_{(n)}$

$$\begin{aligned} n(1/2 + \theta - x_{(1)})^{n-1} &= n(1/2 - \theta + x_{(n)})^{n-1} \implies \\ 1/2 + \theta - x_{(1)} &= 1/2 - \theta + x_{(n)} \implies \\ \theta - x_{(1)} &= -\theta + x_{(n)} \implies \\ \theta &= \frac{x_{(1)} + x_{(n)}}{2} \end{aligned} \quad (3)$$

$$\hat{\theta} \sim \frac{X_{(1)} + X_{(n)}}{2}$$

Exercise 4

Part A

Let the prior be $\theta \sim \text{unif}(0, 1)$ and $X|\theta \sim \text{Bernoulli}(\theta)$

$$\begin{aligned} p(\theta|X) &\propto \prod_{i=1}^n p(X_i|\theta) * p(\theta) = \\ &\prod_{i=1}^n \theta^{x_i} (1 - \theta)^{1-x_i} = \\ &\theta^{\sum_{i=1}^n x_i} (1 - \theta)^{\sum_{i=1}^n 1-x_i} = \\ &\theta^{\sum_{i=1}^n x_i} (1 - \theta)^{n - \sum_{i=1}^n x_i} \end{aligned} \quad (4)$$

Let $y = \sum_{i=1}^n x_i$ so,

$$\theta^y (1 - \theta)^{n-y} = \theta^{(y+1)-1} * (1 - \theta)^{(n-y+1)-1} \propto \text{beta}(y+1, n-y+1) \quad (5)$$

\therefore the Bayes Estimate is $\frac{y+1}{y+1+n-y+1} = \frac{y+1}{n+2}$

Part B

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n \theta^{x_i} (1 - \theta)^{1-x_i} \implies \\ \text{Log}(L(\theta)) &= \sum_{i=1}^n \log(\theta^{x_i} (1 - \theta)^{1-x_i}) = \\ &= \sum_{i=1}^n \log(\theta^{x_i}) + \sum_{i=1}^n \log((1 - \theta)^{1-x_i}) = \\ &= \sum_{i=1}^n x_i \log(\theta) + \sum_{i=1}^n (1 - x_i) \log(1 - \theta) = \\ &= n\bar{x} * \log(\theta) + (n - n\bar{x}) * \log(1 - \theta) \\ \delta / d\theta (\log(L(\theta))) &= \frac{n\bar{x}}{\theta} - \frac{n - n\bar{x}}{1 - \theta} = 0 \implies \\ &= \frac{n\bar{x}}{\theta} = \frac{n - n\bar{x}}{1 - \theta} \implies \\ n\bar{x} * (1 - \theta) &= (n - n\bar{x})(\theta) \implies \\ &= n\bar{x} = n\theta \end{aligned} \quad (6)$$

$\therefore \hat{\theta} = \bar{x} = \frac{y}{n}$

Since the prior distribution is uniform, $E(\theta) = \frac{1}{2}$

$$\begin{aligned} \frac{1}{2}(1 - \omega) + \frac{y}{n}(\omega) &= \frac{y+1}{n+2} \implies \\ \omega &= \frac{n}{n+2} \implies \\ 1 - \omega &= \frac{2}{n+2} \end{aligned} \quad (8)$$

$$\frac{1}{2}(\frac{2}{n+2}) + \frac{y}{n}(\frac{n}{n+2}) = \frac{y+1}{n+2} \quad (9)$$

Exercise 5

Let $f(x; \theta) = e^{-(x-\theta)}$, $Y_1 = \min(X_i)$, and $T = T(X_1, \dots, X_n)$

$$F(x) = \int_{\theta}^x e^{-(x-\theta)} dx = 1 - e^{\theta-x} \quad (10)$$

From Pitman 319, $f_{\min}(x) = n(1 - F(x))^{n-1} f(x)$

$$\begin{aligned} f_{\min}(x) &= n(1 - (1 - e^{\theta-x}))^{n-1} e^{-(x-\theta)} = \\ &= n(e^{\theta-x})^{n-1} e^{-(x-\theta)} = \\ &= n * (e^{\theta-x})^n \end{aligned} \quad (11)$$

$$\begin{aligned}
P(X_1, \dots, X_n | T = t) &= \frac{e^{n\theta - \sum x_i}}{n * (e^{\theta - t})^n} = \\
&\frac{1}{n} e^{n\theta - \sum x_i - (\theta - t)n} = \\
&\frac{1}{n} e^{-\sum x_i - (-t)n}
\end{aligned} \tag{12}$$

Since the conditional density and range do not depend on θ , $Y_1 = \min(X_i)$ is a sufficient Statistic